No. 16-1161

## 

Bev Erly R. GILL, ET AL., APPELLANTS,
v.

WILLIAM Whitford, et AL., APPELLEES.

On Appeal From the United States District Court for the Western District of Wisconsin

## JOINT APPENDIX

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Appeal Docketed March 24, 2017
Jurisdiction Postponed June 19, 2017

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## Relevant Docket Entries

## U.S. District Court <br> Western District of Wisconsin (Madison) <br> CIVIL DOCKET FOR CASE \#: <br> 3:15-cv-00421-bbc

Case Title: Whitford, William, et al. v. Nichol, Gerald et al.
Date Filed: July 8, 2015

Assigned to: District Judge Barbara B. Crabb

## Date \# Docket Text

07/08/2015 1 COMPLAINT against All
Defendants, filed by All Plaintiffs. (Attachments omitted) (Earle, Peter) Modified on 7/8/2015. (Entered: 07/08/2015)

08/18/2015 24 Notice of Motion and MOTION TO DISMISS by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) Modified on 8/18/2015. (kwf) (Entered: 08/18/2015)

08/08/2015 25 Brief in Support of 24 Motion to Dismiss by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) (Entered: 08/18/2015)

09/29/2015 30 ORDER appointing Circuit Judge Kenneth F. Ripple and Chief District Judge William C. Griesbach, of the Eastern District of Wisconsin, as additional members of the three-judge court. Signed by District Judge Diane P. Wood, Chief Judge USCA for the seventh circuit on $9 / 23 / 2015$. (voc) (Entered: 09/29/2015)

09/29/2015 31 Brief in Opposition by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter re: 24 Motion to Dismiss filed by Gerald C. Nichol, Harold V. Froehlich, Timothy Vocke, John

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Franke, Elsa Lamelas, Kevin J. Kennedy, Thomas Barland. (Odorizzi, Michele) (Entered: 09/29/2015)

10/09/2015 32 Brief in Reply by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke in Support of 24 Motion to Dismiss. (Keenan, Brian) (Entered: 10/09/2015)

10/15/2015 33 Pretrial Conference Order - Oral Argument on Motion to Dismiss 24 set for 11/4/2015 at 01:30 PM before the Three Judge Panel. Dispositive Motions due $1 / 4 / 2016$. Final Pretrial Submissions due 4/25/2016. Joint Pretrial Statement and Each Party's Statement of Facts with Proposed Special Verdict Form due 5/9/2016. Trial Brief and Five Complete sets of Premarked Trial Exhibits due $5 / 16 / 2016$. Court Trial set for 5/23/2016 at 09:00 AM before the Three Judge Panel. Signed by Magistrate Judge Stephen L.

Crocker on 10/15/15. (jat) (Entered: 10/15/2015)

11/23/2015 39 Defendant's Supplement to 24 Motion to Dismiss filed by Gerald C. Nichol, Harold V. Froehlich, Timothy Vocke, John Franke, Elsa Lamelas, Kevin J. Kennedy, Thomas Barland. (Russomanno, Anthony) Modified on 11/23/2015. (lak) (Entered: 11/23/2015)

11/23/2015 40 Plaintiff's Supplement to 24 Motion to Dismiss filed by Gerald C. Nichol, Harold V. Froehlich, Timothy Vocke, John Franke, Elsa Lamelas, Kevin J. Kennedy, Thomas Barland. (Odorizzi, Michele) Modified on 11/24/2015. (lak) (Entered: 11/23/2015)

11/30/2015 41 Reply by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter re: 24 Motion to Dismiss. (Plaintiffs

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Reply to Defendants' Supplemental Brief on Standing.) (Odorizzi, Michele) Modified on 12/1/2015. (lak) (Entered: 11/30/2015)

11/30/2015 42 Reply by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke re: 24 Motion to Dismiss. (Response Brief on Standing.) (Keenan, Brian) Modified on 12/1/2015. (lak) (Entered: 11/30/2015)

12/17/2015 43 ORDER denying 24 Motion to Dismiss by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. Signed by District Judge Barbara B. Crabb on 12/17/2015. (voc) (Entered: 12/17/2015)

12/30/2015 44 ANSWER by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa

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Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian)
(Entered: 12/30/2015)

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01/04/2016 46 Brief in Support of 45 Motion for Summary Judgment, by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) (Entered: 01/04/2016)

01/04/2016 49 Declaration of Brian Keenan filed by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke re: 45 Motion for Summary

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Judgment, (Attachments omitted). (Keenan, Brian) Modified on 1/5/2016. (lak) (Entered: 01/04/2016)

01/04/2016 50 Declaration of Nicholas Goedert filed by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke re: 45 Motion for Summary Judgment, (Attachments omitted). (Keenan, Brian) Modified on 1/5/2016. (lak) (Entered: 01/04/2016)

01/05/2016 51 Expert Report of Nicholas Goedert by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke, (Attachments omitted). (Keenan, Brian) Modified on 1/5/2016. (lak) (Entered: 01/05/2016)

01/05/2016 54 Expert Report of Kenneth Mayer by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald

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C. Nichol, Timothy Vocke. (Keenan, Brian) Modified on 1/5/2016: Exhibits/Annex are not attached separately. (lak) (Entered: 01/05/2016)

01/05/2016 55 Declaration of Sean Trende filed by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke re: 45 Motion for Summary Judgment, (Attachments omitted) (Keenan, Brian) Modified on 1/5/2016. (lak) (Entered: 01/05/2016)

01/15/2016 56 AMENDED ANSWER by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) Modified on $1 / 18 / 2016$. (lak) (Entered: 01/15/2016)

01/22/2016 58 Declaration of Simon David Jackman filed by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue

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Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter re: 45 Motion for Summary Judgment, (Attachments omitted) (Odorizzi, Michele) Modified on 1/25/2016: Clarified exhibit descriptions. (lak) (Entered: 01/22/2016)

01/22/2016 59 Declaration of Kenneth Mayer filed by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter re: 45 Motion for Summary Judgment, (Attachments omitted) (Odorizzi, Michele) Modified on 1/25/2016: Clarified exhibit descriptions. (lak) (Entered: 01/22/2016)

01/25/2016 62 Expert Report of Simon David Jackman by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison

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Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Odorizzi, Michele) (Entered: 01/25/2016)

01/25/2016 63 Expert Report of Simon David Jackman (Rebuttal) by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Odorizzi, Michele) (Entered: 01/25/2016)

01/25/2016 64 Expert Report of Kenneth Mayer (Rebuttal) by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Odorizzi, Michele) (Entered: 01/25/2016)

01/25/2016 65 Deposition of Nicholas Goedert taken on 12/15/15, (Attachments

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omitted) (Odorizzi, Michele) (Entered: 01/25/2016)

01/25/2016 68 Brief in Opposition by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter re: 45 Motion for Summary Judgment, filed by Gerald C. Nichol, Harold V. Froehlich, Timothy Vocke, John Franke, Elsa Lamelas, Kevin J. Kennedy, Thomas Barland. (Odorizzi, Michele) (Entered: 01/25/2016)

02/04/2016 73 Brief in Reply by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke in Support of 45 Motion for Summary Judgment. (Keenan, Brian) (Entered: 02/04/2016)

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Crabb, Circuit Judge Kenneth F. Ripple and Chief District Judge William C. Griesbach: Oral Argument Hearing held on $3 / 23 / 2016$ re 45 Motion for Summary Judgment by defendants. [2:12] (Court Reporter LS.) (voc) (Entered: 03/23/2016)

03/25/2016 89 Transcript of Motion Hearing, held 3/23/2016 before Judge Kenneth Ripple, Judge Barbara B. Crabb and Judge William Griesbach. (voc) (Entered: 03/25/2016)

04/07/2016 94 ORDER: IT IS ORDERED that The motion for summary judgment filed by defendants Gerald C. Nichol, Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas, Timothy Vocke and Kevin J. Kennedy, dkt. \# 45 , is DENIED.

The motion filed by plaintiffs William Whitford, Roger Anclam, Emily Bunting, Mary Lynne Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, James Seaton, Allison Seaton, Jerome Wallace and Don

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Winter to exclude the opinions of Sean Trende, dkt. \# 70, is DENIED WITHOUT PREJUDICE to plaintiffs' refiling it at the conclusion of trial.

Trial will begin on Tuesday, May, 24, 2016 and should be completed by Friday, May 27, 2016. If the parties believe that is not a sufficient amount of time, they should explain their concerns in writing no later than April 18, 2016.

Signed by Circuit Judge Kenneth F. Ripple, District Judge Barbara B. Crabb and District Judge William C. Griesbach on 4/7/2016. (voc) (Entered: 04/07/2016)

04/18/2016 95 Amended Expert Report of Kenneth Mayer (Rebuttal), Updated March 31, 2016, by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter.
(Greenwood, Ruth) (Entered: 04/18/2016)

05/09/2016 125 Joint Final Pretrial Conference Report by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Greenwood, Ruth) Modified on 5/10/2016: Requested exhibit lists and deposition designations be filed as separate docket entries. Exhibit lists also filed at 102 and 103 . (lak) (Entered: 05/09/2016)

05/10/2016 130 Exhibit List by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) (Entered: 05/10/2016)

05/16/2016 133 Trial Brief by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol,

Timothy Vocke. (Keenan, Brian) (Entered: 05/16/2016)

05/16/2016 134 Trial Brief by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Greenwood, Ruth) (Entered: 05/16/2016)

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05/26/2016 143 Minute Entry for proceedings held before Circuit Judge Kenneth F. Ripple, District Judge Barbara B. Crabb and District Judge William C. Griesbach: Third Day of Court Trial held on $5 / 26 / 2016$. Evidence entered, trial continues. [7:08] (Entered: 05/26/2016)

05/27/2016 145 Minute Entry for proceedings held before Circuit Judge Kenneth F. Ripple, District Judge Barbara B. Crabb and District Judge William
C. Griesbach: Fourth Day of Court Trial held on 5/27/2016. Trial completed, briefing set: Post-trial briefs due 6/10/2016, Replies due 6/20/2016. [6:35] (arw) (Entered: 05/27/2016)

05/27/2016 146 Court Trial Exhibit List. (arw) (Entered: 05/27/2016)

06/08/2016 147 Transcript of First Day of Court Trial, held 5/24/2016 before Judge Barbara B. Crabb. Court Reporter: LS. (voc) (Entered: 06/08/2016)

06/08/2016 148 Transcript of Second Day of Jury Trial, held 5/25/2016 before Judge Barbara B. Crabb. Court Reporter: LS. (voc) (Entered: 06/08/2016)

06/08/2016 149 Transcript of Third Day of Jury Trial, held 5/26/2016 before Judge Barbara B. Crabb. Court Reporter: LS. (voc) (Entered: 06/08/2016)

06/10/2016 150 Transcript of Fourth Day of Court Trial, held 5/27/16 before Judge Barbara B. Crabb, Judge Kenneth

Ripple, and Judge William Griesbach. Court Reporter:LS. (jat) (Entered: 06/10/2016)

06/10/2016 153 Post Trial Brief by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) (Entered: 06/10/2016)

06/10/2016 155 Post Trial Brief by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Greenwood, Ruth) (Entered: 06/10/2016)

06/20/2016 156 Post Trial Brief by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) (Entered: 06/20/2016)

06/20/2016 157 Post Trial Brief (Reply) by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter. (Harless, Annabelle) (Entered: 06/20/2016)

11/21/2016 166 OPINION and ORDER. Signed by Judges Kenneth F. Ripple, Barbara B. Crabb and William C. Griesbach. Signed by District Judge Barbara B. Crabb on 11/21/2016. (voc) (Entered: 11/21/2016)

12/21/2016 169 Response re: 166 OPINION and ORDER. Brief on Remedy by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) Modified on 12/22/2016. (lak) (Entered: 12/21/2016)

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Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter, (Attachments omitted) Modified on 12/22/2016. (lak) (Entered: 12/21/2016

01/05/2017 173 Response re: 166 OPINION and ORDER. Response Brief on Remedies by Defendants Thomas Barland, John Franke, Harold V. Froehlich, Kevin J. Kennedy, Elsa Lamelas, Gerald C. Nichol, Timothy Vocke. (Keenan, Brian) Modified on 1/5/2017. (lak) (Entered: 01/05/2017)

01/05/2017 174 Response re: 166 OPINION and ORDER. Response Brief on Remedies by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald

Winter. (Harless, Annabelle) Modified on 1/6/2017. (lak) (Entered: 01/05/2017)

01/27/2017 182 OPINION and ORDER. Signed by Judges Kenneth F. Ripple, Barbara B. Crabb and William C. Griesbach on 1/27/2017. (voc) (Entered: 01/27/2017)

01/27/2017 183 JUDGMENT entered in favor of Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter dismissing the case. (voc) (Entered: 01/27/2017)

02/06/2017 185 Motion to Alter or Amend Judgment to Retain Jurisdiction Regarding Remedy by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome

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Wallace, William Whitford, Donald Winter. (Attachments omitted) (Poland, Douglas) (Entered: 02/06/2017)

02/14/2017 187 Brief in Opposition by Defendants Beverly R. Gill, Julie M. Glancey, Ann S. Jacobs, Steve King, Don Millis, Mark L. Thomsen re: 185 Motion to Alter or Amend Judgment, filed by Jerome Wallace, Allison Seaton, Helen Harris, Donald Winter, James Seaton, Emily Bunting, Wayne Jensen, William Whitford, Janet Mitchell, Wendy Sue Johnson, Mary Lynn Donohue, Roger Anclam. (Keenan, Brian) (Entered: 02/14/2017)

02/16/2016 188 Brief in Reply by Plaintiffs Roger Anclam, Emily Bunting, Mary Lynn Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter in Support of 185 Motion to Alter or Amend Judgment. (Harless, Annabelle) (Entered: 02/16/2017)

JA23

02/22/2017 189 ORDER granting 185 Motion to Alter or Amend Judgment to Retain Jurisdiction Regarding Remedy. Signed by Judges Kenneth F. Ripple, Barbara B. Crabb and William C. Griesbach on 2/22/2016. (voc) (Entered: 02/22/2017)

02/22/2017 190 AMENDED JUDGMENT entered in favor of Plaintiffs Roger Anclam, Emily Bunting, Mary Lynne Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, William Whitford, Donald Winter dismissing the case. (voc) (Entered: 02/22/2017)

02/24/2017 191 NOTICE OF APPEAL by Defendants Beverly R. Gill, Julie M. Glancey, Ann S. Jacobs, Steve King, Don Millis, Mark L. Thomsen as to 182 Order, 190 Judgment. Filing fee of $\$ 505$, receipt number 0758-1977883 paid. No Docketing Statement filed. (Keenan, Brian) (Entered: 02/24/2017)

## JA24

03/15/2017 192 Judgment Corrected Pursuant to Rule 60(a) to correct the inadvertent omission of court approval of form as required by Rule 58(b)(2) (BBC /PAO). (voc) (Entered: 03/15/2017)

03/20/2017 193 AMENDED NOTICE OF APPEAL by Defendants Beverly R. Gill, Julie M. Glancey, Ann S. Jacobs, Steve King, Don Millis, Mark L. Thomsen as to 190 Judgment, 192 Judgment, 183 Judgment. Filing fee of \$505, receipt number 0758-1977883 paid. No Docketing Statement filed. (Keenan, Brian) (Entered: 03/20/2017)

## Complaint for Declaratory and Injunctive Relief

## IN THE UNITED STATES DISTRICT COURT FOR THE WESTERN DISTRICT OF WISCONSIN

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WILLIAM WHITFORD, )
ROGER ANCLAM, EMILY )
BUNTING, MARY LYNNE )
DONOHUE,HELEN ) No.
HARRIS, WAYNE JENSEN, )
WENDY SUE JOHNSON, )
JANET MITCHELL, )
ALLISON SEATON, JAMES )
SEATON, JEROME )
WALLACE, and DONALD )
WINTER,
    Plaintiffs,
    v.
GERALD C. NICHOL,
) }28\mathrm{ U.S.C. 2284(a)
THOMAS BARLAND, JOHN )
FRANKE, HAROLD V. )
FROEHLICH, KEVIN J. )
KENNEDY, ELSA )
LAMELAS, and TIMOTHY )
VOCKE,
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Defendants. )
)
)

NOW COME Plaintiffs William Whitford, Roger Anclam, Emily Bunting, Mary Lynne Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, Allison Seaton, James Seaton, Jerome Wallace, and Donald Winter, by their undersigned attorneys, and complain of Defendants Gerald C. Nichol, Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas, Kevin J. Kennedy, and Timothy Vocke, not personally, but solely in their official capacities as members of the Wisconsin Government Accountability Board, as follows:

## INTRODUCTION

1. Plaintiffs seek both a declaratory judgment that the Wisconsin State Assembly district plan adopted in 2012 by Wisconsin Act 43 (the "Current Plan") violates the First and Fourteenth Amendments of the United States Constitution and an order permanently enjoining the implementation of the Current Plan in the 2016 election. As explained in greater detail below, the Current Plan is, by any measure, one of the worst partisan gerrymanders in modern American history. In the first election in which it was in force in 2012, the Current Plan enabled Republican candidates to win sixty of the Assembly's ninety-nine seats even though Democratic candidates won a majority of the statewide Assembly
vote. The evidence is overwhelming that the Current Plan was adopted to achieve precisely that result: indeed, before submitting the map for approval, the Republican leadership retained an expert (at State expense) who predicted the partisan performance of each proposed district-as it turned out, with remarkable accuracy.
2. This kind of partisan gerrymandering is both unconstitutional and profoundly undemocratic. It is unconstitutional because it treats voters unequally, diluting their voting power based on their political beliefs, in violation of the Fourteenth Amendment's guarantee of equal protection, and because it unreasonably burdens their First Amendment rights of association and free speech. Extreme partisan gerrymandering is also contrary to core democratic values because it enables a political party to win more legislative districts-and thus more legislative power-than is warranted by that party's popular support. By distorting the relationship between votes and assembly seats, it causes policies to be enacted that do not accurately reflect the public will. In the end, a political minority is able to rule the majority and to entrench itself in power by periodically manipulating election boundaries.
3. Partisan gerrymandering has increased throughout the United States in recent years as a result of both a rising tide of partisanship and greater technological sophistication, which enables maps to be drawn in ways that are likely to enable the party
in power to remain in power even if it no longer represents the views of the majority of voters. This nationwide trend threatens a "core principle of republican government,' namely, 'that the voters should choose their representatives, not the other way around." Arizona State Legislature v. Arizona Independent Redistricting Comm'n, No. 13-1314 (U.S. June 29, 2015), slip op. at 35.
4. The United States Supreme Court has recognized that partisan gerrymandering can be unconstitutional. Nevertheless, a constitutional challenge has yet to succeed on that ground because plaintiffs have been unable to offer a workable standard to distinguish between permissible political line-drawing and unconstitutional partisan gerrymandering. In this case, plaintiffs propose a new test that is workable, based on the concept of partisan symmetry-the idea that a district plan should treat the major parties symmetrically with respect to the conversion of votes to seats and that neither party should have a systematic advantage in how efficiently its popular support translates into legislative power.
5. One way to measure a district plan's performance in terms of partisan symmetry is to determine whether there is an "efficiency gap" between the performances of the two major parties and, if so, to compare the magnitude of that gap to comparable district plans in the modern era nationwide. The efficiency gap captures in a single
number all of a district plan's cracking and packing-the two fundamental ways in which partisan gerrymanders are constructed. Cracking means dividing a party's supporters among multiple districts so that they fall short of a majority in each one. Packing means concentrating one party's backers in a few districts that they win by overwhelming margins. Both cracking and packing result in "wasted" votes: votes cast either for a losing candidate (in the case of cracking) or for a winning candidate but in excess of what he or she needed to prevail (in the case of packing). The efficiency gap is the difference between the parties' respective wasted votes in an election, divided by the total number of votes cast.
6. When the efficiency gap is relatively small and roughly equivalent to the efficiency gaps that have traditionally existed, the map should not be deemed unconstitutional. In such cases, there may be no intent to treat voters unequally; in any event, the effects of any gerrymandering are likely to be redressable through the political process. But where the efficiency gap is large and much greater than the historical norm, there should be a presumption of unconstitutionality. In such a case, an intent to systematically disadvantage voters based on their political beliefs can be inferred from the severity of the gerrymander alone. And because such severe gerrymanders are likely to be extremely durable as well, it is unlikely that the disadvantaged party's adherents will be able to protect themselves through
the political process. Where partisan gerrymandering is extreme, the process itself is broken: current legislators have no incentive to alter it, and adherents of the disadvantaged party are unable to do so because their votes have been unfairly diluted.
7. Wisconsin's Current Plan is presumptively unconstitutional under this analysis. In the 2012 election, the Current Plan resulted in an efficiency gap of roughly $13 \%$ in favor of Republican candidates. Between 1972 and 2014, fewer than four percent of all state house plans in the country benefited a party to that extent. In the 2014 election, the efficiency gap remained extremely high at $10 \%$. Between 1972 and 2010, not a single plan anywhere in the United States had an efficiency gap as high as the Current Plan in the first two elections after redistricting. A district plan this lopsided is also highly unlikely ever to become neutral over its ten-year lifespan. Indeed, we can predict with nearly $100 \%$ confidence that, absent this Court's intervention, Wisconsin's Current Plan will continue to unfairly favor Republican voters and candidates-and unfairly disadvantage Democratic voters and candidates-throughout the remainder of the decade.
8. There are three additional facts that reinforce the conclusion that the Current Plan is unconstitutional. First, the Current Plan was not the result of an ordinary political process, where a bill is formulated through a give-and-take between political adversaries and subject to open debate. Instead, it
was drawn up in secret by the Legislature's Republican leadership, without consultation with Democratic leaders or rank-and-file members of either party, with the purpose and intent of altering what was already a favorable map to maximize the Republican Party's partisan advantage. Then the proposal was rammed through the Assembly, without any opportunity for real debate.
9. Second, the Current Plan is also an outlier by another measure of partisan symmetry-partisan bias. Partisan bias is the difference in the share of seats that each party would win if they tied statewide, each receiving $50 \%$ of the vote. In 2012, there was a $13 \%$ bias in favor of Republicans; in a tied election, Republicans would have won $63 \%$ of the Assembly seats, with Democrats winning only $37 \%$. In 2014, there was a $12 \%$ bias in favor of Republicans.
10. Third, the Current Plan's extreme partisan skew was entirely unnecessary. Plaintiffs have designed a Demonstration Plan that complies at least as well as the Current Plan with every legal requirement-equal population, the Voting Rights Act, compactness, and respect for political subdivisions-but that is almost perfectly balanced in its partisan consequences. Thus, defendants cannot salvage the Current Plan on the theory that adherence to redistricting criteria or the State's underlying political geography made an unfair plan unavoidable.
11. To be clear, plaintiffs do not seek to replace a pro-Republican gerrymander with a plan that is gerrymandered to be pro-Democratic. Rather, plaintiffs seek as a remedy the creation of a neutral plan that is not gerrymandered to give either side an unfair partisan advantage.

## JURISDICTION AND VENUE

12. This Court has jurisdiction over this action pursuant to 28 U.S.C. §§ 1331, 1343(a)(3) and (4), and 2284. It also has jurisdiction under 28 U.S.C. §§ 2201 and 2202, the Declaratory Judgments Act, to grant the declaratory relief requested.
13. Pursuant to 28 U.S.C. § 2284(a), a three-judge panel should be convened to hear this case.
14. Venue is proper in this judicial district under 28 U.S.C. § 1391(b). At least one of the Defendants resides in the Western District of Wisconsin. In addition, at least six of the plaintiffs reside and vote in this judicial district.

## PARTIES

15. Plaintiffs are qualified, registered voters in the State of Wisconsin, who reside in various counties and legislative districts. Plaintiffs are all supporters of the public policies espoused by the Democratic Party and of Democratic Party candidates. Together with other Democratic voters, plaintiffs have been
harmed by the Current Plan's unlawful partisan gerrymandering because it treats Democrats unequally based on their political beliefs and impermissibly burdens their First Amendment right of association. Some of the plaintiffs have been packed into districts with other Democratic voters, while others live in districts that have been cracked by the Current Plan to disadvantage Democratic candidates in close races. Either way, the purpose and effect of the Current Plan is to dilute their voting strength because of their political affiliations.
16. Regardless of where they reside in Wisconsin and whether they themselves reside in a district that has been packed or cracked, all of the plaintiffs have been harmed by the manipulation of district boundaries in the Current Plan to dilute Democratic voting strength. As a result of the statewide partisan gerrymandering, Democrats do not have the same opportunity provided to Republicans to elect representatives of their choice to the Assembly. As a result, the electoral influence of plaintiffs and other Democratic voters statewide has been unfairly, disproportionately, and undemocratically reduced.
17. Plaintiff William Whitford, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 76th Assembly District in Madison in Dane County, Wisconsin.
18. Plaintiff Roger Anclam, a citizen of the United States and of the State of Wisconsin, is a resident and
registered voter in the 31st Assembly District in Beloit in Rock County, Wisconsin.
19. Plaintiff Emily Bunting, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 49th Assembly District in Richland County, Wisconsin.
20. Plaintiff Mary Lynne Donohue, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 26th Assembly District in Sheboygan in Sheboygan County, Wisconsin. In addition to the injury suffered by all Democrats in Wisconsin, Ms. Donohue was harmed when the City of Sheboygan was split into Districts 26 and 27 and District 26 was cracked and converted from a Democratic to a Republican district. See infra -| 9 63-65.
21. Plaintiff Helen Harris, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 22nd Assembly District in Milwaukee, in Milwaukee County, Wisconsin.
22. Plaintiff Wayne Jensen, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 63rd Assembly District in Rochester, in Racine County, Wisconsin.
23. Plaintiff Wendy Sue Johnson, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 91st Assembly

District in Eau Claire, in Eau Claire County, Wisconsin. In addition to the injury suffered by all Democrats in Wisconsin, Ms. Johnson was harmed when Democratic voters were packed into District 91, wasting their votes and diluting the influence of Ms. Johnson's vote, as part of a gerrymander that reduced the number of Democratic seats in her region. See infra 『\| 69-71.
24. Plaintiff Janet Mitchell, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 66th Assembly District in Racine, in Racine County, Wisconsin. In addition to the injury suffered by all Democrats in Wisconsin, Ms. Mitchell was harmed when Democratic voters were packed into District 66, wasting their votes and diluting the influence of Ms. Mitchell's vote, as part of a gerrymander that reduced the number of Democratic seats in her region. See infra $\mathbb{\|} \mathbb{T}$ 66-68.
25. Plaintiffs James and AllisonSeaton, citizens of the United States and of the State of Wisconsin, are residents and registered voters in the 42nd Assembly District in Lodi, in Columbia County, Wisconsin.
26. Plaintiff Jerome Wallace, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 23rd Assembly District, in Fox Point, in Milwaukee County, Wisconsin. In addition to the injury suffered by all Democrats in Wisconsin, Mr. Wallace was harmed when Democrats in District 22 were cracked so that
his previously Democratic district is now a Republican district. See infra $\| \mathbb{\|}$ 60-62.
27. Plaintiff Don Winter, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 55th Assembly District in Neenah, in Winnebago County, Wisconsin.
28. Defendant Gerald C. Nichol is the Chair of the Wisconsin Government Accountability Board ("G.A.B.") and is named solely in his official capacity as such. The G.A.B. is a state agency under Wis. Stat. $\S 15.60$, which has "general authority" over and "responsibility for the administration of . . . [the State's] laws relating to elections and election campaigns," Wis. Stat. § 5.05(1), including the election every two years of Wisconsin's representatives in the Assembly.
29. Defendants Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas, and Timothy Vocke are all members of the G.A.B. and are named solely in their official capacities as such.
30. Defendant Kevin J. Kennedy is the Director and General Counsel of the G.A.B. and is named solely in his official capacity as such.

## BACKGROUND

The Current Plan Was Intended To
Discriminate Against Democrats
31. The Current Plan was drafted and enacted with the specific intent to maximize the electoral advantage of Republicans and harm Democrats to the greatest possible extent, by packing and cracking Democratic voters and thus wasting as many Democratic votes as possible. Indeed, after a trial in prior litigation, a three-judge court characterized claims by the Current Plan's drafters that they had not been influenced by partisan factors as "almost laughable" and concluded that "partisan motivation . . . clearly lay behind Act 43." Baldus v. Wisconsin Government Accountability Board, 849 F.Supp.2d 840, 851 (E.D. Wis. 2012).
32. The Current Plan was drafted via a secret process run solely by Republicans in the State Assembly and their agents, entirely excluding from participation all Democratic members of the Assembly as well as the public, and preventing public knowledge of and deliberation about the parameters of the Plan.
33. In January 2011, Scott Fitzgerald, Republican member of the Wisconsin State Senate and Wisconsin Senate Majority Leader, and Jeff Fitzgerald, Republican member of the Wisconsin State Assembly and Speaker of the Assembly, hired attorney Eric McLeod ("McLeod") and the law firm of Michael, Best \& Friedrich, LLP ("Michael Best"), ostensibly to represent the entire Wisconsin State Senate and Wisconsin State Assembly in connection with the reapportionment of the state legislative districts after
the 2010 Census. In fact, McLeod and Michael Best were retained to assist the Republican leadership in the Legislature in designing a pro-Republican partisan gerrymander.
34. To accomplish this goal, McLeod and Michael Best supervised the work of the legislative aide to the Republican Speaker of the Assembly, Adam Foltz, and the legislative aide to the Republican Majority Leader of the Senate, Tad Ottman, in planning, drafting, negotiating, and gaining the favorable vote commitments of a majority of Republican legislators sufficient to obtain passage of the Current Plan through Wisconsin Act 43.
35. In creating the Current Plan, McLeod, Michael Best, Foltz, and Ottman used past election results to measure the partisanship of the electorate and to design districts, through packing and cracking, that would maximize the number of districts that would elect a Republican and minimize the number of districts that would elect a Democrat. Thus, they intentionally diluted the electoral influence of Democrats, including that of plaintiffs, and discriminated against Democrats, including plaintiffs, because of their political views.
36. McLeod, Michael Best, Foltz, and Ottman were assisted in their work by Dr. Ronald Keith Gaddie, a professor of political science at the University of Oklahoma. Dr. Gaddie created a model that analyzed the expected partisan performance of
all of the districts established by Act 43. Dr. Gaddie's model forecast that the Assembly plan would have a pro-Republican efficiency gap of $12 \%$. When a common methodology is used to ensure an apples-toapples comparison, this is almost exactly the efficiency gap that the Assembly plan actually exhibited in the 2012 election.
37. Preparation of the Current Plan was done in complete secrecy, excluding Democrats and the public from any part of the process. Indeed, even Republican state legislators were prevented from receiving any information that would allow public discussion or deliberation about the plan. All redistricting work was done in Michael Best's office and the "map room" was located there. A formal written policy provided that only the Senate Majority Leader, the Speaker of the House and their aides Ottman and Foltz, and McLeod and legal staff designated by McLeod would have unlimited access to the map room.
38. The access policy provided for limited access by rank-and-file legislators: "Legislators will be allowed into the office for the sole purpose of looking at and discussing their district. They are only to be present when an All Access member is present. No statewide or regional printouts will be on display while they are present (with the exception of existing districts). They will be asked at each visit to sign an agreement that the meeting they are attending is confidential and they are not to discuss it." But only Republican legislators were allowed even this limited
access. After signing the secrecy agreements contemplated by the policy, Republican legislators were allowed to see only small portions of the map: how their own districts would be affected and details of the partisan performance of voters in their districts in the past, showing that they would be reliable Republican districts.
39. Under the direction and supervision of McLeod, Ottman met with 17 Republican members of the Wisconsin State Senate, identified in Ex. 4 hereto. Each of them signed a secrecy agreement entitled "Confidentiality and Nondisclosure Related to Reapportionment" before being allowed to review and discuss the plan that Michael Best had been hired to develop. The secrecy agreement said that McLeod had "instructed" Ottman to meet with certain members of the Senate to discuss the reapportionment process and characterized such conversations as privileged communications pursuant to the attorney-client and attorney work product privileges-even though the assertion of the privilege was a part of an elaborate "charade" designed "to cover up a process that should have been public from the outset." Baldus v. Wisconsin Government Accountability Board, 843 F.Supp.2d 955, 958-61 (E.D. Wis. 2012).
40. Under the direction and supervision of McLeod, Foltz met with 58 Republican members of the Wisconsin State Assembly, identified in Ex. 4 hereto. Each of them signed the same secrecy
agreement entitled "Confidentiality and Nondisclosure Related to Reapportionment" before being allowed to review and discuss the plan that Michael Best had been hired to develop, which also improperly described their conversations as privileged.
41. On July 11, 2011, the plan was introduced by the Committee on Senate Organization without any Democratic members of the Legislature having previously seen their districts or the plan as a whole. As noted above, all Republican members of the Legislature had previously seen their individual districts along with visual aids demonstrating the partisan performance of these districts, but had not seen the overall map.
42. Act 43 was passed in extraordinarily rushed proceedings with little opportunity for input by the public. A public hearing was held on July 13, 2011. The bill was then passed by the Senate on July 19, 2011, and by the Assembly the next day on July 20, 2011. Act 43 was published on August 23, 2011.
43. McLeod and Michael Best were paid $\$ 431,000$ in State taxpayer funds for their work on the plan, even though they worked solely for Republican leaders of the Legislature and for the benefit of Republicans, and even though they provided no services to Democrats, entirely excluded them from the process, and concealed their work from the public, preventing any public deliberation about the plan.

# The Current Plan Has The Effect of 

Discriminating Against Democrats
The Efficiency Gap Reliably Measures Partisan Gerrymandering
44. The Supreme Court has unanimously agreed that partisan gerrymandering can rise to the level of a constitutional violation. See Vieth v. Jubelirer, 541 U.S. 267, 293 (2004) ("[A]n excessive injection of politics is unlawful") (emphasis added). To date, though, partisan gerrymandering plaintiffs have failed to propose a judicially manageable standard for deciding what constitutes an "excessive" injection of politics into the redistricting process.
45. In the Court's most recent gerrymandering case, LULAC v. Perry, 548 U.S. 399 (2006), a majority of the Justices expressed support for a test based on the concept of partisan symmetry. Partisan symmetry is a "require[ment] that the electoral system treat similarly-situated parties equally." Id. at 466 (Stevens, J., concurring in part and dissenting in part). In other words, a map is symmetrical when it creates a level playing field, giving neither major party a systematic advantage over its opponent in the conversion of electoral votes into legislative seats.
46. In LULAC, the Court considered one particular measure of partisan symmetry, called partisan bias. As described above, partisan bias refers to the divergence in the share of seats that each
party would win given the same share (typically 50\%) of the statewide vote. See id. at 419-20 (opinion of Kennedy, J.); id. at 466 (Stevens, J., concurring in part and dissenting in part).
47. Partisan bias is not the only measure of partisan symmetry. In the last few years, political scientists and legal academics have developed a new symmetry metric, called the efficiency gap, which improves on partisan bias in several respects. See Eric M. McGhee, Measuring Partisan Bias in SingleMember District Electoral Systems, 39 Legis Stud. Q. 55 (2014); Nicholas O. Stephanopoulos \& Eric M. McGhee, Partisan Gerrymandering and the Efficiency Gap, 82 U. Chi. L. Rev. 101 (2015); Expert Report of Prof. Kenneth R. Mayer (July 3, 2015) ("Mayer Report"), attached hereto as Ex. 2; Expert Report of Prof. Simon D. Jackman (July 7, 2015) ("Jackman Report") attached hereto as Ex. 3.
48. The efficiency gap is rooted in the insight that, in a legal regime in which each district must have an approximately equal population, there are only two ways to implement a partisan gerrymander. First, a party's supporters can be cracked among a large number of districts so that they fall somewhat short of a majority in each one. These voters' preferred candidates then predictably lose each race. Second, a party's backers can be packed into a small number of districts in which they make up enormous majorities. These voters' preferred candidates then prevail by overwhelming margins. All partisan gerrymandering
is accomplished through cracking and packing, which enables the party controlling the map to manipulate vote margins in its favor.
49. Both cracking and packing produce so-called "wasted" votes-that is, votes that do not directly contribute to a candidate's election. When voters are cracked, their votes are wasted because they are cast for losing candidates. Similarly, when voters are packed, their votes are wasted to the extent they exceed the $50 \%$-plus-one threshold required for victory (in a two- candidate race). Partisan gerrymandering also can be understood as the manipulation of wasted votes in favor of the gerrymandering party, so that it wastes fewer votes than its adversary.
50. The efficiency gap is the difference between the parties' respective wasted votes in an election, divided by the total number of votes cast. Suppose, for example, that there are five districts in a plan with 100 voters each. Suppose also that Party A wins three of the districts by a margin of 60 votes to 40 , and that Party B wins two of them by a margin of 80 votes to 20. Then Party A wastes 10 votes in each of the three districts it wins and 20 votes in each of the two districts it loses, adding up to 70 wasted votes. Likewise, Party B wastes 30 votes in each of the two districts it wins and 40 votes in each of the three districts it loses, adding up to 180 wasted votes. The difference between the parties' respective wasted
votes is 110 , which, when divided by 500 total votes, yields an efficiency gap of $22 \%$ in favor of Party A.
51. The efficiency gap is not based on the principle that parties have a right to proportional representation based on their share of the statewide vote, nor does it measure the deviation from seat-vote proportionality. Instead, by aggregating all of a plan's cracking and packing into a single number, the efficiency gap measures a party's undeserved seat share: the proportion of seats a party receives that it would not have received under a balanced plan in which both sides had approximately equal wasted votes. In the above example, for instance, the $22 \%$ efficiency gap in favor of Party A means that it won $22 \%$ more seats-in this example, 1 more seat out of 5 -than it would have under a balanced plan.
52. Over the 1972-2014 period-since the end of the reapportionment revolution of the 1960s- the distribution of state house plans' efficiency gaps has been normal and has had a median of almost exactly zero. See Jackman Report at 61; Stephanopoulos \& McGhee, supra, at 140-42. This indicates that neither party has enjoyed an overall advantage in state legislative redistricting during the modern era.
53. However, recently the average absolute efficiency gap (i.e., the mean of the absolute values of all plans' efficiency gaps in a given year) has increased sharply. This metric stayed roughly constant from 1972 to 2010. But in the current cycle,
fueled by rising partisanship and greater technological sophistication, it spiked to the highest level recorded in the modern era: over $6 \%$ for state house plans. See Jackman Report at 47; Stephanopoulos \& McGhee, supra, at 142-45. This means that the severity of today's partisan gerrymandering is historically unprecedented-as is the need for judicial intervention.

## Wisconsin's Current Plan Is an Outlier

54. Between 1972 and the present, the efficiency gaps of Wisconsin's Assembly plans became steadily larger and more pro-Republican. The Current Plan represents the culmination of this trend, exhibiting the largest and most pro-Republican efficiency gap ever recorded in modern Wisconsin history. In the 1970s, the Assembly plan had an average efficiency gap close to zero. In both the 1980s and the 1990s, it had an average pro-Republican gap of $2 \%$. The Republican advantage deepened in the 2000s to an average gap of $8 \%$. And it then surged, thanks to the Current Plan, to an average gap of $11 \%$ in 2012 and 2014. See Jackman Report at 34; Stephanopoulos \& McGhee, supra, at 154-56.
55. More specifically, using the same methodology as for all other states, the Current Plan produced a pro-Republican efficiency gap of $13 \%$ in 2012 and 10\% in 2014. The 2012 figure represents the 28th-worst score in modern American history (out of nearly 800 total plans), placing the Current Plan in the worst 4\%
of this distribution, more than two standard deviations from the mean. Based on this historical data, there is close to a zero percent chance that the Current Plan's efficiency gap will ever switch signs and favor the Democrats during the remainder of the decade. Furthermore, prior to the current cycle, not a single plan in the country had efficiency gaps as high as the Current Plan's in the first two elections after redistricting. See Jackman Report at 63.
56. Using a more detailed methodology available only for Wisconsin, the Current Plan produced a proRepublican efficiency gap of $12 \%$ in 2012. This is a figure nearly identical to the one calculated using the national data. Using the Wisconsin-specific methodology as well as data compiled prior to 2012 by Dr. Gaddie, the expert retained by the Legislature's Republican leadership to assist them in drafting the Current Plan, that Plan was forecast to produce an efficiency gap of $12 \%$. This figure also is nearly identical, and shows that the Current Plan performed precisely as its authors hoped and expected. See Mayer Report at 46.
57. This extraordinary level of partisan unfairness was achieved through the rampant cracking and packing of Wisconsin's Democratic voters, which resulted in their votes being disproportionately wasted. The Mayer Report shows that Democratic voters were cracked so that Republican candidates were far more likely to prevail in close races (where the winner had $60 \%$ or less of the vote): Republicans
were likely to win 42 such districts, while Democrats would win only $17 .{ }^{1}$ Democrats were also packed into a number of districts where they would win overwhelmingly (by getting $80 \%$ or more of the vote): there were eight districts where Democrats would win by this margin, compared to zero districts where Republicans would win such a lopsided victory. Thus, through gerrymandering, Republican votes were used more efficiently than Democratic votes to elect representatives, producing an undemocratic result that does not accurately reflect the preferences of the Wisconsin electorate. See Mayer Report at 38-41.
58. The forecasts of Dr. Gaddie, the Republican consultant, prior to the 2012 election confirm that the Current Plan was expected and intended to crack and pack Wisconsin's Democratic voters to this extent. Dr. Gaddie predicted that Republicans would win 46 Assembly districts by a margin smaller than $60 \%$ $40 \%$, compared to just 20 such victories for Democrats. He also predicted that Democrats would prevail in seven districts by a margin greater than $80 \%$ - $20 \%$, compared to zero such wins for Republicans. See Mayer Report at 38-41. These figures are nearly identical to plaintiffs' estimates, and further demonstrate that the Current Plan was

[^0]intended to disadvantage Democrats and waste Democratic votes to the maximum extent possible.

## Examples of Cracking and Packing in the Current Plan

59. These plan-level statistics are the product of innumerable local cracking and packing decisions. Across Wisconsin, the Current Plan systematically alters prior district configurations to waste larger numbers of Democratic votes and smaller numbers of Republican votes. The following regional examples (depicted in map form in Exhibit 1 hereto) show how the Current Plan deliberately allocates Democratic voters less efficiently and Republican voters more efficiently. These are only illustrative examples; they do not show all of the ways in which Wisconsin's current pro-Republican gerrymander was achieved. In addition, the examples focus on: (1) the 2012 election because it was the first one held after this cycle's redistricting; (2) the 2008 election because it was the most comparable prior election, featuring a similar share of the statewide Assembly vote for each party (53.9\% Democratic in 2008, 51.4\% Democratic in 2012) and also coinciding with a presidential election; and (3) Plaintiffs' Demonstration Plan, because it reveals the fair results that could have been, but were not, attained in 2012.

Milwaukee, Ozaukee, Washington, and Waukesha Counties:
60. Under the prior Assembly plan that was in force from 2002-2010 (the "Prior Plan"), District 22 included part of northeastern Milwaukee County; District 23 included part of northern Milwaukee County (home to Plaintiff Wallace) and part of southern Ozaukee County; and District 24 included part of Washington and Waukesha Counties. In the 2008 election, a Democratic candidate won District 22, and Republican candidates won Districts 23 and 24. Under the Demonstration Plan, a Democratic candidate would win District 22, and Republican candidates would win Districts 23 and 24.
61. As a result of the Current Plan, Democratic voters who were in the old District 22 were cracked into the new Districts 23 and 24. Due to these changes, Districts 22, 23, and 24 were won by Republican candidates in 2012.
62.The shift from one Democratic seat and two Republican seats in the Prior Plan and the Demonstration Plan in Milwaukee, Ozaukee, Washington, and Waukesha Counties, to zero Democratic seats and three Republican seats in the Current Plan, contributed to Wisconsin's current proRepublican efficiency gap. This gerrymandering and its results are shown in the maps attached hereto as Ex. 1.

[^1]63. Under the Prior Plan, District 26 centered on the City of Sheboygan in the central eastern part of Wisconsin (home to Plaintiff Donohue) and District 27 consisted of the northern part of Sheboygan County as well as parts of Fond du Lac, Calumet, and Manitowoc Counties. In the 2008 election, a Democratic candidate won District 26 and a Republican candidate won District 27. Under the Demonstration Plan, a Democratic candidate would win District 26, and a Republican candidate would win District 27.
64. As a result of the Current Plan, Democratic voters who were in District 26 were cracked so that roughly half of that district was distributed to District 27 and additional voters from south of Sheboygan County were added to District 26. Due to these changes, Districts 26 and 27 were won by Republican candidates in 2012.
65. The shift from one Democratic seat and one Republican seat in the Prior Plan and the Demonstration Plan in Sheboygan County and southern Fond du Lac, Manitowoc and Calumet Counties, to zero Democratic seats and two Republican seats in the Current Plan, contributed to Wisconsin's current pro-Republican efficiency gap. This gerrymandering and its results are shown in the maps attached hereto as Ex. 1.

Racine and Kenosha Counties:

JA52
66. Under the Prior Plan, Districts 61, 62, 63, 64, 65, and 66 were almost entirely within Racine and Kenosha Counties in the southeastern edge of Wisconsin (the City of Racine is home to Plaintiff Mitchell). Districts 61 and 62 centered on the City of Racine, with District 63 covering the western side of Racine County. Districts 64 and 65 centered on the City of Kenosha, with District 66 covering the western edge of Kenosha County. In the 2008 election, Democratic candidates won Districts 61, 62, 64, and 65, while Republican candidates won Districts 63 and
66. Under the Demonstration Plan, Democratic candidates would win Districts 62, 63, 64, and 66, while Republican candidates would win Districts 61 and 65.
67. As a result of the Current Plan, Democratic voters who were in the old Districts 61 and 62 were packed into the new District 66, thus wasting more Democratic votes in the region. Due to these changes, Districts 64, 65, and 66 were won by Democratic candidates in 2012, while Districts 61, 62, and 63 were won by Republican candidates.
68. The shift from four Democratic seats and two Republican seats in the Prior Plan and the Demonstration Plan in Racine and Kenosha Counties, to three Democratic seats and three Republican seats in the Current Plan, contributed to Wisconsin's current pro-Republican efficiency gap. This
gerrymandering and its results are shown in the maps attached hereto as Ex. 1.

> Buffalo, Chippewa, Eau Claire, Jackson, La Crosse, Pepin, Pierce, St. Croix, and Trempealeau Counties:
69. Under the Prior Plan, most of seven Districts ( $67,68,91,92,93,94$, and 95) were spread across Buffalo, Chippewa, Eau Claire, Jackson, La Crosse, Pepin, Pierce, St. Croix, and Trempealeau Counties in northwestern Wisconsin (Eau Claire is home to Plaintiff Johnson). In the 2008 election, Democratic candidates won five of the seven Districts (68, 91, 92, 93, and 95), and Republicans won two of them (67 and 94). The district numbers in the Demonstration Plan are slightly different; instead of District 68, District 69 is in Eau Claire County. Under the Demonstration Plan, Democratic candidates would win six of seven Districts (67, 69, 91, 92, 94, and 95) and a Republican candidate would win one of them (93).
70. As a result of the Current Plan, Democratic voters who were in the old District 68 were packed into the new District 91, and Democrats in the rest of old District 68 as well as old Districts 91 and 93 were cracked into the new Districts 68,92 , and 93 . Due to these changes, Democratic candidates won only four of the seven districts in 2012 (91, 92, 94, and 95), and Republican candidates won three of them (67, 68, and 93).
71. The shift from five or six Democratic seats, in the Prior Plan and Demonstration Plan respectively, and two or one Republicanseats in the Prior Plan and Demonstration Plan respectively, to four Democratic seats and three Republican seats in the Current Plan, in Buffalo, Chippewa, Eau Claire, Jackson, La Crosse, Pepin, Pierce, St. Croix, and Trempealeau Counties, contributed to Wisconsin's current pro-Republican efficiency gap. This gerrymandering and its results are shown in the maps attached hereto as Ex. 1.

> Adams, Columbia, Marathon, Marquette, Portage, and Wood Counties:
72. Under the Prior Plan, most of eight Districts (42, 47, 69, 70, 71, 72, 85, and 86) were spread across Adams, Columbia, Marathon, Marquette, Portage, and Wood counties in central Wisconsin (Columbia County is home to Plaintiffs Allison and James Seaton). In the 2008 election, Democratic candidates won five of the eight Districts (42, 70, 71, 72, and 85), and Republicans won three Districts (47, 69, and 86). In the Demonstration Plan the district numbers are different ( $5,40,41,42,71,72,86$, and 87 ), but of these eight Districts, Democratic candidates would win five (71, 86, 40, 41, and 42), and Republican candidates would win three ( 5,72 , and 87 ).
73. As a result of the Current Plan, Democratic voters who were in the old Districts 42 , 70, and 72 were cracked, and the new Districts $41,42,69,70,71$, 72,85 , and 86 were created in areas of Adams,

Columbia, Marathon, Marquette, Portage, and Wood Counties. Due to these changes, Democratic candidates won only three of the eight Districts (70, 71 , and 85) in 2012, and Republican candidates won five of them $(41,42,69,72$, and 86 ).
74. The shift from five Democratic seats and three Republican seats in the Prior Plan and the Demonstration Plan in Adams, Columbia, Marathon, Marquette, Portage, and Wood Counties, to three Democratic seats and five Republican seats in the Current Plan, contributed to Wisconsin's current proRepublican efficiency gap. This gerrymandering and its results are shown in the maps attached hereto as Ex. 1.

## Brown and Manitowoc Counties:

75. Under the Prior Plan, Brown and Manitowoc Counties were split to include parts of Districts 1, 2, $4,5,25,88,89$, and 90 in the Green Bay area of Wisconsin. In the 2008 election, Democratic candidates won Districts 2, 5, 25, and 88, and Republican candidates won Districts 1, 4, 89, and 90. Under the Demonstration Plan, Brown and Manitowoc Counties would include Districts 1, 2, 3, $25,26,88,89$, and 90 . Under the DemonstrationPlan, Democrats would win Districts 2 and 88, and Republicans would win the remaining six districts.
76. As a result of the Current Plan, Democratic voters who were in the old Districts 2,5 and 25 were
cracked into the new Districts 2, 5, 25, and 88. Due to these changes, seven of the eight districts in the Brown and Manitowoc County area (1, 2, 4, 5, 25, 88, and 89) were won by Republican candidates in 2012, and one District (90) was won by a Democratic candidate in 2012.
77. The shift from four or two Democratic seats in the Prior Plan and the Demonstration Plan, respectively, and four or six Republican seats in the Prior Plan and the Demonstration Plan, respectively, to one Democratic seat and seven Republican seats in the Current Plan, in Brown and Manitowoc Counties, contributed to Wisconsin's current pro- Republican efficiency gap. This gerrymandering and its results are shown in the maps attached hereto as Ex.1.

## Wisconsin Does Not Need to Have a Gerrymandered Plan

78. Not only did the Current Plan exhibit extremely large efficiency gaps in 2012 and 2014, but this poor performance was entirely unnecessary and served no legitimate purpose. It would have been possible for Wisconsin to enact an Assembly plan that treated both parties symmetrically and did not disproportionately waste Democratic votes. To prove this point, plaintiffs' expert has designed a Demonstration Plan that would have had an efficiency gap of just $2 \%$ in 2012 (assuming all contested districts and no incumbents). See Mayer Report at 46. This far better score is attributable to
plaintiffs' efforts not to crack and pack Democratic voters, and instead to enable both parties to convert their popular support into legislative seats with equal ease.
79. Plaintiffs' Demonstration Plan performs at least as well as the Current Plan on every other relevant metric. Both plans have total population deviations of less than $1 \%$-far below the courts' $10 \%$ threshold for presumptive constitutionality. Both plans have six African American opportunity districts and one Hispanic opportunity district, and so are identical for Voting Rights Act purposes. The Demonstration Plan splits one fewer municipal boundary than the Current Plan (119 versus 120), and so is superior in that regard. And the Demonstration Plan's districts are substantially more compact than the Current Plan's (average compactness of 0.41 versus 0.28 ). See Mayer Report at 37 .
80. The Demonstration Plan proves that the Current Plan's extreme pro-Republican tilt cannot be blamed on either an effort to comply with legitimate redistricting criteria or Wisconsin's underlying political geography. Both of those factors were perfectly compatible with a neutral map.

## COUNT I - FOURTEENTH AMENDMENT VIOLATION

81. Plaintiffs incorporate and re-allege paragraphs 1-80 of this Complaint as paragraphs 180 of this Count I.
82. The Current Plan is a partisan gerrymander so extreme that it violates Plaintiffs' Fourteenth Amendment right to equal protection of the laws. The Current Plan intentionally and severely packs and cracks Democratic voters, thus disproportionately wasting their votes, even though a neutral map could have been drawn instead. Accordingly, Wisconsin's Act 43 deprives plaintiffs of their civil rights under color of state law in violation of 42 U.S.C. §§ 1983 and 1988.
83. The efficiency gap provides a workable test to identify unconstitutional partisan gerrymandering similar to the two-part approach applied to state legislative reapportionment claims. In a reapportionment challenge, the first issue is whether a district plan's total population deviation exceeds $10 \%$. If so, the plan is presumptively unconstitutional, and if not, it is presumptively valid. The second issue, which is reached only if the total population deviation is greater than $10 \%$, is whether the malapportionment is necessary to achieve a legitimate state goal. The state bears the burden at this stage of rebutting the presumption of unconstitutionality. See Voinovich v. Quilter, 507 U.S. 146, 161-62 (1993); Brown v. Thomson, 462 U.S. 835, 842-43 (1983); Connor v. Finch, 431 U.S. 407, 418 (1977).
84. The same two-part approach should be applied to partisan gerrymandering claims, only with the efficiency gap substituted for total population deviation. The first step in the analysis is whether a plan's efficiency gap exceeds a certain numerical threshold. If so, the plan is presumptively unconstitutional, and if not, it is presumptively valid. The second step, which is reached only if the efficiency gap is sufficiently large, is whether the plan's severe partisan unfairness is the necessary result of a legitimate state policy, or inevitable given the state's underlying political geography. The state would bear the burden at this stage of rebutting the presumption of unconstitutionality.
85. The Current Plan is plainly unlawful under this two-part test. First, it was forecast to produce, and then did produce, an efficiency gap of approximately $13 \%$ in the 2012 election. This is an extraordinarily high level of partisan unfairness, more than two standard deviations from the mean: as noted above, the 2012 figure represents the 28thworst score in modern American history (out of nearly 800 total plans), placing the Current Plan in the worst $4 \%$ of this distribution. This is also not a temporary or transient gerrymander. The Current Plan's efficiency gap means that there is close to a zero percent chance that the Plan will ever favor Democrats during its lifespan. See Jackman Report at 60 . Given its severity and predicted durability, the Current Plan's efficiency gap far exceeds any
plausible threshold for presumptive unconstitutionality.
86. Indeed, even a $7 \%$ efficiency gap should be presumptively unconstitutional. A 7\% efficiency gap is at the edges of the overall distribution of all state house plans in the modern era, making it indicative of uncommonly severe gerrymandering. See Jackman Report at 61. Historical analysis shows that with a $7 \%$ efficiency gap, the gerrymandering is also likely to be unusually durable-over its lifespan, a plan with an efficiency gap of that magnitude is unlikely ever to favor the opposing party. See Jackman Report at 61. However, this Court need not decide at what point an efficiency gap is large enough to trigger a presumption of unconstitutionality. In the state legislative reapportionment context, the applicable cutoff (10\%) emerged over a series of cases, in which extreme population deviations (of $34 \%$, then $26 \%$, then $20 \%$ ) were struck down and deviations of $8 \%$ and $10 \%$ were upheld before the $10 \%$ threshold was adopted. Here too the Current Plan's extreme efficiency gap should be deemed presumptively unconstitutional, without the need to decide what the cut-off should be.
87. Second, the State cannot rebut the presumption that the Current Plan is unlawful. Plaintiffs' Demonstration Plan would have had an efficiency gap of just $2 \%$ in 2012 while complying with all federal and state criteria at least as well as the Current Plan. See Mayer Report at 46. Accordingly,
neither an attempt to achieve legitimate redistricting goals nor Wisconsin's underlying political geography could have necessitated the Current Plan's partisan imbalance.
88. In addition to its extreme efficiency gap, the Current Plan exhibits a severe partisan bias. The Current Plan produced a partisan bias of $13 \%$ in 2012 and $12 \%$ in 2014 - scores that in and of themselves demonstrate the unconstitutional effects produced by the Current Plan.
89. Finally, there is no doubt that the Current Plan was specifically intended and indeed designed to benefit Republican candidates, and to disadvantage Democratic candidates, to the greatest possible extent. Thus, the Current Plan had both the purpose and effect of subordinating the adherents of one political party and entrenching a rival party in power, in violation of their right to equal protection under the law.

## COUNT II-FIRST AMENDMENT VIOLATION

90. Plaintiffs incorporate and re-allege paragraphs 1-89 of this Complaint as paragraphs 189 of this Count II.
91. Plaintiffs and other Democratic voters in the state of Wisconsin have a First Amendment right to freely associate with each other without discrimination by the State based on that association;
to participate in the political process and vote in favor of Democratic candidates without discrimination by the State because of the way they vote; and to express their political views without discrimination by the State because of the expression of those views or the content of their expression.
92. Wisconsin Act 43 violates the First and Fourteenth Amendments because it intentionally uses voters' partisan affiliation to affect the weight of their votes. By taking the actions described above, the drafters of the Current Plan deliberately discriminated against plaintiffs and other Democratic voters because they are Democrats and have voted for and will vote for Democratic candidates and because of the positions they have expressed and will take on public affairs - that is, because of their views and the content of their expression.
93. By excessively and unreasonably cracking and packing groups of Democratic voters to intentionally weaken their voting power, the State of Wisconsin discriminated against Democratic voters, including the plaintiffs, on the basis of their voting choices, their political views, and the content of their expression.
94. The unusual extent of the partisan gerrymandering in this case, as shown by the extremely high efficiency gap and the factors described above, indicates that the gerrymandering in this case is so high that the Current Plan denies to
plaintiffs and other Democratic voters in Wisconsin their rights to free association and freedom of expression guaranteed by the First and Fourteenth Amendments.
95. For these reasons, and because Act 43 and the Current Plan have the purpose and effect of subjecting Democrats to disfavored treatment by reason of their views, Act 43 and the Current Plan are subject to strict scrutiny and cannot be upheld absent a compelling government interest, which is not present in this case.
96. Accordingly, Wisconsin's Act 43 deprives plaintiffs of their civil rights under color of state law in violation of 42 U.S.C. §§ 1983 and 1988.

## RELIEF REQUESTED

WHEREFORE, Plaintiffs respectfully request that this Court:
97. Declare Wisconsin's 99 State Assembly Districts, established by Act 43, unconstitutional and invalid, and the maintenance of these districtsfor any primary, general, special, or recall election a violation of plaintiffs' constitutional rights;
98. Enjoin Defendants and the G.A.B.'s employees and agents, including the county clerks in each of Wisconsin's 72 counties, from administering, preparing for, and in any way permitting the
nomination or election of members of the State Assembly from the unconstitutional districts that now exist;
99. In the absence of a state law establishing a constitutional district plan for the Assembly districts, adopted by the Legislature and signed by the Governor in a timely fashion, establish a redistricting plan that meets the requirements of the U.S. Constitution and federal statutes and the Wisconsin Constitution and state statutes;
100. Award plaintiffs their reasonable attorneys' fees, costs, and litigation expenses incurred in bringing this action; and
101. Grant such further relief as the Court deems just and proper.

By: /s/ Peter G. Earle<br>Peter G. Earle<br>One of the attorneys for<br>plaintiffs

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Applications for admission pro hac vice pending

## Preliminary Pre-trial Conference Order

IN THE UNITED STATES DISTRICT COURT FOR THE WESTERN DISTRICT OF WISCONSIN

WILLIAM WHITFORD, et
al.,
Plaintiffs, v.

GERALD NICHOL, et al.,

PRELIMINARY
PRETRIAL
CONFERENCE
ORDER

Defendants.

This court held a telephonic preliminary pretrial conference on October 13, 2015. Plaintiffs appeared by Peter Earle and Ruth Greenwood. Defendants appeared by Brian Keenan and Anthony Russomanno. The court set the schedule for this case and advised the parties that their conduct throughout this case is governed by this pretrial conference order and the attachments to it.

The parties and their attorneys must at all times treat everyone involved in this lawsuit with courtesy and consideration. The parties must attend diligently to their obligations in this lawsuit and must reasonably accommodate each other in all matters so as to secure the just, speedy and inexpensive resolution of each proceeding in this matter as
required by Fed. R. Civ. Pro. 1. Failure to do so shall have consequences.

1. Oral Argument on Pending Motion to Dismiss: November 4, 2015 at 1:30 p.m.

Argument shall be to the three-judge panel. Each side will have 30 minutes to present it [sic] arguments.

## 2. Disclosure of Experts: Plaintiffs: October 23,

 2015Defendants: December 2, 2015

Rebuttal: December 16, 2015

Absent the parties' agreement to a different procedure, all disclosures mandated by this paragraph must comply with the requirements of Rule 26(a)(2). Given the tight calendar in this lawsuit, there shall be no supplementation under Rule 26(e) without leave of court.

## 3. Deadline for Filing Dispositive Motions:

 January 4, 2016All dispositive motions must be accompanied by supporting briefs. All responses to any dispositive motion must be filed and served within 21calendar days of service of the motion. Any reply by the movant must be filed and served within 10 calendar days of
service of the response. Given the tight calendar in this lawsuit, there shall be no extensions of these deadlines.

All parties must follow this court's procedure governing summary judgment motions, a copy of which is attached to this order. The court will not consider any document that does not comply with its summary judgment procedure.

Parties are to undertake discovery in a manner that allows them to make or respond to dispositive motions within the scheduled deadlines. The fact that the general discovery deadline cutoff occurs after the deadlines for filing and briefing dispositive motions is not a ground for requesting an extension of the motion and briefing deadlines.

## 2. Discovery Cutoff: April 1, 2016

All discovery in this case must be completed not later than the date set forth above, absent written agreement of all parties to some other date. Absent written agreement of the parties or a court order to the contrary, all discovery must conform with the requirements of Rules 26 through 37 and 45. Rule 26(a)(1) governs initial disclosures unless the parties agree in writing to the contrary.

The following discovery materials shall not be filed with the court unless they concern a motion or other matter under consideration by the court:
interrogatories; responses to interrogatories;requests for documents; responses to requests for documents; requests for admission; and responses to requests for admission.

A party need not file a deposition transcript with the court until that party is using the deposition in support of some other submission, at which time the entire deposition must be filed. Note well: as detailed later in this order, any deposition that has not been filed with the court by May 2, 2016 shall not be used by any party for any purpose at trial. All deposition transcripts must be in compressed format. The court will not accept duplicate transcripts. The parties must determine who will file each transcript.

A party may not file a motion regarding discovery until that party has made a good faith attempt to resolve the dispute. All efforts to resolve the dispute must be set forth in any subsequent discovery motion filed with this court. By this order, the court requires all parties to a discovery dispute to attempt to resolve it quickly and in good faith. Failure to do so could result in cost shifting and sanctions under Rule 37.

This court also expects the parties to file discovery motions promptly if self-help fails. Parties who fail to do so may not seek to change the schedule on the ground that discovery proceeded too slowly to meet the deadlines set in this order.

All discovery-related motions must be accompanied by a supporting brief, affidavit, or other document showing a prima facie entitlement to the relief requested. Any response to a discovery motion must be served and filed within seven calendar days of service of the motion. Replies may not be filed unless requested by the court.

## 4. Final Pretrial Submissions

Not later than April 25,2016, each party shall file and serve all materials specified in Rule 26(a)(3), unless a different procedure is directed below.

Not later than May 2, 2016, counsel are to confer for the following purposes:
A. To enter into comprehensive written stipulations of all uncontested facts in such form that they can be offered at trial as the first evidence presented by the party desiring to offer them. If there is a challenge to the admissibility of some uncontested facts that one party wishes included, the party objecting and the grounds for objection must be stated.
B. To make any deletions from their previouslyexchanged lists of potential trial witnesses.
C. To enter into written stipulations setting forth the qualifications of expert witnesses.
D. To examine, mark, and list all exhibits that any party intends to offer at trial. (A copy of this court's procedures for marking exhibits is contained in this packet.)
E. To agree as to the authenticity and admissibility of such exhibits so far as possible and note the grounds for objection to any not agreed upon.
F. To agree so far as possible on the contested issues of law.
G. To examine and prepare a list of all depositions and portions of depositions to be read into evidence and agree as to those portions to be read. If any party objects to the admissibility of any portion, the name of the party objecting and the grounds shall be set forth.

It shall be the responsibility of plaintiffs' counsel to convene the conference between counsel and, following that conference, to prepare the Pretrial Statement described in the next paragraph.

Not later than May 9, 2016, the parties jointly shall submit a Pretrial Statement containing the following:

> A. The parties' comprehensive written stipulations of all uncontested facts.
B. An updated prediction on the probable length of the trial.
C. The names of all prospective witnesses. Only witnesses so listed will be permitted to testify at the trial except for good cause shown.
D. The parties' written stipulation setting forth the qualifications of all expert witnesses.
E. Schedules of all exhibits that will be offered in evidence at the trial, together with an indication of those agreed to be admissible and a summary statement of the grounds for objection to any not agreed upon. Only exhibits so listed shall be offered in evidence at the trial except for good cause shown.
F. An agreed statement of the contested issues of law supplemented by a separate statement by each counsel of those issues of law not agreed to by all parties.
G. A list of all depositions and portions of depositions to be offered in evidence, together with an indication of those agreed to be admissible and summary statements of the grounds for objections to any not so agreed upon. If only portions of a deposition are to be offered, counsel should mark the deposition itself with colored markers identifying the portions each party will rely upon.
H. Complete copies of all deposition transcripts to be used at trial, in compressed format.

Not later than May 9, 2016, counsel for each side shall file and serve a statement of all the facts that counsel will request the court to find at the conclusion of the trial. In preparing these statements, counsel should have in mind those findings that will support a judgment in their client's favor. The proposed findings should be complete. They should be organized in the manner in which counsel wish them to be entered. They should include stipulated facts, as well as facts not stipulated to but which counsel expect to be supported by the record at the conclusion of the trial. Those facts that are stipulated shall be so marked.

Along with these proposed findings of fact, counsel for each side shall file and serve a proposed form of special verdict, as if the case were to be tried to a jury.

Not later than May 16, 2016, counsel for each side shall file and serve:
A. Five complete sets of counsel's pre-marked trial exhibits to be used by the judges, the court clerk and the court reporter as working copies at trial.
B. A trial brief.

Final pretrial submissions are to be filed as stated above with no exceptions.

## 4. Bench Trial: May 23, 2016 at 9:00 a.m.

Trial shall be to a three-judge panel. The parties currently estimate that this case will take eight days to try.

This case will be tried in an electronically equipped courtroom and the parties shall present their evidence using this equipment. Counsel are responsible for timely ensuring the compatibility of any of their personal equipment with the court's system.

Entered this $15^{\text {th }}$ day of October, 2015.

## BY THE COURT:

/s/
STEPHEN L. CROCKER
Magistrate Judge

# Opinion and Order Denying Defendants' Motion to Dismiss 

## IN THE UNITED STATES DISTRICT COURT FOR THE WESTERN DISTRICT OF WISCONSIN

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WILLIAM WHITFORD,
ROGER ANCLAM, EMILY
BUNTING, MARY LYNNE
DONOHUE, HELEN
HARRIS, WAYNE JENSEN,
WENDY SUE JOHNSON,
JANET MITCHELL,
ALLISON SEATON, JAMES
SEATON, JEROME
WALLACE and DONALD
WINTER,
OPINION AND ORDER
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Plaintiffs,
$15-c v-421-b b c$
v.

GERALD C. NICHOL, THOMAS BARLAND, JOHN FRANKE, HAROLD V FROEHLICH, KEVIN J KENNEDY, ELSA LAMELAS and TIMOTHY VOCKE, Defendants.

In this civil action brought under 42 U.S.C. § 1983, plaintiffs are Wisconsin residents and Democratic voters who are challenging the 2012 districting plan for the Wisconsin Assembly on the ground that the plan is an example of "extreme partisan gerrymandering." Cpt. 『| 2, dkt.\#1. Plaintiffs contend that the plan violates the First and Fourteenth Amendments to the United States Constitution because the plan "treats voters unequally, diluting their voting power based on their political beliefs, in violation of the Fourteenth Amendment's guarantee of equal protection" and "unreasonably burdens their First Amendment rights of association and free speech." Id.

Defendants have filed a motion to dismiss, dkt. \#24, which is ready for review. Although we believe that plaintiffs face significant challenges in prevailing on their claims, we conclude that plaintiffs' complaint is sufficient to state a claim upon which relief may be granted. Accordingly, we are denying defendants' motion to dismiss.

In their complaint, plaintiffs allege the following facts.

## ALLEGATIONS OF FACT

A. Parties

Plaintiffs William Whitford, Roger Anclam, Emily
Bunting, Mary Lynne Donohue, Helen Harris, Wayne

Jensen, Wendy Sue Johnson, Janet Mitchell, James Seaton, Allison Seaton, Jerome Wallace and Don Winter are United States citizens registered to vote in Wisconsin. They reside in various counties and legislative districts throughout the state. All of them are "supporters of the public policies espoused by the Democratic Party and of Democratic Party candidates." Cpt. ब 15, dkt. \#1.

Defendant Gerald C. Nichol is the chair of the Wisconsin Government Accountability Board, which is responsible for the administration of Wisconsin's laws relating to elections and election campaigns. Defendants Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas and Timothy Vocke are all members of the board. Defendant Kevin J. Kennedy is the director and general counsel for the board.

## B. Passage of Wisconsin Act 43

In January 2011, Scott Fitzgerald, Republican member of the Wisconsin State Senate and Senate Majority Leader, and Jeff Fitzgerald, Republican member of the Wisconsin State Assembly and Speaker of the Assembly, hired lawyer Eric McLeod and the law firm of Michael, Best \& Friedrich, LLP, to assist with the reapportionment of the state legislative districts after the 2010 Census. The intent of the speaker and majority leader was to design a pro-Republican partisan gerrymander. To accomplish this goal, the firm supervised the work of legislative aides in planning, drafting and negotiating Wisconsin

Act 43, which contains the 2012 Assembly districting plan.

The law firm and the aides used past election results to measure the partisanship of the electorate and to design districts that would maximize the number of districts that would elect a Republican and minimize the number of districts that would elect a Democrat. This would be accomplished in two ways, by "cracking" or "packing." Cracking means dividing a party's supporters among multiple districts so that they fall short of a majority in each one. Packing means concentrating one party's backers in a few districts that they win by overwhelming margins. Both cracking and packing result in "wasted" votes, that is, votes cast either for a losing candidate (in the case of cracking) or for a winning candidate but in excess of what he or she needed to prevail (in the case of packing).

The firm and the aides received assistance from Dr. Ronald Keith Gaddie, a professor of political science at the University of Oklahoma. Gaddie created a model that analyzed the expected partisan performance of all of the districts established by Act 43. Gaddie's model forecast that the Assembly plan would have a pro-Republican "efficiency gap" of 12 percent. The efficiency gap is the difference between the parties' respective wasted votes in an election, divided by the total number of votes cast.

All redistricting work was done in the firm's office. Only the speaker, the majority leader, their aides, McLeod and legal staff designated by McLeod would have unlimited access to the plan while it was prepared. The access policy provided for limited access by other Republican legislators:

Legislators will be allowed into the office for the sole purpose of looking at and discussing their district. They are only to be present when an All Access member is present. No statewide or regional printouts will be on display while they are present (with the exception of existing districts). They will be asked at each visit to sign an agreement that the meeting they are attending is confidential and they are not to discuss it.

Cpt. ब 38, dkt. \#1. Democratic legislators were not granted any access to the office. They had no involvement in drafting the plan.

After signing the secrecy agreements contemplated by the policy, Republican legislators were allowed to see only small portions of the map. This included information regarding how their own districts would be affected. Under the direction and supervision of McLeod, the aides met with Republican members of both houses. Each of the members signed a secrecy agreement entitled "Confidentiality and Nondisclosure Related to Reapportionment" before being allowed to review and discuss the plan.

On July 11, 2011, the plan was introduced by the Committee on Senate Organization. At that time, no Democratic members of the legislature had seen their districts or the plan as a whole.

On July 13, 2011, a public hearing was held. On July 19, 2011, the Senate passed the bill; on July 20, 2011, the Assembly passed it. On August 23, 2011, Act 43 was published.

The firm received $\$ 431,000$ from public funds for their work on the plan.

## C. Comparison of Wisconsin Act 43 to Other Plans

From 1972 to 2014, the median efficiency gap for state house plans across the country was close to zero. This indicates that neither party has enjoyed an overall advantage in state legislative redistricting during the modern era. However, recently the average absolute efficiency gap, that is, the mean of the absolute values of all plans' efficiency gaps in a given year, has increased sharply. This metric stayed roughly constant from 1972 to 2010 , but in the current cycle, it spiked to the highest level recorded in the modern era, more than 6 percent for state house plans.

Between 1972 and the present, the efficiency gaps of Wisconsin's Assembly plans became steadily larger and more pro-Republican. The current plan represents the culmination of this trend, exhibiting
the largest and most pro-Republican efficiency gap ever recorded in modern Wisconsin history. In the 1970s, the Assembly plan had an average efficiency gap close to zero. In both the 1980s and the 1990s, it had an average pro-Republican gap of 2 percent. The Republican advantage deepened in the 2000s to an average gap of 8 percent. Under the current plan, the average gap is 11 percent.

A 7 percent efficiency gap is at the edges of the overall distribution of all state house plans in the modern era, making it indicative of uncommonly severe gerrymandering. Historical analysis shows that with a 7 percent efficiency gap, the gerrymandering is also likely to be unusually durable. Over its lifespan, a plan with an efficiency gap of that magnitude is unlikely ever to favor the opposing party.

In 2012, the current plan produced a proRepublican efficiency gap of 13 percent. In 2014, it was 10 percent. The 2012 figure represents the 28th largest score in modern American history (out of nearly 800 total plans), placing the current plan in the most partisan 4 percent of this distribution, more than two standard deviations from the mean. This historical data suggests that there is close to a zero percent chance that the current plan's efficiency gap will ever favor the Democrats during the remainder of the decade. Prior to the current cycle, not a single plan in the country had efficiency gaps as high as the
current plan's in the first two elections after redistricting.

Using a more detailed methodology available only for Wisconsin, the current plan produced a proRepublican efficiency gap of 12 percent in 2012. This is a figure nearly identical to the one calculated using the national data. It is also the same efficiency gap predicted by Dr. Gaddie when the plan was being drafted.

Under the current plan, Republican candidates have been far more likely to prevail in close races. In addition, there were eight districts in which Democrats won with more than 80 percent of the vote. There were no districts in which Republicans won by such a wide margin. Across Wisconsin, the current plan systematically alters prior district configurations to waste larger numbers of Democratic votes and smaller numbers of Republican votes.

## D. Possible Alternatives to Wisconsin Act 43

It would have been possible for Wisconsin to enact an Assembly plan that treated both parties symmetrically. Under a plan prepared by plaintiffs, the efficiency gap would have been 2 percent in 2012 (assuming that races were contested and that no races included an incumbent). This score is attributable to plaintiffs' efforts not to crack and pack Democratic voters and instead to enable both parties to convert
their popular support into legislative seats with equal ease.

Plaintiffs' plan performs at least as well as the current plan on every other metric used by courts to evaluate the validity of a districting plan. Both plans have total population deviations of less than 1 percent. Both plans have six African American opportunity districts and one Hispanic opportunity district. Plaintiffs' plan splits one fewer municipal boundary than the current plan. The districts in plaintiffs' plan are substantially more compact than the current plan (average compactness of 0.41 versus 0.28).

## OPINION

## A. Standard of Review

To satisfy federal pleading standards, a plaintiff need only draft a complaint that provides the defendants adequate notice and "state[s] a claim to relief that is plausible on its face." Bell Atlantic Corp. v. Twombly, 550 U.S. 544, 570 (2007); see also Ashcroft v. Iqbal, 556 U.S. 662, 678 (2009). "Plausible" does not mean "probable." Iqbal, 556 U.S. at 678. See also Alexander v. United States, 721 F.3d 418, 422 (7th Cir. 2013) ("[I]t is not.. . necessary (or appropriate) to stack up inferences side by side and allow the case to go forward only if the plaintiff's inferences seem more compelling than the opposing inferences.") (internal quotations omitted). Rather,
"plausible" means that the plaintiffs'_allegations are sufficient to raise a right to relief above the speculative level. Twombly, 550_U.S. at 555 . The same standard applies to both the merits and jurisdiction. Silha v. Act, Inc., No. 15-1083, - F.3d. —, 2015 WL 7281602, at *4 (7th Cir. Nov. 18, 2015) ("When_evaluating a facial challenge to subject matter jurisdiction under Rule 12(b)(1), a court should use Twombly-Iqbal's 'plausibility' requirement."). We must accept all well-pleaded factual allegations as true and view them in the light most favorable to the plaintiffs. Luevano v. Wal-Mart Stores, Inc., 722 F.3d 1014, 1027 (7th Cir. 2013).

## B. Political Question Doctrine

In their opening brief, defendants ask this court to grant their motion to dismiss on the ground that plaintiffs' claims are not justiciable, or, more specifically, that partisan gerrymandering claims raise political questions that only other branches of government can resolve because the claims lack a judicially manageable standard. Zivotofskyex rel. Zivotofsky v. Clinton, 132 S. Ct. 1421, 1427 (2012) ("[A] controversy involves a political question where there is a textually demonstrable constitutional commitment of the issue to a coordinate political department; or a lack of judicially discoverable and manageable standards for resolving it.") (internal quotations and alterations omitted). We decline this request because defendants' position has not been adopted by a majority of the justices on the Supreme

Court. In Davis v. Bandemer, 478 U.S. 109, 123 (1986), the Court rejected the argument that partisan gerrymandering claims are nonjusticiable political questions. In Vieth v. Jubelirer, 541 U.S. 267, 305 (2004), four justices expressed the view that partisan gerrymandering is a political question, but the other five justices rejected that view. In League of United Latin American Citizens v. Perry, 548 U.S. 399, 420 (2006), the Court declined to revisit the issue. Since LULAC, the Court has not considered the merits of a partisan gerrymandering claim, so we conclude that Bandemer still controls on the narrow question whether partisan gerrymandering claims are barred under the political question doctrine. See also Shapiro v. McManus, No. 14-990, - U.S. - 2015 WL 8074453, at *5 (U.S. Dec. 8, 2015) (acknowledging that a majority of the Court has declined to find partisan gerrymandering claims nonjusticiable). Until a majority of the Supreme Court rules otherwise, lower courts must continue to search for a judicially manageable standard.

## C. Standing

In an order dated November 17, 2015, dkt. \#38, we asked the parties to submit supplemental briefs on the threshold question whether plaintiffs have standing to sue under the test articulated in Lujan v. Defenders of Wildlife, 504 U.S. 555, 560 (1992), which requires the plaintiffs to show that they have suffered a concrete and particularized injury that is fairly traceable to the defendants' conduct and that is likely
to be redressed by winning the lawsuit. In particular, we asked the parties to discuss whether a voter has a legal interest in the election results outside his or her own district. Having reviewed the parties' submissions, we are persuaded that plaintiffs have met their burden at the pleading stage to allege that they have standing.

In their supplemental briefs, plaintiffs say that their injury is set forth in paragraph 16 of their complaint:

Regardless of where they reside in Wisconsin and whether they themselves reside in a district that have been cracked or packed, all of the plaintiffs have been harmed by the manipulation of district boundaries in the Current Plan to dilute Democratic voting strength. As a result of the statewide partisan gerrymandering, Democrats do not have the same opportunity provided to Republicans to elect representatives of their choice to the Assembly. As a result, the electoral influence of plaintiffs and other Democratic voters statewide has been unfairly, disproportionately, and undemocratically reduced

In other words, we understand plaintiffs to identify their injury as not simply their inability to elect a representative in their own districts, but also their reduced opportunity to be represented by Democratic
legislators across the state. Plts.' Supp. Br., dkt. \#41, at 5 ("The Current Plan's enormous (and intentional) pro-Republican efficiency gap injures all Democrats in Wisconsin by diluting the collective value of their individual votes on a statewide basis.").

In arguing that plaintiffs do not have standing to bring a statewide challenge, defendants point to Justice Stevens' dissent in Vieth, 541 U.S. at 327-28, in which he suggested that a plaintiff's standing in a political gerrymandering case should be governed by the same standard as a racial gerrymandering case under the equal protection clause, which generally limits a plaintiff's standing to the district in which he or she lives. Alabama Legislative Black Caucus v. Alabama, 135 S. Ct. 1257, 1265 (2015) ("[A] voter who lives elsewhere in the State . . . normally lacks standing to pursue a racial gerrymandering claim."). Because plaintiffs have not joined a Democrat from each of the 99 Assembly districts, defendants argue that plaintiffs lack standing to bring a statewide challenge. Lower courts and commentators have also raised questions about whether the type of injury plaintiff allege is sufficiently concrete and particularized to confer standing. E.g., Radogno v. Illinois State Bd. of Elections, No. 1:11-CV-04884, 2011 WL 5025251, at * 4 (N.D. Ill. Oct. 21, 2011) ("The standing analysis for political gerrymandering claims is complicated by the largely unresolved status of political gerrymandering claims in general. That is, even if such claims are theoretically viable . . . it is not particularly clear who would have standing to
bring them."); Heather K. Gerken, Lost in the Political Thicket: The Court, Election Law, and the Doctrinal Interregnum, 153 U. Pa. L. Rev. 503, 509 n. 31 (2004) (noting "the difficulty that those endorsing a purely individualist approach to the partisan gerrymander would encounter in describing the injury in sufficiently concrete terms to confer standing").

Although the answer is not free from doubt, we conclude that plaintiffs' alleged injury is sufficiently concrete and particularized under current law to satisfy Lujan with respect to a statewide challenge to the districting plan, even without a plaintiff from every legislative district. In each of the three cases in which the Supreme Court considered partisan gerrymandering claims, the plaintiffs were challenging the plan statewide, yet only one justice (Justice Stevens) questioned the plaintiffs' standing. LULAC, 548 U.S. at 419-20 (opinion of Kennedy, J.) (discussing statewide scope of plaintiffs' partisan gerrymandering claims); Vieth, 541 U.S. at 285 (plurality opinion) ("[A]ppellants propose a test that is satisfied only when partisan advantage was the predominant motivation behind the entire statewide plan.") (internal quotations omitted); Bandemer, 478 U.S. at 127 ("[T]he claim made by the appellees in this case is a claim that the 1981 apportionment discriminates against Democrats on a statewide basis."). Arguably, Justice O'Connor raised a similar concern in Bandemer, although she treated the issue as one related to the merits rather than standing.

Bandemer, 478 U.S. at 153 ("To treat the loss of candidates nominated by the party of a voter's choice as a harm to the individual voter, when that voter cannot vote for such candidates and is not represented by them in any direct sense, clearly exceeds the limits of the Equal Protection Clause.").

As we noted in the November 17 order, the Supreme Court's failure to address standing in Bandemer, Vieth and LULAC is not dispositive because "assumptions-even on jurisdictional issues-are not binding." Domino's Pizza, Inc. v. McDonald, 546 U.S. 470, 478-79 (2006). However, it seems telling that the Supreme Court has not rested its determinations on this threshold issue. Even in Vieth, in which the district court found expressly that the plaintiffs had standing to raise a statewide claim, Vieth v. Pennsylvania, 188 F. Supp. 2d 532, 539-40 (M.D. Pa. 2002), and Justice Stevens raised standing concerns, the other justices did not think it necessary to consider the issue, even though dismissing the case for lack of standing would have involved a more straightforward analysis than a discussion of the political question doctrine or the merits. Again, when the Court decided LULAC two years later, only Justice Stevens discussed standing.

In other cases, the Supreme Court has recognized injuries similar to those alleged by plaintiffs in this case. For example, in cases challenging the drawing of legislative districts under Section 2 of the Voting Rights Act, the harm may include the dilution of a
racial minority's political power through "cracking" and "packing" that minority in order to minimize the number of districts in which that minority may elect the candidate of its choice. E.g., Johnson v. De Grandy, 512 U.S. 997, 107 (1994) ("[M]anipulation of district lines can dilute the voting strength of politically cohesive minority group members, whether by fragmenting the minority voters among several districts where a bloc-voting majority can routinely outvote them, or by packing them into one or a small number of districts to minimize their influence in the districts next door."). See also LULAC, 548 U.S. at 496 (Roberts, C.J., concurring in part, concurring in the judgment and dissenting in part) ("[A] § 2 plaintiff must at least show an apportionment that is likely to perform better for minority voters, compared to the existing one."). Under Section 2, the scope of the claim is tied to the scope of the injury, so standing to sue is limited to "[p]laintiffs [who] reside in a reasonably compact area that could support additional" majority-minority districts. Pope v. Cty. of Albany, No. 1:11-CV-0736 LEK/CFH, 2014 WL 316703, at *5-6 (N.D.N.Y. Jan. 28, 2014) (citing Johnson, 512 U.S. at 1013-15).

Because plaintiffs' alleged injury in this case relates to their statewide representation, it follows that they should be permitted to bring a statewide claim. As plaintiffs point out, the Supreme Court has found that individual plaintiffs have standing to bring challenges to the entire state's districting map in "one-person, one-vote" cases, in which the plaintiffs
allege that population differences among legislative districts violate the equal protection clause. Baker v. Carr, 369 U.S. 186, 204-08 (1962). At this stage, this is further support for a more general view that plaintiffs challenging legislative districts have standing to challenge the entire state plan when the nature of the injury is statewide. Defendants say that Baker is only about the dilution of an individual vote, but that is not necessarily true. Bandemer, 478 U.S. at 166-67 ("While population disparities do dilute the weight of individual votes, their discriminatory effect is felt only when those individual votes are combined.") (emphasis added) (opinion of Powell, J.).

We acknowledge that the Supreme Court's limited discussion of standing in the context of gerrymandering claims leaves some questions unanswered.E.g., Richard H. Fallon, Jr., The Fragmentation of Standing, 93 Tex. L. Rev. 1061, 1117 (2015) (stating that injury recognized in "oneperson, one-vote" cases does "not fit comfortably within the conceptual bounds" of Lujan framework); Timothy G. O'Rourk, Shaw v. Reno: The Shape of Things to Come, 26 Rutgers L.J. 723, 773 (1995) (questioning whether standing in "one- person, onevote" cases should be treated differently from racial gerrymandering cases). Although it may be that ultimately the Supreme Court decides to limit standing in all gerrymandering cases the same way it has limited racial gerrymandering claims under the equal protection clause, we believe that, under
current law, plaintiffs have adequately alleged an injury in fact.

We reach the same conclusion with respect to the second and third elements of standing, which are causation and redressability. Plaintiffs have alleged that defendants' districting plan has denied them a fair chance to elect representatives across the state and that adopting a new plan that complies with their theory of partisan symmetry would make it easier for them to gain representation. At this stage of the proceedings, we must accept those allegations as true.

Our conclusion that plaintiffs have adequately alleged standing is supported by defendants' failure to cite any cases in which a court found in a partisan gerrymandering case that the plaintiffs did not have standing to bring a statewide challenge. Although the cases plaintiffs cite contain little discussion of standing, we are hesitant to dismiss a case for lack of standing based solely on the pleadings when other courts considering partisan gerrymandering consistently have assumed that standing exists to challenge a statewide plan. E.g., Perez v. Perry, 26 F. Supp. 3d 612 (W.D. Tex. 2014); Baldus v. Members of Wisconsin Government Accountability Board, 849 F. Supp. 2d 840 (E.D. Wis. 2012); Committee for a Fair and Balanced Map v. Illinois State Board of Elections, 835 F. Supp. 2d 563 (N.D. Ill._2011); Fletcher v. Lamone, 831 F. Supp. 2d 887, 903-04 (D. Md. 2011); Perez v. Texas, 2011 WL 9160142, at *9 (W.D. Tex. 2011); Radogno v. Illinois State B of Elections, 2011

WL 5868225 (N.D. Ill. 2011); Radogno v. Illinois State Board of Elections, 2011 WL 5025251, at *4 (N.D. Ill. 2011).

Accordingly, we are denying defendants' motion to dismiss for lack of standing. However, defendants are free to raise this issue again on a more developed record in the context of a motion for summary judgment.

## D. Merits

With respect to the merits, the parties focus on plaintiffs' equal protection claim, so we will do the same. (The parties debate whether defendants' motion seeks dismissal of plaintiffs' First Amendment claim, but we need not resolve that issue because neither side identifies an analytical difference between the two claims.) Generally, an equal protection claim requires a showing of a discriminatory intent and a discriminatory effect. Bandemer, 478 U.S. at 127 (plurality opinion) (citing City of Mobile, Alabama v. Bolden, 446 U.S. 55, 67-68 (1980)). However, because the Supreme Court has stated that some amount of partisan bias is inevitable in redistricting, e.g., id. at 129 (plurality opinion), the challenge in partisan gerrymandering claims has been in determining "how much is too much" and in choosing the appropriate standard for making that determination. Vieth, 541 U.S. at 298-99 (plurality opinion); id. at 344 (Souter, J., dissenting). Thus far, the Supreme Court has notidentified a standard for
reviewing a partisan gerrymandering claim, but it has left open the possibility that an appropriate standard may be found. Vieth, 541 U.S. at 306 (Kennedy, J., concurring in the judgment) ("I would not foreclose all possibility of judicial relief if some limited and precise rationale were found to correct an established violation of the Constitution in some redistricting cases.").

Plaintiffs set forth a three-part test for establishing a constitutional violation. In step one, the plaintiffs must show that the defendants intended to discriminate against an "identifiable political group" of which the plaintiffs are a member. Plts.' Br., dkt. \#31, at 9 (quoting Bandemer, 478 U.S. at 127 (plurality opinion)). In step two, the plaintiffs must show a discriminatory effect through a metric called the "efficiency gap," which is discussed more below. If the plaintiffs make the first two showings, the burden shifts to the defendants in step three to show that the efficiency gap was "the necessary result of either a legitimate state policy or the state's underlying political geography." Id. at 10. Plaintiffs say that they modeled steps two and three after the standard for "one-person, one-vote" gerrymandering cases, under which the state must show that population deviations over ten percent are justified by a legitimate state interest. E.g., Brown v. Thomson, 462 U.S. 835, 842-43 (1983).

With respect to the first element, plaintiffs point to their allegations that Republican state legislators
hired lawyers and an expert for the purpose of redrawing all district lines to maximize Republican victories and minimize wins for Democratic candidates. Cpt. $\mathbb{T}$ 8, $31,33-36$, dkt. \#1. The plan was drafted in secret and without any input from Democrats. Id. at ब| ${ }^{\text {I }}$ 8, 31-32, 37-40. Defendants do not challenge this part of plaintiffs'standard and they do not deny that plaintiffs have adequately alleged discriminatory intent against an identifiable political group (Democratic voters).

The parties focus on whether plaintiffs have adequately pleaded a discriminatory effect and, more generally, whether plaintiffs have identified a judicially discernible and manageable standard for making a showing of discriminatory effect. Plaintiffs' theory of equal representation comes from a concept called "partisan symmetry," which plaintiffs define as "the idea that a district plan should treat the major parties symmetrically with respect to the conversion of votes to seats and that neither party should have a systematic advantage in how efficiently its popular support translates into legislative power." Cpt. ब 4, dkt. \#1. See also LULAC, 548 U.S. at 466-67 (Stevens, J., concurring in part and dissenting in part) ("The symmetry standard requires that the electoral system treat similarly-situated parties equally.").

Plaintiffs measure partisan symmetry through what they call the "efficiency gap," which plaintiffs describe as follows:
[t]he efficiency gap captures in a single number all of a district plan's cracking and packingthe two fundamental ways in which partisan gerrymanders are constructed. Cracking means dividing a party's supporters among multiple districts so that they fall short of a majority in each one. Packing means concentrating one party's backers in a few districts that they win by overwhelming margins. Both cracking and packing result in "wasted" votes: votes cast either for a losing candidate (in the case of cracking) or for a winning candidate but in excess of what he or she needed to prevail (in the case of packing). The efficiency gap is the difference between the parties' respective wasted votes in an election, divided by the total number of votes cast.

Cpt. ब 5, dkt. \#1. Plaintiffs provide an example to demonstrate how the efficiency gap is calculated:

Suppose, for example, that there are five districts in a plan with 100 voters each. Suppose also that Party A wins three of the districts by a margin of 60 votes to 40 , and that Party B wins two of them by a margin of 80 votes to 20 . Then Party A wastes 10 votes in each of the three districts it wins and 20 votes in each of the two districts it loses, adding up to 70 wasted votes. Likewise, Party B wastes 30 votes in each of the two districts it wins and 40 votes in each of the three districts it loses,
adding up to 180 wasted votes. The difference between the parties' respective wasted votes is 110 , which, when divided by 500 total votes, yields an efficiency gap of $22 \%$ in favor of Party A.

Id. at ब 50. (Another measure of partisan symmetry is "partisan bias," which plaintiffs define as "the difference between the shares of seats that the parties would win if they each received the same share of the statewide vote." Plts.' Br., dkt. \#31, at 9. Although plaintiffs allege that the 2012 Assembly plan demonstrates a high level of partisan bias, the parties focus on the efficiency gap, so we will do the same.)

According to plaintiffs, the efficiency gap accurately measures discriminatory effect because it shows the extent to which a "party . . . enjoy[s] a significant advantage in how efficiently its votes convert into seats." Plts.' Br., dkt. \#31, at 18. Plaintiffs say that such an advantage violates "every voter['s]constitutional right to equal treatment in the electoral system-and the right not to be treated differently based on the voter's political beliefs." Id. Thus, plaintiffs argue, if they can show that the defendants acted with partisan intent and that the efficiency gap exceeds a "reasonable threshold," then the plan is presumptively unconstitutional. Plts.' Br., dkt.\#31, at 9. In determining the threshold, the court looks at the efficiency gap from other elections over time and across the country. Id. at 4. Plaintiffs contend that a gap of more than 7 percent is a strong
indicator that the bias in favor of a particular party is likely to endure for the life of the districting plan. Id.

Plaintiffs assert that the 2012 Assembly Plan meets their test because the efficiency gap for the 2012 election was 12 percent and the efficiency gap for the 2014 election was 10 percent, both of which are greater than the threshold. In their motion, defendants do not challenge the sufficiency of plaintiffs' allegations that a district plan with an efficiency gap as high as Wisconsin's Assembly plan is "highly unlikely ever to become neutral over its tenyear lifespan" or that plaintiffs "can predict with nearly $100 \%$ confidence that. . . Wisconsin's Current Plan will continue to unfairly favor Republican voters and candidates-and unfairly disadvantage Democratic voters and candidates-throughout the remainder of the decade." Cpt. © 7, dkt. \#1. In addition, defendants do not challenge the sufficiency of plaintiffs' allegations that the Assembly plan's efficiency gap cannot be justified by traditional districting criteria or any other legitimate factor.

Defendants' primary argument is that partisan symmetry is no different from the theories that the Supreme Court has rejected in the past. In particular, defendants say that partisan symmetry is simply a form of proportional representation, which the Supreme Court has said repeatedly is not required by the Constitution. E.g., Bandemer, 478 U.S. at 129-30 (plurality opinion) ("Our cases, however, clearly foreclose any claim that the Constitution requires
proportional representation or that legislatures in reapportioning must draw district lines to come as near as possible to allocating seats to the contending parties in proportion to what their anticipated statewide vote will be.").

Plaintiffs argue that the efficiency gap is about comparing the wasted votes of each party, not determining whether the party's percentage of the statewide vote share is reflected in the number of representatives that party elects successfully. At this stage, we must accept as true the allegation that an election's results may have a small efficiency gap without being proportional or they may be proportional and still have a large efficiency gap. Plts.' Br., dkt. \#31, at 24 (citing Nicholas O. Stephanopoulos \& Eric M. McGhee, Partisan Gerrymandering and the Efficiency Gap, 82 U. Chi.L. Rev. 831, 854 \& n. 118 (2015)). ${ }^{1}$ Further, the plaintiffs
${ }^{1}$ Plaintiffs provide the following example in their brief:
[A]ssume that a state has ten districts, each with a hundred voters, and two parties, Party A and Party B. Assume also that Party A wins two districts by a margin of 80 to 20 and four districts by a margin of 70 to 30 , and that Party $B$ wins four districts by a margin of 60 to 40 . Then there is perfectly proportional representation; Party A receives 600 of the 1000 votes in the state ( $(2 \mathrm{x}$ $80)+(4 \times 70)+(4 \times 40))$ and wins six of the ten seats. But the efficiency gap here is not zero. It is actually $10 \%$, the difference between Party A's 300 wasted votes $((2 \times 30)+(4 \times 20)+(4 \times 40))$
in Bandemer, Vieth and LULAC did not rely on partisan symmetry in their arguments before the Court and the Court did not reject partisan symmetry as a tool in determining whether partisan gerrymandering is unconstitutional. LULAC, 548 U.S. at 417 (opinion of Kennedy, J.) (rejecting standard requiring plaintiffs to show "single-minded purpose . . . to gain partisan advantage"); Vieth, 541 U.S. at 284-87 (plurality opinion) (rejecting standard requiring plaintiffs to show "predominant intent to achieve partisan advantage," "systematically 'pack[ing]' and 'crack[ing]' the rival party's voters" and "thwart[ing] plaintiffs' ability to translate a majority of votes into a majority of seats"); Bandemer, 478 U.S. at 129-30 (rejecting proportional representation requirement).

In fact, some of the justices have pointed to partisan symmetry as a theory with promise. LULAC, 548 U.S. at 465-66 (Stevens, J., concurring in part and dissenting in part) ("[T]he symmetry standard, a measure social scientists use to assess partisan bias . . . is undoubtedly a reliable standard for measuring a burden on the complainants'
and Party B's 200 wasted votes $((2 \times 20)+(4 \times$ $30)+(4 \times 10))$, divided by the 1000 total votes cast.

Plts.' Br., dkt. \#31, at 24. See also Stephanopoulos \& McGhee, supra, 82 U. Chi. L. Rev. at 854 n. 118 ("According to the efficiency gap equation, ... [i]f a party receives 60 percent of the vote and 60 percent of the seats, for example, a plan would have an efficiency gap of 10 percent against the party.").
representative rights.") (internal quotations and alterations omitted); id. at 483-84 (Souter, J., concurring in part and dissenting in part) ("[N]or do I rule out the utility of a criterion of symmetry as a test. Interest in exploring this notion is evident. Perhaps further attention could be devoted to the administrability of such a criterion at all levels of redistricting and its review.") (citations omitted). Justice Kennedy's support for partisan symmetry is tepid at best, but he left room for partisan symmetry to play some role in the analysis. Id. at 419-20 (opinion of Kennedy, J.)) ("Without altogether discounting its utility in redistricting planning and litigation, I would conclude asymmetry alone is not a reliable measure of unconstitutional partisanship."). See also id. at 468 n. 9 (Stevens, J., concurring in part and dissenting in part) ("I appreciate Justice Kennedy's leaving the door open to the use of the [partisan symmetry] standard in future cases.").

Much of defendants' remaining argument is devoted to mischaracterizations of plaintiffs' proposed standard. For example, defendants argue that plaintiffs' test does not take into account traditional districting principles or the reasons unrelated to partisan intent that voters of a particular party might be "cracked" or "packed," such as the natural concentration of Democrats into urban areas. Dfts.' Br., dkt. \#25, at 22-23. In addition, defendants say that, under plaintiffs' proposed standard, the 2002 Wisconsin Assembly plan would be unconstitutional
because it had a large efficiency gap, even though it was drawn by a court. Id. at 24 .

These arguments rely on the assumption that plaintiffs' proposed standard consists of nothing except a calculation of the efficiency gap. Defendants simply have ignored step one and step three of plaintiff's standard. Even if the plaintiffs were able to establish that the efficiency gap is a sufficiently strong pillar to support a constitutional violation, the plaintiffs still must prove partisan intent (step one). The defendants also might be able to show that a large efficiency gap is justified by a legitimate state interest, which may include traditional districting criteria such as equal population, compliance with the Voting Rights Act, compactness, respect for political subdivisions or respect for communities of interest (step three).

We have reviewed defendants' remaining arguments and conclude that they are unpersuasive or premature. A determination whether plaintiffs' proposed standard is judicially manageable relies at least in part on the validity of plaintiffs' expert opinions, which we must accept as true in the context of a motion a dismiss. A more developed record may show that plaintiffs' claims cannot be legally distinguished from the partisan gerrymandering claims that the Supreme Court has rejected in the past. However, current law does not foreclose plaintiffs' claims and those claims are modeled after a standard that the Supreme Court has adopted in

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other contexts. Accordingly, we conclude that plaintiffs have stated a claim for relief that is plausible on its face and we are denying defendants' motion to dismiss.

## ORDER

IT IS ORDERED that the motion to dismiss filed by defendants Gerald C. Nichol, Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas,Timothy Vocke and Kevin J. Kennedy, dkt. \#24, is DENIED.

Entered this 17th day of December, 2015.

## BY THE COURT:

/s/

KENNETH F. RIPPLE
Circuit Judge
/s/

BARBARA B. CRABB
District Judge
/s/

WILLIAM C. GRIESBACH
District Judge

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## Cross-References to Supplemental Appendix

Use of Efficiency Gap in Analyzing Partisan Gerrymandering, Professor Nicholas Goedert (ECF No.51) appears at: SA1-25

Analysis of the Efficiency Gaps of Wisconsin's Current Legislative District Plan and Plaintiffs' Demonstration Plan, Dr. Kenneth R. Mayer (ECF No. 54) appears at: SA26-98

Declaration of Sean Trende (ECF No. 55) appears at: SA99-146

Rebuttal Report: Response to Expert Reports of Sean Trende and Nicholas Goedert, Dr. Kenneth R. Mayer (ECF No. 59-2) appears at: SA147-198

Assessing the Current Wisconsin State Legislative Districting Plan, Professor Simon Jackman (ECF No. 62) appears at: SA179-254

Rebuttal Report, Professor Simon Jackman (ECF No. 63) appears at: SA255-281

# Opinion and Order Denying Defendants' Motion for Summary Judgment 

IN THE UNITED STATES DISTRICT COURT FOR THE WESTERN DISTRICT OF WISCONSIN

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WILLIAM WHITFORD, ROGER
ANCLAM, EMILY BUNTING,
MARY LYNNE DONOHUE,
HELEN HARRIS, WAYNE
JENSEN, WENDY SUE
JOHNSON, JANET MITCHELL,
ALLISON SEATON, JAMES
SEATON, JEROME WALLACE
and DONALD WINTER,
                    Plaintiffs, OPINION AND
                                    ORDER
v.
    15-cv-421-bbc
GERALD C. NICHOL, THOMAS BARLAND, JOHN FRANKE, HAROLD V. FROEHLICH, KEVIN J. KENNEDY, ELSA LAMELAS and TIMOTHY VOCKE,
Defendants.
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The question in this case is whether Wisconsin Act 43-the 2012 districting plan for the Wisconsin Assembly-is an unconstitutional partisan gerrymander. Plaintiffs are Wisconsin residents and Democratic voters who allege that the plan is "one of the worst partisan gerrymanders in modern American history." Cpt. ब 1, dkt. \#1. In particular, plaintiffs allege that Republican legislators drew the plan in secret, in consultation with a political scientist and without any input from Democrats, in an attempt to maximize Republican wins and minimize Democratic influence over the political process for as long as the plan was in place. In addition, plaintiffs allege that Republicans were successful in their attempt, gaining significantly more Assembly seats in 2012 and 2014 than their level of public support suggests. As proof that Republicans unfairly manipulated district lines, plaintiffs created their own plan, which they say satisfies traditional districting criteria such as compactness, contiguity and respect for political subdivisions as well or better than Act 43 but treats Democrat and Republican voters much more equally.

In an order dated December 17, 2015, dkt.\#43, we denied defendants' motion to dismiss after concluding that plaintiffs' allegations were sufficient to state a plausible claim for relief. Now defendants have filed a motion for summary judgment, dkt. \#45, which is ready for review. In addition, plaintiffs have filed what they call a "motion in limine" to exclude the
opinions of one of defendants' named experts, Sean Trende. Dkt. \#70.

Defendants raise many important points in their summary judgment submissions. It may be that one or more of these objections carries the day in the end. However, we believe that deciding the case now as a matter of law would be premature because there are factual disputes regarding the validity of plaintiffs' proposed measurement for determining the existence of a constitutional violation. Accordingly, we deny defendants' motion for summary judgment and allow the case to proceed to trial.

We are also denying plaintiffs' motion in limine without prejudice to plaintiffs' renewing the motion at the conclusion of trial. Plaintiffs raise significant objections in their motion. However, because it is not necessary to consider Trende's opinions in order to resolve the motion for summary judgment and because the trial will be to a court rather than to a jury, we believe the prudent course of action is to rule on the admissibility of Trende's opinions after he has an opportunity to testify. Metavante Corp. v. Emigrant Savings Bank, 619 F.3d 748, 760 (7th Cir. 2010) ("[T]he court in a bench trial need not make reliability determinations [regarding experts] before evidence is presented."); In re Salem, 465 F.3d 767, 777 (7th Cir. 2006) ("[W]here the factfinder and the gatekeeper are the same, the court does not err in admitting the [expert] evidence subject to the ability later to exclude it or disregard it if it turns out not to
meet the standard of reliability established by Rule 702.").

To accommodate a court scheduling conflict, the trial will begin on Tuesday, May 24, 2016, at 9:00 a.m. The parties should be prepared to finish the trial in four days.

## OPINION

In the order denying the motion to dismiss, we considered three issues: (1) whether challenges to a partisan gerrymander were justiciable; (2) whether plaintiffs had standing to sue; and (3) whether plaintiffs stated a plausible claim for relief. We answered each of these questions in the affirmative. Because defendants do not raise any new arguments about justiciability or standing in their summary judgment submissions, we see no reason to discuss those issues in this opinion. Instead, we will focus on whether plaintiffs have raised any genuine issues of material fact with respect to the various objections raised in defendants' motion for summary judgment. Fed. R. Civ. P. 56.

## A. Legal Background

As the parties well know, there is much uncertainty in the law regarding partisan gerrymandering. Although the Supreme Court has well-established tests for analyzing alleged gerrymanders with respect to race, e.g., Miller v.

Johnson, 515 U.S. 900, 916-17 (1995), and equal population, e.g., Evenwel v. Abbott, No. 14-940, 2016 WL 1278477, at *3 (U.S. Apr. 4, 2016); Brown v. Thomson, 462 U.S. 835, 842-43 (1983), the Court has struggled to determine the appropriate test for gerrymanders based on political affiliation. In Davis v. Bandemer, 478 U.S. 109, 118-27 (1986), a majority of the Court agreed that partisan gerrymander claims are justiciable under the equal protection clause and that the plaintiffs must prove a discriminatory intent and a discriminatory effect. However, the Court could not agree on a specific standard to apply, particularly with respect to determining a discriminatory effect. Compare Bandemer, 478 U.S. at 133 (plurality opinion) ("[U]nconstitutional discrimination occurs only when the electoral system is arranged in a manner that will consistently degrade a voter's or a group of voters' influence on the political process as a whole."), with id. at 161 (Powell, J., concurring in part and dissenting in part) (question is whether legislature acted solely for partisan ends to the exclusion of "all other neutral factors relevant to the fairness of redistricting").

In Vieth v. Jubelirer, 541 U.S. 267 (2004), four Justices concluded that Bandemer should be overruled because partisan gerrymanders present political questions that cannot be answered by federal courts. Id. at 305 (plurality opinion). Four other Justices agreed that the Bandemer plurality did not provide a workable standard, but they disagreed with the plurality regarding justiciability and they

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proposed alternative standards for reviewing a partisan gerrymandering claim. Compare Vieth, 541 U.S. at 339 (Stevens, J., dissenting) (question is "whether the legislature allowed partisan considerations to dominate and control the lines drawn, forsaking all neutral principles"), with Vieth, 541 U.S. at 346-51 (Souter, J., dissenting) (proposing burden-shifting framework modeled after McDonnell Douglas Corp. v. Green, 411 U.S. 792 (1973)), and Vieth, 541 U.S. at 360-61 (Breyer, J., dissenting) (question is whether there was "unjustified use of political factors to entrench a minority in power").

In the middle, Justice Kennedy concluded that neither the Justices nor the parties had provided a workable standard, but he declined to close the door on future partisan gerrymandering claims. Id. at 30608 (Kennedy, J., concurring in the judgment). Rather, he stated that "courts should be prepared to order relief" if"workable standards do emerge." Id. at 317. He suggested that future cases could be guided not just by the equal protection clause but also by the First Amendment, focusing on the question whether a plan "burden[s] or penaliz[es] citizens because of their participation in the electoral process, their voting history, their association with a political party, or their expression of political views." Id. at 314. See also Baldus, 849 F. Supp. 2d at 853 ("[P]erhaps the Court will find some day that the First Amendment also protects persons against state action that intentionally uses their partisan affiliation to affect the weight of their vote.").

Finally, in League of United Latin American Citizens v. Perry, 548 U.S. 399 (2006), the Court assumed that partisan gerrymanders are justiciable, but a majority concluded that the plaintiffs had failed to identify "a manageable, reliable measure of fairness for determining whether a partisan gerrymander violates the Constitution." Id. at 414. Again, the dissenting Justices proposed alternative standards in line with those they proposed in Vieth. Compare LULAC, 548 U.S. at 474 (Stevens, J., concurring in part and dissenting in part), with LULAC, 548 U.S. at 391-92 (Breyer, J., concurring in part and dissenting in part).

Since LULAC, the Supreme Court has not considered a partisan gerrymandering claim. Thus, it is left to parties bringing those claims and the lower courts considering them to continue to search for a workable standard that reflects a voter's right to "fair and effective representation." Reynolds v. Sims, 377 U.S. 533, 565 (1964). See also Baldus, 849 F. Supp. 2 d at 853 ("Justice Kennedy's pivotal opinion [in Vieth] appeared to throw the ball to the litigating parties to come up with a manageable legal standard.").

## B. Plaintiffs' Proposed Standard

In this case, plaintiffs' proposed test adopts the basic structure of a claim brought under the equal protection clause, which generally requires a showing of discriminatory intent and discriminatory effect.

Bandemer, 478 U.S. at 127 (plurality opinion) (citing City of Mobile, Alabama v. Bolden, 446 U.S. 55, 6768 (1980)). Perhaps in response to Justice Kennedy's opinion in Vieth, plaintiffs' complaint includes a claim under the First Amendment as well, but at this point, neither side has developed a separate argument under the First Amendment or identified any analytical differences between plaintiffs' First Amendment and equal protection claims.

Plaintiffs' proposed test has three parts. First, the plaintiffs must show that the defendants acted with discriminatory intent. More specifically, plaintiffs frame the question as whether the "plan was designed with the intention of benefiting one party and disadvantaging its adversary." Plts.' Br., dkt.\#68, at 58. At oral argument, plaintiffs summarized this element as an intent to disadvantage on the basis of political affiliation and they said that they modeled the element on the standard in Bandemer. Trans., dkt. \#89, at 47 .

Plaintiffs' most significant innovation in their test is the second part, with respect to discriminatory effect. Under this part, the plaintiffs must show that the plan "exhibited a high and durable level of partisan asymmetry in the first election after redistricting." Id. at 59. Plaintiffs define "partisan symmetry" as "the idea that the electoral system should treat similarly-situated parties equally, so that they are able to convert their popular support into legislative representation with approximately
equal ease." Id. at 49 (internal quotations omitted). Plaintiffs say that partisan symmetry provides an appropriate basis for evaluating discriminatory effect because several Justices in LULAC relied on it or otherwise discussed it favorably. E.g., LULAC, 548 U.S. at 466 (Stevens, J., concurring in part and dissenting in part) (partisan symmetry is "undoubtedly a reliable standard for measuring a burden on the complainants' representative rights") (internal quotations omitted); id. at 483-84 (Souter, J., concurring in part and dissenting in part) ("[N]or do I rule out the utility of a criterion of symmetry as a test. Interest in exploring this notion is evident. Perhaps further attention could be devoted to the administrability of such a criterion at all levels of redistricting and its review.") (internal citations omitted). See also id. at 420 (opinion of Kennedy, J.) (declining to "altogether discount[] [partisan symmetry's] utility in redistricting planning and litigation").

In addition, plaintiffs say that partisan symmetry reflects the Supreme Court's description of partisan gerrymandering in other cases. Arizona State Legislature v. Arizona Independent Redistricting Commission, 135 S. Ct. 2652, 2658 (2015) (partisan gerrymandering is "the drawing of legislative district lines to subordinate adherents of one political party and entrench a rival party in power"); Vieth, 541 U.S. at 271 n. 1 (plurality opinion) (gerrymandering is "giv[ing] one political party an unfair advantage by diluting the opposition's voting strength") Bandemer,

478 U.S. at 127 (plurality opinion) (gerrymandering is "the manipulation of individual district lines" causing a party's "voters over the State as a whole" to be "subjected to unconstitutional discrimination."). See also Vieth, 541 U.S. at 335 (Stevens, J., dissenting) ("Gerrymandering always involves the drawing of district boundaries to maximize the voting strength of the dominant political faction and to minimize the strength of one or more groups of opponents.").

Finally, plaintiffs say that partisan symmetry is widely accepted among scholars as the most appropriate way to measure partisan fairness. Plts.' Br., dkt. \#68, at 50 (citing Bernard Grofman \& Gary King, The Future of Partisan Symmetry As A Judicial Test for Partisan Gerrymandering After Lulac $v$. Perry, 6 Election L.J. 2, 6 (2007) ("We are aware of no published disagreement or even clear misunderstanding in the scholarly community about partisan symmetry as a standard for partisan fairness in plurality-based American elections since [1987.]")).

Plaintiffs measure partisan symmetry with a metric they call the "efficiency gap," which is a figure that represents the difference between the parties' "wasted votes" in an election. A vote is "wasted" under this analysis if it is either (1) cast for a candidate who lost the election or (2) cast for the winning candidate, but in excess of what the candidate needed to win. Plts.' PFOF ब 6, dkt. \#79. The efficiency gap for a particular election is the
difference between the parties' total wasted votes among all of the districts, divided by the total number of votes cast.

In the December 17, 2015 order, we noted the following example of an efficiency gap calculation provided in plaintiffs' complaint:

Suppose, for example, that there are five districts in a plan with 100 voters each. Suppose also that Party A wins three of the districts by a margin of 60 votes to 40 , and that Party B wins two of them by a margin of 80 votes to 20. Then Party A wastes 10 votes in each of the three districts it wins and 20 votes in each of the two districts it loses, adding up to 70 wasted votes. Likewise, Party B wastes 30 votes in each of the two districts it wins and 40 votes in each of the three districts it loses, adding up to 180 wasted votes. The difference between the parties' respective wasted votes is 110 , which, when divided by 500 total votes, yields an efficiency gap of $22 \%$ in favor of Party A.

Cpt. ब 50 , Dkt. \#1. ${ }^{1}$

[^2]The purpose of the efficiency gap is to capture in one number the extent to which voters of a particular party are "packed" and "cracked." Packing means concentrating one party's supporters in a few districts so that they win by overwhelming margins. Cracking means dividing a party's supporters among multiple districts so that they fall short of a majority in each one. Vieth, 541 U.S. at 287 n.7. Plaintiffs say that a high level of cracking and packing (and thus a large efficiency gap) is indicative of discriminatory effect because, all things being equal, the number of wasted votes for both parties should be about the same. Moreover, plaintiffs say that if a plan produces an efficiency gap of greater than 7 percent after the first election, subsequent elections under the same plan are highly likely to continue to be skewed in favor of the same party, even if another party significantly increases its vote share. Plts.' PFOF © © 12, 85-93, 114-18, 154, 170, dkt. \#79. Thus, plaintiffs believe that an efficiency gap of more than 7 percent, combined with a showing of discriminatory intent, should trigger a presumption that the districting plan is unconstitutional.

Plaintiffs identify an alternative measure of partisan symmetry called "partisan bias," which they defined previously as "the difference between the

[^3]shares of seats that the parties would win if they each received the same share of the statewide vote." Plts.' Br., dkt.\#31, at 9. However, neither side develops an argument in their briefs regarding the application of partisan bias to this case. At oral ar gument, plaintiffs suggested that partisan bias could be used as a kind of "robustness check" on the accuracy of the efficiency gap. Trans., dkt. \#89, at 70. Because the parties did not explore this issue in their briefs, we decline to consider it at this time.

Finally, under the third part of plaintiffs' proposed test, if plaintiffs prove both discriminatory intent and discriminatory effect, the burden shifts to defendants. In particular, the defendants must show that the plan's "severe asymmetry" was "unavoidable" in light of "the state's political geography and legitimate redistricting objectives." Plts.' Br., dkt. \#68, at 1, 59. Plaintiffs say that they modeled this part of their test after the equal apportionment cases, in which the burden shifts to the state to justify a plan if the plaintiffs show more than a ten percent population deviation among the districts. E.g., Brown, 462 U.S. 835 at 842-43.

## C. Application of Plaintiffs' Standard

For the purpose of their motion for summary judgment, defendants do not deny that plaintiffs could prove their claim under their proposed standard. With respect to the first element, discriminatory intent, plaintiffs allege that

Republican leaders in the state legislature hired a law firm and a political scientist to design an Assembly plan that would maximize the electoral advantage of Republicans. In particular, plaintiffs allege that the Republicans used past election results to measure the partisanship of the electorate and then to design districts that would either "crack" Democratic voters (dividing them into multiple districts to prevent them from reaching a majority) or "pack" those voters (concentrating them into a small number of districts). In this way, Republicans hoped to maximize the number of districts that would elect a Republican and minimize the number of districts that would elect a Democrat. Republican leaders drafted the plan in secret, without any input from Democrats, and then enacted the plan as Act 43 with little debate. Baldus, 849 F. Supp. 2d at 845,851 (summarizing process of enacting Act 43 and finding statements that drafters were not influenced by partisan factors "to be almost laughable"). During oral argument, defendants conceded that plaintiffs can prove this element of the test as plaintiffs have framed it. Trans., dkt. \#89, at 9, $88 .{ }^{2}$

With respect to the second element, discriminatory effect, plaintiffs' expert Simon Jackman, a political scientist, measured a 13 percent efficiency gap in the Republicans' favor for the 2012 Assembly election; plaintiff's other expert, Kenneth

[^4]Mayer, also a political scientist, calculated a 12 percent pro-Republican efficiency gap, using a more elaborate method. Plts.' PFOF \|Tll 10 and 15, dkt. \#79. (Mayer used the "full form" method, which means that he tallied wasted votes district by district. Id. at - 120. Jackman used the "simplified" method, using the formula $(\mathrm{S}-0.5)-2(\mathrm{~V}-0.5)$, where S was a party's statewide seat share and V was a party's statewide vote share. Id. at © 121.) These election results, plaintiffs say, were consistent with what the legislature's consultant predicted when he aided the Republicans in drafting the plan. Id. at 9 97. It is undisputed that, from 1972 to 2010, not a single legislative map in the country was as asymmetric in its first two elections as those generated in 2012 and 2014 Wisconsin Assembly elections. Id. at \| 11. According to Jackman, the map is so skewed in favor of the Republicans that there is a nearly 100 percent chance that the plan will continue to disadvantage Democrats, as measured by the efficiency gap, throughout the life of the plan. Id. at $\mathbb{\| \|} 11,84$.

With respect to the third element, whether the Republican advantage can be justified by neutral reasons, defendants have made no effort in their summary judgment submissions to defend Act 43 on neutral grounds. However, as evidence that Act 43 cannot be justified by neutral measures, plaintiffs submitted their own proposed plan, which plaintiffs say has a much smaller efficiency gap of 2 percent in favor of Republicans, but still satisfies other

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legitimate districting criteria at least as well as Act 43. Plts.' PFOF $\mid$ \| $16,142$.

## D. Defendants' Challenges to Plaintiffs' Standard

Rather than challenge plaintiffs' ability to meet the standard, defendants challenge the standard itself. However, a review of defendants' objections show that there are fact issues that need to be resolved at trial.

## 1. Efficiency gap

a. Efficiency gap as a measure of discriminatory effect

The bulk of defendants' objections relate to plaintiffs' proposed measure of discriminatory effect, the efficiency gap. Of these, the primary objection seems to be that the efficiency gap is not a good measure of discriminatory effect because even seemingly neutral plans can have a large efficiency gap. For example, defendants point to Wisconsin's 2002 Assembly plan. Although a federal court drew that plan (based on plans submitted by the political parties), Baumgart v. Wendelberger, No. 01-C-0121, 2002 WL 34127471, at *4 (E.D. Wis. May 30, 2002) amended, No. 01-C-0121, 2002 WL 34127473 (E.D. Wis. July 11, 2002), the efficiency gap for the plan was 7.5 percent in favor of the Republicans in 2002 and then fluctuated between 4 percent and 12 percent in favor of the Republicans for the remainder of the
decade. Dfts.' PFOF बी 212-216, dkt. \#74. Defendants also point to other states that have had pro-Republican efficiency gaps of more than 5 percent in recent years, even when the plan was drawn by a neutral body. Dfts.' Br., dkt. \#46, at 38. More generally, defendants rely on the opinion of one of their experts, Sean Trende, to argue that legislative plans in Wisconsin and around the country are more likely to favor Republicans in recent years because of political geography, not partisan intent. Dfts.' Br., dkt. \#48, at 26-30; Dfts.’ PFOF \|\| 234-45, dkt. \#74. In other words, defendants argue that Democrats are naturally packed into a smaller number of districts, which makes it more likely that their share of the votes statewide will be greater than their share of the legislative seats.

In response, plaintiffs say that Wisconsin's 2002 plan is an anomaly. The average efficiency gap for the Wisconsin Assembly in the 1970s, 1980s and 1990s ranged between 0.3 percent and 2.4 percent. Plts.' PFOF $9 \| 44-46$, dkt. \#79. Further, plaintiffs say that the efficiency gap may have been high in the 2002 plan because the court adopted a plan more similar to the one proposed by Republicans. Plts.' Br., dkt. \#68, at 18. As a general rule, plaintiffs say, it is much more common for plans drafted by one party to have a significantly larger efficiency gap than plans drafted through a nonpartisan or bipartisan process. Plts.' PFOF ब 174, dkt. \#79.

In any event, plaintiffs say that political geography does not explain why efficiency gaps in Wisconsin and elsewhere have become increasingly pro-Republican in recent decades. Rather, according to Mayer, Democrats and Republicans in Wisconsin have comparable spatial distributions. Plts.' PFOF【T 51-58, dkt. \#79. More generally, plaintiffs cite evidence that there is no national trend of increasing Democratic clustering. Id. at $\mathbb{T \|}$ 26-27. Rather, plaintiffs say, the reason for larger efficiency gaps favoring Republicans is increasing Republican control of state legislatures. Id. at $\mathbb{\|} \| 49-50,58,156$. If control over state legislatures had remained constant, efficiency gaps across the country would have remained relatively constant as well, including in Wisconsin. Id. at $\mathbb{T} \mid$ 23-24, $31,49-50,58,156$.

Defendants disagree with some of plaintiffs' conclusions, but they do not object to the admissibility of their experts' opinions. Accordingly, we conclude that there is a genuine dispute on the question whether a large efficiency gap is a strong indicator of a discriminatory effect.

In their reply brief and at oral argument, defendants seemed to concede that there is a genuine dispute on this issue, but they argued that the dispute is not material because the mere existence of large efficiency gaps in plans adopted by neutral bodies is sufficient to discredit the efficiency gap as a tool for measuring a constitutional violation. We are not willing to go that far, at least not in the context of a
motion for summary judgment. Plaintiffs are not arguing that voters have a right to equal results such that any plan with a large efficiency gap must be invalidated. Rather, they are arguing that they have a right to be free from being intentionally disadvantaged when they vote. This is consistent with case law under both the First Amendment and the equal protection clause, which recognizes that there are many instances in which a government act or policy may have a disparate impact even in the absence of intentional discrimination and that disparate impact alone is not enough to sustain a constitutional claim. Bond v. Atkinson, 728 F.3d 690, 692-93 (7th Cir. 2013) ("[T]he Supreme Court held in Washington v. Davis, 426 U.S. 229 (1976), that disparate impact does not violate the equal protection clause of the fourteenth amendment."); Christian Legal Society Chapter of the University of California, Hastings College of the Law v. Martinez, 561 U.S. 661, 700 (2010) (Stevens, J., concurring) ("[I]t is a basic tenet of First Amendment law that disparate impact does not, in itself, constitute viewpoint discrimination.") (citing cases). Defendants cite no authority for the view that discriminatory intent and discriminatory effect must be borne out by the same evidence.

As an alternative to their broader argument that the efficiency gap is inherently a poor measure of discriminatory effect, defendants say that what is considered a neutral efficiency gap should not be zero. This is because using zero as a baseline does not
isolate the portion of the efficiency gap that is attributable to partisan bias. Dfts.' Br., dkt. \#46, at 36. Rather, defendants say that the baseline should incorporate whatever natural advantage a party has as a result of political geography.

Defendants raise an interesting point that may be worth exploring at trial, but we do not believe that it is a ground for granting summary judgment. At most, this is a suggestion to alter the threshold of the plaintiffs' test and, perhaps, shift the burdens of production or proof. Because it is genuinely disputed whether Wisconsin's political geography has played a significant role in contributing to Act 43's efficiency gap, an adjustment to the baseline would not be dispositive at this stage of the case. However, if the facts show at trial that political geography can and does have an impact on Wisconsin's and other states' efficiency gaps, then that would support a view that some burden should be placed on plaintiffs to show as part of their prima facie case the extent to which political geography cannot explain the efficiency gap generated by Act 43.
b. Other objections to the efficiency gap

Defendants raise various other objections, both to the efficiency gap as a general concept for measuring discriminatory effect and to the way that plaintiffs have chosen to implement the efficiency gap in this case: (1) plaintiffs' experts made assumptions about incumbency and voter turnout that undermine the
accuracy of the efficiency gap; (2) by calculating the efficiency gap using the results of only one election, plaintiffs cannot show that the efficiency gap will be a predictable and reliable indicator of discriminatory effect throughout the life of a districting plan; (3) plaintiffs' experts failed to come up with a consistent way to calculate the efficiency gap; (4) plaintiffs' experts did not include all the data that they should have in their analyses; (5) plaintiffs' standard implies that they have a constitutional right to an efficiency gap favoring the Democrats; (6) the efficiency gap constitutionalizes a proportionality standard; and (7) a large number of districting plans around the country have what plaintiffs view as an unreasonably large efficiency gap.

The first four of these objections require little discussion because it is clear that plaintiffs have raised factual disputes requiring a trial. In particular, with respect to the use of assumptions about incumbency and voter turnout, plaintiffs' experts conducted additional analysis (what they call "robustness checks") to make sure their assumptions did not have a significant effect on their results. Plts.' PFOF 9\| 94-113. Defendants are free to argue at trial that plaintiffs' methods are not sufficiently reliable to be helpful in determining a constitutional violation.

With respect to the reliability of the efficiency gap to predict whether the same party will have an unfair advantage in future elections, plaintiffs cite expert evidence that historically a large initial efficiency gap

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has been a very strong indicator of a large efficiency gap throughout the life of the districting plan. Plts.' PFOF \|ी $40-84,89-93,166-69$, dkt. \#79. In addition, plaintiffs' experts conducted "sensitivity testing" in order to control for swings in elections; the results of that testing did not undermine their conclusions regarding the reliability of the efficiency gap. Id. at \|\|f $85-88,114-18,154,170$. With respect to the differences between plaintiffs' "full form" and "simplified" methods for calculating an efficiency gap, plaintiffs cite evidence that there is little practical difference between the result generated by the methods, so the choice of method does not affect the measure's viability. Id. at I 122-35. Finally, defendants' arguments about data that plaintiffs' experts should have included in their analysis are classic examples of issues that can be raised during cross examination at trial. Manpower, Inc. v. Insurance Co. of Pennsylvania, 732 F.3d 796, 809 (7th Cir. 2013) ("Assuming a rational connection between the data and the opinion-as there was here-an expert's reliance on faulty information is a matter to be explored on cross-examination; it does not go to admissibility. Our system relies on crossexamination to alert the jury to the difference between good data and speculation.") (internal quotations omitted).

Defendants may be able to show at trial that the court should not accept plaintiffs' version of the facts. Again, however, defendants do not object to the
admissibility of plaintiffs' evidence, so we cannot resolve these issues on summary judgment.

Defendants' last three objections require more analysis. These are discussed below.
a. Implications of plaintiffs' durability threshold

As noted above, plaintiffs argue that an efficiency gap of seven percent or greater should qualify as a discriminatory effect under their test. Plaintiffs chose seven percent as a threshold in part because of their experts' opinion that a plan with such a large gap is "durable," meaning that the plan is likely to continue to give the majority party an advantage in subsequent elections under the plan, even if the minority party increases its vote share. Plts.' PFOF बी $12,85-93$, 114-18, 154, 170, dkt. \#79.

Defendants argue that plaintiffs do not have a right to an efficiency gap that favors Democrats, Dfts.' Reply Br., dkt. \#73, at 23, but this appears to be a misinterpretation of plaintiffs' position. Plaintiffs are not saying that they have a right to regain control of the legislature. Rather, plaintiffs say that they picked a threshold that was durable in an attempt to answer the question raised repeatedly by the Supreme Court, which is how extreme the discriminatory effects of the gerrymander must be, or, in other words, "how much is too much." Vieth, 541 U.S. at 298-99 (plurality opinion). Justice Breyer echoed this view when he said that court intervention
should be limited to "the unjustified use of political factors to entrench a minority in power." Vieth, 541 U.S. at 360 (Breyer, J., dissenting) (emphasis added). However, other members of the Court want more specificity. In Vieth, 541 U.S. at 307-08, Justice Kennedy expressed the need "to define clear, manageable, and politically neutral standards for measuring the particular burden a given partisan classification imposes on representational rights." This is exactly what plaintiffs are attempting to do with the efficiency gap.

Focusing on durability makes some sense because it is an indication that ordinary political processes cannot fix the problem, so court intervention is needed. Reynolds v. Sims, 377 U.S. 533, 553-54 (1964) (in context of gerrymandering claim for population deviations, recognizing that " $[\mathrm{n}]$ o effective political remedy to obtain relief against the alleged malapportionment . . appears to have been available"); Vieth, 541 U.S. at 361 (Breyer, J., dissenting) ("Where unjustified entrenchment takes place, voters find it far more difficult to remove those responsible for a government they do not want; and these democratic values are dishonored."). Focusing specifically on the life span of the plan also makes obvious sense because the political landscape changes each time a new plan is enacted. Defendants do not challenge plaintiffs' view that durability is an appropriate measure of discriminatory effect, so we need not resolve that issue in this opinion. It is enough to say that a judgment in plaintiffs' favor

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would not give plaintiffs or anyone else a constitutional right to gain control of the legislature or to draw a plan that is biased in their favor.
b. Efficiency gap as a constitutional requirement for "hyper-proportional" representation

Defendants say that the efficiency gap is an inadequate measurement of a plan's partisan effect because it is "a measure of proportionality," Dfts.' Br., dkt. \#46, at 47, which the Supreme Court has said repeatedly is not required by the Constitution. E.g., LULAC, 548 U.S. at 419 ("[T]here is no constitutional requirement of proportional representation."); Bandemer, 478 U.S. at 132 (plurality opinion) ("[T]he mere lack of proportional representation will not be sufficient to prove unconstitutional discrimination."). Defendants seem to acknowledge that plaintiffs' test does not require proportional representation per se, in the sense that a party's seat share must be the same as that party's share of votes. Rather, defendants say that plaintiffs' standard requires what defendants call "hyper-proportionality." Dfts.' Br., dkt. \#46, at 48. This is because, under plaintiffs' "simplified method" for calculating the efficiency gap, the efficiency gap remains zero only if the party receiving more than 50 percent of the vote receives a 2 percent increase in its share of the seats for every 1 percent increase in its share of the votes. Plts.' PFOF - 136, dkt. \#79. For example, 51 percent of the votes would translate into 52 percent of the seats, 52 percent of the votes would translate into 54 percent of
the seats and 75 percent of the votes would translate to 100 percent of the seats. Perhaps "hypermajoritarianism" would be a more accurate name for defendants' objection because the formula suggests that a majority of voters should have an even larger majority of seats.

Defendants' argument is important, but it would be premature to conclude that precedent forecloses plaintiffs' claim because of this formula. For one thing, plaintiffs say that the ratio is not a normative requirement of their test; it is simply what happens when a districting plan treats the parties symmetrically. Plts.' PFOF © 145, dkt. \#79. This seems to be borne out by history, which shows that a 1 percent increase in vote share generally leads to a two percent increase in seat share. Plts.' PFOF $\mathbb{\|} \|$ 137-39, 146, dkt. \#79. See also LULAC, 548 U.S. at 464-65 (Stevens, J., concurring in part and dissenting in part) ("[O]ur electoral system tends to produce a 'seat bonus' in which a party that wins a majority of the vote generally wins an even larger majority of the seats.").

Further, plaintiffs' standard does not require a $2: 1$ ratio between seat share and vote share. The efficiency gap is only part of plaintiffs' test, so no claim can prevail simply because a districting plan produces a particular vote share to seat share ratio. Even without considering the other elements of the standard, the $2: 1$ ratio appears in plaintiffs' formula only when the efficiency gap is zero. Plts.' PFOF $\mathbb{I}$

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136, dkt.\#79. Because plaintiffs' standard allows for a significant deviation from a zero efficiency gap, it also allows for a significant deviation from the $2: 1$ ratio. Id. at $\mathbb{\|} 148$.

Perhaps defendants mean to make a more subtle point, which is that the efficiency gap is an improper measure simply because it treats a particular vote share to seat share ratio as the "ideal" result. Again, however, the "ideal" result proposed by plaintiffs is the situation in which no voter has an unfair advantage over another in obtaining representation by the party of his or her choice. Defendants have not cited any authority that forecloses plaintiffs' view, but both parties should be prepared to present evidence on this point at trial.

Further, it is likely that any objective standard for measuring partisan gerrymandering will have so me connection to the basic principle that the collective will of the people should not be subverted indefinitely by an entrenched minority, a principle long recognized by the Supreme Court. Reynolds, 377 U.S. at 565 ("[L]egislatures . . . should be bodies which are collectively responsive to the popular will."). As the plurality in Bandemer recognized, "a preference for a level of parity between votes and representation sufficient to ensure that significant minority voices are heard and that majorities are not consigned to minority status, is hardly an illegitimate extrapolation fromour general majoritarian ethic and the objective of fair and adequate representation
recognized in Reynolds." Bandemer, 478 U.S. at 126 n.6. Opinions by other Justices reflect the same basic understanding. LULAC, 548 U.S. at 419 (opinion of Kennedy, J.) ("[A] congressional plan that more closely reflects the distribution of state party power seems a less likely vehicle for partisan discrimination than one that entrenches an electoral minority."); LULAC, 548 U.S. at 467-68 (Stevens, J., concurring in part and dissenting in part) (districting plan is presumptively unconstitutional ifequal share of votes for two parties produces large seat differential because in that case the plan "imposes ... a significant disadvantage on a politically salient group of voters"); Vieth, 541 U.S. at 352 n. 7 (Souter, J., dissenting) ("[T]he Constitution guarantees no right to proportional representation... It does not follow that the Constitution permits every state action intended to achieve any extreme form of disproportionate representation."); Vieth, 541 U.S. at 360-61 (Breyer, J., dissenting) (gerrymandering that allows "a party that enjoys only minority support among the populace... to take, and hold, legislative power . . . violates basic democratic norms").

Perhaps at trial it will become clear that the efficiency gap cannot be reconciled with Supreme Court precedent. At this stage, however, we are not persuaded that defendants have made that showing.
c. Potential breadth of plaintiffs' standard

Defendants say that plaintiffs' test is not "limited and precise" as Justice Kennedy suggested it should be in Vieth, 541 U.S. at 306 (Kennedy, J., concurring in the judgment), because such a large number of state districting plans across the country have an efficiency gap of at least seven percent. According to plaintiffs' own experts, approximately 20 to 25 percent of plans adopted by a party with unified control of the state government (both houses and the governorship) have an initial efficiency gap of seven percent or more. Plts.' PFOF 【 โ 99,74, dkt.\#79. (The parties agree that unified control of the government generally leads to an attempt to manipulate districts for partisan gain, though plaintiffs point to examples in which that is not the case. Id. at $9 \mathbb{T} 76$ and 172 (citing plans enacted under unified party control in California, Maine and Vermont that did not lead to partisan gerrymanders).)

As an initial matter, plaintiffs say that the 20 to 25 percent figure is inflated because it does not take into consideration plans that can be justified with neutral reasons. More generally, plaintiffs say that, to the extent there is a large number of suspect plans, that is not evidence of a weakness of their test, but evidence that "the practice of partisan gerrymandering is ubiquitous and very severe." Trans., dkt. \#89, at 76. They also argue that federal courts invalidated many districting plans after recognizing other types of gerrymandering claims, so the potential effect of plaintiffs' proposed standard on current districting plans should not be a reason to
reject the standard. Plts.' PFOF \|\| 77-78, dkt. \#79 (citing Gary W. Cox \& Jonathan N. Katz, Elbridge Gerry's Salamander (2002), and Ellen D. Katz et al., Documenting Discrimination in Voting: Judicial Findings under Section 2 of the Voting Rights Act, 39 U. Mich. J.L. Reform 643, 655 (2006)).

Of course, plaintiffs are correct that courts cannot decline their duty to enforce the Constitution simply because a ruling may have far reaching effects. However, we agree with defendants that the usefulness of the efficiency gap as a tool for measuring partisan effect may be lessened if a large efficiency gap is a common feature of districting plans. A theme in a number of opinions by Supreme Court Justices is that court intervention in partisan gerrymandering cases should be limited to rare and extreme circumstances. E.g., Bandemer, 478 U.S. at 133 (plurality opinion) (raising concern that "a low threshold for legal action would invite attack on all or almost all reapportionment statutes"); Vieth, 541 U.S. at 339 (Stevens, J., dissenting) (proposing what he described as "a narrow test [that] would cover only a few meritorious claims, but . . . would preclude extreme abuses"); Vieth, 541 U.S. at 354 (Souter, J., dissenting) (courts should be able to "identify at least the worst cases of gerrymandering"); Vieth, 541 U.S. at 362 (Breyer, J., dissenting) ("Courts need not intervene often to prevent the kind of abuse I have described."). This view could be undermined if we were to adopt a standard that rendered suspect a large swath of districting plans around the country.

Again, however, this objection is not a ground for granting summary judgment. As discussed above, the extent to which Wisconsin's and other states' efficiency gaps are caused by partisan bias is a disputed fact. If the facts at trial show that Wisconsin's efficiency gap is caused by neutral factors, then it will not be necessary to determine the potential implications of a ruling in plaintiffs' favor.

Further, this seems to be another objection that relates less to the validity of the efficiency gap as a general matter and more to the choice of how large an efficiency gap must be to sustain a constitutional claim. If plaintiffs' proposed formulation is not sufficiently demanding, this may support raising the threshold necessary to support a claim. Another possibility would be to incorporate into plaintiffs' prima facie case a requirement to show that any large efficiency gap cannot be justified by legitimate interests, possibilities the panel has not foreclosed.

Even if this court were to grant relief to plaintiffs, it might not be necessary to establish a threshold in this case. As plaintiffs point out, in the equal apportionment cases, the Supreme Court did not determine at first how large a population deviation must be in order to trigger a presumption of unconstitutionality. Rather, the Court proceeded on a case by case basis, settling on ten percent as the threshold only after several years. Nicholas 0 . Stephanopoulos \& Eric M. McGhee, Partisan Gerrymandering and the Efficiency Gap, 82 U. Chi. L.

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Rev. 831, 890-91 (2015). Because plaintiffs allege in this case that the efficiency gap created by Act 43 is one of the largest in recent history, determining a threshold may be something that can wait for another day. Bandemer, 478 U.S. at 123 ("arithmetic presumption" not necessary to adjudicate partisan gerrymandering claim).

## 2. Intent element

[^5]Because defendants did not raise this issue until their reply brief, we are not required to consider it. Narducci v. Moore, 572 F.3d 313, 324 (7th Cir. 2009) ("[T] he district court is entitled to find that an argument raised for the first time in a reply brief is forfeited."). However, we will discuss some of the potential issues raised by this element to provide guidance at trial.

In attempting to craft an intent element in a partisan gerrymandering case, a litigant or a court must navigate the minefield of Supreme Court precedent on this issue. In Vieth, 541 U.S. at 284, the plurality rejected a partisan gerrymandering standard that required a showing that the defendants acted with "a predominant intent to achieve partisan advantage." In LULAC, 548 U.S. at 418, the Court rejected a standard that required a showing that partisan gain was the "sole motive" for the map's design.

Perhaps cognizant of the Court's skepticism of heightened intent requirements, plaintiffs went back to Bandemer for their intent element. In that case, the plurality required only a showing of an intent to discriminate against an identifiable political group. Bandemer, 478 U.S. at 127 (plurality opinion). The plurality declined to adopt a more demanding intent requirement, even though it acknowledged that, "[a]s long as redistricting is done by a legislature, it should not be very difficult to prove that the likely political consequences of the reapportionment were intended."

Id. at 128-29. In other words, the assumption is that members of a particular party generally will try to benefit themselves and hurt their adversaries.

In their opening brief, defendants seem to agree with the view that a heightened intent requirement would be inconsistent with Supreme Court precedent. Dfts.' Br., dkt. \#46, at 41 ("If the intent element calls for a more searching inquiry, then the standard fails under Vieth" because "[t]he Vieth plurality and Justice Kennedy both rejected a standard that incorporated a 'predominant intent' standard."). Further, defendants did not directly criticize the intent requirement in Bandemer anywhere in their briefs or during oral argument. However, in their reply brief, defendants seem to suggest that a heightened intent element is required by Vieth. Thus, defendants' position now seems to be that there is no viable intent element for a partisan gerrymandering claim. Defendants reiterated that position during oral argument. When asked by the court what defendants believed the intent requirement should be, counsel stated, "I'm not sure that this is something that can be solved." Trans., dkt. \#89, at 7.

As discussed above, a majority of the Supreme Court has directed litigants and lower courts to continue searching for an appropriate standard for deciding partisan gerrymandering claims. In light of that directive, it would be inappropriate to interpret prior case law as rejecting all formulations of the intent requirement for those claims.

During oral argument, plaintiffs' counsel stated her view that the Bandemer holding regarding intent remains controlling precedent, even after Veith and LULAC. Trans., dkt. \#89, at 44. That view may be debatable, but the parties have not fully addressed that issue, so we believe that it would be premature to decide it now.

At least one Justice has questioned the constitutionality of any districting plan that disadvantages members of a particular party. Vieth, 541 U.S. at 324 (Stevens, J., dissenting) ("[T]he plurality errs in assuming that politics is 'an ordinary and lawful motive.' We have squarely rejected the notion that a 'purpose to discriminate on the basis of politics' is never subject to strict scrutiny.") (citation omitted). However, a majority of the Justices in Vieth appeared to accept the view that "[a] determination that a gerrymander violates the law must rest on something more than the conclusion that political classifications were applied," id. at 307 (Kennedy, J., concurring in the judgment). The plurality in Vieth identified the "excessive injection of politics" as the basis for a constitutional violation. Id. at 293 (plurality opinion). Justice Kennedy however, was more circumspect; he noted that "[e]xcessiveness is not easily determined." Id. at 316 (Kennedy, J., concurring in the judgment), and suggested focusing on evidence that a legislature's plan is unrelated to neutral districting criteria. Id. at 312-13.

At oral argument, other alternative formulations of intent emerged. One suggestion was that plaintiffs show that defendants had the intent to prevent the minority party from regaining control throughout the life of the districting plan. Trans., dkt \#89, at 5-6.

Accordingly, it would be inappropriate to grant summary judgment on this ground. That being said, plaintiffs will have the burden at trial to prove that defendants acted with discriminatory intent, so they should be prepared to present the strongest evidence that they have on this issue-including comparative evidence of prior redistricting plans in the State of Wisconsin-in order to meet even the most demanding intent requirement. Specifically, the parties should be prepared to address the evidence bearing on intent in light of the Justices' concerns in Vieth, the discussion with this court at argument, and the parties own formulations on that element.

## 3. Burden shifting

If plaintiffs prove discriminatory intent and effect in steps one and two, plaintiffs' proposed test then shifts the burden to defendants to show that the large efficiency gap was "unavoidable" in light of the state's political geography and legitimate districting objectives. In their opening brief, defendants' primary objection to this portion of plaintiffs' test was really another objection to the efficiency gap. In particular, defendants argued that it was "fundamentally unfair" to shift the burden to
defendants because the efficiency gap was not an adequate measure of discriminatory effect. Dfts.' Br., dkt. \#46, at 42. Because this is simply a repackaging of arguments that we have said we cannot resolve on a motion for summary judgment, it is unnecessary to consider this issue further.

In their reply brief, defendants argue that plaintiffs' standard is unfair because it will be impossible to show that a particular efficiency gap was "unavoidable." Rather, with the near-infinite number of ways to draw a plan, there will always be a way to "reverse-engineer a plan that has a better political result for one side while coming close in population deviation, compactness and municipal splits." Dfts.' Reply Br., dkt. \#73, at 10. At oral argument, plaintiffs addressed this objection by stating that defendants would not "have to show that these particular district lines were absolutely necessary." Trans., dkt. \#89, at 60. Rather, defendants would have to show that any alternative plan would have "roughly the same kind of excessive . . . efficiency gap." Id. We understand plaintiffs to mean that defendants would retain some flexibility in choosing how to draw district lines.

When asked at oral argument whether anyone on the Supreme Court had proposed a similar burdenshifting scheme as part of a partisan gerrymandering claim, plaintiffs' counsel's initial response was that no one had. Id. at 58. Instead, counsel stated that plaintiffs had adapted the burden-shifting portion of
their standard from cases involving equal apportionment, such as Voinovich v. Quilter, 507 U.S. 146 (1993), Brown, 462 U.S. 835, and Connor v. Finch, 431 U.S. 407, 414 (1977). Id. at 63. However, later in the argument, plaintiffs' counsel stated that a similar burden-shifting standard could be found in the partisan gerrymandering context in the plurality's opinion in Bandemer, in Justice Stevens's opinion in Karcher v. Daggett, 462 U.S. 725 (1983), and in Justice Souter's opinion in Vieth. Trans., dkt. \#89, at 65.

The cases plaintiffs cite may support an argument that some type of burden-shifting is appropriate, but they do not support plaintiffs' view that defendants must show that their plan was "unavoidable." In Bandemer, 478 U.S. at 127-43, the plurality focused most of its opinion on the issue of discriminatory effect. Because the plurality found that the plaintiffs had not met their burden on that element, it did not have to go any further. However, in responding to Justice Powell's dissenting opinion, the plurality stated that the various factors he proposed in his test "might well be relevant to an equal protection claim." Bandemer, 478 U.S. at 141. The plurality elaborated, "[t]he equal protection argument would proceed along the following lines: If there were a discriminatory effect and a discriminatory intent, then the legislation would be examined for valid underpinnings." Id. However, because the plurality "found that there was insufficient discriminatory effect to constitute an equal protection violation," it "did not reach the
question of the state interests (legitimate or otherwise) served by the particular districts as they were created by the legislature." Id. at 141-42. Thus, although the plurality suggested that it would consider the state's interests as part of any test, the plurality did not specify which party should shoulder the burden on that issue.

Plaintiffs are correct that Justice Stevens and Justice Souter both proposed a burden-shifting standard for partisan gerrymandering claims, but neither of them proposed placing a burden on defendants as demanding as the one plaintiffs propose. In Karcher, 462 U.S. at 751, Justice Stevens stated that he "would consider whether the plan has a significant adverse impact on an identifiable political group, whether the plan has objective indicia of irregularity, and then, whether the State is able to produce convincing evidence that the plan nevertheless serves neutral, legitimate interests of the community as a whole." Under Justice Souter's standard, after the plaintiffs met their prima facie case, Justice Souter "would then shift the burden to the defendants to justify their decision by reference to objectives other than naked partisan advantage." Vieth, 541 U.S. at 351 (Souter J., dissenting).

Neither Justice suggested that the defendants should be required to show that a plan was "unavoidable" in light of traditional districting criteria. In fact, under Justice Souter's test, the plaintiffs would have to show as part of their prima
facie case both that the legislature "paid little or no heed to those traditional districting principles whose disregard can be shown straightforwardly" and that the legislature could have drawn a fairer plan that "deviated less from traditional districting principles." Vieth, 541 U.S. at 348 (Souter, J., dissenting). See also LULAC, 548 U.S. at 491 (Breyer, J., concurring in part and dissenting in part) (including evidence of "a radical departure from traditional boundarydrawing criteria" as part of plaintiffs' prima facie case).

The equal apportionment cases plaintiffs cite are similar. After a plaintiff challenging population disparities in state legislative districts establishes her prima facie case, the burden shifts to the defendants to show that their plan is "justified." Brown, 462 U.S. at 843 . In particular, the question is "whether the legislature's plan 'may reasonably be said to advance [a] rational state policy." Id. (quoting Mahan v. Howell, 410 U.S. 315, 328 (1973)). Again, there is no requirement to show that the plan was "unavoidable."

Plaintiffs' proposed standard seems to be most similar to the one that applies to equal apportionment requirements in congressional redistricting. In those cases, if the plaintiffs meet their prima facie case, "the burden shifts to the State to 'show with some specificity' that the population differences 'were necessary to achieve some legitimate state objective." Tennant v. Jefferson County Commission, 133 S. Ct.

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3, 5 (2012) (quoting Karcher, 462 U.S. at 740-41) (emphasis added). However, the Supreme Court has expressly declined to adopt the "necessity" standard in the context of state legislative districting because the two types of challenges are governed by different constitutional provisions, Article I, § 2 (with respect to congressional districts) and the equal protection clause (with respect to state legislative districts). Mahan, 410 U.S. at 321 ("[M]ore flexibility [i]s constitutionally permissible with respect to state legislative reapportionment than in congressional redistricting."). Because plaintiffs in this case are relying on the equal protection clause rather than Article I, § 2, the more lenient standard is more instructive.

Further, even with respect to congressional districts, the plaintiffs are required to show as part of their prima facie case that "the population differences among districts could have been reduced or eliminated altogether by a good-faith effort to draw districts of equal population." Karcher, 462 U.S. at 730. Under their proposed test, plaintiffs have no burden to show that defendants could have drafted a better plan.

In sum, we believe that plaintiffs have overstated defendants' burden in part three of their proposed test. However, this conclusion does not require summary judgment in defendants' favor. As noted above, defendants have made no effort to justify the plan using neutral criteria. Thus, to the extent that

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defendants have any burden to prove the legitimacy of the plan, this element must be resolved at trial. Further, to the extent that plaintiffs have an initial burden to show that defendants' plan cannot be justified using neutral criteria, we believe that plaintiffs have met that burden for the purpose of defendants' motion for summary judgment by drafting a plan with a dramatically lower efficiency gap while still satisfying neutral criteria.

Again, because the parties have not fully briefed the question of how this element should be formulated, it would be premature to answer the question in this order. At trial, both sides should be prepared to submit whatever evidence they have to show whether Act 43 can be justified by neutral criteria.

## ORDER

## IT IS ORDERED that

1. The motion for summary judgment filed by defendants Gerald C. Nichol, Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas, Timothy Vocke and Kevin J. Kennedy, dkt. \#45, is DENIED.
2. The motion filed by plaintiffs William Whitford, Roger Anclam, Emily Bunting, Mary Lynne Donohue, Helen Harris, Wayne Jensen, Wendy Sue Johnson, Janet Mitchell, James Seaton, Allison Seaton, Jerome Wallace and Don Winter to exclude the opinions of Sean Trende, dkt. \#70, is DENIED WITHOUT PREJUDICE to plaintiffs' refilingit at the conclusion of trial.
3. Trial will begin on Tuesday, May, 24, 2016 and should be completed by Friday, May 27, 2016. If the parties believe that is not a sufficient amount of time, they should explain their concerns in writing no later than April 18, 2016.

Entered this 7th day of April, 2016.

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## BY THE COURT:

/s/
KENNETH F. RIPPLE
Circuit Judge
/s/

BARBARA B. CRABB
District Judge
/s/

WILLIAM C.
GRIESBACH
District Judge

## Cross-Reference to Supplemental Appendix

Amended Rebuttal Report: Response to Expert Reports of Sean Trende and Nicholas Goedert, Dr. Kenneth R. Mayer (ECF No. 95) appears at: SA282314

# Joint Final Pre-Trial Report <br> IN THE UNITED STATES DISTRICT COURT FOR THE WESTERN DISTRICT OF WISCONSIN 

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WILLIAM WHITFORD, )
ROGER ANCLAM, EMILY )
BUNTING, MARY LYNNE )
DONOHUE, HELEN ) No. 15-cv-421-bbc
HARRIS, WAYNE JENSEN, )
WENDY SUE JOHNSON, )
JANET MITCHELL, )
ALLISON SEATON, JAMES )
SEATON, JEROME )
WALLACE, and DONALD )
WINTER,
Plaintiffs,
v.
GERALD C. NICHOL, THOMAS BARLAND, JOHN ) FRANKE, HAROLD V. ) FROEHLICH, KEVINJ. ) KENNEDY, ELSA ) LAMELAS, and TIMOTHY ) VOCKE,
Defendants.
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## JOINT FINAL PRETRIAL REPORT

This action for declaratory relief challenges 2011 Wisconsin Act 43, which adopted new boundaries for the state's legislative districts, and codified them in Chapter 4 of the Wisconsin Statutes. The case is scheduled for trial commencing Tuesday, May 24, 2016 and is expected to last four days. In accordance with the Court's October 15, 2015 Scheduling Order (Dkt. 33) and Civil L.R. 16(c)(1), the parties, through their respective counsel, submit the following pretrial report.

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## JOINT STATEMENT OF STIPULATED FACTS

## Plaintiffs

1. Plaintiffs are qualified, registered voters in the State of Wisconsin, who reside in various counties and legislative districts.
2. Plaintiffs are all supporters of the Democratic party and of Democratic candidates, and they almost always vote for Democratic candidates in Wisconsin elections.
3. Plaintiff William Whitford, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 76th Assembly District in Madison, in Dane County, Wisconsin.
4. Plaintiff Roger Anclam, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 31st Assembly District in Beloit, in Rock County, Wisconsin.
5. Plaintiff Emily Bunting, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 49th Assembly District in Viola, Richland County, Wisconsin.
6. Plaintiff Mary Lynne Donohue, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 26th Assembly District in Sheboygan, in Sheboygan County, Wisconsin.
7. Plaintiff Helen Harris, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 22nd Assembly District in Milwaukee, in Milwaukee County, Wisconsin.
8. Plaintiff Wayne Jensen, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 63rd Assembly District in Rochester, in Racine County, Wisconsin.
9. Plaintiff Wendy Sue Johnson, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 91st Assembly District in Eau Claire, in Eau Claire County, Wisconsin.
10. Plaintiff Janet Mitchell, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 66th Assembly District in Racine, in Racine County, Wisconsin.
11. Plaintiffs James and Allison Seaton, citizens of the United States and of the State of Wisconsin, are residents and registered voters in the

42nd Assembly District in Lodi, in Columbia County, Wisconsin.
12. Plaintiff Jerome Wallace, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 23rd Assembly District, in Fox Point, in Milwaukee County, Wisconsin.
13. Plaintiff Don Winter, a citizen of the United States and of the State of Wisconsin, is a resident and registered voter in the 55th Assembly District in Neenah, in Winnebago County, Wisconsin.

## Defendants

14. Defendant Gerald C. Nichol is the Chair of the Wisconsin Government Accountability Board ("G.A.B."), and is named solely in his official capacity as such. The G.A.B. is a state agency under Wis. Stat. § 15.60, which has "general authority" over and "responsibility for the administration of . . . [the State's] laws relating to elections and election campaigns," Wis. Stat. § 5.05(1), including the election every two years of Wisconsin's representatives in the Assembly.
15. Defendants Thomas Barland, John Franke, Harold V. Froehlich, Elsa Lamelas, and Timothy Vocke are all members of the G.A.B., and are named solely in their respective official capacities as such.
16. Defendant Kevin J. Kennedy is the Director and General Counsel of the G.A.B., and is named solely in his official capacity as such.

## The Redistricting Process in 2011

17. In 2011, Adam Foltz was a legislative aide to the Republican then-Speaker of the Wisconsin Assembly.
18. In 2011, Tad Ottman was a legislative aide to Republican Majority Leader of the Wisconsin Senate.
19. In 2011, Adam Foltz and Tad Ottman worked with consultants, including Joseph Handrick and Professor Keith Gaddie, as well as others, to develop a redistricting plan for Wisconsin's legislative districts.
20. In January 2011, Scott Fitzgerald, Republican member of the Wisconsin State Senate and Wisconsin Senate Majority Leader, and Jeff Fitzgerald, Republican member of the Wisconsin State Assembly and Speaker of the Assembly, hired attorney Eric McLeod ("McLeod") and the law firm of Michael Best to represent the entire Wisconsin State Senate and Wisconsin State Assembly in connection with the reapportionment of the state legislative districts after the 2010 Census.

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21. On January 3, 2011, the Committee on Senate Organization approved the following motion with all three Republican members of the Committee (Senator Scott Fitzgerald, Senator Michael Ellis, and Senator Glenn Grothman) voting "Aye" and the single Democrat member (Senator Mark Miller) voting "No":
[MOTION] To authorize the hiring of the law firms of Michael Best \& Friedrich, LLP and Troupis Law Office, LLC for services related to redistricting of legislative and congressional districts for the 2012 elections. The law firms shall perform work at the direction of the Majority Leader. This authorization includes the authority to provide the law firms with any redistricting software applications procured or developed by the Legislature that are necessary to facilitate participation in the redistricting drafting process. Upon adoption of this motion, the retention of the law firm of O'Neil, Cannon, Hollman, DeJong, S.C. is terminated. The Chief Clerk may pay the law firm of O'Neil, Cannon, Hollman, DeJong, S.C. for services rendered through the date on which this ballot is adopted but not for services rendered on any date thereafter." [The Motion/Ballot was part of the record in Baldus (2:11-cv-00562-JPS-DPW-RMD, filed 12/16/11 Doc. 81-2) and is subject to judicial notice pursuant to FRE Rule 201(b)(2)].

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22. On January 4, 2011, the Assembly Organization Committee approved the following motion to:
"Authorize the Speaker of the Assembly, Jeff Fitzgerald, to retain legal counsel for the purpose of apportioning and redistricting the Legislative and Congressional Districts following the 2010 decennial Census as required by Article IV, Section 3 of the Wisconsin Constitution. Such counsel will be compensated under s. 20.765(1)(a)." [The Motion was part of the record in Baldus (2:11-cv-00562-JPS-DPW-RMD, filed 12/16/11 Doc. 81-3) and is subject to judicial notice pursuant to FRE Rule 201(b)(2)].
23. All redistricting work was done in Michael Best's office before the file (the redistricting plan that became Act 43) was sent to the Legislative Reference Bureau for drafting, and the "map room" where all redistricting work was done was located in Michael Best's office.
24. A formal written policy provided that only the Senate Majority Leader, the Speaker of the House, and their aides Tad Ottman and Adam Foltz, and Michael Best attorney Eric Mcleod and legal staff designated by Mr. McLeod, would have unlimited access to the "map room."
25. The access policy provided for limited access by rank and file legislators:
"Legislators will be allowed into the office [mapping room] for the sole purpose of looking at and discussing their district. They are only to be present when an All Access member is present. No statewide or regional printouts will be on display while they are present (with the exception of existing districts). They will be asked at each visit to sign an agreement that the meeting they are attending is confidential and they are not to discuss it." But only Republican legislators were allowed even this limited access.
26. Three computers were deployed by the Legislative Technology Services Bureau ("LTSB") to the "map room" at Michael Best \& Friedrich for use in drafting the redistricting plan. Each computer contained two mirrored internal hard drives and one external hard drive. On July 15, 2010, a computer coded for identification purposes as WRK32587 was deployed to Michael Best \& Friedrich for use by Tad Ottman. Computer WRK32587 was deployed with an external hard drive with the identification code of HDD32575. On June 4, 2012, computer WRK32587 was moved from Michael Best \& Friedrich to the legislative office of Senator Scott Fitzgerald in the Capitol Building. On May 21, 2015, the hard drives from computer WRK32587 and its external hard drive HDD32575 were shredded pursuant to the
established policy and procedures for disposal established by the LTSB. Ylvisaker Dep. (Dkt. 106), at 14:18-15:12, 23:7-26:17, 28:7-31:17; Ex. 49, Ex. 50 at 12 .
27. Also on July 15, 2010, a computer coded WRK32586 was deployed to Michael Best \& Friedrich for use by Adam Foltz. Computer WRK32586 was deployed with an external hard drive with the identification code of HDD32574. On September 13, 2012, computer WRK32586 was returned to the LTSB. On May 21, 2015, the hard drives from computer WRK32586 and its external hard drive HDD32574 were shredded pursuant to the established policy and procedures for disposal established by the LTSB. Ylvisaker Dep. (Dkt. 106), at 14:18-15:12, 23:7-26:17, 28:7-31:17; Ex. 49, Ex. 50 at 12 .
28. On March 21, 2011, a third computer coded WRK32864 was deployed to Michael Best \& Friedrich for use by Joseph Handrick. Computer WRK32864 was deployed with an external hard drive with the identification code of HDD32579. On June 4, 2012, computer WRK32864 was moved from Michael Best \& Friedrich to the legislative office of Senator Scott Fitzgerald in the Capitol Building. On May 21, 2015, the hard drives from computer WRK32864 and its external hard drive HDD32579 were shredded pursuant to the established policy and procedures for disposal established by the LTSB. Ylvisaker Dep.
(Dkt. 106), at 14:18-15:12, 23:7-26:17, 28:7-31:17; Ex. 49, Ex. 50 at 12.
29. In the course of drafting the redistricting plan enacted by Act 43 (the Current Plan) for Wisconsin's legislative districts, Adam Foltz, Tad Ottman, and Keith Gaddie examined the past partisan performance of voters in the existing legislative districts, as well as the expected future partisan performance of voters in various configurations of potential new districts.
30. Specifically, in the course of developing the Current Plan for Wisconsin's legislative districts, Adam Foltz, Tad Ottman, and Keith Gaddie examined whether past districts were likely to vote majority Republican or majority Democratic, and whether various configurations of potential new districts were likely to vote majority Republican or majority Democratic.
31. On April 11, 2011, Professor Ronald Keith Gaddie entered into a Consulting Services Agreement with Michael Best \& Friedrich. The agreement stated that Professor Gaddie was to serve as a consultant to Michael Best \& Friedrich in connection with its representation of the Wisconsin State Senate and the Wisconsin State Assembly on "matters relating to the reapportionment of the Wisconsin Senate, Assembly and Congressional Districts arising out of the 2010 census." The agreement described Professor Gaddie's "duties" as including "service as an independent

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advisor on the appropriate racial and/or political make-up of legislative and congressional districts in Wisconsin," and would include "providing advice based on certain statistical and demographic information and on election data or information." Additionally, the Consulting Services Agreement stated, "Any work papers or materials prepared by you, or under your direction, belong to the Senate pursuant to the Representation, and every page must be sealed or otherwise stamped "Attorney/Client Work-Product Privilege Confidential."
32. On April 17, 2011, Keith Gaddie drafted a note to himself while he was in Madison, Wisconsin, providing consulting services for the development of a redistricting plan. The document stated in full:
"The measure of partisanship should exist to establish the change in the partisan balance of the district. We are not in court this time; we do not need to show that we have created a fair, balanced, or even a reactive map. But, we do need to show to lawmakers the political potential of the district.

I have gone through the electoral data for state office and built a partisan score for the assembly districts. It is based on a regression analysis of the Assembly vote from 2006, 2008, and 2010, and it is based on prior election indicators of future election performance.

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I am also building a series of visual aides to demonstrate the partisan structure of Wisconsin politics. The graphs will communicate the top-to-bottom party basis of the state politics. It is evident, from the recent Supreme Court race and also the Milwaukee County executive contest, that the partisanship of Wisconsin is invading the ostensibly nonpartisan races on the ballot this year." Gaddie Dep. (Dkt. 108), at 95: 6-96:2.
33. On March 9, 2016, during his deposition, Keith Gaddie was asked the following question:

> "Q: You said something to the effect that is important to understand the partisan effect. Why is it important to understand the partisan effect?"

Professor Gaddie responded to that question:
"A: Well, again, I was writing as a political scientist. If you're going to redistrict it's important to understand the consequences of it. Lawmakers are going to be concerned about a variety of different consequences of a redistricting. The impact on their constituency, the impact on other constituencies.

If a lawmaker comes in and wants to know what you did to his district, it would be nice to be able to tell him we've got an estimate of what your district
used to look like in terms of partisanship and here's what it looks like now. So this kind of technique allows us to generate a measure that you can show to somebody and explain to them, this is what we think the net electoral impact is on your constituency.

In the aggregate, it means you can look at an entire map and ascertain the extent to which you have moved the partisan balance one way or the other."

Gaddie Dep. (Dkt. 108), at 98:24-99:24.
"Q: And you use the word "potential" there. What did you mean by the word potential?
A. If you had an election in the future, how might it turn out. So when I say potential, what I'm saying is that if we ran an election, this is our best estimate of what a non-incumbent election would look like given a particular set of circumstances, depending on whether one party is stronger or weaker.
Q. And that's what your regression model was designed to do, to show that potential of the district?
A. Yeah, it was designed to tease out a potential estimated vote for the legislator in the district and then allow you to also look at that and say, okay, what if the Democrats have a good year? What if
the Republicans have a good year? How does it shift? Okay?

The other thing is we know that districts don't correspond precisely to our statistical models all the time. So we're not concerned just with the crafting of the district or a point estimate of the vote. It's only an estimate. There's error. Right? There's going to be a range within which the outcome might occur.

The idea was to give to those people that were mapping, those people that were making choices, as much knowledge as we could glean about each district by giving them the most leverage on the least amount of data." Gaddie Dep. (Dkt. 108) at 100:22-102:3.
34. On March 9, 2016, during his deposition, Keith Gaddie was asked the following question:
"Q: But a significant part of your work that you were retained to do and that you did perform in 2011 had to do with the - with building a regression model to be able to test the partisan makeup and performance of districts as they might be configured in different ways, correct?"

Professor Gaddie responded to that question:
A. "Yes, that's correct."

Gaddie Dep. (Dkt. 108) at 46:12-19.
35. Professor Gaddie identified two measures to estimate the partisan change that would occur due to redistricting:
"There are basically two ways you can measure or you can estimate a partisan change when you redistrict. One is to use what's called a reconstituted election technique where we take either one or an index with several statewide elections, exongenous elections, which are elections that occur outside a district. Right? Higher levels of office. And we attempt to get a sense of a partisan average from that.

Or what you can do is you can take the actual election results, okay, the actual outcomes of previous elections, you turn those into a dependent variable, an outcome of interest, and then you regress using linear regression those results on these larger statewide measures.

The other thing you do is you attempt to take into account whether or not there's an incumbent running so that you can account for the incumbency impact. Again, it's been four years since I did this. But what we did is I had proposed to the map drawers that if they wanted to present a best estimate of partisan impact so the lawmakers can understand the consequence of different maps, that a regressions driven technique is the best
approach. So I set about building a regression equation using data that should have been produced to generate estimates of partisanship, partisan behavior in those districts for different district proposals.

So what this - what this spreadsheet is, is the consequence of applying one of those models. If it is what I think it is, it's the consequence of applying one of those models to a map generated by a map maker where what we know is, we know the statewide election results, and we then put those data for each district into the regression equation and that gives us an estimated vote value for each district. And that's what reported here, assuming no incumbent.

Gaddie Dep. (Dkt. 108) at 43:16-45:8.
36. "joe base map numbers.xlsx" is a document saved on the disc, Amended Lanterman Decl., Ex. B (Dkt. 97-2), and located in the "WRK32864 Responsive Spreadsheets Deduplicated file," and is a true and correct copy of a spreadsheet found by Mark Lanterman on the computer deployed to Michael Best \& Friedrich for use by Joseph Handrick. Amended Lanterman Decl., Ex. B (Dkt. 97-2).
37. The metadata for "joe base map numbers" is shown here:

JA168

| File Name | joe base map numbers.xlsx |
| :---: | :---: |
| Extension | xlsx |
| Created (Central) | $\begin{aligned} & \text { 4/11/2011 5:09:21 PM (2011-04- } \\ & 11 \text { 22:09:21 UTC) } \end{aligned}$ |
| Accessed (Central) | $\begin{aligned} & \text { 5/12/2011 7:06:05 PM (2011-05- } \\ & 13 \text { 00:06:05 UTC) } \end{aligned}$ |
| Modified (Central) | 5/12/2011 7:06:05 PM (2011-0513 00:06:05 UTC) |
| File Path | /Users/tad/Documents/joe base map numbers.xlsx |
| File Size | 22.91 KB |
| Author | tad |
| Last Saved By | tad |
| Office Created Date | $\begin{aligned} & \text { 4/11/2011 4:35:26 PM (2011-04- } \\ & 11 \text { 21:35:26 UTC) } \end{aligned}$ |
| Office Last Printed Date | $\begin{aligned} & \text { 5/12/2011 7:04:21 PM (2011-05- } \\ & 13 \text { 00:04:21 UTC) } \\ & \hline \end{aligned}$ |
| Office Last Saved Date | 5/12/2011 7:06:05 PM (2011-05- 13 00:06:05 UTC) |
| Hidden Columns or Rows | FALSE |
| Track Changes | FALSE |
| MD5 Hash Value | $9697 f 259 c b 6 d e 2 e 7 e 838 a 4 d e 973 f 2$ 481 |

Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32684 Responsive Spreadsheets File Detail Report."
38. The "joe base map numbers" spreadsheet lists district-by-district partisanship scores developed by Handrick, Foltz, and Ottman. Gaddie Dep. (Dkt. 108) at 40:12-24, 223:7-12.
39. The "joe base map numbers" spreadsheet lists district-by-district partisan scores for three Assembly district plans: the "current map," "basemap BASIC," and "basemap assertive." Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32864 Responsive Spreadsheets Deduplicated file."
40. "TADOTTMANSUPPPROD000094" is a true and correct copy of a spreadsheet created by Tad Ottman in 2011 and produced to the Court as part of the Legislature's supplemental production in Baldus v. Brennan (2:11-cv-00562-JPS-DPW-RMD; dated January 10, 2012).
41. "TADOTTMANSUPPPROD000094" lists district-by-district partisan scores developed by Handrick, Foltz, and Ottman. Gaddie Dep. (Dkt. 108) at 40:12-24, 223:7-12.
42. "TADOTTMANSUPPPROD000097" is a true and correct copy of a spreadsheet created by Tad Ottman in 2011 and produced to the Court as part of the Legislature's supplemental production in Baldus v. Brennan (2:11-cv-00562-JPS-DPW-RMD; dated January 10, 2012).
43. "TADOTTMANSUPPPROD000097" lists district-by-district partisan scores developed by Handrick, Foltz, and Ottman. Gaddie Dep. (Dkt. 108) at 40:12-24, 223:7-12.
44. "Plancomparisons.xlsm," a document saved on the disc, Amended Lanterman Decl., Ex. B (Dkt. 97-2), and located in the WRK32864 Responsive Spreadsheets Deduplicated file, is a true and correct copy of a spreadsheet found by Mark Lanterman on the computer deployed to Michael Best \& Friedrich for use by Joseph Handrick.
45. The metadata for "PlanComparisons" is shown here:

| File Name | PlanComparisons.xlsm |
| :---: | :---: |
| Extension | xlsm |
| Created (Central) | $\begin{aligned} & \text { 5/13/2011 12:58:51 PM (2011- } \\ & \text { 05-13 17:58:51 UTC) } \end{aligned}$ |
| Accessed (Central) | $\begin{aligned} & \text { 7/14/2011 1:32:51 PM (2011-07-14 } \\ & \text { 18:32:51 UTC) } \end{aligned}$ |
| Modified (Central) | $\begin{aligned} & \text { 7/14/2011 1:32:51 PM (2011-07-14 } \\ & \text { 18:32:51 UTC) } \end{aligned}$ |
| File Path | /Users/tad/Desktop/PlanComparisons .xlsm |
| File Size | 69.10 KB |
| Author | afoltz |
| Last Saved By | tad |

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| Office <br> Created <br> Date | $\begin{aligned} & \text { 5/2/2011 } 6: 13: 18 \text { PM (2011-05-02 } \\ & \text { 23:13:18 UTC) } \\ & \hline \end{aligned}$ |
| :---: | :---: |
| Office Last Printed Date | $\begin{aligned} & \text { 6/15/2011 3:28:17 PM (2011-06-15 } \\ & \text { 20:28:17 UTC) } \end{aligned}$ |
| Office Last Saved Date | $\begin{aligned} & \text { 7/14/2011 1:32:51 PM (2011-07-14 } \\ & \text { 18:32:51 UTC) } \end{aligned}$ |
| Hidden Columns or Rows | FALSE |
| Track <br> Changes | FALSE |
| MD5 Hash Value | 8d0b9118f01010be5b553b0306e6003 7 |

Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32684 Responsive Spreadsheets File Detail Report."
46. The "PlanComparisons" spreadsheet lists district-by-district partisan scores developed by Handrick, Foltz, and Ottman. Gaddie Dep. (Dkt. 108) at 40:12-24, 223:7-12.
47. The "PlanComparisons" spreadsheet lists district-by-district partisan proxy scores for four Assembly district plans: each tab includes an
identical column for a "Current" plan, and there are three tabs labeled as "Joe Aggressive," "Joe Aggressive (2)," and "TeamMap 6-15-11." Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32864 Responsive Spreadsheets Deduplicated file." Gaddie Dep. (Dkt. 108) at 215:22-217-20.
48. A spreadsheet labeled "Final Map" is a true and correct copy of a spreadsheet created by Adam Foltz. Gaddie Dep. (Dkt. 108), Ex. 39 at 3; Foltz. Dep. (Dkt 109) at 128:14-16.
49. The metadata associated with the "Final Map" is written on Exhibit 39, as follows:
"Plan Comparisons.xlsm"
created 5/9/11 5:39 PM
accessed 4/27/12 4:50 PM
modified 4/27/12 4:50 PM
file path:
/users/afoltz/Desktop/projects/PlanComparisons.x lsm

Gaddie Dep. (Dkt. 108), Ex. 39 at 1; Amended Lanterman Decl., Ex. B (Dkt. 97-2).
50. The "Final Map" spreadsheet lists district-by-district partisan scores developed by Handrick,

Foltz, and Ottman. Gaddie Dep. (Dkt. 108) at 40:1224, 223:7-12.
51. The spreadsheets shown in "joe base map numbers,"
"PlanComparisons,"
TADOTTMANSUPPPROD000094,"
"TADOTTMANSUPPPROD000097," and "Final Map" all include district-by-district partisan scores for both the "current map" and a different version of a potential future plan. Gaddie Dep. (Dkt. 108) at 220:25-221:13.
52. The "current map" referred to in "joe base map numbers," "PlanComparisons," "TADOTTMANSUPPPROD000094," "TADOTTMANSUPPPROD000097," and "Final Map," denotes the existing map, the maps as constituted in the State of Wisconsin before the 2012 re- map. Gaddie Dep. (Dkt. 108) at 234:22-24.
53. The district-by-district partisan scores for the "Current map" column in "joe base map numbers," and the "Current" column for the Assembly in "PlanComparisons,"
"TADOTTMANSUPPPROD000094," "TADOTTMANSUPPPROD000097," and "Final Map" are identical for all 99 districts.
54. "joe base map" is a document saved on the disc, Amended Lanterman Decl., Ex. B (Dkt. 97-2), and located in the WRK32864 Responsive Spreadsheets Deduplicated file, and is a true and
correct copy of a spreadsheet found by Mark Lanterman on the computer deployed to Michael Best \& Friedrich for use by Joseph Handrick. Amended Lanterman Decl., Ex. B (Dkt. 97-2).
55. The district-by-district partisan scores for the "base map BASIC" columns (columns F and P) in "joe base map numbers" are identical to the district-by-district partisan scores listed in the column "ALL0410" (column AU) in "joe base map."
56. "Final Map" was "probably the final map," and at minimum, "it's a safe assumption that [the map is] very near the completion of the process." Foltz Dep. (Dkt. 113) at 140:6-11, referring to Gaddie Dep. (Dkt. 108), Ex 39 at 3.
57. Professor Gaddie produced "S-curves" for draft Assembly redistricting plans prepared by Adam Foltz, Tad Ottman, and Joe Handrick. Gaddie Dep. (Dkt. 108) at 126:2-10.
58. Professor Gaddie agreed "with Joe Handrick to provide these types of spreadsheets to Adam Foltz, to himself and Adam Foltz and Tad Ottman, for the legislature in the drafting process. So one thing we do, they would create a map, then there would be part - there's electoral history data attached to it. Those data were used to generate spreadsheets of this sort that indicated how a district would perform on a partisan measure under different scenarios." Gaddie Dep. (Dkt. 108) at 40:14-24.
59. S-curves show "based upon an expected statewide vote for one party of the other which seats are going to tend more Democratic shaded in blue, more Republican shaded in red. Light blue means that they're Democratic tending, but competitive. Orange means they're Republican tending but competitive." Gaddie Dep. (Dkt. 108) at 128:10-16.
60. S-curves show "as you move the value of the vote for one party either up or down, you can see the responsiveness of the districts and how they shift and the number of seats that come into play for one party or fall away." Gaddie Dep. (Dkt. 108) at 129:6-11.
61. S-curves provide "a visualization of both the distribution of partisanship in the districts and the sensitivity of individual districts to changes and partisan strength across the state, assuming that the entire state shifts in the same direction one way or the other." Gaddie Dep. (Dkt. 108) at 129:12-18.
62. "Composite_Current_Curve.xlsx" is located in the WRK32586 Responsive Spreadsheets Deduplicated file, and is a true and correct copy of an "S-Curve" found by Mark Lanterman on the computer deployed to Michael Best \& Friedrich for use by Adam Foltz. Amended Lanterman Decl., Ex. B (Dkt. 97-2).
63. The metadata for "Composite_Current_Curve" is as follows:

JA176

| File Name | Composite_Current_Curve.xlsx |
| :---: | :---: |
| Extension | Xlsx |
| Created (Central) | $\begin{aligned} & \text { 5/28/2011 12:03:01 PM (2011-05- } \\ & 28 \text { 17:03:01 UTC) } \end{aligned}$ |
| Accessed (Central) | $\begin{aligned} & \text { 6/1/2011 11:48:33 AM (2011-06- } \\ & 01 \text { 16:48:33 UTC) } \end{aligned}$ |
| Modified (Central) | $\begin{aligned} & \text { 6/1/2011 11:48:33 AM (2011-06- } \\ & 01 \text { 16:48:33 UTC) } \end{aligned}$ |
| File Path | /Users/afoltz/Desktop/Projects/C omposite_Current_Curve.xlsx |
| File Size | 447.98 KB |
| Author | Ronald Keith Gaddie |
| Last Saved By | Afoltz |
| Office Created Date | $\begin{aligned} & \text { 5/28/2011 8:12:17 AM (2011-05- } \\ & 28 \text { 13:12:17 UTC) } \end{aligned}$ |
| Office Last Printed Date | 6/1/2011 10:46:26 AM (2011-06- <br> 01 15:46:26 UTC) |
| Office Last Saved Date | $\begin{aligned} & \text { 6/1/2011 11:48:33 AM (2011-06- } \\ & \text { 01 16:48:33 UTC) } \end{aligned}$ |
| Hidden Columns or Rows | FALSE |
| Track Changes | FALSE |
| MD5 Hash Value | $\begin{array}{\|l} \hline \text { 2acd25783c0be60bbe563ab3240 } \\ 24556 \\ \hline \end{array}$ |

JA177

Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32586 Responsive Spreadsheets File Detail Report."
64. In "Composite_Current_Curve," the total number of seats for which Republicans have a baseline over 50\%, using Professor Gaddie's regression model, for statewide Republican vote shares between $46 \%$ and $52 \%$ is as follows:

| $\mathbf{4 6 \%}$ | $\mathbf{4 7 \%}$ | $\mathbf{4 8 \%}$ | $\mathbf{4 9 \%}$ | $\mathbf{5 0} \%$ | $\mathbf{5 1} \%$ | $\mathbf{5 2 \%}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 36 | 42 | 46 | 53 | 58 | 62 | 64 |

Amended Lanterman Decl., Ex. B (Dkt. 97-2).
65. "Composite_Joe_Assertive_Curve.xlsx" is located in the WRK32586 Responsive Spreadsheets Deduplicated file, and is a true and correct copy of an "S-Curve" found by Mark Lanterman on the computer deployed to Michael Best \& Friedrich for use by Adam Foltz. Amended Lanterman Decl., Ex. B (Dkt. 97-2).
66. The metadata for
"Composite_Joe_Assertive_Curve" is as follows:

| File Name | Composite_Joe_Assertive_Curve.xlsx |
| :--- | :--- |
| Extension | Xlsx |
| Created <br> (Central) | $5 / 28 / 2011$ <br> $17: 03: 01 ~ 12: 03: 01 ~ P M ~(2011-05-28 ~$ |
| Accessed <br> (Central) | $5 / 28 / 2011$ <br> $17: 49: 55$ <br> UTC) |

JA178

| Modified (Central) | $\begin{aligned} & \text { 5/28/2011 12:49:56 PM (2011-05-28 } \\ & \text { 17:49:56 UTC) } \end{aligned}$ |
| :---: | :---: |
| File Path | /Users/afoltz/Desktop/Projects/Composi te_Joe_Assertive_Curve.xlsx |
| File Size | 440.42 KB |
| Author | Ronald Keith Gaddie |
| Last Saved By | Afoltz |
| Office Created Date | 5/28/2011 8:12:17 AM (2011-05-28 13:12:17 UTC) |
| Office Last Printed Date |  |
| Office Last Saved Date | $\begin{aligned} & \text { 5/28/2011 12:49:56 PM (2011-05-28 } \\ & \text { 17:49:56 UTC) } \end{aligned}$ |
| Hidden Columns or Rows | FALSE |
| Track Changes | FALSE |
| MD5 Hash Value | 4a25a4cc8403f9c9ffb61b1eb0bb0de5 |

Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32586 Responsive Spreadsheets File Detail Report."
67. In "Composite_Joe_Assertive_Curve," the total number of seats for which Republicans have a baseline over 50\%, using Professor Gaddie's regression model, for statewide Republican vote shares between $46 \%$ and $52 \%$ is as follows:

| $\mathbf{4 6 \%}$ | $\mathbf{4 7 \%}$ | $\mathbf{4 8 \%}$ | $\mathbf{4 9 \%}$ | $\mathbf{5 0 \%}$ | $\mathbf{5 1} \%$ | $\mathbf{5 2 \%}$ |
| :--- | ---: | :--- | :--- | :--- | :--- | ---: |
| 44 | 50 | 55 | 58 | 60 | 62 | 63 |

Amended Lanterman Decl., Ex. B (Dkt. 97-2).
68. "Team_Map_Curve.xlsx" is located in the WRK32586 Responsive Spreadsheets Deduplicated file, and is a true and correct copy of an "S-Curve" found by Mark Lanterman on the computer deployed to Michael Best \& Friedrich for use by Adam Foltz. Amended Lanterman Decl., Ex. B (Dkt. 97-2).
69. The metadata for "Team_Map_Curve" is as follows:

| File Name | Team_Map_Curve.xlsx |
| :---: | :---: |
| Extension | Xlsx |
| Created <br> (Central) | $\begin{aligned} & \text { 6/14/2011 1:56:03 PM (2011-06-14 } \\ & \text { 18:56:03 UTC) } \end{aligned}$ |
| Accessed (Central) | $\begin{aligned} & \text { 6/14/2011 1:56:03 PM (2011-06-14 } \\ & \text { 18:56:03 UTC) } \end{aligned}$ |
| Modified (Central) | $\begin{aligned} & \text { 6/14/2011 1:56:03 PM (2011-06-14 } \\ & \text { 18:56:03 UTC) } \end{aligned}$ |

JA180

| File Path | /Users/afoltz/Desktop/Projects/Team_ Map_Curve.xlsx |
| :---: | :---: |
| File Size | 35.70 KB |
| Author | Ronald Keith Gaddie |
| Last Saved By | Afoltz |
| Office <br> Created <br> Date | 6/14/2011 12:06:15 PM (2011-06-14 17:06:15 UTC) |
| Office Last Printed Date | 6/14/2011 1:47:35 PM (2011-06-14 18:47:35 UTC) |
| Office Last Saved Date | $\begin{aligned} & \text { 6/14/2011 1:56:03 PM (2011-06-14 } \\ & \text { 18:56:03 UTC) } \end{aligned}$ |
| Hidden Columns or Rows | FALSE |
| Track Changes | FALSE |
| MD5 Hash Value | 5a79df0e25b95605c14ca7824dbb8614 |

Amended Lanterman Decl., Ex. B (Dkt. 97-2), "WRK32586 Responsive Spreadsheets File Detail Report."
70. In "Team_Map_Curve," the total number of seats for which Republicans have a baseline over 50\%,

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using Professor Gaddie's regression model, for statewide Republican vote shares between $46 \%$ and $52 \%$ is as follows:

| $\mathbf{4 6 \%}$ | $\mathbf{4 7 \%}$ | $\mathbf{4 8 \%}$ | $\mathbf{4 9 \%}$ | $\mathbf{5 0 \%}$ | $\mathbf{5 1 \%}$ | $\mathbf{5 2 \%}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 46 | 50 | 54 | 56 | 58 | 60 | 64 |

71. On March 9, 2016, during his deposition, Keith Gaddie was asked the following question:
Q. Is the Team Map Curve a more pro Republican map than a pro Democrat map?

Professor Gaddie responded to that question:
A. Let me look at it for a minute. Okay. At $50 \%$ of the expected vote statewide, of the 99 assembly districts it appears that 55 of them are either safely or leaning Republican with 21 of those seats being competitive Republican districts. At 53\% Republican statewide vote of the 99 assembly districts, 46 of them appear to be districts that we would term safely Republican based upon the estimate. So there is a Republican lean in this map, yes.

Gaddie Dep. (Dkt. 108) at 167:6-17.
72. No Democrats participated in the drafting process that led to the creation of the redistricting plan that was enacted in Act 43.
73. Prior to the legislative introduction of Act 43, no Democrat was given an opportunity to see the boundaries of any legislative districts in the proposed map.
74. Prior to the legislative introduction of Act 43, Republican legislators who had not been involved in drafting the plan were allowed to see the boundaries of their own district, but were not allowed to see the boundaries of any other district in the map.
75. Prior to the passage of Act 43, when Republican legislators were shown the boundaries of what would be their new legislative district, they were given information about the expected partisan voting patterns in the district, i.e., what percentage of voters were likely to vote for a Republican candidate and what percentage of voters were likely to vote for a Democratic candidate.
76. Under the direction and supervision of Eric McLeod, Tad Ottman met with 17 Republican members of the Wisconsin State Senate, identified in Exhibit 4 attached to the Complaint. Each of the 17 Republican Senators signed a secrecy agreement entitled "Confidentiality and Nondisclosure Related to Reapportionment" before being allowed to review and discuss their districts.
77. The secrecy agreement stated that Eric McLeod had "instructed" Tad Ottman to meet with certain members of the Senate to discuss the reapportionment process and characterized such conversations as privileged communications pursuant to the attorney-client and attorney work product privileges.
78. Under the supervision of Eric McLeod, Adam Foltz met with 58 Republican members of the Wisconsin State Assembly, identified in Exhibit 4 attached to the Complaint. Each of the 58 Republican Representatives signed a secrecy agreement entitled "Confidentiality and Nondisclosure Related to Reapportionment" before being allowed to review and discuss their districts, which also improperly described their conversations as privileged.
79. After each of the 58 Republican members of the Wisconsin State Assembly signed the secrecy agreement entitled "Confidentiality and Nondisclosure Related to Reapportionment," they gave it to Adam Foltz and none kept a copy for themselves. Foltz Dep. (Dkt. 110) at 357:16-358:3.
80. Robin Vos participated in each of the meetings that Adam Foltz had with each of the 58 Republican members of the Wisconsin State Assembly listed in Exhibit 4 of the Complaint. Foltz Dep. (Dkt. 110) at 263:6-265:5.

## JA184

81. Exhibit 100 to the deposition of Adam Foltz, dated $2 / 1 / 12$, is an authentic copy (within the meaning of Fed. Evid. Rule 901(a)) of a one-page memo addressed to Representative Garey Bies from Adam Foltz, dated June 19, 2011, with copies to Speaker Jeff Fitzgerald, Majority Leader Scott Suder, and Representative Robin Vos, which is captioned "New Map for the 1st District" and which had attached to it a map of the new 1st Assembly District that became part of Act 43. The information contained in the memo identified the partisan performance of the new 1st Assembly District based on data from five prior elections (Scott Walker in 2010, J.B. Van Hollen in 2010, John McCain in 2008, J.B. Van Hollen in 2008, and George W. Bush in 2004). Similar one-page memos with analogous partisan performance data with attached copies of the member's new district were sent to each of the 58 Republican members of the Wisconsin State Assembly on the same date, June 19, 2011. Foltz Dep. (Dkt.110) at 266:10-267:15.
82. Exhibit 113 to the deposition of Adam Foltz, dated $2 / 1 / 12$, is an authentic copy (within the meaning of Fed. Evid. Rule 901(a)) of a one-page memo created by Adam Foltz on June 20, 2011, at 12:34 p.m., and which was last saved on Adam Foltz's computer on July 7, 2011, at 2:40 p.m. and was a WORD document captioned "General Talking Points for Robin." Foltz Dep. (Dkt.110) at 337:6-16, 347:22351:4.
83. Exhibit 114 to the deposition of Adam Foltz, dated $2 / 1 / 12$, is an authentic copy (within the meaning of Fed. Evid. Rule 901(a)) of a printout of the metadata associated with Exhibit 113 to the same deposition, which was a WORD document created on June 20, 2011, at 12:34 p.m. and which was last saved on Adam Foltz's computer on July 7, 2011, at 2:40 p.m. Foltz Dep. (Dkt.110) at 337:6-16, 347:22-351:4.
84. In Baldus v. Wisconsin Government Accountability Board, 843 F. Supp. 2d 955, 959 (E.D. Wis. 2012), the Court held that the Legislature improperly asserted attorney-client and work product privileges to prevent discovery of information regarding the redistricting process.
85. On July 11, 2011, the Current Plan was introduced by the Committee on Senate Organization without any Democratic members of the Legislature having previously seen their districts or the plan as a whole. All Republican members of the Legislature had previously seen their individual districts along with visual aids demonstrating the partisan performance of their districts, but had not seen the overall map.
86. A public hearing was held on July 13, 2011. The bill was then passed by the Senate on July 19, 2011, and by the Assembly the next day on July 20, 2011. Act 43 was published on August 23, 2011.
87. Eric McLeod and Michael, Best \& Friedrich, LLP, were paid $\$ 431,000.00$ in State taxpayer funds for their work on the Current Plan.
88. "ADAMFOLZSUPPPROD000431" is true and correct copy of a page from Adam Foltz's calendar for June 20, 2011 - June 24, 2011.
89. "ADAMFOLZSUPPPROD000431" shows meetings with twenty-nine individual Republican legislators during the week of June 20, 2011 - June 24, 2011.
90. "ADAMFOLZSUPPPROD000424" is a true and correct copy of a document titled "General Talking Points" drafted by Adam Foltz in 2011 in advance of the individual meetings held with Republican legislators in June 2011, to discuss the redistricting plan that would become Act 43.
91. "ADAMFOLZSUPPPROD000119" is a true and correct copy of a series of 59 memos addressed to each Republican Assembly member, and CCed to Speaker Jeff Fitzgerald, Majority Leader Scott Suder, and Rep. Robin Vos, from Adam Foltz - Assembly Redistricting Coordinator, dated 6/19/2011 with the subject lines "New Map for the 1st District," "New Map for the 2nd District," and so on until "New Map for the 99th District."
92. Page 62 of 63 in document 156-1 filed on 2/14/12 in Baldus v. Brennan, 2:11-cv- 00562-JPS-

DPW-RMD, is a true and correct copy of an email from Tad Ottman to Jim Troupis, Raymond Taffora, Eric M. McLeod, and Adam Foltz, sent on July 12, 2011 at 10:00PM with the subject line "Hearing memos" and listing attachment titled "sb148 committee memos.docx."
93. Page 63 of 63 in document 156-1 filed on $2 / 14 / 12$ in Baldus v. Brennan, 2:11-cv- 00562-JPS-DPW-RMD, is a true and correct copy of an email from Tad Ottman to Adam Foltz, sent on July 12, 2011 at 8:52PM with the subject line "committee memos" and listing attachment titled "sb146 committee memos.docx."
94. "ADAMFOLZSUPPPROD000446.PDF" is a true and correct copy of an email from Dana Wolff to Tad Ottman and Adam Foltz and CCed to Tony Van Der Wielen sent on Monday May 9, 2011 at 12:32PM, with the subject line "Letter" and listing attachment titled "MCD_Letter.pdf."
95. Page 56 of 63 in document 156-1 filed on 2/14/12 in Baldus v. Brennan, 2:11-cv-00562-JPS-DPW-RMD, is a true and correct copy of an email from Tad Ottman to Jim Troupis and Eric M McLeod, CCed to Adam Foltz, sent on Friday February 25, 2011 at 2:31PM, with the subject line "Redistricting timeline."
96. "MBF000217" is a true and correct copy of an email from Jim Troupis to Tad Ottman and Adam Folz, CCed to Eric M McLeod and Sarah Troupis, sent

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on Monday, June 13, 2011 at 8:25AM, with the subject line "Gaddie \& Hispanic."
97. Page 3 of 63 in document 156-1 filed on $2 / 14 / 12$ in Baldus v. Brennan, 2:11-cv- 00562-JPS-DPW-RMD, is a true and correct copy of an email from Tad Ottman to Jim Troupis, Eric M. McLeod, Raymond Taffora, and Adam Foltz sent on Wednesday July 13, 2011 at 1:45PM with the subject line "Latino voices will be there."
98. "Foltz001075" is a true and correct copy of a chart prepared by Adam Foltz in 2011.
99. "Foltz001075" sets out the population deviations for the seats that were held following the 2010 elections by the "GOP," by "Indp" and by "Dem" in separate categories.

## Professor Jackman's Reports

100. The efficiency gap indicates the extra proportion of seats that an advantaged party wins relative to a baseline where the parties are wasting equal numbers of votes. Jackman Rpt. (Dkt. 62) at 19.
101. Defendants' expert, Professor Goedert, "concur[s] that this shortcut is an appropriate and useful summary measure." Goedert Rpt. (Dkt. 51) at 5; Goedert Dep. (Dkt. 65) at 70:17-71:1.
102. Defendants' expert, Sean Trende, noted that in 2012 Professor Mayer calculated that the Current Plan had an efficiency gap of $-11.7 \%$ using the full method and Mr. Trende calculated the efficiency gap for 2012 as $-9.9 \%$ using the simplified method, a difference of 1.8 percentage points. Mayer Rpt. (Dkt. 54) at 46; Jackman Rpt. (Dkt. 62) at 71; Trende Rpt. (Dkt. 55) at 59.
103. Similarly, Mr. Trende noted that Professor Mayer calculated that the DemonstrationPlan had an efficiency gap of $-2.2 \%$ using the full method and Mr . Trende calculated the efficiency gap for 2012 as -0.8\% using the simplified method, a difference of 1.4 percentage points. Mayer Rpt. (Dkt. 54) at 46; Jackman Rpt. (Dkt. 62) at 71; Trende Rpt. (Dkt. 55) at 60 .
104. Under the simplified method only, the (S -$0.5)-2(\mathrm{~V}-0.5)$ formula implies that for the efficiency gap to be zero, there must be a $2: 1$ relationship between seat share and vote share (also known as "responsiveness"). Jackman Rpt. (Dkt. 62) at 17-18.
105. As Professor Goedert has explained in his report and other work, a responsiveness of 2 "conform[s] with the observed average seat/votes curve in historical U.S. congressional and legislative elections." Goedert Rpt. (Dkt. 51) at 6; Goedert Dep. (Dkt. 65) at 95:17-21.
106. At the congressional level, the seat/vote curve had "an average slope of 2.02 for the past 40 years." During "the preceding 70 years," it had an "average of 2.09." Goedert Dep., Ex. 20 (Dkt. 65-2) at 7.
107. Professor Jackman's dataset used for his calculations of the efficiency gap in state legislative elections spans the period 1972 to 2014, representing the post-malapportionment era. Jackman Rpt. (Dkt. 62) at 19.
108. Professor Jackman's calculations of the efficiency gap rely on a dataset widely used in political science and freely available from the Inter-University Consortium for Political and Social Research (ICPSR study number 34297). The release of the dataset utilized by Professor Jackman covers state legislative election results from 1967 to 2014, updated by Carl Klarner (Indiana State University and Harvard University). Jackman Rpt. (Dkt. 62) at 20; Jackman Dep. (Dkt. 53) at 46:23-47:14.
109. Professor Jackman uses a subset of the original dataset for general elections since 1972 in states whose lower houses are elected via singlemember districts, or where single- member districts are the norm. Professor Jackman treats multimember districts "with positions" as if they are singlemember districts. Jackman Rpt. (Dkt. 62) at 20; Jackman Dep (Dkt. 53) at 44:24-46:22.

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110. The total dataset used by Professor Jackman spans 83,260 district-level state legislative races, from 786 elections across 41 states. Jackman Rpt. (Dkt. 62) at 20-21, and Figure 5. Jackman Dep. (Dkt. 53) 48:1-3.
111. Professor Jackman groups the efficiency gap scores across the series of elections held under the same districting plan, using the unique identifier for the districting plan in place for each state legislative election provided by Stephanopoulos and McGhee, as shown in the following chart:


Figure 6: 786 state legislative elections available for analysis, 1972-2014, by state, grouped by districting plan (horizontal line).
Jackman Rpt. (Dkt. 62) at 22-23.
112. Professor Jackman calculated the efficiency gap for every state house election for which data was available over the period from 1972 to 2014, using actual election results. Professor Jackman did not

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aggregate wasted votes district by district, but rather used a simplified computation method based on statewide electoral data, with the formula EG $=(S-$ $0.5)-2(V-0.5)$, where EG is the efficiency gap, S is the statewide Democratic seat share, and $V$ is the statewide Democratic vote share. Jackman Rpt. (Dkt. 62) at 16-17.
113. Professor Jackman's analysis found that for a plan with an initial efficiency gap of $-7 \%$, the average efficiency gap over the life of the plan is estimated to be -5.3\%.
114. Similarly, Professor Jackman's analysis found that for a plan with an initial efficiency gap of $7 \%$, the average efficiency gap over the life of the plan is estimated to be $3.7 \%$.
115. The average net efficiency gap (i.e., the mean of the actual values of all plans' efficiency gaps in a given year) has recently trended in a Republican direction. This metric was mildly pro-Democratic from the early 1970s to the mid-1990s, but has been moderately pro-Republican from the mid-1990s to the present. Jackman Rpt. (Dkt. 62) at 44-45; Stephanopoulos \& McGhee, supra, at 873.
116. There are 206 distinct plans in Professor Jackman's database. Of these, 70 plans (or 34\%) had an initial efficiency gap greater than $7 \%$ in magnitude, and 32 plans (or 16\%) had an initial efficiency gap greater than $10 \%$ in magnitude.

Jackman Rpt. (Dkt. 62) at 7; Jackman Rebuttal Rpt. (Dkt. 63) at 18-20; Jackman Decl. Ex. F (Dkt. 58-6).
117. Of the 70 plans that had an initial efficiency gap greater than $7 \%$ in magnitude, 43 plans (or $21 \%$ of the 206 total plans) were designed by a single party that had unified control over redistricting. Jackman Rpt. (Dkt. 62) at 7; Jackman Rebuttal Rpt. (Dkt. 63) at 18-20; Jackman Decl. Ex. F (Dkt. 58-6).
118. Of the 32 plans that had an initial efficiency gap greater than $10 \%$ in magnitude, 20 plans (or $10 \%$ of the 206 total plans) were designed by a single party that had unified control over redistricting. Jackman Rpt. (Dkt. 62) at 7; Jackman Rebuttal Rpt. (Dkt. 63) at 18-20; Jackman Decl. Ex. F (Dkt. 58-6).
119. Of the 43 plans from the current redistricting cycle in Professor Jackman's database, 16 (or $37 \%$ of the 43 plans) had initial efficiency gaps above $7 \%$ in magnitude, and of these, 11 plans (or $26 \%$ of the 43 plans) were designed by a single party that had unified control over redistricting. Jackman Rpt. (Dkt. 62) at 7; Jackman Rebuttal Rpt. (Dkt. 63) at 1820; Jackman Decl. Ex. F (Dkt. 58-6).
120. Of the 43 plans from the current redistricting cycle in Professor Jackman's database, 11 plans (or $26 \%$ of the 43 plans) had initial efficiency gaps greater than $10 \%$ in magnitude and of these, 7 plans (or $16 \%$ of the 43 plans) were designed by a single party that had unified control over
redistricting. Jackman Rpt. (Dkt. 62) at 7; Jackman Rebuttal Rpt. (Dkt. 63) at 18-20; Jackman Decl. Ex. F (Dkt. 58-6).
121. The following chart identifies: (i) the number of plans, historically and currently, in Professor Jackman's database that had an initial efficiency gap above 7\%; (ii) the number of plans with an initial efficiency gap above $7 \%$ and unified party control; (iii) the number of plans with an initial efficiency gap above $10 \%$; and (iv) the number of plans with an initial efficiency gap above $10 \%$ and unified party control:

| Historical |  | Current |  |
| :--- | :--- | :--- | :--- |
| All plans | 206 | Current plans | 43 |
| All plans with initial <br> $E G$ above 7\% | 70 | Current plans with <br> initial $E G$ above 7\% | 16 |
| All plans with initial <br> $E G$ above 7\% and <br> unified party control <br> over redistricting | Current plans with <br> initial $E G$ above 7\% <br> and unified party <br> control over <br> redistricting | 11 |  |
| All plans with initial <br> $E G$ above 10\% | 32 | Current plans with <br> initial $E G$ above 10\% | 11 |


| All plans with initial |
| :--- | :--- | :--- |
| $E G$ above $10 \%$ and |
| unified party control |
| over redistricting |$\quad$| Current plans with |
| :--- |
| initial $E G$ above $10 \%$ |
| and unified party |
| control over |
| redistricting |$\quad 7$| 7 |
| :--- |

Jackman Rpt. (Dkt. 62) at 7; Jackman Rebuttal Rpt. (Dkt. 63) at 18-20; Jackman Decl. Ex. F (Dkt. 58-6).
122. The proportion of plans created by Republicans in full control of the state government increased from about $10 \%$ in the 1990s, to about $20 \%$ in the 2000 s, to about $40 \%$ in the 2010s (in 49 states, excluding Nebraska). By comparison, fewer than $20 \%$ of current plans were designed by Democrats in full control of the state government. Jackman Rebuttal Rpt. (Dkt. 63) at 19; Trende Dep. (Dkt. 66) at 79:1123.
123. The reapportionment revolution of the 1960s resulted in the invalidation of almost every state house, state senate, and congressional plan in the country. Jackman Decl. Ex. J (Dkt. 58-10) at 4.
124. Wisconsin does not have equal turnout across Assembly districts.
125. In Wisconsin's 2012 Assembly elections, the turnout in individual districts varied from just over 8,000 votes in District 8 to over 37,000 votes in District 14.
126. In Wisconsin's 2014 elections, the turnout in individual districts varied from approximately 6,400 votes in District 8 to over 31,400 votes in District 23.
127. The presence of imputed vote totals leads to uncertainty in Professor Jackman's calculation of vote share, which "generates uncertainty in determining how far each point lies above or below the orange, zero efficiency gap benchmark."
128. Professor Jackman expresses his $E G$ calculations as "point estimates" with lines indicating a $95 \%$ level of confidence.
129. Professor Jackman has less confidence in the "point estimate" of his $E G$ as the number of uncontested seats increases.
130. Professor Jackman found that "[t]he distribution of $E G$ measures trends in a proRepublican direction through the 1990s, such that by the 2000 s, $E G$ measures were more likely to be negative (Republican efficiency over Democrats)."
131. Professor Jackman plotted the efficiency gap of each plan in each year from lowest to highest (from most favorable to Republicans to least) and then overlaying estimates of the smoothed weighted quantiles (with blue lines showing the 25th percentile, 50 th percentile, and 75 th percentile plan).

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132. The median efficiency gap has been negative (favorable to the Republicans) since the mid1990s.
133. The most favorable median toward Democrats since 2000 was in 2010.
134. The 25 th percentile has been below $5 \%$ since the mid-1990s and even approached $7 \%$ in 2004, 2010, and 2012.
135. The 75 th percentile has been below $5 \%$ since the mid-1990s and has hovered between $1 \%$ and $2 \%$ since 2000 .
136. Professor Jackman's calculation of the "the probability that a given efficiency gap number from a given election year is positive or negative" also shows a trend in favor of Republicans.
137. Professor Jackman finds that in every election year since 1996, more plans have had negative efficiency gaps than positive ones with the exception of 2010.
138. In 2010, Professor Jackman found that the proportion of plans having a positive efficiency gap was slightly more than 0.5 .
139. In 2006, $75 \%$ of plans produced a negative efficiency gap while only $25 \%$ of plans produced a

JA199
positive efficiency gap, with similar results in 2000 and 2012.
140. Since 1996, the year with the greatest proportion of efficiency gap measures favoring Democrats was 2010, in which there was a slightly more than a 50-50 probability of a plan being positive (favorable to Democrats).
141. Professor Jackman chose to look at the first election in the plan because he "tried to put [himself] in the shoes of litigants" who would have to "intervene early before we've seen much data all from the plan, the election results the plan is throwing off."
142. For all plans Professor Jackman studied since 1972, he finds that $36 \%$ of all plans produced an efficiency gap of $7 \%$ or greater in the first election: $18 \%$ on the positive side and $18 \%$ on the negative side.
143. For all plans Professor Jackman studied since 1991, $34 \%$ of all plans produced an efficiency gap greater than $7 \%$ in magnitude in the first election: $22 \%$ produced a gap of at least $-7 \%$ in magnitude and $12 \%$ percent produced a gap of at least $+7 \%$ in magnitude.
144. For all plans since 1972 that Professor Jackman studied, he finds that $18 \%$ of plans that had an EG of at least $-7 \%$ in magnitude go on to produce an election with a positive $E G$.

## JA200

145. For all plans Professor Jackman studied since 1991, he finds that $40 \%$ of plans that produce an EG of at least $+7 \%$ in magnitude in the first election go on to produce an election with a negative $E G$.
146. For all plans Professor Jackman studied since 1991, he finds that $18 \%$ of plans that produce an EG of at least $-7 \%$ in magnitude in the first election go on to produce an election with a positive $E G$.
147. For all plans Professor Jackman studied since 1991, he finds that $60 \%$ of plans that produce an EG of at least $+7 \%$ in magnitude in the first election go on to produce an election with a negative $E G$.
148. Professor Jackman finds that "we seldom see a plan in the 1990s or later that commence with a large-pro Democratic efficiency gap."
149. In the 1990s and later, Professor Jackman finds that the probability the first election has an efficiency gap greater than $+5 \%$ (favorable to Democrats) "is only about $11 \%$."
150. Negative efficiency gaps "are much more likely under the first election in post- 1990 plans: almost $40 \%$ of plans open with $E G<-.05$ and about $20 \%$ of plans open with $E G<-.10$."
151. Jackman finds that "plans with at least one election" of an efficiency gap of $7 \%$ or greater "are reasonably common."

JA201
152. Jackman finds that $53 \%$ of plans since 1972 have one election with an $E G$ of $7 \%$ or greater in magnitude, with $29 \%$ of plans having a gap of $-7 \%$ or greater in magnitude and $25 \%$ of plans having a gap of $+7 \%$ or greater.
153. When looking at plans since 1991, $47 \%$ of plans have had at least one election with an $E G$ greater than $7 \%$ in magnitude, with $38 \%$ of plans having an election with a gap of $-7 \%$ or greater in magnitude and $19 \%$ of plans having an election with a gap of $+7 \%$ or greater.
154. Since 1972, $33 \%$ of plans have had an election with an $E G$ of $10 \%$ or greater in magnitude, with $18 \%$ having an election with a gap of $-10 \%$ in magnitude and $15 \%$ having an election with a gap of +10\% or greater.
155. When looking just at elections since 1991, $35 \%$ of plans have had an election with an $E G$ of at least $10 \%$ in magnitude: $24 \%$ of plans have had an election with a gap of $-10 \%$ in magnitude and $11 \%$ of plans having an election with a gap of $+10 \%$.
156. Professor Jackman found that 17 of the 141 plans for which he could calculated three or more efficiency gaps (12\%) were"utterly unambiguous with respect to the sign of the efficiency gap,"i.e., that even the confidence level bar did not cross over to the other sign.
157. Of these seventeen plans, sixteen of them were favorable to the Republicans and only one was favorable to the Democrats.
158. One of the "utterly unambiguous" plans was the Wisconsin 2002 Plan put in place by the federal court in Baumgart v. Wendelberger, No. 01-C-0121, 2002 WL 34127471, at *1 (E.D. Wis. May 30, 2002), amended, 2002 WL 34127473 (E.D. Wis. July 11, 2002).
159. Professor Jackman calculated EGs for the 2012 and 2014 elections for 39 states.
160. Fifty point estimates were negative (64.1\%) while twenty-eight point estimates were positive (35.9\%).
161. Eighteen states (46\%) had point estimates for 2012 and 2014 that were both negative.
162. Included among this eighteen were Minnesota, Missouri, New York, and Kansas.
163. With respect to the entire country, Professor Jackman found that " $[\mathrm{t}]$ he distribution of $E G$ measures trends in a pro-Republican direction through the 1990s, such that by the 2000s, $E G$ measures were more likely to be negative."
164. The median plan has been negative since the mid-1990s and the 25th percentile has been below
$5 \%$ since the mid-1990s and even approached $7 \%$ in 2004, 2010, and 2012.
165. Meanwhile the seventy-fifth percentile has only favored Democrats by $1 \%-2 \%$.
166. In every election year since 1996, more plans have had negative efficiency gaps than positive ones with about $75 \%$ of plans producing a negative efficiency gap in 2000, 2006 and 2012.
167. In 2012, the Republicans won five seats (Districts 1, 26, 50, 72 and 93) with no more than $51.3 \%$ of the total vote.
168. The margin of victory across all of these races was about 3,200 votes, each less than 900 votes and one at only 109 votes (District 93).
169. For 2012 and 2014, Professor Jackman calculates that Illinois had one negative efficiency gap and one narrowly positive efficiency gap.

## Professor Mayer's Reports

170. To generate his baseline partisanship estimates, Professor Mayer assumed that all districts were contested and that no incumbents were running. This method removes the effect of incumbents, who may or may not be running in an alternative plan. The consultant retained by the state legislature, Professor Gaddie, used the same method. Mayer Rpt.
(Dkt. 54) at 31; Mayer Dep. (Dkt. 52) at 63:15-24, 70:4-17; Gaddie Dep. (Dkt. 108) at 43:9-44:22.
171. Professor Mayer's regression model used wards as the unit of analysis to increase the number of observations and allow for more precise estimates. Mayer Rpt. (Dkt. 54) at 8.
172. Professor Mayer's regression model relied on demographic and electoral data provided by the LTSB and the G.A.B., both online and in the 2013 edition of the Wisconsin Blue Book. Mayer Rpt. (Dkt. 54) at 10.
173. The full specification for the regression model that Professor Mayer used is:

JA205

$$
\begin{aligned}
\text { Assembly }_{\text {Vote }_{i}=} & \alpha+\beta_{1} \text { Total VEP }_{i}+\beta_{2} \text { BlackVEP }_{i}+\beta_{3} \text { Hispanic VEP }_{i} \\
& +\beta_{4} \begin{array}{c}
\text { Democratic }
\end{array}+\beta_{5} \begin{array}{c}
\text { Republican }^{\text {Presidential Vote }}{ }_{i}
\end{array}{ }^{\text {Presidential Vote }} i
\end{aligned}
$$

Where

| Assembly Vote | Number of votes cast for the Republican or Democratic <br> candidate in the 2012 Assembly election in ward $i$. I <br> estimate separate equations for the Democratic and <br> Republican candidates <br> Votal VEP |
| :--- | :--- |
| Census eligible population in ward $i$, as measured in the 2010 |  |
| Black VEP | Voting eligible Black population in ward $i$ |

Mayer Rpt. (Dkt. 54) at 10-11.
174. The full specification for the regression model that Professor Mayer used includes the Assembly vote by ward as the dependent variable and the following as independent variables (each by ward): total voting eligible population; black voting eligible population; Hispanic voting eligible
population; Democratic presidential vote; Republican presidential vote; Democratic incumbent; Republican incumbent; and a set of fixed effect dummy variables for each county, with Dunn County as the excluded value. Mayer Rpt. (Dkt. 54) at 10-11.
175. Professor Keith Gaddie used a regression model "very similar" to the one used by Professor Mayer in 2002 in the Baumgart litigation, stating that he "basically replicated [Professor Mayer's] model," to predict the Current Plan's partisan consequences prior to the Plan's enactment. Gaddie Dep. (Dkt. 108) at 53:3-7, 47:10-14, 43:9-44:22; Mayer Rpt. (Dkt. 54) at 29.
176. In Table 2, Professor Mayer's regression model incorrectly predicted the outcomes of only two extremely competitive districts: District 51 (actual Republican vote: 51.9\%; predicted Republican vote: 49.9\%) and District 70 (actual Republican vote: 49.7\%; predicted Republican vote: 50.1\%). Mayer Rpt. (Dkt. 54) at 24-25; Mayer Dep. (Dkt. 52) at 87:22-23.
177. According to Table 2, these incorrect predictions are balanced, one for each party, meaning that in the aggregate, Professor Mayer's model estimated the partisan distribution of contested districts in 2012 ( 56 Republican, 16 Democratic) with perfect accuracy. Mayer Rpt. (Dkt. 54) at 24-25.

JA207
178. Professor Mayer's baseline partisanship model produces the following vote totals and twoparty vote percentages:

| City | Dem. Votes | Rep. Votes | Total |
| :---: | :---: | :---: | :---: |
| Milwaukee | 193,940 <br> $(77.9 \%)$ | 54,992 <br> $(22.1 \%)$ | 248,932 |
| Madison | 109,466 <br> $(78.0 \%)$ | 30,928 <br> $(22.0 \%)$ | 140,394 |
| Green Bay | $23,403(55.2 \%)$ | 18,998 <br> $(44.8 \%)$ | 42,402 |
| Kenosha | $26,515(62.6 \%)$ | 15,828 <br> $(37.4 \%)$ | 42,342 |
| Racine | $22,614(70.4 \%)$ | $9,517(29.6 \%)$ | 32,131 |
| Appleton | $18,232(51.6 \%)$ | 17,129 <br> $(48.4 \%)$ | 35,361 |
| Waukesha | $15,257(37.6 \%)$ | 25,273 <br> $(62.4 \%)$ | 40,530 |
| Oshkosh | $17,364(52.1 \%)$ | 15,945 <br> $(47.9 \%)$ | 33,309 |
| Eau Claire | $20,601(59.2 \%)$ | 14,202 <br> $(40.8 \%)$ | 34,803 |
| Janesville | $20,208(58.9 \%)$ | 14,080 <br> $(41.1 \%)$ | 34,288 |
| La Crosse | $17,554(67.4 \%)$ | $8,485(32.6 \%)$ | 26,039 |
| Sheboygan | $14,573(56.5 \%)$ | 11,215 <br> $(43.5 \%)$ | 25,787 |
| Beloit | $11,440(63.3 \%)$ | $6,623(36.7 \%)$ | 18,062 |

179. Professor Mayer's baseline partisanship model for Act 43 produces 197 wasted votes for the Republicans and 16,235 wasted votes for the Democrats in District 1.
180. In the actual 2012 election, in District 1 the Republican won with 16,993 votes and the Democrat lost with 16,124 votes.
181. In the actual election, in District 1, there were 435 wasted votes for the Republicans and 16,124 wasted votes for the Democrats.
182. In the actual 2012 election, the Republican candidate won District 50 with 12,842 votes to the Democratic candidate's 11,945 votes.
183. In the actual election, the Republican candidate won District 51 with 10,642 votes to the Democratic candidate's 10,577 votes.
184. In the actual election, the Republican candidate won District 68 with 13,758 votes to the Democratic candidate's 12,482 votes.
185. In the actual election, the Democratic candidate won District 70 with 13,518 votes to the Republican candidate's 13,374.
186. For his model, Professor Mayer admits that "the average absolute error in the vote margin is 1.49\%."
187. Professor Mayer's baseline partisanship model of Act 43 contains 42 districts with at least a 50\% Democratic baseline.
188. Professor Mayer's baseline partisanship model of Act 43 contains 17 seats that have a baseline between 50-55\% Republican. These districts and percentages are shown in the chart below, from the least Republican to the most Republican:

| District | Mayer Baseline Rep. \% |
| :---: | :---: |
| 93 | $50.2 \%$ |
| 1 | $50.6 \%$ |
| 67 | $51.6 \%$ |
| 29 | $52.2 \%$ |
| 88 | $52.3 \%$ |
| 4 | $52.3 \%$ |
| 49 | $52.5 \%$ |
| 27 | $52.7 \%$ |
| 42 | $53.0 \%$ |
| 26 | $53.3 \%$ |
| 62 | $53.9 \%$ |
| 31 | $54.1 \%$ |
| 70 | $54.1 \%$ |
| 40 | $54.2 \%$ |
| 28 | $54.6 \%$ |
| 30 | $54.7 \%$ |
| 21 | $54.9 \%$ |
|  |  |

Comparison of Act 43 with Prior Plans
189. In the 1980s, a federal court drew the State Assembly districts. Wisc. State AFL-CIO v. Elections

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Bd., 543 F. Supp. 630 (E.D. Wis. 1982). The districts were amended by a legislature and Governor with unified Democratic control in 1983 and used for the period 1984-1990.
190. The average efficiency gap of the Wisconsin State Assembly redistricting plan from 1992-2000 was $-2.4 \%$. Jackman Rpt. (Dkt. 62) at 72; Jackman Decl. Ex. F (Dkt. 58-6) at 18.
191. In the 1990s, a federal court drew the State Assembly districts. Prosser v. Elections Bd., 793 F. Supp. 859 (W.D. Wis. 1992). The Prosser court took into account likely electoral effects and designed the map that was the "least partisan" and "create[d] the least perturbation in the political balance of the state." Id. at 871.
192. The average efficiency gap of the Wisconsin State Assembly redistricting plan from 2002-2010 was $-7.6 \%$. Jackman Rpt. (Dkt. 62) at 72; Jackman Decl. Ex. F (Dkt. 58-6) at 25.
193. In the 2000s, a federal court drew the State Assembly districts. See Baumgart v. Wendelberger, 2002 WL 34127471 (E.D. Wis. May 30, 2002).
194. A summary of the average efficiency gap for each decade, and the list of who was in control of the redistricting process is shown in this table:

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| Decade | Control of <br> government | Average <br> efficiency gap |
| :---: | :---: | :---: |
| $1972-1980$ | Divided | $-0.3 \%$ |
| $1982-1990$ | Court drawn, <br> then unified <br> Democratic <br> control | $-1.9 \%$ |
| $1992-2000$ | Court drawn | $-2.4 \%$ |
| $2002-2010$ | Court drawn | $-7.6 \%$ |

195. Between 1972 and 2014, fewer than four percent of all state house plans nationwide had an efficiency gap with an absolute value of $13 \%$ or higher. Jackman Rpt. (Dkt. 62) at 7; Defs. Adnission to RFA \#20.
196. Between 1972 and 2010, no state house plan anywhere in the United States had an efficiency gap as large as the Current Plan in the first two elections after redistricting. Jackman Rpt. (Dkt. 62) at 4; Defs. Admission to RFA \#21.
197. The Current Plan created six blackmajority districts (districts 10-12 and 16-18), ranging from $56.7 \%$ to $67.6 \%$ black population, and from $51.1 \%$ to $61.8 \%$ black voting age population. The Demonstration Plan retains six black-majority districts, ranging from $60.0 \%$ to $63.4 \%$ black

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population, and from $56.2 \%$ to $60.5 \%$ black voting age population. Mayer Rpt. (Dkt. 54) at 37.
198. In Baldus v. Wisc. Gov't Accountability Bd., 849 F. Supp. 2d 840 (E.D. Wis. 2012), a federal court created a Latino-majority district in Milwaukee (District 8). The Demonstration Plan retains the boundaries of this district. Mayer Rpt. (Dkt. 54) at 38.
199. According to the 2010 Census, Wisconsin is $70.2 \%$ urbanized, and according to the 2014 update to the Census, Wisconsin is $6.6 \%$ black and $6.5 \%$ Hispanic.
200. The 1992 Assembly map entered by the Prosser court plan had an overall range of population deviation of 0.91 percent, with 48 districts below the ideal and 51 above the ideal. Only one district was more than a half point away from the ideal. In the Senate, the 1992 plan had an overall deviation range 0.52 percent, with 15 districts above the ideal population and 18 below the ideal.
201. The 2002 Assembly map entered by the Baumgart court had an overall range of 1.59 percent deviation, with 47 districts above the ideal, 51 below the ideal, and one exactly apportioned district. In the Senate, the overall deviation range of the 2002 map was 0.98 percent, with 15 districts above the ideal population, 17 below, and one perfectly apportioned. Of the 99 Assembly districts in 2002, 77 districts were

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within +/- 0.5 percent of the ideal population; in the Senate, 32 of 33 districts fell in this range.
202. Act 43 creates 99 Assembly districts with populations falling within a range of 0.76 percent (+0.39 percent to -0.37 percent) of the ideal population; 56 districts are above the ideal population, 41 are below the ideal, and two districts are perfectly apportioned. In the Senate, population variations fall within a range of 0.62 percent $(+0.35$ percent to -0.27 percent); 17 districts are above the ideal population, 14 are below the ideal, and two districts are perfectly apportioned.
203. The population deviation in Act 43 from the ideal for each Assembly and Senate district (using 2010 Census data) is described in the Appendix to Act 43 and Tables 2 and 3 to the pretrial report filed in the Baldus case on February 14, 2012.
204. A summary of population deviation in Assembly districts in Act 43, the 1992 plan, and the 2002 plan is in Table 4 of the pretrial report filed in the Baldus case on February 14, 2012.
205. Each state Senate district is composed of three entire state Assembly districts.
206. Assembly members serve two-year terms. Senators serve four-year, staggered terms with half elected in presidential years and the other half coincident with gubernatorial elections.
207. The 1992 Federal Court map for the Assembly split 72 municipalities.
208. In 2002, the Federal Court's Assembly map split 50 municipalities.
209. Act 43 splits 62 municipalities in the Assembly.
210. The 1992 Federal Court map split 47 counties in the Assembly.
211. In 2002, the Federal Court divided 51 counties in the Assembly
212. Act 43 splits 58 counties in the Assembly.
213. Two widely-used measures of compactness applied to legislative districts are the Perimeter-toArea measure and the Smallest Circle score.
214. The Perimeter-to-Area measure compares the relative length of the perimeter of a district to its area. It represents the area of the district as the proportion of the area of a circle with the same perimeter. The score ranges from 0 to 1 , with a value of 1 indicating perfect compactness. This score is achieved if a district is a circle. Most redistricting software generates this measure as the PolsbyPopper statistic.
215. Smallest Circle scores measure the space occupied by the district as a proportion of the space of the smallest encompassing circle, with values ranging from 0 to 1 . A value of 1 indicates perfect compactness and is achieved if a district is a circle. This statistic is often termed the Reock measure by redistricting applications. Ernest C. Reock, Jr. 1961, "A Note: Measuring Compactness as a Requirement of Legislative Apportionment," Midwest Journal of Political Science 5: 70-74.
216. The average Smallest Circle score for the entire Assembly map is 0.39 (range from 0.20 to 0.61 ).
217. The average Smallest Circle score for the entire Assembly map drawn by the Baumgart court in 2002 was 0.41 (range from 0.18 to 0.63 ).
218. The average Perimeter To Area score for the Assembly map is .28 (range of .05 to .56).
219. The average Perimeter To Area score for the Assembly map drawn by the Baumgart court in 2002 was 0.29 (range of 0.06 to 0.58 ).
220. The average Assembly compactness scores are marginally lower for Act 43 than for the 2002 court-crafted plan.
221. The following chart contains a summary of municipal splits, county splits and compactness scores for Act 43 and prior plans.

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|  | Municipa <br> l Splits | County <br> Splits | Reock <br> (mean) | Polsby- <br> Popper <br> (mean) |
| :---: | :---: | :---: | :---: | :---: |
| 1972 Plan |  | 49 |  |  |
| 1982 Plan |  | 41 |  |  |
| 1992 Plan | 72 | 47 |  |  |
| 2002 Plan | 50 | 51 | 0.41 | 0.29 |
| Act 43 | 62 | 58 | 0.39 | 0.28 |

222. The average efficiency gap of the Wisconsin State Assembly redistricting plan from 1972-1980 was $-0.3 \%$, and it was drawn by divided government. Jackman Rpt. (Dkt. 62) at 72; Jackman Decl. Ex. F (Dkt. 58-6) at 3.
223. The average efficiency gap of the Wisconsin State Assembly redistricting plan from 1982-1990 was $-1.9 \%$. Jackman Rpt. (Dkt. 62) at 72; Jackman Decl. Ex. F (Dkt. 58-6) at 11.

## The Demonstration Plan

224. There are eighteen districts in Professor Mayer's Demonstration Plan that are 50\%-55\% Democratic under his baseline partisanship model, assuming all seats were contested and no incumbents were running, including sixteen districts between $50 \%-53.4 \%$. The following table shows these districts ordered from least Democratic to most Democratic.

JA217

| Demonstration Plan <br> District | Predicted Dem. Vote \% |
| :---: | :---: |
| 49 | $50.3 \%$ |
| 92 | $50.5 \%$ |
| 86 | $50.7 \%$ |
| 96 | $51.5 \%$ |
| 91 | $51.7 \%$ |
| 81 | $51.8 \%$ |
| 40 | $51.9 \%$ |
| 42 | $51.9 \%$ |
| 67 | $51.9 \%$ |
| 71 | $52.1 \%$ |
| 20 | $52.3 \%$ |
| 29 | $52.3 \%$ |
| 51 | $52.6 \%$ |
| 64 | $52.8 \%$ |
| 54 | $53.4 \%$ |
| 57 | $53.4 \%$ |
| 2 | $54.1 \%$ |
| 45 | $54.6 \%$ |

225. In the 2014 election environment the statewide vote for Democratic candidates for the Assembly fell 3.4 percentage points, from $51.4 \%$ down to $48.0 \%$.
226. On the criteria listed below, the Demonstration Plan performs as shown in the table below:

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|  |  | Demonstration Plan | Act 43 |
| :---: | :---: | :---: | :---: |
| Population Deviation |  | 0.86\% | 0.76\% |
| Average Compactness (Reock) |  | 0.41 | 0.39 |
| Number of Municipal Splits | County | 55 | 58 |
|  | City Town Village | 64 | 62 |

Mayer Rpt. (Dkt. 54) at 37.
227. The Demonstration Plan has a marginally larger population deviation than the Current Plan ( $0.86 \%$ versus $0.76 \%$ ), but is well below even the strictest standards applied to state legislative plans. Mayer Rpt. (Dkt. 54) at 37.
228. The Demonstration Plan's districts are slightly more compact on average than the Current Plan's, with an average Reock score of 0.41 , compared to 0.39 for the Current Plan. Mayer Rpt. (Dkt. 54) at 37.
229. The Demonstration Plan has one fewer municipal split than the Current Plan (119 versus 120). Mayer Rpt. (Dkt. 54) at 37.

## History of Elections in Wisconsin

230. The Government Accountability Board's official election results are authoritative for Wisconsin elections dating back to the year 2000.
231. For elections in years prior to 2000, the Wisconsin Blue Book's election results are authoritative.
232. The City of Milwaukee Election Commission maintains election results dating back to 1997 on its website. These results are authoritative for election results in the City of Milwaukee.
233. The following chart contains the number of seats won by Democratic, Republican and Independent candidates in the November general elections from 1972 to 2014. The party with the majority is listed in bold.

| Year | Democrat | Republican | Independent |
| :---: | :---: | :---: | :---: |
| 1972 | $\mathbf{6 2}$ | 37 |  |
| 1974 | $\mathbf{6 3}$ | 36 |  |
| 1976 | $\mathbf{6 6}$ | 33 |  |
| 1978 | $\mathbf{6 0}$ | 39 |  |
| 1980 | $\mathbf{5 9}$ | 40 |  |
| 1982 | $\mathbf{5 9}$ | 40 |  |
| 1984 | $\mathbf{5 2}$ | 47 |  |
| 1986 | $\mathbf{5 4}$ | 45 |  |
| 1988 | $\mathbf{5 6}$ | 43 |  |
| 1990 | $\mathbf{5 8}$ | 41 |  |
| 1992 | $\mathbf{5 2}$ | 47 |  |
| 1994 | 48 | $\mathbf{5 1}$ |  |
| 1996 | 47 | $\mathbf{5 2}$ |  |


| 1998 | 44 | $\mathbf{5 5}$ |  |
| :--- | :--- | :--- | :--- |
| 2000 | 43 | $\mathbf{5 6}$ |  |
| 2002 | 41 | $\mathbf{5 8}$ |  |
| 2004 | 39 | $\mathbf{6 0}$ |  |
| 2006 | 47 | $\mathbf{5 2}$ |  |
| 2008 | $\mathbf{5 2}$ | 46 | 1 |
| 2010 | 38 | $\mathbf{6 0}$ | 1 |
| 2012 | 39 | $\mathbf{6 0}$ |  |
| 2014 | 36 | $\mathbf{6 3}$ |  |

234. The Democrats won a majority of seats in the Wisconsin Assembly in each general election from 1972 through 1994.
235. The Republicans won a majority of seats in the Wisconsin Assembly in each general election from 1994 through 2014, with the exception of the 2008 election.
236. The Assembly map in place for the 1972 , 1974, 1976, 1978 and 1980 plans was enacted by the Democratic Assembly and Republican Senate and signed by a Democratic Governor.
237. The Assembly map in place for the 1982 election was put in place by the federal court in Wisconsin State AFL-CIO v. Elections Bd., 543 F. Supp. 630 (E.D. Wis. 1982).
238. The Assembly map in place for the 1982 election was amended and enacted by the Democratic Assembly and Democratic Senate and signed by a

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Democratic Governor and was then in place for the 1984, 1986, 1988 and 1990 elections.
239. The Assembly map in place for the 1992 , 1994, 1996, 1998 and 2000 elections was drawn by the federal court in Prosser v. Elections Board, 793 F. Supp. 859 (W.D. Wis. 1992).
240. The Assembly map in place for the 2002, 2004, 2006, 2008 and 2010 elections was drawn by the federal court in Baumgart v. Wendelberger, No. 01-C-0121, 2002 WL 34127471, at *1 (E.D. Wis. May 30, 2002), amended, 2002 WL 34127473 (E.D. Wis. July 11, 2002).
241. Professor Jackman analyzed each Wisconsin Assembly elections since 1972 and found that Wisconsin's $E G$ has ranged from a high (most favorable to Democrats) of $+2.48 \%$ in 1994 to a low (most favorable to Republicans) of $-13.31 \%$ in 2012.
242. Disregarding results from the current plan, the lowest $E G$ was $-11.83 \%$ in 2006.
243. The most favorable $E G$ towards Democrats notably occurred in 1994 when the Republicans gained control of the Assembly for the first time since the 1968 election.
244. Professor Jackman finds that "Wisconsin has recorded an unbroken run of negative $E G$ estimates from 1998 to 2014."
245. The last positive $E G$ that Professor Jackman found in Wisconsin was the $2.48 \%$ from 1994.
246. With respect to the 2002 Plan, Professor Jackman calculated an average efficiency gap of $7.6 \%$, with $-4.0 \%$ as the most favorable year to Democrats and $-11.8 \%$ as the most favorable year to Republicans.
247. In 1992, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $52.5 \%$. Given that Professor Jackman calculates an $E G$ of $-2 \%$, the Democratic vote share was $52.25 \%$ because the implied seat share if the efficiency gap was zero is 54.5\%.
248. In 1994, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $48.5 \%$. Given that Professor Jackman calculates an $E G$ of $+2 \%$, the Democratic vote share was $48.25 \%$ because the implied seat share if the efficiency gap was zero is 46.5\%.
249. In 1996, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $47.5 \%$. Given that Professor Jackman calculates an $E G$ of $0 \%$, the Democratic vote share was $48.75 \%$ because the implied seat share if the efficiency gap was zero is 47.5\%.
250. In 1998, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $44.5 \%$. Given that Professor Jackman calculates an $E G$ of $-7.5 \%$, the Democratic vote share was $51 \%$ because the implied seat share if the efficiency gap was zero is $52 \%$.
251. In 2000, the Democrats' seat share, rounded to the nearest $0.25 \%$, was $43.5 \%$. Given that Professor Jackman calculates an $E G$ of $-6 \%$, the Democratic vote share was $49.75 \%$ because the implied seat share if the efficiency gap was zero is 49.5\%.
252. In 2002, the Democrats'seat, share rounded to the nearest $0.25 \%$, was $41.5 \%$.Given that Professor Jackman calculates an EG of $-7.5 \%$, the Democratic vote share was $49.5 \%$ because the implied seat share if the efficiency gap was zero is $49 \%$.
253. In 2004, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $40 \%$. Given that Professor Jackman calculates an $E G$ of $-10 \%$, the Democratic vote share was $50 \%$ because the implied seat share if the efficiency gap was zero is $50 \%$.
254. In 2006, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $47.5 \%$. Given that Professor Jackman calculates an $E G$ of $-12 \%$, the Democratic vote share was $54.75 \%$ because the implied seat share if the efficiency gap was zero is 59.5\%.
255. In 2008, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $53 \%$. Given that Professor Jackman calculates an $E G$ of $-5 \%$, the Democratic vote share was $54 \%$ because the implied seat share if the efficiency gap was zero is $58 \%$.
256. In 2010, the Democrats'seat share, rounded to the nearest $0.25 \%$, was $39 \%$. Given that Professor Jackman calculates an $E G$ of $-4 \%$, the Democratic vote share was $46.5 \%$ because the implied seat share if the efficiency gap was zero is $43 \%$.
257. In 2012, Professor Jackman calculates that the Democrats' vote share was $51.4 \%$. This yields an implied seat share of $52.8 \%$ if the efficiency gap was zero. The Democrats' actual seat share was $39.4 \%$, yielding an efficiency gap of $-13.4 \%$.
258. In 2014, Professor Jackman calculates that the Democrats' vote share was $48.0 \%$. This yields an implied seat share of $46.0 \%$ if the efficiency gap was zero. Their actual seat share was $36.4 \%$, which yields an efficiency gap of $-9.6 \%$.
259. In 1988, Michael Dukakis, the Democratic candidate for President, won $1,126,794$ votes in Wisconsin to Republican George H.W. Bush's $1,047,499$ votes, winning $51.8 \%$ of the two-party vote.
260. In the presidential election nationwide, George H.W. Bush won 53.9\% of the two-party vote and Dukakis won 46.1\%.

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261. The following chart shows the vote totals for Dukakis and Bush in each county in Wisconsin.

| County | Dukakis <br> Vote | Bush Vote | Two Party <br> Total |
| :--- | ---: | ---: | ---: |
| Adams | 3,598 | 3,258 | 6,856 |
| Ashland | 4,526 | 2,926 | 7,452 |
| Barron | 8,951 | 8,527 | 17,478 |
| Bayfield | 4,323 | 3,095 | 7,418 |
| Brown | 41,788 | 43,625 | 85,413 |
| Buffalo | 3,481 | 2,783 | 6,264 |
| Burnett | 3,537 | 2,884 | 6,421 |
| Calumet | 6,481 | 8,107 | 14,588 |
| Chippewa | 11,447 | 9,757 | 21,204 |
| Clark | 6,642 | 6,296 | 12,938 |
| Columbia | 9,132 | 10,475 | 19,607 |
| Crawford | 3,608 | 3,238 | 6,846 |
| Dane | 105,414 | 69,143 | 174,557 |
| Dodge | 12,663 | 17,003 | 29,666 |
| Door | 5,425 | 6,907 | 12,332 |
| Douglas | 13,907 | 6,440 | 20,347 |
| Dunn | 9,205 | 7,273 | 16,478 |
| Eau Claire | 21,150 | 17,664 | 38,814 |
| Florence | 1,018 | 1,106 | 2,124 |
| Fond du Lac | 15,887 | 21,985 | 37,872 |
| Forest | 2,142 | 1,845 | 3,987 |
| Grant | 9,421 | 10,049 | 19,470 |
| Green | 5,153 | 6,636 | 11,789 |
| Green Lake | 3,033 | 5,205 | 8,238 |
| Iowa | 4,268 | 4,240 | 8,508 |
|  |  |  |  |

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| County | Dukakis <br> Vote | Bush Vote | Two Party <br> Total |
| :--- | ---: | ---: | ---: |
| Iron | 2,090 | 1,599 | 3,689 |
| Jackson | 3,924 | 3,555 | 7,479 |
| Jefferson | 11,816 | 14,309 | 26,125 |
| Juneau | 3,734 | 4,869 | 8,603 |
| Kenosha | 30,089 | 21,661 | 51,750 |
| Kewaunee | 4,786 | 4,330 | 9,116 |
| La Crosse | 22,204 | 21,548 | 43,752 |
| Lafayette | 3,521 | 3,665 | 7,186 |
| Langlade | 4,254 | 4,884 | 9,138 |
| Lincoln | 5,819 | 5,257 | 11,076 |
| Manitowoc | 19,680 | 16,020 | 35,700 |
| Marathon | 24,658 | 24,482 | 49,140 |
| Marinette | 8,030 | 9,637 | 17,667 |
| Marquette | 2,463 | 3,059 | 5,522 |
| Menominee | 1,028 | 381 | 1,409 |
| Milwaukee | 268,287 | 168,363 | 436,650 |
| Monroe | 6,437 | 7,073 | 13,510 |
| Oconto | 6,549 | 7,084 | 13,633 |
| Oneida | 7,414 | 8,130 | 15,544 |
| Outagamie | 27,771 | 33,113 | 60,884 |
| Ozaukee | 12,661 | 22,899 | 35,560 |
| Pepin | 1,906 | 1,311 | 3,217 |
| Pierce | 8,659 | 6,045 | 14,704 |
| Polk | 8,981 | 6,866 | 15,847 |
| Portage | 16,317 | 12,057 | 28,374 |
| Price | 3,987 | 3,450 | 7,437 |
| Racine | 39,631 | 36,342 | 75,973 |
|  |  |  |  |

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| County | Dukakis <br> Vote | Bush Vote | Two Party <br> Total |
| :--- | ---: | ---: | ---: |
| Richland | 3,643 | 4,026 | 7,669 |
| Rock | 29,576 | 28,178 | 57,754 |
| Rusk | 3,888 | 3,063 | 6,951 |
| St. Croix | 11,392 | 9,960 | 21,352 |
| Sauk | 8,324 | 10,225 | 18,549 |
| Sawyer | 3,231 | 3,260 | 6,491 |
| Shawano | 6,587 | 8,362 | 14,949 |
| Sheboygan | 23,429 | 23,471 | 46,900 |
| Taylor | 3,785 | 4,254 | 8,039 |
| Trempealeau | 6,212 | 4,902 | 11,114 |
| Vernon | 5,754 | 5,226 | 10,980 |
| Vilas | 3,781 | 5,842 | 9,623 |
| Walworth | 12,203 | 18,259 | 30,462 |
| Washburn | 3,393 | 3,074 | 6,467 |
| Washington | 15,907 | 24,328 | 40,235 |
| Waukesha | 57,598 | 90,467 | 148,065 |
| Waupaca | 7,078 | 11,559 | 18,637 |
| Waushara | 3,535 | 4,953 | 8,488 |
| Winnebago | 28,508 | 35,085 | 63,593 |
| Wood | 16,074 | 16,549 | 32,623 |
|  | $1,126,794$ | $1,047,499$ | $2,174,293$ |

262. In 1992, Bill Clinton, the Democratic candidate for President, won 1,041,066 votes in Wisconsin to Republican George H.W. Bush's 930,855 , winning $52.8 \%$ of the two-party vote share.

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263. In the presidential election nationwide, Clinton won $53.5 \%$ of the two-party vote share to Bush's 46.5\%.
264. The following chart shows the vote totals for Clinton and Bush in each county in Wisconsin.

| County | Clinton <br> Vote | Bush <br> Vote | Two Party <br> Total |
| :--- | ---: | ---: | ---: |
| Adams | 3,539 | 2,465 | 6,004 |
| Ashland | 4,213 | 2,372 | 6,585 |
| Barron | 8,063 | 6,572 | 14,635 |
| Bayfield | 3,873 | 2,393 | 6,266 |
| Brown | 37,513 | 42,352 | 79,865 |
| Buffalo | 2,996 | 2,029 | 5,025 |
| Burnett | 3,172 | 2,340 | 5,512 |
| Calumet | 5,701 | 7,541 | 13,242 |
| Chippewa | 10,487 | 8,215 | 18,702 |
| Clark | 5,540 | 4,977 | 10,517 |
| Columbia | 9,348 | 9,099 | 18,447 |
| Crawford | 3,540 | 2,390 | 5,930 |
| Dane | 114,724 | 61,957 | 176,681 |
| Dodge | 11,438 | 14,971 | 26,409 |
| Door | 4,735 | 5,468 | 10,203 |
| Douglas | 12,319 | 5,679 | 17,998 |
| Dunn | 7,965 | 5,283 | 13,248 |
| Eau Claire | 21,221 | 15,915 | 37,136 |
| Florence | 978 | 942 | 1,920 |
| Fond du Lac | 13,757 | 19,785 | 33,542 |

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| County | Clinton <br> Vote | Bush <br> Vote | Two Party <br> Total |
| :--- | ---: | ---: | ---: |
| Forest | 1,904 | 1,393 | 3,297 |
| Grant | 8,914 | 7,678 | 16,592 |
| Green | 5,467 | 4,887 | 10,354 |
| Green Lake | 2,772 | 3,897 | 6,669 |
| Iowa | 4,467 | 3,288 | 7,755 |
| Iron | 1,762 | 1,273 | 3,035 |
| Jackson | 3,681 | 2,644 | 6,325 |
| Jefferson | 11,593 | 13,072 | 24,665 |
| Juneau | 4,177 | 4,051 | 8,228 |
| Kenosha | 27,341 | 19,854 | 47,195 |
| Kewaunee | 4,050 | 3,570 | 7,620 |
| La Crosse | 22,838 | 18,891 | 41,729 |
| Lafayette | 3,143 | 2,582 | 5,725 |
| Langlade | 3,630 | 3,890 | 7,520 |
| Lincoln | 5,297 | 4,321 | 9,618 |
| Manitowoc | 15,903 | 14,008 | 29,911 |
| Marathon | 21,482 | 20,948 | 42,430 |
| Marinette | 7,626 | 7,984 | 15,610 |
| Marquette | 2,533 | 2,322 | 4,855 |
| Menominee | 691 | 244 | 935 |
| Milwaukee | 235,521 | 151,314 | 386,835 |
| Monroe | 6,427 | 6,118 | 12,545 |
| Oconto | 5,898 | 5,720 | 11,618 |
| Oneida | 7,160 | 6,725 | 13,885 |
| Outagamie | 23,735 | 30,370 | 54,105 |
| Ozaukee | 11,879 | 22,805 | 34,684 |
| Pepin | 1,673 | 1,098 | 2,771 |

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| County | Clinton <br> Vote | Bush <br> Vote | Two Party <br> Total |
| :--- | ---: | ---: | ---: |
| Pierce | 7,824 | 4,844 | 12,668 |
| Polk | 7,746 | 5,446 | 13,192 |
| Portage | 15,553 | 10,914 | 26,467 |
| Price | 3,575 | 2,654 | 6,229 |
| Racine | 34,875 | 32,310 | 67,185 |
| Richland | 3,458 | 3,144 | 6,602 |
| Rock | 31,154 | 21,942 | 53,096 |
| Rusk | 3376 | 2,430 | 3,376 |
| St. Croix | 10281 | 8,114 | 10,281 |
| Sauk | 9128 | 8,886 | 9,128 |
| Sawyer | 2796 | 2,658 | 2,796 |
| Shawano | 6,062 | 7,253 | 13,315 |
| Sheboygan | 20,568 | 22,526 | 43,094 |
| Taylor | 3,305 | 3,415 | 6,720 |
| Trempealau | 6,218 | 3,577 | 9,795 |
| Vernon | 5,673 | 4,072 | 9,745 |
| Vilas | 3,764 | 4,616 | 8,380 |
| Walworth | 11,825 | 15,727 | 27,552 |
| Washburn | 3,080 | 2,586 | 5,666 |
| Washington | 13,339 | 22,739 | 36,078 |
| Waukesha | 50,270 | 91,461 | 141,731 |
| Waupaca | 6,666 | 10,252 | 16,918 |
| Waushara | 3,402 | 4,045 | 7,447 |
| Winnebago | 27,234 | 33,709 | 60,943 |
| Wood | 13,208 | 13,843 | 27,051 |
|  | $1,041,066$ | 930,855 | $1,971,921$ |

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265. In 1996, Bill Clinton, the Democratic candidate for President, won $1,071,971$ votes in Wisconsin to Republican Bob Dole's 845,029 votes, winning $55.9 \%$ of the two-party vote share.
266. In the presidential election nationwide, Clinton won $54.7 \%$ of the two-party vote to Dole's 45.3\%.
267. Bill Clinton won Milwaukee, Dane and Rock Counties with $64 \%$ of the two-party vote and carried the rest of the state with $52 \%$ of the vote, a difference of twelve percentage points.
268. The following chart shows the vote totals for Clinton and Dole in each county in Wisconsin.

| County | Clinton <br> Vote | Dole <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Adams | 4,119 | 2,450 | 6,569 |
| Ashland | 3,808 | 1,863 | 5,671 |
| Barron | 8,025 | 6,158 | 14,183 |
| Bayfield | 3,895 | 2,250 | 6,145 |
| Brown | 42,823 | 38,563 | 81,386 |
| Buffalo | 2,681 | 1,800 | 4,481 |
| Burnett | 3,625 | 2,452 | 6,077 |
| Calumet | 6,940 | 7,049 | 13,989 |
| Chippewa | 9,647 | 7,520 | 17,167 |
| Clark | 5,540 | 4,622 | 10,162 |
| Columbia | 10,336 | 8,377 | 18,713 |

JA232

| County | Clinton <br> Vote | Dole <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Crawford | 3,658 | 2,149 | 5,807 |
| Dane | 109,347 | 59,487 | 168,834 |
| Dodge | 12,625 | 12,890 | 25,515 |
| Door | 5,590 | 4,948 | 10,538 |
| Douglas | 10,976 | 5,167 | 16,143 |
| Dunn | 7,536 | 4,917 | 12,453 |
| Eau Claire | 20,298 | 13,900 | 34,198 |
| Florence | 869 | 927 | 1,796 |
| Fond du Lac | 15,542 | 16,488 | 32,030 |
| Forest | 2,092 | 1,166 | 3,258 |
| Grant | 9,203 | 7,021 | 16,224 |
| Green | 6,136 | 4,697 | 10,833 |
| Green Lake | 3,152 | 3,565 | 6,717 |
| Iowa | 4,690 | 2,866 | 7,556 |
| Iron | 1,725 | 1,260 | 2,985 |
| Jackson | 3,705 | 2,262 | 5,967 |
| Jefferson | 13,188 | 12,681 | 25,869 |
| Juneau | 4,331 | 3,226 | 7,557 |
| Kenosha | 27,964 | 18,296 | 46,260 |
| Kewaunee | 4,311 | 3,431 | 7,742 |
| La Crosse | 23,647 | 16,482 | 40,129 |
| Lafayette | 3,261 | 2,172 | 5,433 |
| Langlade | 4,074 | 3,206 | 7,280 |
| Lincoln | 6,166 | 4,076 | 10,242 |
| Manitowoc | 16,750 | 13,239 | 29,989 |
| Marathon | 24,012 | 19,874 | 43,886 |

JA233

| County | Clinton <br> Vote | Dole <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Marinette | 2,859 | 2,208 | 5,067 |
| Marquette | 2,859 | 2,208 | 5,067 |
| Menominee | 992 | 230 | 1,222 |
| Milwaukee | 216,620 | 119,407 | 336,027 |
| Monroe | 6,924 | 5,299 | 12,223 |
| Oconto | 6,723 | 5,389 | 12,112 |
| Oneida | 7,619 | 6,339 | 13,958 |
| Outagamie | 28,815 | 27,758 | 56,573 |
| Ozaukee | 13,269 | 22,078 | 35,347 |
| Pepin | 1,585 | 1,007 | 2,592 |
| Pierce | 7,970 | 4,599 | 12,569 |
| Polk | 8,334 | 5,387 | 13,721 |
| Portage | 15,901 | 9,631 | 25,532 |
| Price | 3,523 | 2,545 | 6,068 |
| Racine | 38,567 | 30,107 | 68,674 |
| Richland | 3,502 | 2,642 | 6,144 |
| Rock | 32,450 | 20,096 | 52,546 |
| Rusk | 2941 | 2,219 | 2,941 |
| St. Croix | 11384 | 8,253 | 11,384 |
| Sauk | 9889 | 7,448 | 9,889 |
| Sawyer | 2773 | 2,603 | 2,773 |
| Shawano | 6,850 | 6,396 | 13,246 |
| Sheboygan | 22,022 | 20,067 | 42,089 |
| Taylor | 3,253 | 3,108 | 6,361 |
| Trempealau | 5,848 | 3,035 | 8,883 |
| Vernon | 5,572 | 3,796 | 9,368 |
|  |  |  |  |

JA234

| County | Clinton <br> Vote | Dole <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Vilas | 4,226 | 4,496 | 8,722 |
| Walworth | 13,283 | 15,099 | 28,382 |
| Washburn | 3,231 | 2,703 | 5,934 |
| Washington | 17,154 | 25,829 | 42,983 |
| Waukesha | 57,354 | 91,729 | 149,083 |
| Waupaca | 7,800 | 8,679 | 16,479 |
| Waushara | 3,824 | 3,573 | 7,397 |
| Winnebago | 29,564 | 27,880 | 57,444 |
| Wood | 14,650 | 12,666 | 27,316 |
|  | $1,071,971$ | 845,029 | $1,917,000$ |

269. In 2000, Albert Gore, the Democratic candidate for President, won $1,242,987$ votes in Wisconsin to Republican George W. Bush's 1,237,279 votes, winning $50.1 \%$ of the two-party vote.
270. In the presidential election nationwide, Gore won $50.27 \%$ of the two-party vote to Bush's 49.73\%.
271. The following chart shows the vote totals for Gore and Bush in each county in Wisconsin, as well as a subtotal for votes in the City of Milwaukee.

JA235

| County | Gore <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Adams | 4,826 | 3,920 | 8,746 |
| Ashland | 4,356 | 3,038 | 7,394 |
| Barron | 8,928 | 9,848 | 18,776 |
| Bayfield | 4,427 | 3,266 | 7,693 |
| Brown | 49,096 | 54,258 | 103,354 |
| Buffalo | 3,237 | 3,038 | 6,275 |
| Burnett | 3,626 | 3,967 | 7,593 |
| Calumet | 8,202 | 10,837 | 19,039 |
| Chippewa | 12,102 | 12,835 | 24,937 |
| Clark | 5,931 | 7,461 | 13,392 |
| Columbia | 12,636 | 11,987 | 24,623 |
| Crawford | 4,005 | 3,024 | 7,029 |
| Dane | 142,317 | 75,790 | 218,107 |
| Dodge | 14,580 | 21,684 | 36,264 |
| Door | 6,560 | 7,810 | 14,370 |
| Douglas | 13,593 | 6,930 | 20,523 |
| Dunn | 9,172 | 8,911 | 18,083 |
| Eau Claire | 24,078 | 20,921 | 44,999 |
| Florence | 816 | 1,528 | 2,344 |
| Fond du Lac | 18,181 | 26,548 | 44,729 |
| Forest | 2,158 | 2,404 | 4,562 |
| Grant | 10,691 | 10,240 | 20,931 |
| Green | 7,863 | 6,790 | 14,653 |
| Green Lake | 3,301 | 5,451 | 8,752 |
| Iowa | 5,842 | 4,221 | 10,063 |
| Iron | 1,620 | 1,734 | 3,354 |
|  |  |  |  |

JA236

| County | Gore <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Jackson | 4,380 | 3,670 | 8,050 |
| Jefferson | 15,203 | 19,204 | 34,407 |
| Juneau | 4,813 | 4,910 | 9,723 |
| Kenosha | 32,429 | 28,891 | 61,320 |
| Kewaunee | 4,670 | 4,883 | 9,553 |
| La Crosse | 28,455 | 24,327 | 52,782 |
| Lafayette | 3,710 | 3,336 | 7,046 |
| Langlade | 4,199 | 5,125 | 9,324 |
| Lincoln | 6,664 | 6,727 | 13,391 |
| Manitowoc | 17,667 | 19,358 | 37,025 |
| Marathon | 26,546 | 28,883 | 55,429 |
| Marinette | 8,676 | 10,535 | 19,211 |
| Marquette | 3,437 | 3,522 | 6,959 |
| Menominee | 949 | 225 | 1,174 |
| Milwaukee | 252,329 | 163,491 | 415,820 |
| City of |  |  |  |
| Milwaukee | 165,598 | 69,075 | 234,673 |
| subtotal |  |  |  |
| Monroe | 7,460 | 8,217 | 15,677 |
| Oconto | 7,260 | 8,706 | 15,966 |
| Oneida | 8,339 | 9,512 | 17,851 |
| Outagamie | 32,735 | 39,460 | 72,195 |
| Ozaukee | 15,030 | 31,155 | 46,185 |
| Pepin | 1,854 | 1,631 | 3,485 |
| Pierce | 8,559 | 8,169 | 16,728 |
| Polk | 8,961 | 9,557 | 18,518 |
|  |  |  |  |

JA237

| County | Gore <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Portage | 17,942 | 13,214 | 31,156 |
| Price | 3,413 | 4,136 | 7,549 |
| Racine | 41,563 | 44,014 | 85,577 |
| Richland | 3,837 | 3,994 | 7,831 |
| Rock | 40,472 | 27,467 | 67,939 |
| Rusk | 3161 | 3,758 | 3,161 |
| St. Croix | 13077 | 15,240 | 13,077 |
| Sauk | 33035 | 11,586 | 13,035 |
| Sawyer | 7,335 | 3,972 | 3,333 |
| Shawano | 23,569 | 29,648 | 16,883 |
| Sheboygan | 3,254 | 5,278 | 53,217 |
| Taylor | 6,678 | 5,002 | 11,680 |
| Trempealau | 6,577 | 5,684 | 12,261 |
| Vernon | 4,706 | 6,958 | 11,664 |
| Vilas | 15,492 | 22,982 | 38,474 |
| Walworth | 3,695 | 3,912 | 7,607 |
| Washburn | 18,115 | 41,162 | 59,277 |
| Washington | 64,319 | 133,105 | 197,424 |
| Waukesha | 8,787 | 12,980 | 21,767 |
| Waupaca | 4,239 | 5,571 | 9,810 |
| Waushara | 33,983 | 38,330 | 72,313 |
| Winnebago | 15,936 | 17,803 | 33,739 |
| Wood | $1,242,987$ | $1,237,279$ | $2,480,266$ |
|  |  |  |  |

272. In 2004, John Kerry, the Democratic candidate for President, won $1,489,504$ votes in Wisconsin to Republican George W. Bush's 1,478, 120 votes, winning $50.2 \%$ of the two-party vote.
273. In the presidential election nationwide, Bush won $51.24 \%$ of the two-party vote to Kerry's 48.76\%.
274. The following chart shows the vote totals for Kerry and Bush in each county in Wisconsin, along with a subtotal for votes in the City of Milwaukee.

| County | Kerry <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Adams | 5,447 | 4,890 | 10,337 |
| Ashland | 5,805 | 3,313 | 9,118 |
| Barron | 11,696 | 12,030 | 23,726 |
| Bayfield | 5,845 | 3,754 | 9,599 |
| Brown | 54,935 | 67,173 | 122,108 |
| Buffalo | 3,998 | 3,502 | 7,500 |
| Burnett | 4,499 | 4,743 | 9,242 |
| Calumet | 10,290 | 14,721 | 25,011 |
| Chippewa | 14,751 | 15,450 | 30,201 |
| Clark | 6,966 | 7,966 | 14,932 |
| Columbia | 14,300 | 14,956 | 29,256 |
| Crawford | 4,656 | 3,680 | 8,336 |
| Dane | 181,052 | 90,369 | 271,421 |
| Dodge | 16,690 | 27,201 | 43,891 |
| Door | 8,367 | 8,910 | 17,277 |

JA239

| County | Kerry <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Douglas | 16,537 | 8,448 | 24,985 |
| Dunn | 12,039 | 10,879 | 22,918 |
| Eau Claire | 30,068 | 24,653 | 54,721 |
| Florence | 993 | 1,703 | 2,696 |
| Fond du Lac | 19,216 | 33,291 | 52,507 |
| Forest | 2,509 | 2,608 | 5,117 |
| Grant | 12,864 | 12,208 | 25,072 |
| Green | 9,575 | 8,497 | 18,072 |
| Green Lake | 3,605 | 6,472 | 10,077 |
| Iowa | 7,122 | 5,348 | 12,470 |
| Iron | 1,956 | 1,884 | 3,840 |
| Jackson | 5,249 | 4,387 | 9,636 |
| Jefferson | 17,925 | 23,776 | 41,701 |
| Juneau | 5,734 | 6,473 | 12,207 |
| Kenosha | 40,107 | 35,587 | 75,694 |
| Kewaunee | 5,175 | 5,970 | 11,145 |
| La Crosse | 33,170 | 28,289 | 61,459 |
| Lafayette | 4,402 | 3,929 | 8,331 |
| Langlade | 4,751 | 6,235 | 10,986 |
| Lincoln | 7,484 | 8,024 | 15,508 |
| Manitowoc | 20,652 | 23,027 | 43,679 |
| Marathon | 30,899 | 36,394 | 67,293 |
| Marinette | 10,190 | 11,866 | 22,056 |
| Marquette | 3,785 | 4,604 | 8,389 |
| Menominee | 1,412 | 288 | 1,700 |
| Milwaukee | 297,653 | 180,287 | 477,940 |

JA240

| County | Kerry <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :---: | ---: | ---: | ---: |
| City of <br> Milwaukee <br> subtotal | 198,907 | 75,746 | 274,653 |
| Monroe | 8,973 | 10,375 | 19,348 |
| Oconto | 8,534 | 11,043 | 19,577 |
| Oneida | 10,464 | 11,351 | 21,815 |
| Outagamie | 40,169 | 48,903 | 89,072 |
| Ozaukee | 17,714 | 34,904 | 52,618 |
| Pepin | 2,181 | 1,853 | 4,034 |
| Pierce | 11,176 | 10,437 | 21,613 |
| Polk | 11,173 | 12,095 | 23,268 |
| Portage | 21,861 | 16,546 | 38,407 |
| Price | 4,349 | 4,312 | 8,661 |
| Racine | 48,229 | 52,456 | 100,685 |
| Richland | 4,501 | 4,836 | 9,337 |
| Rock | 46,598 | 33,151 | 79,749 |
| Rusk | 3820 | 3,985 | 3,820 |
| St. Croix | 18784 | 22,679 | 18,784 |
| Sauk | 15708 | 14,415 | 15,708 |
| Sawyer | 4411 | 4,951 | 4,411 |
| Shawano | 8,657 | 12,150 | 20,807 |
| Sheboygan | 27,608 | 34,458 | 62,066 |
| Taylor | 3,829 | 5,582 | 9,411 |
| Trempealau | 8,075 | 5,878 | 13,953 |
| Vernon | 7,924 | 6,774 | 14,698 |
| Vilas | 5,713 | 8,155 | 13,868 |

JA241

| County | Kerry <br> Vote | Bush <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Walworth | 19,177 | 28,754 | 47,931 |
| Washburn | 4,705 | 4,762 | 9,467 |
| Washington | 21,234 | 50,641 | 71,875 |
| Waukesha | 73,626 | 154,926 | 228,552 |
| Waupaca | 10,792 | 15,941 | 26,733 |
| Waushara | 5,257 | 6,888 | 12,145 |
| Winnebago | 40,943 | 46,542 | 87,485 |
| Wood | 18,950 | 20,592 | 39,542 |
|  | $1,489,504$ | $1,478,120$ | $2,967,624$ |

275. In 2008, Barack Obama, the Democratic candidate for President, won 1,677,211 votes in Wisconsin to Republican John McCain's 1,262,393 votes, winning $57.05 \%$ of the two- party vote.
276. In the presidential election nationwide, Obama won $53.69 \%$ of the two-party vote to McCain's 46.31\%.
277. The following chart shows the vote totals for Obama and McCain in each county in Wisconsin including a subtotal of votes in the City of Milwaukee.

| County | Obama <br> Vote | McCain <br> Vote | Two <br> Party <br> Total |
| :---: | ---: | ---: | ---: |
| Adams | 5,806 | 3,974 | 9,780 |
| Ashland | 5,818 | 2,634 | 8,452 |

JA242

| County | Obama <br> Vote | McCain <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Barron | 12,078 | 10,457 | 22,535 |
| Bayfield | 5,972 | 3,365 | 9,337 |
| Brown | 67,269 | 55,854 | 123,123 |
| Buffalo | 3,949 | 2,923 | 6,872 |
| Burnett | 4,337 | 4,200 | 8,537 |
| Calumet | 13,295 | 12,722 | 26,017 |
| Chippewa | 16,239 | 13,492 | 29,731 |
| Clark | 7,454 | 6,383 | 13,837 |
| Columbia | 16,661 | 12,193 | 28,854 |
| Crawford | 4,987 | 2,830 | 7,817 |
| Dane | 205,984 | 73,065 | 279,049 |
| Dodge | 19,183 | 23,015 | 42,198 |
| Door | 10,142 | 7,112 | 17,254 |
| Douglas | 15,830 | 7,835 | 23,665 |
| Dunn | 13,002 | 9,566 | 22,568 |
| Eau Claire | 33,146 | 20,959 | 54,105 |
| Florence | 1,134 | 1,512 | 2,646 |
| Fond du | 23,463 | 28,164 | 51,627 |
| Forest | 2,673 | 1,963 | 4,636 |
| Grant | 14,875 | 9,068 | 23,943 |
| Green | 11,502 | 6,730 | 18,232 |
| Green Lake | 4,000 | 5,393 | 9,393 |
| Iowa | 7,987 | 3,829 | 11,816 |
| Iron | 1,914 | 1,464 | 3,378 |
| Jackson | 5,572 | 3,552 | 9,124 |
| Jefferson | 21,448 | 21,096 | 42,544 |
|  |  |  |  |

JA243

| County | Obama Vote | McCain Vote | Two <br> Party <br> Total |
| :---: | :---: | :---: | :---: |
| Juneau | 6,186 | 5,148 | 11,334 |
| Kenosha | 45,836 | 31,609 | 77,445 |
| Kewaunee | 5,902 | 4,711 | 10,613 |
| La Crosse | 38,524 | 23,701 | 62,225 |
| Lafayette | 4,732 | 2,984 | 7,716 |
| Langlade | 5,182 | 5,081 | 10,263 |
| Lincoln | 8,424 | 6,519 | 14,943 |
| Manitowoc | 22,428 | 19,234 | 41,662 |
| Marathon | 36,367 | 30,345 | 66,712 |
| Marinette | 11,195 | 9,726 | 20,921 |
| Marquette | 4,068 | 3,654 | 7,722 |
| Menominee | 1,257 | 185 | 1,442 |
| Milwaukee | 319,819 | 149,445 | 469,264 |
| City of Milwaukee subtotal | 213,436 | 57,665 | 271,101 |
| Monroe | 10,198 | 8,666 | 18,864 |
| Oconto | 9,927 | 8,755 | 18,682 |
| Oneida | 11,907 | 9,630 | 21,537 |
| Outagamie | 50,294 | 39,677 | 89,971 |
| Ozaukee | 20,579 | 37,172 | 57,751 |
| Pepin | 2,102 | 1,616 | 3,718 |
| Pierce | 11,803 | 9,812 | 21,615 |
| Polk | 10,876 | 11,282 | 22,158 |
| Portage | 24,817 | 13,810 | 38,627 |
| Price | 4,559 | 3,461 | 8,020 |

JA244

| County | Obama <br> Vote | McCain <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Racine | 53,408 | 45,954 | 99,362 |
| Richland | 5,041 | 3,298 | 8,339 |
| Rock | 50,529 | 27,364 | 77,893 |
| Rusk | 3855 | 3,253 | 3,855 |
| St. Croix | 21177 | 22,837 | 21,177 |
| Sauk | 18617 | 11,562 | 18,617 |
| Sawyer | 4765 | 4,199 | 4,765 |
| Shawano | 10,259 | 9,538 | 19,797 |
| Sheboygan | 30,395 | 30,801 | 61,196 |
| Taylor | 4,563 | 4,586 | 9,149 |
| Trempealau | 8,321 | 4,808 | 13,129 |
| Vernon | 8,463 | 5,367 | 13,830 |
| Vilas | 6,491 | 7,055 | 13,546 |
| Walworth | 24,177 | 25,485 | 49,662 |
| Washburn | 4,693 | 4,303 | 8,996 |
| Washington | 25,719 | 47,729 | 73,448 |
| Waukesha | 85,339 | 145,152 | 230,491 |
| Waupaca | 12,952 | 12,232 | 25,184 |
| Waushara | 5,868 | 5,770 | 11,638 |
| Winnebago | 48,167 | 37,946 | 86,113 |
| Wood | 21,710 | 16,581 | 38,291 |
|  | $1,677,211$ | $1,267,393$ | $2,944,604$ |

278. In 2008, Democratic candidates for the Assembly ran about three points behind Obama in the statewide two-party vote.
279. In 2012, Barack Obama, the Democratic candidate for President, won $1,620,985$ votes in Wisconsin to Republican Mitt Romney's 1,407,966 votes, winning $53.5 \%$ of the two- party vote.
280. In the presidential election nationwide, Obama won $51.96 \%$ of the two-party vote to Romney's 48.04\%.
281. The following chart shows the vote totals for Obama and Romney in each county in Wisconsin along with a subtotal for the votes in the City of Milwaukee.

| County | Obama <br> Vote | Romney <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Adams | 5,542 | 4,644 | 10,186 |
| Ashland | 5,399 | 2,820 | 8,219 |
| Barron | 10,890 | 11,443 | 22,333 |
| Bayfield | 6,033 | 3,603 | 9,636 |
| Brown | 62,526 | 64,836 | 127,362 |
| Buffalo | 3,570 | 3,364 | 6,934 |
| Burnett | 3,986 | 4,550 | 8,536 |
| Calumet | 11,489 | 14,539 | 26,028 |
| Chippewa | 15,237 | 15,322 | 30,559 |
| Clark | 6,172 | 7,412 | 13,584 |
| Columbia | 17,175 | 13,026 | 30,201 |
| Crawford | 4,629 | 3,067 | 7,696 |
| Dane | 216,071 | 83,644 | 299,715 |

JA246

| County | Obama Vote | Romney Vote | Two <br> Party <br> Total |
| :---: | :---: | :---: | :---: |
| Dodge | 18,762 | 25,211 | 43,973 |
| Door | 9,357 | 8,121 | 17,478 |
| Douglas | 14,863 | 7,705 | 22,568 |
| Dunn | 11,316 | 10,224 | 21,540 |
| Eau Claire | 30,666 | 23,256 | 53,922 |
| Florence | 953 | 1,645 | 2,598 |
| Fond du Lac | 22,379 | 30,355 | 52,734 |
| Forest | 2,425 | 2,172 | 4,597 |
| Grant | 13,594 | 10,255 | 23,849 |
| Green | 11,206 | 7,857 | 19,063 |
| Green Lake | 3,793 | 5,782 | 9,575 |
| Iowa | 8,105 | 4,287 | 12,392 |
| Iron | 1,784 | 1,790 | 3,574 |
| Jackson | 5,298 | 3,900 | 9,198 |
| Jefferson | 20,158 | 23,517 | 43,675 |
| Juneau | 6,242 | 5,411 | 11,653 |
| Kenosha | 44,867 | 34,977 | 79,844 |
| Kewaunee | 5,153 | 5,747 | 10,900 |
| La Crosse | 36,693 | 25,751 | 62,444 |
| Lafayette | 4,536 | 3,314 | 7,850 |
| Langlade | 4,573 | 5,816 | 10,389 |
| Lincoln | 7,563 | 7,455 | 15,018 |
| Manitowoc | 20,403 | 21,604 | 42,007 |
| Marathon | 32,363 | 36,617 | 68,980 |
| Marinette | 9,882 | 10,619 | 20,501 |
| Marquette | 4,014 | 3,992 | 8,006 |

JA247

| County | Obama <br> Vote | Romney <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Menominee | 1,191 | 179 | 1,370 |
| Milwaukee | 332,438 | 154,924 | 487,362 |
| City of <br> Milwaukee <br> subtotal | 227,384 | 56,553 | 283,937 |
| Monroe | 9,515 | 9,675 | 19,190 |
| Oconto | 8,865 | 10,741 | 19,606 |
| Oneida | 10,452 | 10,917 | 21,369 |
| Outagamie | 45,659 | 47,372 | 93,031 |
| Ozaukee | 19,159 | 36,077 | 55,236 |
| Pepin | 1,876 | 1,794 | 3,670 |
| Pierce | 10,235 | 10,397 | 20,632 |
| Polk | 10,073 | 12,094 | 22,167 |
| Portage | 22,075 | 16,615 | 38,690 |
| Price | 3,887 | 3,884 | 7,771 |
| Racine | 53,008 | 49,347 | 102,355 |
| Richland | 4,969 | 3,573 | 8,542 |
| Rock | 49,219 | 30,517 | 79,736 |
| Rusk | 3397 | 3,676 | 3,397 |
| St. Croix | 19910 | 25,503 | 19,910 |
| Sauk | 18736 | 12,838 | 18,736 |
| Sawyer | 4486 | 4,442 | 4,486 |
| Shawano | 9,000 | 11,022 | 20,022 |
| Sheboygan | 27,918 | 34,072 | 61,990 |
| Taylor | 3,763 | 5,601 | 9,364 |
| Trempealau | 7,605 | 5,707 | 13,312 |
|  |  |  |  |

JA248

| County | Obama <br> Vote | Romney <br> Vote | Two <br> Party <br> Total |
| :--- | ---: | ---: | ---: |
| Vernon | 8,044 | 5,942 | 13,986 |
| Vilas | 5,951 | 7,749 | 13,700 |
| Walworth | 22,552 | 29,006 | 51,558 |
| Washburn | 4,447 | 4,699 | 9,146 |
| Washington | 23,166 | 54,765 | 77,931 |
| Waukesha | 78,779 | 162,798 | 241,577 |
| Waupaca | 11,578 | 14,002 | 25,580 |
| Waushara | 5,335 | 6,562 | 11,897 |
| Winnebago | 45,449 | 42,122 | 87,571 |
| Wood | 18,581 | 19,704 | 38,285 |
|  | $1,620,985$ | $1,407,966$ | $3,028,951$ |

282. In 2012, Obama won Milwaukee, Dane and Rock Counties with $69 \%$ of the two-party vote but won only $47 \%$ of the two-party vote in the rest of the state (to Mitt Romney's 53\%), a difference of twentytwo percentage points.
283. In the November 2010 election, Republican candidates won the Governor's office, a majority in the State Senate and retook the majority in the Assembly.
284. In the November 2010 election, Scott Walker won the Governor's office with $52.25 \%$ of the total vote ( $52.9 \%$ of the two-party vote).
285. In the November 2010 election, Republicans won 60 seats in the Assembly.
286. Professor Jackman calculates that the Republican candidates for the Assembly won $53.5 \%$ of the statewide two-party vote share in the November 2010 election.
287. On June 5, 2012, Governor Walker survived a recall attempt with $53.08 \%$ of the vote ( $53.4 \%$ of the two-party vote).
288. In November of 2012, President Obama won Wisconsin in the presidential election with $52.83 \%$ of the total vote ( $53.5 \%$ of the two-party vote).
289. Wisconsin's Democratic candidates for the Assembly ran about two points behind the President's vote share: Professor Jackman calculates that Democrats had a two-party vote share of $51.4 \%$.
290. In November of 2014, the Republicans increased their control of the Assembly by winning 63 seats, equating to a $63.6 \%$ seat share. Professor Jackman calculates that Republican candidates for the Assembly won $52 \%$ of the statewide two-party vote share in the November 2014 elections.
291. In 2010, Bob Ziegelbauer won assembly district 25 , and even though he ran as an independent, he typically voted with Republicans. Jason Stein \& Patrick Marley, More than They

Bargained For: Scott Walker, Unions, and the Fight for Wisconsin, Earle Decl. Ex. G (Dkt. 57-7) at 119.
292. Mr. Trende admitted that there are no "peer-reviewed studies that have analyzed the geographic clustering of Democratic and Republican voters by examining trends in counties won by each part[y's] presidential candidate." Trende Dep. (Dkt. 66) at 51:6-11.
293. Mr. Trende admitted that the maps he relied upon make no adjustment for counties' very different populations. Trende Dep.(Dkt. 66) at 52:2553:3; Goedert Dep. (Dkt. 65) at 186:5-7.
294. Mr. Trende admitted that the maps he relied on do not display each party's margin of victory in each county. Trende Dep. (Dkt. 66) at 52:3-6.
295. Mr. Trende admitted that the maps he relied on are based on presidential rather than state legislative election results. Trende Dep. (Dkt. 66) at 53:25-54:13, 56:9-58:9.

*     *         * 


# STIPULATIONS OF WITNESS QUALIFICATIONS 

## Professor Kenneth Mayer, Ph.D.

299. Kenneth Mayer is a Professor of Political Science at the University of Wisconsin- Madison, and a faculty affiliate at the University's La Follette School of Public Affairs.
300. Dr. Mayer teaches courses on American politics, the presidency, Congress, campaign finance, election law, and electoral systems.
301. From 1996 to 2000, Dr. Mayer served as an Associate Professor in the Department of Political Science at the University of Wisconsin-Madison.
302. From 1989 through 1996, Dr. Mayer was an Assistant Professor in the Department of Political Science at the University of Wisconsin-Madison.
303. Dr. Mayer received a Ph.D. in Political Science from Yale University in 1988, where his graduate training included courses in econometrics and statistics.
304. Dr. Mayer received a M.A., M.Phil. in Political Science from Yale University in 1987.
305. Dr. Mayer received a B.A. in Political Science from the University of California, San Diego
in 1982, where he majored in Political Science and minored in Applied Mathematics.
306. Dr. Mayer has testified at trial or at deposition in the following cases, among others: Baldus et al. v. Brennan et al., 849 F. Supp. 2d 840 (E.D. Wis. 2012); Milwaukee Branch of the NAACP et al. v. Walker et al., 2014 WI 98, 357 Wis. 2d 469, 851 N.W. 2d 262; McComish et al. v. Brewer et al., No.CV-08-1550, 2010 WL 2292213 (D. Ariz. June 23, 2010); and Kenosha County v. City of Kenosha, No. 11-CV1813 (Kenosha County Circuit Court, Kenosha, WI, 2011).
307. Dr. Mayer served as a consultant and expert witness in Baumgart et al. v. Wendelberger et al., No. 01-C-0121, 2002 WL 34127471 (E.D. Wis. May 30, 2002).
308. From 2003 to 2009, Dr. Mayer was Co-Chair of the Committee on Redistricting for the Supreme Court of Wisconsin.
309. Dr. Mayer served as an expert consultant for Prosser for Supreme Court (2011 Wisconsin Supreme Court recount).
310. In 2011, Dr. Mayer served as an expert consultant for Voces de la Frontera in the Milwaukee aldermanic redistricting process.
311. Dr. Mayer is currently serving as an expert witness in the ongoing voting rights case One Wisconsin Institute, Inc. et al. v. Nichol, et al., 3:15-cv324 (W.D. Wis.).
312. Dr. Mayer was part of a research group that consulted for the G.A.B., where he reviewed the G.A.B.'s compliance with federal mandates and reporting systems and surveyed local election practices throughout the state of Wisconsin, resulting in a 2009 report to the G.A.B.
313. Dr. Mayer serves on the Steering Committee of the Wisconsin Elections Research Center, a part of the University of Wisconsin-Madison College of Letters and Science.
314. Dr. Mayer served on the Education and Social Behavioral Sciences Institutional Review Board from 2009-2014, holding the position of Acting Chair in 2011 and Chair from 2012-2014.
315. The U.S. Department of Justice retained Dr. Mayer in 2012 to analyze data and methods regarding election practices in the state of Florida.
316. In 2006, Dr. Mayer was the Fulbright-ANU Distinguished Chair in Political Science at Australian National University.
317. From 1996-2003, Dr. Mayer served as the Director of the Data and Computation Center at the

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College of Letters and Science at the University of Wisconsin-Madison.
318. Dr. Mayer served as a consultant to the RAND Corporation from 1988-1994.
319. From 1985-1986, Dr. Mayer was a Contract Specialist for the Naval Air Systems Command in Washington, D.C.
320. Dr. Mayer has published numerous articles on American politics, the presidency, Congress, campaign finance, election law, and electoral systems in the following peer-reviewed journals: Journal of Politics, American Journal of Political Science, Election Law Journal, Legislative Studies Quarterly, Presidential Studies Quarterly, American Politics Research, Congress and the Presidency, Public Administration Review, and PS: Political Science.
321. Dr. Mayer has also published in several law reviews, including the Richmond Law Review, UCLA Pacific Basin Law Journal, and University of Utah Law Review.
322. An article written by Dr. Mayer and several colleagues, titled "Election Laws, Mobilization, and Turnout," won the award Best Journal Article Published in the American Journal of Political Science in 2014, from the American Political Science Association, State Politics and Policy Section.
323. In 2013, an article written by Dr. Mayer and colleagues titled "Election Laws and Partisan Gains," won the Robert H. Durr Award from the Midwest Political Science Association for the Best Paper Applying Quantitative Methods to a Substantive Problem.
324. Dr. Mayer has won several other honors and awards, including Leo Epstein Faculty Fellow, College of Letters and Science (2012-2015), the Jerry J. and Mary M. Cotter Award, College of Letters and Science (2011-2012), the Alliant Underkofler Excellence in Teaching Award, University of Wisconsin System (2006), and the Pi Sigma Alpha Teaching Award (2006), among others.
325. Dr. Mayer has published and edited numerous books, including The 2012 Presidential Election: Forecasts, Outcomes, and Consequences (2014), The Enduring Debate: Classic and Contemporary Reading in American Government (7th ed. 2013), Faultlines: Readings in American Government (4th ed. 2013), and With the Stroke of a Pen: Executive Orders and Presidential Power (2001), among others.
326. From 2001-2006, Dr. Mayer served as a Book Review Editor for Congress and the Presidency.
327. From 2001-2007, Dr. Mayer was on the Editorial Board of the American Political Science Review.

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328. Dr. Mayer is the recipient of a number of research grants including, among others, the Graduate School Research Committee at the University of Wisconsin (2015-2016), Wisconsin Government Accountability Board (2011-2012), Open Society Institute (2010), Pew Charitable Trusts (2008-2009), Joyce Foundation (2008), JEHT Foundation (2006-2007), National Science Foundation (1995-1998), and the McArthur Foundation (1992-1995).
329. Dr. Mayer has also presented at numerous conferences and events, including the American Political Science Association Annual Meeting, Midwest Political Science Association Meeting, Foreign Fulbright Enrichment Seminar, Reed College Public Policy Lecture Series, Southern Political Science Association Meeting, Miller Center for Public Affairs at the University of Virginia, and the American Politics Seminar at George Washington University, among others.

## Professor Simon Jackman, Ph.D.

330. Simon Jackman is a Professor in the Department of Political Science and (by courtesy) the Department of Statistics at Stanford University.
331. Dr. Jackman teaches courses on American politics and statistical methods in social sciences.
332. Dr. Jackman also currently serves as Chief Executive Officer of the United States Studies Centre at the University of Sydney.
333. From 2002 through 2007, Dr. Jackman was an Associate Professor in the Department of Political Science and (by courtesy) the Department of Statistics at Stanford University.
334. From 1996 through 2002, Dr. Jackman was an Assistant Professor in the Department of Political Science at Stanford University.
335. Dr. Jackman was a Visiting Professor at the United States Studies Centre at the University of Sydney from 2008 to 2009 and 2010 to 2013.
336. From 1994 to 1996, Dr. Jackman was an Assistant Professor in the Department of Political Science at the University of Chicago.
337. Dr. Jackman received his Ph.D. in Political Science from the University of Rochester in 1995, where his graduate training included courses in econometrics and statistics.
338. From 1991-1994, Dr. Jackman was a Visiting Doctoral Student at the Woodrow Wilson School of International and Public Affairs at Princeton University.
339. Dr. Jackman received his B.A. (with first class Honours in Government) from the University of Queensland in 1988.
340. Dr. Jackman has published numerous articles on American politics, election law, and electoral systems in the following peer-reviewed journals: The Journal of Politics, Electoral Studies, The American Journal of Political Science, Legislative Studies Quarterly, Election Law Journal, Public Opinion Quarterly, Journal of Elections, Public Opinion and Parties, and PS: Political Science and Politics.
341. Dr. Jackman authored the articles "Bayesian Analysis for Political Research," Annual Reviews of Political Science (2004), and "Estimation and Inference via Bayesian Simulation: an Introduction to Markov Chain Monte Carlo," American Journal of Political Science (2002), among other articles on political science and quantitative methods.
342. Dr. Jackman is the author of Bayesian Analysis for the Social Sciences (2009).
343. In 2014, Dr. Jackman served as a Program Chair at the Annual Meeting of the American Political Science Association.
344. Dr. Jackman served as a Principal Investigator for the American National Election Studies from 2009 to 2013.
345. From 2007-2008, Dr. Jackman was a Principal Investigator for the Co-Operative Campaign Analysis Project.
346. From 2003 to 2005, Dr. Jackman served as President of the Society for Political Methodology.
347. From 2003 to 2006, Dr. Jackman was the Director of Graduate Studies from the Department of Political Science at Stanford University.
348. Dr. Jackman was elected as a Fellow to the American Academy of Arts and Sciences in 2013.
349. Dr. Jackman has received numerous other awards and honors, including, among others: the Gregory M. Luebbert Prize for Best Article in Comparative Politics Published in 2008 or 2009, from the Comparative Politics Section of the American Political Science Association, the Journal of Politics 2006 Best Paper Award, at the Southern Political Science Association, the New South Wales Residency Expatriate Researchers Award, University of Sydney, and the Dean's Award for Distinguished Teaching at Stanford University, School of Humanities and Sciences at Stanford University (2001).
350. Dr. Jackman has received several prestigious research grants from the National Science Foundation, including in 2010, 2001, and 1999.
351. In 2014, Dr. Jackman served as a consultant to Facebook on the design and analysis of surveys.
352. From 2012 to 2013, Dr. Jackman consulted for the Huffington Post on the matters of tracking and forecasting public opinion leading up to the 2012 presidential campaign.
353. Dr. Jackman served as a consultant for the Federal Communications Commission from 2010 to 2011, assessing how media impacts public opinion and public engagement using Bayesian modeling.
354. Dr. Jackman has been an Associate Editor for several editorial journals, including the Annual Review of Political Science (2005-2013) and Political Analysis (2010 to the present).
355. Dr. Jackman has provided editorial board service to several journals, including the American Political Science Review (current), American Journal of Political Science, Journal of Politics, Electoral Studies, Australian Journal of Political Science (current), Public Opinion Quarterly (current), and Political Analysis.

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356. Dr. Jackman has been invited to speak at numerous lectures, seminars, and workshops, including the Asian Political Methodology Conference, the ACSPRI Social Science Methodology Conference, the Australian Political Studies Association Conference, the Society for Political Methodology, the Munk School of Global Affairs, the Massachusetts Institute of Technology, the Research Triangle Institute, Nuffield College, TEDx Sydney, the International Political Science Association, Stanford University Law School, Princeton University, Harvard University, Yale University, and Vanderbilt University.
357. Dr. Jackman helped develop the software package pscl, a package of classes and methods for $R$ developed in the Political Science Computational Laboratory at Stanford University.
358. Dr. Jackman has served as a Reviewer for the National Research Council, Chair for the Emerging Scholar Committee at the University of Sydney, on the James Madison Awards Committee at the American Political Science Association, Chair of the Distinguished Career Achievement Award Committee for the Society for Political Methodology, and President of the Society for Political Methodology and the Political Methodology Section of the American Political Science Association, among other services to the political science field.

## Sean Trende

359. Trende received a B.A. from Yale University in 1995, with distinction, with a double major in history and political science.
360. Trende received a J.D. from Duke University in 2001, cum laude.
361. Trende received an M.A. from Duke University in 2001, cum laude, in political science.
362. Trende joined RealClearPolitics in January of 2009 as its Senior Elections Analyst. He assumed a fulltime position with RealClearPolitics in March of 2010 and continues as its Senior Elections Analyst.
363. RealClearPolitics is one of the most heavily trafficked political websites in the world.
364. RealClearPolitics provides political analysis and poll aggregation.
365. RealClearPolitics has a readership in excess of 1 million.
366. Trende's work has been cited by David Brooks of The New York Times, Brit Hume of Fox News, Michael Barone of The Almanac of American Politics, Paul Gigot of The Wall Street Journal, and Peter Beinart of The Atlantic.

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367. Trende's responsibilities with RealClearPolitics consist of tracking, analyzing, and writing about elections. Trende is in charge of rating the competitiveness of House of Representatives races, and he collaborates in rating the competitiveness of Presidential, Senate and gubernatorial races.
368. Trende's responsibilities also include studying and writing about legislative redistricting, and supervising and editing the work of RealClearPolitics' elections analyst David Byler.
369. Trende regularly writes columns for RealClearPolitics and has written on partisan gerrymandering and geographic clustering. He has hundreds of articles available online.
370. Trende's readers include political science professors, members of the media, elected representatives, and others.
371. Trende is a Senior Columnist for Dr. Larry Sabato's "Crystal Ball" and has written for the Crystal Ball since January 2014. Dr. Sabato is a professor of political science at the University of Virginia and serves as the director of the University of Virginia Center for Politics.
372. Trende authored a chapter in Dr. Larry Sabato's Barack Obama and the New America: The 2012 Election and the Changing Face of Politics, ch.

12 (2013), which discussed the demographic shifts accompanying the 2012 elections.
373. Trende authored a chapter in Dr. Sabato's The Surge: 2014's Big GOP Win and What It Means for the Next Presidential Election, ch. 12 (2015), which discusses demographics and Electoral College shifts.
374. Trende is the author of The Lost Majority: Why the Future of Government is up For Grabs and Who Will Take It (2012). It includes analysis of demographic and political trends beginning around 1920 and continuing through the modern times.
375. Trende co-authored the Almanac of American Politics 2014 (2013). Trende's focus was researching the history of and writing descriptions for many of the newly-drawn congressional districts.
376. Trende has served as a peer reviewer for articles for the political science journals Party Politics and PS.
377. Trende has spoken before the Heritage Foundation, the American Enterprise Institute, the CATO Institute, the Bipartisan Policy Center, and the Brookings Institution.
378. In 2012, Trende was invited to Brussels to speak about American elections to the European External Action Service, which is the European Union's diplomatic corps.
379. Trende's presentations have included: "The Lost Majorities: 2008, 2010 and America's Political Future," Bradley Lecture, American Enterprise Institute, January 2012; Panelist, "The Future of Red and Blue," Bipartisan Policy Center, Washington, DC, April 2012; "The 2012 Elections: Trends, Prognostications and What's at Stake," 3rd Annual Family Office Wealth Management Forum, Greensboro, Georgia, May 2012; "2012 U.S. Election Series," with Bruce Stokes and Alexandra de Hoop Scheffer, German Marshall Fund, Brussels, Belgium, Oct. 4, 2012
380. Trende has appeared on Fox News and MSNBC to discuss electoral and demographic trends.
381. Trende has spoken on radio shows including First Edition with Sean Yoes, the Diane Rehm Show, the Brian Lehrer Show, the John Batchelor Show, the Bill Bennett Show, Beijing Radio, CNN Radio, NPR, and Fox News Radio.
382. Trende has been cited in publications including The New York Times, The Washington Post, The Los Angeles Times, The Wall Street Journal, and USA Today.
383. Trende sits on the advisory panel for the "States of Change: Demographics and Democracy" project, which is a three-year project sponsored by the Hewlett Foundation involving the Brookings Institution, the American Enterprise Institute, and
the Center for American Progress. The group looks at trends among eligible voters and the overall population, both nationally and in some states.
384. Trende has drawn, using Adobe Illustrator, complete maps of every congressional district ever drawn, dating back to 1789 .
385. Trende authored an expert report in Dickson v. Rucho, No. 11-CVS-16896 (N.C. Super Ct., Wake County), regarding partisanship of various districts, and that report was accepted without objection.
386. Trende authored two expert reports in $N A A C P$ v. McCrory, No. 1:13CV658 (M.D.N.C.), which involves challenges to North Carolina's voter laws, and also testified.
387. Trende authored an expert report in NAACP v. Husted, No. 2:14-cv-404 (S.D. Ohio), and in a later iteration of that litigation, Ohio Democratic Party v. Husted, No. 2:15-CV- 1802 (S.D. Ohio), and testified at trial.

## Professor Nicholas Goedert, Ph.D.

388. Dr. Goedert is currently a Visiting Assistant Professor of political science at Lafayette College in Easton, Pennsylvania.
389. Dr. Goedert has accepted a tenure track professor position in political science at the Virginia Polytechnic Institute and State University (Virginia Tech) starting next school year.
390. In 2012, Dr. Goedert received a Ph.D. from the Department of Politics, Princeton University.
391. Dr. Goedert's dissertation regarding congressional redistricting is titled: "Gerrymandering, Electoral Uncertainty, and Representation." His advisors were Brandice CanesWrone (chair), Nolan McCarty, and Adam Meirowitz.
392. Dr. Goedert's graduate training included coursework on quantitative methods and statistics.
393. In 2009, Dr. Goedert received a M.A. from the Department of Politics, Princeton University.
394. His examination fields were American Politics (Public Opinion, Political Psychology, and Legislative Politics), Formal and Quantitative Methodology.
395. In 2006, Dr. Goedert received a J.D. (cum laude) from Georgetown University Law Center. He specialized in election law.
396. In 2001, Dr. Goedert received a B.A. (magna cum laude) from the Department of Social Studies, Harvard University.

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397. From 2014 to the present, Dr. Goedert is employed as Visiting Assistant Professor, Department of Government and Law, Lafayette College.
398. From 2012 to 2014, Dr. Goedert was a Postdoctoral Research Associate, Department of Political Science at Washington University in St. Louis.
399. Dr. Goedert's peer-reviewed publications include:
a. "The Pseudo-Paradox of Partisan Mapmaking and Congressional Competition," conditionally accepted at State Politics and Policy Quarterly (2016).
b. "The Case of the Disappearing Bias: A 2014 Update to the 'Gerrymandering or Geography' Debate," forthcoming in Research \& Politics (2016 research note).
c. "Redistricting, Risk, and Representation: How Five State Gerrymanders Weathered the Tides of the 2000's." Election Law Journal 13(3): 406-418 (2014).
d. "Gerrymandering or Geography?: How Democrats Won the Popular Vote but Lost the Congress in 2012." Research \& Politics 1(1): 2053168014528683 (2014).
400. Dr. Goedert's working papers include:
a. "Redistricting Institutions, Partisan Tides, and Congressional Competition"
b. "Southern Redistricting under the VRA: A Model of Partisan Tides"
c. "Gerrymandering and Competing Norms of Representation"
d. "Democratic Incumbent Resilience in the Post-1980 Senate: A Theory of Partisan Issue Competence"
e. "The Impact of Geographic Constituencies on Regional Parties: Evidence from Six Nations"
401. Dr. Goedert's conference presentations include:
a. Gerrymandering, Polarization, and Competing Norms of Representation," presented at the Annual Meeting of the American Political Science Association, Washington, DC (2014).
b. "Democratic Incumbent Resilience in the Post-1980 Senate: A Theory of Partisan Issue Competence," presented at the Annual

Conference of the Midwest Political Science Association, Chicago, IL (2014).
c. "Gerrymandering and Competing Norms of Representation," presented at the Annual Conference of the Midwest Political Science Association, Chicago, IL (2012).
d. "Southern Redistricting under the VRA: A Model of Partisan Tides," presented at the State Politics and Policy Conference, Houston, TX (2012).
e. "Redistricting Institutions under Electoral Uncertainty," presented at the Annual Meeting of the American Political Science Association, Seattle, WA (2011).
f. "Redistricting Institutions, Partisan Tides, and Congressional Turnover," presented at the State Politics and Policy Conference, Hanover, NH (2011), the Annual Conference of the MPSA, Chicago, IL, and the Society for Political Methodology Summer Meeting, Princeton, NJ.
402. Dr. Goedert is a contributor to political science blogs at The Washington Post, The Monkey Cage and Wonkblog.
403. Dr. Goedert has written a non-peerreviewed short article titled "Not Gerrymandering,

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but Districting: More Evidence on How Democrats Won the Popular Vote but Lost the Congress" for The Monkey Cage (Nov. 15, 2012).
404. Dr. Goedert's teaching experience includes, as a Visiting Professor, "Introduction to United States Politics" (Fall 2014); "Political Opinion and Participation in the United States" (Fall 2014 and Spring 2016); "Campaigns and Elections" (Spring 2015 and Fall 2015); "Congress and the Legislative Process" (Fall 2015); "Constitutional Law and Politics in the United States" (Spring 2016 (scheduled)); "Representation, Apportionment, and Democratic Participation" (Spring 2015 and Spring 2016).
405. Dr. Goedert has served as a Legislative Analyst for the Maryland General Assembly, Department of Legislative Services, from 2006-2007.
406. Dr. Goedert has served as a manuscript reviewer for Legislative Studies Quarterly; State Politics and Policy Quarterly; Election Law Journal; and Social Influence.

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## Cross-References to Supplemental Appendix

Plaintiffs' Trial Exhibit 93: Expert Analysis by Professor Simon Jackman, Sensitivity of the Efficiency Gap to Uniform Swing appears at: SA315- 320

Plaintiffs' Trial Exhibit 122: Expert Analysis by Professor Simon Jackman, Average Efficiency Gaps for Wisconsin Plans (1970s-2010s) appears at: SA321

Plaintiffs' Trial Exhibit 134: Memo by Keith Gaddie, dated April 17, 2011, Wisconsin_Partisanship appears at: SA322

Plaintiffs' Trial Exhibit 172: Plan Comparisons spreadsheet appears at: SA323-327

Plaintiffs' Trial Exhibit 237: Memo by Tad Ottman, Questions and Responses appears at: SA328-329

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Composite_Adam_Assertive_Curve (Excerpt:
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Plaintiffs' Trial Exhibit 280: TadAggressiveCurve appears at: SA338

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Plaintiffs' Trial Exhibit 284: Summary appears at: SA344-345

Plaintiffs' Trial Exhibit 325B: Expert Analysis by Professor Simon Jackman, EG and Partisan Bias appears at: SA346

Plaintiffs' Trial Exhibit 329: Expert Analysis by Professor Simon Jackman, EG and PB on Same Chart appears at: SA347

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Plaintiffs' Trial Exhibit 342: Memoranda to Republican Legislators from Adam Foltz, dated June 19, 2011, re: New Map (Excerpted for District 1) appears at: SA348-351

Plaintiffs' Trial Exhibit 348: Email from Jim Troupis to Eric McLeod, dated June 21, 2011, re: Experts appears at: SA352

## Plaintiffs' Trial Exhibit 364: Tad MayQandD appears

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Plaintiffs' Trial Exhibit 463: Proposed Map Room Access Policy appears at: SA355

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Plaintiffs' Trial Exhibit 495: Expert Analysis by Professor Simon Jackman, Wisconsin Sensitivity Testing appears at: SA360

Defendants' Trial Exhibit 502: Map of Act 43 appears at: SA361

Defendants' Trial Exhibit 505: Map Showing Prior Plan (2002-2010) Assembly Districts' Deviation from Ideal Population Following 2010 Census appears at: SA362

Defendants' Trial Exhibit 515: Prior Plan (2002-2010) Assembly District Map appears at: SA363

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## Order Postponing Consideration of the Question of Jurisdiction

No. 16-1161
Title: Beverly R. Gill, et al., Appellants
v.

William Whitford, et al.
Docketed:
March 24, 2017
Linked with
16A1149
Lower Ct:
United States District Court
for the Western District of Wisconsin
Case Nos.:
(15-cv-421-bbc)
Decision Date: November 21, 2016
$\sim \sim \sim$ Date~~~ $\sim \sim \sim \sim \sim$ Proceedings and Orders $\sim \sim \sim \sim \sim$

Jun 192017
Further consideration of the question of jurisdiction is POSTPONED to the hearing of the case on the merits.

# In The Supreme Court of the $\mathfrak{H z n t e d}$ Sates 

Beverly R. GLLL, ET AL., APPELLANTS,
v.

WILLIAM Whitford, et AL., AppelLees

On Appeal From the United States District Court for the Western District of Wisconsin

## JOINT APPENDIX VOLUME II

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The following documents have been omitted in printing this joint appendix because they appear on the following pages in the appendix to the Jurisdictional Statement:

Opinion and Order finding 2011 Wisconsin Act 43 Unconstitutional, Whitford v. Gill, No. 3:15-cv-00421-bbc (W.D. Wis. November 21, 2016), ECF No. 166 1a

Opinion and Order Regarding Remedy, Whitford v. Gill, No. 3:15-cv-00421-bbc (W.D. Wis. January 27, 2017), ECF No. 182 316a

Judgment Entered, Whitford v. Gill, No. 3:15-cv-00421-bbc (W.D. Wis. January 27, 2017), ECF No. 183 325a

Order Amending Judgment, Whitford v. Gill, No. 3:15-cv-00421-bbc (W.D. Wis. February 22, 2017), ECF No. 189 327a

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# Use of Efficiency Gap in Analyzing Partisan Gerrymandering <br> Report for State of Wisconsin, Whitford v. Nichol 

Nicholas Goedert

December 2, 2015

## I. Introduction

My name is Nicholas Goedert, and I am currently a Visiting Assistant Professor of Government and Law at Lafayette College in Easton, Pennsylvania. I teach classes in American electoral politics, voter behavior, the United States Congress, constitutional law, and representation theory.

I have been retained by the defendants in this lawsuit to provide expert opinions in the case titled above. More specifically, I have been asked to offer opinions on using the efficiency gap to measure partisan gerrymandering as done by the plaintiffs' experts Professor Kenneth Mayer and Professor Simon Jackman.

## II. Qualifications and compensation

I received a Ph.D. in Politics from Princeton University in 2012, where I completed a dissertation on congressional redistricting, and my graduate training included courses in quantitative methods and statistics. I received my undergraduate degree in Social Studies from Harvard University in 2001, and a J.D. from Georgetown University Law Center in 2006, where I specialized in election law. My curriculum vitae is attached to this report.

All my publications that I have authored or published appear in my curriculum vitae. Those publications include peer-reviewed journals such as: The American Journal of Political Science, State Politics and Policy Quarterly, PS: Political Science and Politics, Election Law Journal, and Research and Politics.

I have published, or have forthcoming publications, specifically on the effects of districting methods on competition in congressional elections in State Politics and Policy Quarterly and Election Law Journal, and on the effects of geographic bias in congressional
districting in Research and Politics and in The Monkey Cage political science blog at The

## Washington Post.

I am being compensated at a rate of $\$ 175$ per hour.

## III. Summary

1.) Despite claims in the plaintiffs' complaint, a large efficiency gap does not necessarily imply an unbalanced map. Instead, a large efficiency gap implies deviation from a predetermined seats/votes curve representing "hyper-proportionate" or "hyperresponsive" representation. Thus, using an efficiency gap standard creates the same constitutional issues as the proportional representation standard the Court has previously rejected. Moreover, requiring adherence to a specific seats/votes curve may discourage legislatures from drawing maps that would fulfill normatively desirable objectives, such as maximizing competitive elections or achieving proportional representation, but do not conform to this expected seats/votes curve. (Section IV)
2.) The plaintiffs' complaint alleges that an efficiency gap of $7 \%$ in a single election is sufficient for presumptive unconstitutionality. But evidence in both the academic literature and the plaintiffs' expert report show that efficiency gaps of the size proposed in the complaint are highly unstable and not particularly informative of future or durable gaps. In fact, as many as half of all maps that exceed this threshold in one election during a decade will be biased in favor of the opposite party in another election during the same decade. And even those few maps that are significantly and durably biased in favor of one party are mostly not even drawn with clear partisan intent. (Section V)
3.) The plaintiffs' complaint lacks a crucial addition "sensitivity testing" prong suggested in the academic literature. Without an additional test of durability, a majority of single election results exceeding the predetermined threshold would be false positives, because they are either not drawn with partisan motivation, or they would be biased in favor of the opposite party in another election during the same decade. The test of durability in Jackman's report is somewhat unclear and arbitrarily conditions durability on the results of small handful of elections. Additionally, even including the sensitivity testing prong as detailed in the literature would be potentially constitutionally problematic. (Section VI)
4.) The expert report of Mayer purports to show that an alternate map (i.e. the Demonstration Plan) could have been drawn with much lower efficiency gap in 2012. However, the map created by Mayer was generated based on significant information, the overall 2012 electoral environment, that was unknowable to the legislature at the time the map needed to be drawn. The Demonstration Plan is also deliberately drawn to exclude information that legislators would likely incorporate into their districting decisions, in the form of incumbency and anticipated uncontested races. Additionally, the report does not provide data on what bias we should expect to observe under the Demonstration Plan given the range of possible future election results. (Section VII)
5.) Any judgment about the partisan motivation behind pro-Republican bias in a map should be made in the context of bias due to the asymmetric geographic dispersion of partisans. This dispersion has generated Republican bias in many states' maps across the nation over the last few decades, growing in the most recent election cycles, as observed in both the academic literature and the plaintiffs' expert report. It has also generated Republican bias in two different non-partisan maps drawn in Wisconsin, in a few cases in excess of
the bias observed in the most recent election cycle under the Republican-drawn map. Evidence of this bias is also observed in an analysis of the distribution of Wisconsin wards. (Section VIII)

## IV. General Properties of Efficiency Gap

A. Efficiency gap demands codification of a specific relationship of seats to votes that amounts to hyper-proportional representation

Efficiency gap is defined both by the plaintiffs and in the academic literature as the ratio of one major party's wasted votes to the other major party's wasted votes. In a single-member, majority rule district, all votes for a losing candidate are wasted, and votes for a winning candidate in excess of the $50 \%$ threshold needed for victory are also wasted. Thus in all individual seat elections with two candidates, exactly half of the votes are counted as wasted, with the losing candidate accounting for a greater share of wasted vote the closer the election is.

Although a precise calculation of efficiency gap across a collection of races requires knowing the total number of votes cast for each major party candidate in each race, this can be simplified into a linear seats/votes curve with zero bias and a slope of 2 if one assumes equal turnout in all districts. (McGhee 2014, p. 80). This simplification is used in the Stephanopoulos and McGhee article originally advocating for efficiency gap as a standard for adjudicating partisan gerrymanders (p. 853), as well as the historical analysis in the plaintiffs' expert report by Jackman (section 6.1, p. 18). I concur that this shortcut is an appropriate and useful summary measure of efficiency gap and also use it in subsequent examples in this report. However, the fact that efficiency gap under basic assumptions simplifies to a single linear seats/votes curve also displays its drawbacks for use as a standard for a Court to judge the constitutionality of a map.

The Supreme Court has stated on multiple occasions that the Constitution does not guarantee a right to proportional representation of any particular group, a sentiment echoed in both the majority and dissent in Vieth (Vieth v. Jubelirer, 541 U.S. at 288; Vieth v. Jubelirer 541 U.S. at 338 (Stevens, J., dissenting); Davis v. Bandemer, 478 U.S. at 111). And the Court has additionally been wary of adopting a standard for partisan gerrymanders that would amount to
proportional representation (Davis v. Bandemer 478 U.S. at 155). Yet the efficiency gap test would codify a very specific translation of seats to votes that is essentially "hyper-proportional" representation. Every 1 percentage point increase in vote would be expected to translate into a 2 percentage point increase in seats in order for a map to be measured as fair.

This formula does have the advantage of roughly conforming with the observed average seat/votes curve in historical U.S. congressional and legislative elections (see e.g. Tufte 1973, Goedert 2014). But this correlation is coincidental and not connected to the theory behind EG. Moreover, the correlation is not guaranteed to hold up over time, especially as populations become more polarized in their partisanship. Codifying this relationship between seats and votes would constrain states wishing to reform their voting or districting systems. There are several ways in which states might wish to draw districts for normatively good reasons that would be seen as highly biased, and thus potentially unconstitutional, when measured under EG, especially when taking into account unpredictable electoral tides.

## B.) An efficiency gap standard may discourage drawing of competitive districts

Because they are highly sensitive to tides, implementing an efficiency gap standard may discourage legislatures from drawing maps with too many competitive seats. During a wave election favoring either party, competitive districts may all fall in one direction, causing an extreme EG measurement favoring that party despite the balanced intent behind drawing these districts.

For example, suppose a state with 20 districts contained a roughly even number of Democrats and Republicans, but that the state's mapmakers chose to draw half these districts to be evenly balanced, and half to clearly favor one party. So ten districts are drawn to be $50 \%$ Democratic and 50\% Republican, while five districts are drawn to be overwhelmingly (e.g 75\%) Democratic, while the last five districts are drawn to be similarly overwhelmingly Republican.

Now, suppose in one election the Democrats win $55 \%$ of the two-party vote overall (a wave slightly smaller than 2008 at the national congressional level), and that this gain in vote share is spread approximately evenly across the state. Each party would still win the five seats that were drawn to be safe for them, but the Democrats would also win all ten seats drawn to be most competitive. Thus, the Democrats would win $75 \%$ of the seats with $55 \%$ of the vote. Efficiency Gap would prescribe that a fair map would assign Democrats only $60 \%$ of the seats with this vote share, and so this map would be measured as $15 \%$ biased in favor of the Democrats. Of course, if Republicans won $55 \%$ of the vote, evenly spread across the state, the map would have a $15 \%$ Republican bias under efficiency gap. But the test suggested by the plaintiffs asks the Court to evaluate the constitutionality of a map based only on the bias measured in one election.

Moreover, during a time in which several states are moving to reform their redistricting process and incorporating the value of political competition into reform considerations, we do observe real maps that efficiency gap would judge too sensitive to shifting tides on both sides. For example, Arizona congressional districts are drawn by a nonpartisan commission that since 2001 is required by state law to try to drawn competitive districts when possible. After the 2000 Census, this commission drew half the state's 8 districts in a balance within $6 \%$ of the national average presidential vote share throughout the decade (as measured by Cook's PVI, a measure of the partisanship of congressional districts relative to the nation based on recent presidential election results). The result has been a great deal of competition and partisan turnover since 2002, but large fluctuations in efficiency gap. As shown in Table 1 below, the map had an efficiency gap of $14 \%$ in favor of Republicans in 2002, but this switched signs twice during the decade, favoring Democrats in 2006 and 2008, and switching back to Republicans in 2010. Under a new, but still nonpartisan map, this switched back a third time in 2012, with an efficiency gap favoring Democrats of $14 \%$. This Commission has at various times been accused
by both parties of acting with partisan intent; efficiency gaps may yield spurious evidence of partisan bias even when motivated only by desire to enhance competition.

Table 1. Arizona Congressional Results, 2002-2012

| Year | GOP seats | GOP Vote | Eff. Gap |
| :---: | :---: | :---: | :---: |
| 2002 | 75.0\% | 55.7\% | 13.6\% |
| 2004 | 75.0\% | 60.9\% | 3.1\% |
| 2006 | 50.0\% | 52.1\% | -4.2\% |
| 2008 | 37.5\% | 46.4\% | -5.4\% |
| 2010 | 62.5\% | 53.1\% | 6.2\% |
| 2012 | 44.4\% | 54.3\% | -14.1\% |

## C.) An efficiency gap standard may discourages enactment of proportional representation.

While the Court has held that the Constitutional does not require it, proportional representation of political parties is a permissible goal that a state may choose to adopt (Vieth $v$. Jubelirer, 541 U.S. 338 (Stevens, J., dissenting); Gaffney v. Cummings, 412 U.S. at 754). But because the efficiency gap requires a $2: 1$ "hyper-proportional" relationship between seats and votes, it may also discourage the drawing of districts to achieve 1:1 proportional representation. For example, suppose a state's partisan identification is 60\% Democrat and $40 \%$ Republican and has 20 districts. The state wishes to achieve fair proportional representation, and so draws 12 districts to be $100 \%$ Democratic and 8 districts to be $100 \%$ Republican. If Democrats do get $60 \%$ of the vote, they will win $60 \%$ of seats, but EG requires that a fair map would award $70 \%$ of seats to Democrats in this scenario. Thus, the map that was both proportional and virtually guaranteed to yield a Democratic majority would be measured by EG to be biased by $10 \%$ in favor of Republicans.

Note that the above hypothetical would create a map completely resistant to shifts in partisan tides, which may be normatively undesirable. But one might also imagine a map drawn
to achieve proportional representation and still be responsive to change. For example, imagine a state with 20 districts, evenly balanced between Democrats and Republicans in an election without tides favoring either party. Suppose District 1 is drawn to be $97.5 \%$ Democratic, and then each subsequent district is drawn to be $5 \%$ more Republican than the last. So District 2 is 92.5\% Democratic; District 10 is $52.5 \%$ Democratic; District 11 is $47.5 \%$ Democratic; and District 20 is $2.5 \%$ Democratic and $97.5 \%$ Republican (see Table 2 below).

## Table 2. Efficiency Gap Under Hypothetical Map Designed to Create Proportional Representation

| District | Partisan Baseline (\% Dem under 50/50 Statewide Party Split) | Winning Party under statewide vote: |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | 50\% Dem | 55\% Dem | 60\% Dem |
| 1 | 97.5\% | D | D | D |
| 2 | 92.5\% | D | D | D |
| 3 | 87.5\% | D | D | D |
| 4 | 82.5\% | D | D | D |
| 5 | 77.5\% | D | D | D |
| 6 | 72.5\% | D | D | D |
| 7 | 67.5\% | D | D | D |
| 8 | 62.5\% | D | D | D |
| 9 | 57.5\% | D | D | D |
| 10 | 52.5\% | D | D | D |
| 11 | 47.5\% | R | D | D |
| 12 | 42.5\% | R | R | D |
| 13 | 37.5\% | R | R | R |
| 14 | 32.5\% | R | R | R |
| 15 | 27.5\% | R | R | R |
| 16 | 22.5\% | R | R | R |
| 17 | 17.5\% | R | R | R |
| 18 | 12.5\% | R | R | R |
| 19 | 7.5\% | R | R | R |
| 20 | 2.5\% | R | R | R |
| Statewide Total | 50\% | $10 \mathrm{D} / 10 \mathrm{R}$ | $11 \mathrm{D} / 9 \mathrm{R}$ | $12 \mathrm{D} / 8 \mathrm{R}$ |
| Efficiency Gap |  | 0\% | -5\% | -10\% |

Under an election that is split 50/50 in the vote, Democrats will likely win districts 1 though 10, and Republicans districts 11 through 20, yielding no net efficiency gap. But if the balance of the electorate changes, either permanently or through a single wave election, the seat
share for each party will likely shift proportionately, create efficiency gap bias. If Democrats win $60 \%$ of the vote statewide, they will now win districts 1 though 12 , or $60 \%$ of the seats. Yet efficiency gap prescribes that a party should win $70 \%$ of the seats with this vote share, so the map would be judged as $10 \%$ biased (and thus presumptively unconstitutional) in favor of the Republicans.

We can also observe anecdotal evidence of large efficiency gaps in real maps designed to draw safe and roughly proportional districts by bipartisan agreement. In the 2000's decade, Democrats controlled all branches of state government in California, but instead of crafting an aggressively partisan congressional map, worked closely with Republicans in the legislature to draw districts that would protect incumbents of both parties and thus create almost entirely safe seats. In 2008, Democrats won $64 \%$ of the congressional seats in California with approximately $64 \%$ of the statewide vote share. But efficiency gap would judge this map to be biased in favor of Republicans by $14 \%$ that year, and thus presumptively unconstitutional were this the first year after redistricting, despite being drawn under Democratic control, and passed by large majorities of both parties in the legislature.

Contrary to the plaintiffs' assertion in complaint paragraph 51, a large efficiency gap does not imply a map is unbalanced, as shown in the above examples. Even a "balanced" map can show extreme EG bias under some (or even all) electoral tides conditions and varying normative definitions of balance.

## V. Historical Instability and Fluctuations in Efficiency Gap

A. Past results demonstrate enormous instability even within a given decade and sensitivity to very realistic partisan tides

The plaintiffs' complaint alleges that a districting plan should be considered presumptively unconstitutional if an efficiency gap of 7\% is observed in a single election (paragraph 86) (though they also propose that the Court could declare the specific Wisconsin plan unconstitutional without setting an exact threshold). In doing so, they rely on the Jackman report (p. 56), and also cite research by Stephanopoulos \& McGhee (2015) suggesting an 8\% threshold for state house plans. The complaint alleges that "where the efficiency gap is large and much greater than the historical norm...intent to systematically disadvantage voters based on their political beliefs can be inferred by the severity of the gerrymander alone" (Plaintiffs' complaint, paragraph 6). Yet both the academic research and data presented by the plaintiffs' expert show that such intent cannot be inferred.

Indeed, as both the Jackman report and the Stephanopoulos \& McGhee article comprehend, merely observing a given threshold gap in a single election is not very informative as to the gap that we might expect over the lifetime of a plan. Indeed, Jackman acknowledges that "Conditional on observing an election with $E G>.07$, there is a $45 \%$ chance that under the same plan we will observe $E G<0$. ." (p.56). In other words, about half of all plans over the past 40 years that crossed the threshold for presumptive unconstitutionality in one election are also biased in favor of the opposing party in at least one election during the same decade. As measured by Stephanopoulos and McGhee, this is also true of 5 out of 14 state house plans crossed their 8\% threshold for Republican bias during the 2000's decade (Stephanopoulos and McGhee 2015, p. 882).

And several iconic examples of Republican gerrymanders did not even display a consistent efficiency gap through the decade. Perhaps most famously, the Pennsylvania congressional map drawn by Republicans and upheld in Vieth v. Jubilirer elected Republicans to just 7 of 19 seats based on about $44 \%$ of the major-party in 2008, resulting in an efficiency gap bias EG bias in favor of Democrats in 2008. Similar backfires occurred the same year in other states districted by Republicans such as Virginia and Ohio. By the estimates of Stephanopoulos and McGhee, 18 of the 23 congressional or state legislative plans that were alleged in suits prior to 2010 to be unlawful partisan gerrymanders were actually measured as being biased in both directions during the decade of their existence. And the only plans definitively biased in favor of Republicans occurred in Florida in the 2000's, a state that served as an iconic example of bias created from geographic dispersion rather than intentional gerrymander, as discussed in Section VIII below.

## B. Very few plans are unambiguous as to sign, and they are usually not even partisan gerrymanders

Indeed, it is rare that a map is clearly is biased in favor or one party or another over the course of an entire decade, and the few plans that are clearly biased are not even necessarily partisan gerrymanders. On p. 53, the Jackman report mentions that only $12 \%$ (17 out of 141) of state legislative plans analyzed over four decades are unambiguous as to the direction of their bias, based on his measurement of confidence over imputations in uncontested races; these 17 plans are listed on Table 1 on p. 55. 16 of the 17 plans are biased in favor of the Republicans, suggesting natural geographic bias favoring Republicans discussed further below. But more importantly, most of these plans are not partisan gerrymanders. Of the 16 most Republican plans, only six or seven would plausibly be called partisan gerrymanders from the standpoint of partisan control of the districting process. Instead, they include such plans as the New York
legislature in every decade (usually under split control), an example used by Rodden and Chen to demonstrate asymmetric geography. Additionally, the short list also includes the Wisconsin map from 2001-2010 that was drawn by a court. So a durable bias in favor of Republicans is not even a sign of deliberate partisan intent in even the strongest anecdotal evidence.

## VI. Testing the Aensitivity and Durability of Efficiency Gap

A. The plaintiffs' complaint does not include a crucial second part to the empirical test for presumptive unconstitutionality, sensitivity testing for future results

Stephanopoulos and McGhee also allow that "most redistricting plans are volatile enough that their precise consequences cannot be forecast with great accuracy. Specifically, a plan's efficiency gap in one election is a relatively weak predictor of its gap in the next election" (p. 864). Therefore, observing a certain gap in one election is not a sufficient test of presumptive unconstitutionality for Stephanopoulos and McGhee. Instead, they suggest that for a map to be presumed unconstitutional, it should not only reach a specified level of bias in a particular election, but also be very unlikely to switch signs in bias over the foreseeable elections in the future (p. 889). "(W)e recommend setting the bar at... 8 percent or state house plans, with the further proviso that sensitivity testing show that the efficiency gaps are unlikely to hit zero over the plans' lifetime." (p. 887, emphasis mine).

Stephanopoulos and McGhee evaluate the second criteria through "sensitivity testing", shifting the actual election results by $7.5 \%$ in each direction for congressional plans, and $5.5 \%$ in each direction for legislative plans, and calculating the gaps for each shift (p. 864). Under this second test, most of the instances of efficiency gaps beyond the initial threshold would not be judged presumptively unconstitutional because the simulated gap is too unstable.

The plaintiffs' complaint includes no such second part to the test for presumptive unconstitutionality. Without this second part to the test, almost any plan could be judged
presumptively unconstitutional under some election conditions. Thus, the EG standard could come down to a pure subjective evaluation of partisan intent, combined with a well-time fluke election result.
B. Jackman's report contains testing of robustness of EG measures over time, but it is unclear how these are to be incorporated into the test

In place of an explicit sensitivity testing prong to be applied to each map at issue, the Jackman report implies that sensitivity testing though modeling a future range of possible election results is unnecessary because efficiency gaps of a certain magnitude are historically unlikely to switch signs when observed in the first elections after redistricting. But conditioning one's observations only on particular election results is rather arbitrary, and in this case, likely biases toward a finding of EG durability. This is because among the notable national "wave elections" during in the period from 1972-2014 (e.g. 1974, 1994, 2008, 2010), none occurred immediately following a redistricting year. Instead, most post-redistricting elections occurred in years of relative partisan balance at the legislative level. The lack of notable wave elections among those picked to condition on is probably coincidental, but likely does result in less instability than if the durability of EG measurements were observed after such a wave election. There is no guarantee that in the future, a wave election will not occur immediately after redistricting, and thus applying this standard to future cycles would inappropriately imply durability. A more accurate test would be how often a gap of a certain magnitude in any cycle implied consistency across an entire decade. As previously noted, this test gives us much less confidence about the durability of a single EG measurement.
C. Even the Stephanopoulos \& McGhee sensitivity testing is a flawed way to judge constitutionality after a single election

But even the sensitive test as proposed by Stephanopoulos and McGhee is problematic. The Stephanopoulos and McGhee sensitivity testing prong is an important acknowledgement of the fluctuations observed in efficiency gap as electoral tides shift. Yet as the authors themselves concede, this test involves simulating future election results assuming a hypothetical uniform swing across all districts, a method found problematic in evaluations of partisan bias by Justice Kennedy in LULAC v. Perry (548 U.S. at 420). The authors justify the use of this method nevertheless by saying it is not used to calculate the point estimate of bias, only the uncertainty. But given the overwhelming number of false positives generated from reliance on the point estimate alone, this underestimates the importance of the sensitivity testing prong in the final determination of constitutionality.

Additionally, Stephanopoulos and McGhee argue that their sensitivity test involves hypothetical swings much smaller than needed to evaluate the symmetry of partisan bias, as they only swing results in either direction 7.5 percentage points in the case of congressional maps, and 5.5 points in the case of state legislative maps. Yet this shift may not be sufficient to simulate the plausible range of election results than may be observed with a decade. For example, the Republican share of the two-party aggregated national popular vote in congressional elections jumped from $44.5 \%$ in 2008 to $53.5 \%$ in 2010 (a nine point swing). So shifting the 2008 national result 7.5 points in both directions would have been insufficient to encompass the actual national result two years later. And within a single state, where small variations in incumbency and candidate choice may have greater impact on aggregated results, fluctuations across elections could be even larger.

## VII. Discussion of Mayer Demonstration Plan and Data Imputation

Both the expert reports of Jackman and Mayer rely on imputing votes for counterfactual electoral situations. Most frequently, this is done in case of past election results where a candidate was running without major party opposition. When measuring the bias in a map from an academic standpoint, imputing vote share in unopposed races seems entirely appropriate, as do the specific methods used in both reports to make these imputations. However, this seems more problematic in the context of a legal challenge to a map asserting that a particular individual's constitutional rights have been violated. Specifically, if an individual votes for party A in an election with no major party opposition, it would be curious to allege that individual's right to political representation has been violated because they hypothetically may have voted for party B had a different district been drawn to induce party B to run a candidate. And it would be even more curious to blame that hypothetical lack of representation on the mapmaker as opposed to the party that chose to run no candidate in the district or the voter who nevertheless voted for the opposing party.

But the most concerning imputation decisions come in the case of the demonstration plan presented in the Mayer expert report. The plaintiffs claim that this demonstration plan shows that these alternate districts would have produced an efficiency gap bias of only $2 \%$. However, this calculation is made not by assuming that any of the existing candidates in the 2012 elections ran in new districts, but by imputing a baseline partisanship for each new district, and adjusting this baseline for 2012 electoral conditions, assuming all districts are contested by both major parties and no districts are contested by incumbents (Mayer report, p. 31 and 45). As with the previous discussion of imputation of votes in uncontested races, this technique seems appropriate in studying the baseline characteristics of a map for academic purposes. But legislators will of course not draw a map assuming that no incumbents will run, or that all races will be contested. Instead, the actual mapmakers will probably have a fair idea of which districts will be contested
by which incumbents, and which districts are likely to be uncontested. So while it may have been possible to draw a map with a low baseline bias in partisanship absent the effects of incumbency or uncontested elections, this would not be the most accurate data that legislators would be able to access in terms of predicting actual election outcomes.

Moreover, the Mayer plan sets out to predict bias using the actual 2012 election outcome, a narrow statewide victory for the Democrats in terms of aggregated vote totals. But this particular outcome in unknowable to mapmakers at the time maps must be drawn. Mayer points out that this outcome was close to the projection produced by Gaddie or district baseline partisanship prior to the election. But this outcome (where the statewide vote in 2012 closely matched baseline partisanship) was mostly coincidental. It could just as easily have happened that this cycle produced a wave election in favor of either the Democrats or the Republicans, strongly deviating from all baseline estimates. Mayer provides no estimates for the efficiency gap of the demonstration plan under the range of plausible election outcomes facing legislators at the time they were drawing the map.

## VIII. Geographical Bias in Wisconsin and the Nation

## A. Bias from Geographic Dispersion of Partisan: General Arguments

The test proposed in the plaintiffs' complaint allows that a map exceeding the predetermined threshold for bias may rebut the presumption of unconstitutionality by showing such bias is "inevitable given the state's underlying political geography" (paragraph 84). The plaintiffs propose to show that such bias should not be deemed "inevitable" by presenting one specific demonstration plan that, through a series of imputations, would have displayed much lower bias in 2012.

But creating a hypothetical plan with lower bias after knowing the result of a particular election is not a reasonable way to evaluate the propensity of a state's underlying geography to
generate bias, or ability of a nonpartisan actor to anticipate a particular election result prior to the election happening. Instead, evaluation of whether political geography substantially contributed to bias is more appropriately measured by any of several other techniques, including: (1) comparing bias observed in Wisconsin to other comparable states during the same time period; (2) comparing the current map in Wisconsin to previous maps in the same state drawn without partisan motivation; and (3) simulating nonpartisan districts. Any of these methods would suggest that the asymmetric geographic dispersion of partisans makes it much easier and more natural for even a nonpartisan or bipartisan regime to draw a map biased in favor of Republicans in Wisconsin, particularly when the statewide electorate is evenly balanced.

This report does not attempt to simulate nonpartisan districts beyond a simple analysis of ward distribution, but recent research suggests such simulations create substantial Republican bias in state legislatures in several states with similar political geography. Chen and Rodden (2013) show how recent political geography generates substantial Republican bias in legislative elections in states across the nation, even when districts are drawn randomly, while still incorporating values of contiguity and compactness. Chen and Rodden use the geography of Florida as a detailed example, with several very compact urban areas of very concentrated Democratic strength, surrounded by much more sprawling regions of more modest Republican advantage. Yet they simulate random state legislative district in more than 15 additional state (Wisconsin not among the states where data for such simulation was available), and find "that Florida is not an outlier...average bias in favor of Republicans is substantial - surpassing 5\% of state legislative seats - around half the states for which simulations were possible" (Chen \& Rodden 2013, p. 262).

## B. Evidence of growing geographic in nation as a whole

Under multiple different measures, overall bias has been found to be shifting increasingly toward Republicans across the nation in recent decades. Using a very simple methodology, I also find that geography generated an average of $7 \%$ bias in the 2012 congressional elections in states even controlling for the partisanship of districting (Goedert 2014, p. 4). And the Jackman report notes that while the overall average efficiency gap in all state house elections from 1972 to 2014 is very close to zero, the average was significantly more likely to be biased in favor of Democrats in 1970s and 1980s, and more likely to be biased toward Republicans in later decades, especially the 2010s. On p. 44, the Jackman report states that while 5 of the 10 most pro-Republican efficiency gap estimates from the past 40 years were observed in the two most recent cycles (none being in Wisconson), all of the 10 most pro-Democratic estimates occurred prior to 2002.

Additionally, Stephanopoulos and McGhee find that Republican bias in the average state house plan has gradually grown from $-1.5 \%$ in the 1970s and 1980 's to $2.1 \%$ in the 2000 's, peaking at $3.7 \%$ in 2012 (Stephanopoulos \& McGhee 2015, p. 871-2; graph on p. 873). While Stephanopoulos and McGhee attribute much of this growth in prior decades to "favorable trends in voters' residential patterns", they also claim the "spike" in 2012 was caused by more extreme partisan gerrymanders. Nevertheless, this overall bias in favor of Republicans is largely a continuation of a recent trend in political geography. Regardless of how it is measured, geography appears to play a potentially significant role in biasing election results. If the Court is insistent on using efficiency gap as a standard to measure partisan intent, it would seem clear that an adjustment for geography, which is not the result of such intent, should be made in lieu of a predetermined hard-and-fast threshold.

## C. Evidence of asymmetric bias in historical Jackman data

On p. 60-61 of his report, Jackman describes Republican bias as more durable and certain than Democratic bias of the same magnitude. This is apparently noted to suggest that Republican bias observed in a single election should be viewed by the Court as especially dangerous dues to its potential to perpetuate across cycles. But this same observation would also suggest that Republican bias, where observed, is more likely to be due to a more permanent geographic distribution of partisans, rather than more temporary considerations of legislators in anticipation of a single election cycle. This is further supported in Table 1 on p. 55 of the Jackman report. As mentioned above, of the 17 plans that Jackman claims are unambiguous as to sign throughout an entire decade, 16 are biased toward Republicans, and most of these 16 are not Republican gerrymanders.

## D. Specific evidence from Wisconsin

We can see the overall trend toward Republican bias even without partisan intent specifically in the efficiency gap measurements in Wisconsin. From Figure 35 of the Jackman report, Wisconsin saw a larger negative efficiency gap in 2012 than any election in the last 40 years. However, this is just one of nine consecutive cycles of negative efficiency gaps, including seven cycles under two different bipartisan or court-drawn maps, gaps which with slight exceptions at the end of the 2000s, have grown steadily larger over two decades.

And the efficiency gap observed in the most recent 2014 cycle is not at all unusual for recent electoral history in Wisconsin. This is to be expected from geographical bias when tides shift strongly in favor of one party. Using a slightly different but analogous measure of bias, I find in two articles published in Research and Politics that average bias across several congressional maps drawn by Republicans declined from 19\% to 9\% between 2012 and 2014. This decline in bias under the somewhat stronger Republican tide in 2014 is echoed in Jackman's
efficiency gap measurements from Wisconsin, which declines from $14 \%$ to just under $9 \%$. As mentioned above, the efficiency gap found is Wisconsin in 2014 is actually lower than the bias observed under the court drawn state legislative map in Wisconsin in two cycles of the previous decade: 2004 and 2006.

## E. Analysis of Wisconsin Districts at the ward level

Even without regard to a specific district map, we can see the bias inherent in Wisconsin's geography at the ward level. Chen and Rodden posit that bias in several states comes out of a surplus of lean-Republican and safe Democratic pockets of population, compared to relative lack of lean-Democratic and safe Republican pockets. And mapping the distribution of Wisconsin wards confirms this exact pattern.

Based on the 2012 presidential election results, we can estimate what share of the twoparty vote a Democrat would project to win in each ward in an election where each party won $50 \%$ of the statewide vote (data drawn from supplemental attachment to Mayer expert report). Since President Obama won 53.5\% of the two-party statewide vote in 2012, this is most simply done by shifting each ward's actual Democratic vote share down by $3.5 \%$. So a ward that voted $56 \%$ for Obama in 2012 would be estimated to vote $52.5 \%$ Democratic in an evenly balanced election. Figure 1 below shows the proportion of wards, as well as the share of statewide vote these wards comprise, at each level of Democratic support, demonstrating a clear geographic bias favoring lean-Republican wards.


Figure 1. Wisconsin Ward Projections in Evenly Divided Statewide Election (Based on uniform swing from 2012 Presidential Election Results)

The number of wards in Figure 1 peaks at 40-50\% Democratic vote, indicating the surplus of areas that marginally favor Republicans. At the same time, while there are virtually no wards voting overwhelmingly Republican there are several wards that vote overwhelmingly Democratic, and these wards are larger than most other wards in the state.

In an election evenly divided between the parties statewide, Republicans would win $60.2 \%$ of wards, comprising $54.4 \%$ of the voting population. In fact, a majority of all wards in the state ( $50.8 \%$ of wards, comprising $44.3 \%$ of voting population) would be won by Republicans with less than $70 \%$ of the vote. In contrast, less than a third of wards would be won by Democrats with less than $70 \%$ of the vote. Meanwhile, there are many more wards, comprising a much larger share of the population, that were extremely Democratic. In the evenly balanced election, $4 \%$ of wards, comprising $7 \%$ of voting population, would be won by
the Democrat with more $80 \%$ of the vote. Less than $1 \%$ of wards comprising less than $1 \%$ of population would be Republicans by a similarly huge margin.

Overall, it would appear that the recent results in Wisconsin are in line with both a national trend over the past two decades of greater natural Republican bias due to the increasing concentration of Democratic voters in compact urban areas. Republican control of the redistricting process does increase bias toward Republicans in election cycles where the vote share is close to even, but this is highly sensitive to very realistic shifts in the vote share, and should also be considered the context of geographic bias in the same direction.

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Expert report of Professor Simon Jackman
Expert report of Professor Kenneth Mayer
Supplemental Data File: LTSB 2012 Election Data (excel format)

## Cases

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# Analysis of the Efficiency Gaps of Wisconsin's Current Legislative District Plan and Plaintiffs' Demonstration Plan 

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## I. Introduction

My name is Kenneth Mayer and I currently am a Professor of Political Science at the University of Wisconsin-Madison, and a faculty affiliate at the Lafollette School of Public Affairs, at the University. I joined the faculty in 1989. I teach courses on American politics, the presidency, Congress, campaign finance, election law, and electoral systems.

I have been retained by counsel representing the plaintiffs in this lawsuit (the "Plaintiffs") to analyze and provide expert opinions. I have been asked to determine whether, in my opinion, it is possible to create a Wisconsin state legislative map that does not result in systemic partisan advantage, by drawing a legislative district plan that has an efficiency gap as close to zero as possible while complying with federal and state requirements at least as well as the plan enacted by the Wisconsin legislature in Act 43. ${ }^{1}$

I submit this report, which contains the opinions that I intend to give in this matter. I describe my methods for estimating the state Assembly vote in actual and hypothetical state legislative redistricting plans, and for calculating the efficiency gap for Act 43 and for the alternative demonstration plan I drew.

My opinions, which are based on the technical and specialized knowledge that I have gained from my education, training and experience, are premised on commonly used, widely accepted and reliable methods of analysis, the application of the legal requirements of redistricting, and are based on my review and analysis of the following information and materials:

- Redistricting materials available from the Wisconsin legislature at http://legis.wisconsin.gov/gis/data, including Geographic Information System (GIS)

[^6]files for Act 43 districts, and ward level election data for 2012

- Census Bureau data on population, citizenship, and location of institutionalized populations as explained below
- Election data from the 2013-2014 Wisconsin Blue Book for the 2012 State Assembly and presidential elections
- Election data from the Government Accountability Board, including ward level 2012 election results for State Assembly and presidential elections.
- GIS data, including Census population figures, block assignments, and shape files for Wisconsin, available in the GIS program Maptitude for Redistricting
- Files submitted by defendants in Baldus et al. v. Brennan et al.

I conducted my analysis using Stata, Excel, R, and Maptitude for Redistricting.

## II. Qualifications, Publications, Testimony, and Compensation

I have a Ph.D. in political science from Yale University, where my graduate training included courses in econometrics and statistics. My undergraduate degree is from the University of California, San Diego, where I majored in political science and minored in applied mathematics. My curriculum vitae is attached to this report as Exhibit 1.

All publications that I have authored and published in the past ten years appear in my curriculum vitae, attached as Exhibit 1. Those publications include the following peerreviewed journals: Journal of Politics, American Journal of Political Science, Election Law Journal, Legislative Studies Quarterly, Presidential Studies Quarterly, American Politics Research, Congress and the Presidency, Public Administration Review, and PS: Political Science and Politics. I have also published in law reviews, including the Richmond Law Review, the UCLA Pacific Basin Law Journal, and the University of Utah Law Review. My work on campaign finance has been published in Legislative Studies Quarterly, Regulation,

PS: Political Science and Politics, Richmond Law Review, the Democratic Audit of Australia, and in an edited volume on electoral competitiveness published by the Brookings Institution Press. My research on campaign finance has been cited by the Government Accountability Office, and by legislative research offices in Connecticut and Wisconsin.

My work on election administration has been published in the Election Law Journal, American Journal of Political Science, Public Administration Review, and American Politics Research. I was part of a research group retained as a consultant by the Wisconsin Government Accountability Board to review their compliance with federal mandates and reporting systems, and to survey local election officials throughout the state. I serve on the Steering Committee of the Wisconsin Elections Research Center, a unit with the UW-Madison College of Letters and Science. In 2012 I was retained by the U.S. Department of Justice to analyze data and methods regarding Florida's efforts to identify and remove claimed ineligible noncitizens from the statewide file of registered voters.

In the past eight years, I have testified as an expert witness in trial or deposition in the following cases: Baldus et al. v. Brennan et al., 849 F. Supp. 2d 840 (E.D. Wis. 2012);

Milwaukee Branch of the NAACP et al. v. Walker et al., 2014 WI 98, 357 Wis. 2d 469, 851
N.W. 2d 262; McComish et al. v. Brewer et al., No.CV- 08-1550, 2010 WL 2292213 (D. Ariz. June 23, 2010); and Kenosha County v. City of Kenosha, No. 11-CV-1813 (Kenosha County Circuit Court, Kenosha, WI, 2011).

I am being compensated at a rate of $\$ 300$ per hour.

## III. Opinions

## A. Summary

My opinions may be summarized as follows.

- Using a model that estimates baseline ward-level partisanship, I conclude that the redistricting plan enacted by Act 43 is significantly biased against Democrats, with an efficiency gap of $11.69 \%$. The plan achieves this via the use of classic "packing and cracking" gerrymandering techniques: concentrating Democratic voters into districts where they have overwhelming majorities (packing), and drawing other districts so that Democrats constitute partisan minorities well below $50 \%$ and unlikely to win legislative seats (cracking). In doing so, Republicans guarantee a strong majority of legislative seats, even if they obtain well below $50 \%$ of the statewide legislative vote. In 2012, Republicans won $61 \%$ of State Assembly seats ( 60 of 99 ) while achieving only $46.5 \%$ of the statewide vote (as measured by the presidential vote, a common proxy for statewide partisanship).
- Using the same measure of partisan strength that the Wisconsin state legislature used in assessing partisan impact of proposed districts in Act 43, Act 43 has an efficiency gap of $12.36 \%$.
- I created a demonstration redistricting plan (the "Demonstration Plan") that is equivalent to Act 43 on population deviation, has fewer political subdivision splits, and has better compactness scores, with a much lower efficiency gap score of $2.20 \%$. This is less than one-fifth of the Act 43 efficiency gap.
- The Demonstration Plan shows that the partisan advantage secured in Act 43 was in no sense required in order to adhere to the constitutional and statutory requirements of legislative redistricting.


## B. Measuring Partisanship in Actual and Hypothetical Districting Plans

The efficiency gap is a measure of "wasted votes" that fall into two categories: those votes cast for a losing candidate in a district (lost votes), and votes cast for the winning candidate above what is necessary to win (surplus votes). In an existing set of districts, the calculation is based on the actual vote in each district, with adjustments for uncontested races (Stephanopoulos and McGhee 2015). Larger imbalances in the number of wasted votes signify a degree of partisan unfairness against the political party with more wasted votes.

Calculating the efficiency gap in the Demonstration Plan requires estimating what the underlying partisan vote would be in each newly drawn (and hypothetical) district. The gap cannot be estimated by simply rearranging the votes cast in actual Assembly contests into a new
district configuration, as the votes cast for specific Assembly candidates in each district are a function of the electoral environment in that district and whether a race is even contested by both parties. A large literature has developed around the problem of estimating the likely election results in redistricting plan alternatives and calculating summary statistics that characterize existing and hypothetical plans (Gelman and King 1994; Cain 1985).

In most applications, the partisan consequences of a redistricting plan are expressed in terms of the effect on future elections: using prior election results to predict outcomes in subsequent election cycles, or estimating the statewide vote swing required to significantly change the partisan composition of the legislature from one election to the next (Gelman and King 1990; Cain 1985). The results are typically expressed as the estimated two-party vote percentages in each new district (Gelman and King 1994), which are sufficient to forecast who will win an election and calculate swing ratios and seats-votes curves. ${ }^{2}$

My aim is different. Instead of estimating future election results for an existing or proposed hypothetical plan, my goal was to determine whether it was possible to draw a district plan following the 2010 Census that minimized the efficiency gap while maintaining strict fidelity to the federal and state constitutional requirements of population equality, contiguity, compactness, respect for political subdivisions, and compliance with the Voting Rights Act. The efficiency gap is a function of the number of wasted votes, and therefore requires a model that generates predictions of how many votes would have been cast for Democratic and Republican candidates in 2012 in a different district configuration, rather than simply vote

[^7]percentages. My methods provide a way of estimating what the 2012 Assembly election results would have been in such a Demonstration Plan.

Given appropriate data, it is possible to generate reliable and accurate vote count predictions that can be aggregated to any district boundaries. What is required is a set of independent variables that accurately predict the vote in state Assembly elections but which are to the greatest extent possible exogenous to that vote, meaning that the independent variables have underlying values that do not themselves depend on the district vote. If this condition is met, we can estimate what the district vote would have been in an alternative district configuration, since the independent variables do not depend on any particular district configuration. This is not an issue in models that predict future election results, since by definition variables measured today are exogenous to outcomes that occur several years in the future. Because I use one set of election results (the 2012 presidential vote) as part of a model that predicts another set of contemporaneous election results (the 2012 Assembly vote), it is an important but manageable methodological issue.

My method consists of two steps. The first is the construction of a regression model that predicts the 2012 Assembly vote as a function of partisanship, population, demographics, incumbency, and fixed geographic boundaries in Wisconsin's roughly 6,600 wards. In doing so, I establish the empirical relationships between a set of exogenous variables independent of any specific district configurations and the actual Assembly vote in existing wards. In the second step, I use this model to generate a forecast of Assembly vote preferences as a function of these independent variables, and disaggregate this forecast to the Census block level. Using these block level estimates of the Assembly vote, I draw a Demonstration Plan and estimate the Assembly vote and efficiency gap in the resulting districts.

## 1. Step One: A Model of Voting in Assembly Elections

Estimating the Assembly vote in alternative district configurations requires a model that can generate accurate estimates of the underlying partisanship of a district. As I noted above, the most common models regress the observed Assembly vote on measures of district partisan preferences and other variables known to affect the vote, and generate a predicted value of the vote based on the values of the independent variables. Changing district boundaries will change the values of the independent variables as new voters are moved into the district and others moved out, which in turn allows forecasts of what the vote would be in those new districts.

What I am interested in estimating is how many votes will be cast for Democratic and Republican candidates in each district in a demonstration district plan. This involves a different set of variables than is typical in models that evaluate the percentage of votes each party receives, since I require a measure that accounts for both differences in ward populations and variation in turnout.

I use ward level vote totals as the unit of analysis to increase the number of observations available and allow for more precise estimates. Wisconsin's 99 Assembly districts are composed of roughly 6,600 wards, with districts containing between 24 and 153 wards. While the ideal population of an Assembly district is 57,444 , wards have an average population of approximately 869 people, and are far more demographically homogeneous. ${ }^{3}$

[^8]There are four reasons analysis at the ward level is preferable to analysis at the district level. The first is a matter of sheer numbers: the precision of coefficient estimates, forecasting accuracy, and overall statistical power are all strongly related to the number of observations (or sample size). An $n$ of 6,600 is far preferable to an $n$ of 99 , all other things being equal. ${ }^{4}$

The second is the amount of information lost when smaller units are ignored. From a statistical standpoint, using district data when ward data are available imposes the assumption that the values of all of the ward-level variables are equal to the district level variables, when we know this to be untrue immediately upon inspection. Assembly district 1, for example, has 110 populated wards, ranging in population between one and 999 people. In 2012, 73.4\% of the voting age population cast ballots in the Assembly contest, and the victorious Republican Assembly candidate received $51.3 \%$ of the vote. At the ward level, however, there was considerable variation, with the Republican vote percentage ranging from a low of $38.4 \%$ to a high of $75 \%$, and turnout ranging from $50 \%$ to over $90 \%$. Ignoring this information and variation will lead to less accurate estimates and forecasts.

Third, in the second step of the analysis I disaggregate ward level estimates to the block level. Minimizing the differences in size and maximizing the homogeneity across that disaggregation will lead to more accurate block level estimates.

And fourth, each Census block is assigned to a single ward, ${ }^{5}$ with a unique numerical code that identifies the block's location. ${ }^{6}$ These codes allow for disaggregating ward level data

[^9]into blocks and generating inputs for the redistricting software I use in the second step of my analysis.

I use two main sources of data. The first is redistricting data prepared by the Wisconsin Legislative Technology Services Bureau (LTSB), which consists of spreadsheets with ward level Census population data and election results, as well as ward and district shape files containing this data that can be imported into GIS software. ${ }^{7}$ The second source is official election results published by the Government Accountability Board (GAB), both online and in the 2013 edition of the Wisconsin Blue Book.

In my experience working with large data sets, and especially when dealing with complex GIS data, I have found data errors to be a common problem. I assessed the reliability of the LTSB data by checking it against the GAB election data, and found numerous errors that required correction, as well some errors that could not be corrected. ${ }^{8}$ I describe these errors and my corrections in greater detail in an annex to this report. All subsequent references to ward level vote or population counts uses these corrected vote totals.

The regression model used to predict Assembly vote totals takes the standard form of

$$
\mathrm{Y}_{i}=\alpha+\beta \mathrm{X}_{i}+\varepsilon_{i},
$$

where $\mathrm{Y}_{i}$ is the dependent variable in ward $i, \mathrm{X}_{i}$ is a set of independent variables in ward $i$, and $\alpha$, $\beta$, and $\varepsilon_{i}$ are parameters estimated as a function of the variables. The full model is:

$$
\begin{aligned}
& \text { Assembly } \\
& \text { Vote }_{i}
\end{aligned}=\alpha+\beta_{1} \text { Total VEP } P_{i}+\beta_{2} \text { Black VEP }_{i}+\beta_{3}{\text { Hispanic } V E P_{i}}^{\text {A }}
$$

[^10]\[

$$
\begin{gathered}
+\beta_{4} \begin{array}{c}
\text { Democratic } \\
\text { Presidential Vote }_{i}
\end{array}+\beta_{5} \begin{array}{c}
\text { Republican } \\
\text { Presidential Vote }_{i}
\end{array} \\
+\beta_{6} \text { Democratic }_{\text {Incumbent }_{i}}+\beta_{7}^{\text {Republican }_{\text {Incumbent }_{i}}^{\text {Requ }}+\sum_{j=1}^{71} \gamma_{j} \text { County }_{j}+\varepsilon_{i}}
\end{gathered}
$$
\]

Where
Number of votes cast for the Republican or Democratic

Assembly Vote

Black VEP Voting eligible Black population in ward $i$

Hispanic VEP Voting eligible Hispanic population in ward $i$

Democratic Number of votes cast for Barack Obama in the 2012
Presidential Vote

Republican Number of votes cast for Mitt Romney in the 2012
Presidential Vote
Democratic
Incumbent

Republican $\quad 1$ if the Assembly election in ward $i$ has a Republican Incumbent

County
Total VEP presidential election in ward $i$ presidential election in ward $i$

1 if the Assembly election in ward $i$ has a Democratic incumbent, 0 otherwise, multiplied by the VEP in ward $i$ incumbent, 0 otherwise, multiplied by the VEP in ward $i$

Set of fixed effects dummy variables for each county. Dunn County is the excluded value. ${ }^{9}$

The model explains the Assembly vote as a function of four types of variables: district demographics, underlying partisanship, incumbency, and fixed geographic effects.

[^11]
## a. The Dependent Variable: Ward level Assembly Vote

The key quantity of interest in this analysis is the number of Assembly votes for each party, and it is the dependent variable in the model, using LTSB ward data that I corrected using the process outlined above. Since I am interested in estimating actual vote counts and not the percentage of the two party vote, I estimate separate equations for votes received by each party. ${ }^{10}$ Estimating vote counts provides more accuracy than vote percentages, as it controls for variations in turnout across districts. ${ }^{11}$

## b. Independent Variables: Demographic Data

The first three independent variables - Total Voting Age Population (VEP), Black VEP, and Hispanic VEP - are the 2010 Census voting age population counts by ward, adjusted to remove ineligible voters. ${ }^{12}$ Total VEP constitutes a baseline of the size of the voting population, reflecting the fact that the number of votes will be a function of total population. Black and Hispanic VEP are additional controls that reflect the partisan tendencies of key subpopulations as

[^12]well as turnout likelihood. Traditionally, both African American and Hispanic populations vote at lower rates that whites, although in 2012 African American turnout was comparable to white turnout. Hispanic populations vote at lower rates than other demographic groups, in part because of a higher noncitizen population, but also because of socioeconomic factors known to reduce turnout.

I expect weak relationships for these measures because of the importance of the next set of variables, which reflect actual voting in the 2012 presidential election.

## c. Independent Variables: Measures of Partisanship

The next two variables are the number of votes cast for the Democratic and Republican candidates for president in the 2012 election. The presidential vote is widely used as an exogenous measure of district level partisanship (Ansolabehere, Snyder and Stewart 2000, 2001; Gelman and King 1994; Glazier, Grofman, and Robbins 1987; McDonald 2014; Jacobson 2003, 2009), and it correlates very strongly with other more complex measures of partisan strength (Levendusky, Pope, and Jackman 2008).

The presidential vote is, not surprisingly, an extremely strong predictor of the legislative vote. If we know how many votes were cast for the Republican presidential candidate in a ward we will have a very good idea, subject to some conditions, of how many votes will be cast for the Republican candidate in the legislative election in that ward. While not everyone who votes for the Republican presidential candidate will vote for the Republican state legislative candidate, nearly all will, and we can precisely quantify the nature of that relationship.

The strength of the relationship between presidential and Assembly votes is clear in Figures 1 through 3, which plot the total Assembly vote, Republican Assembly vote, and Democratic Assembly vote in 2012 by the respective presidential vote in each contested ward (where voters have an opportunity to express a preference for either party in the legislative race).



Figure 3: Presidential Vote and Assembly Vote 2012


Figure 1 shows that the number of presidential votes cast in a ward is very strongly related to the number of Assembly votes, although almost all wards show a "roll off" as some presidential voters opt not to mark the ballot in the assembly race (the reference line shows where the number of presidential and Assembly votes would be equal). Such drop-offs are ubiquitous in down-ticket races, because voters have less information about lower-level candidates and often have weaker or nonexistent preferences (Wattenberg, McAllister, and Salvanto 2000).

The graphs for the Republican (Figure 2) and Democratic (Figure 3) votes show more variance around this reference line, indicating that some voters are splitting their tickets by voting for a presidential candidate of one party and an Assembly candidate of the other. Nevertheless, the relationship between the number the Republican and Democratic presidential and Assembly votes is apparent. Taken together, these figures indicate that the presidential vote is a very strong predictor of the Assembly vote.

An important property of the presidential vote as an independent variable in this model is that it can be treated as exogenous to (i.e., not caused by) the legislative vote. Exogeneity can be described in two ways. The first is in causal terms. Most voters will vote for the same party for the president and state Assembly, as the above graphs show. These voters are consistent because they are Democrats or Republicans, and partisanship is the factor that explains both vote choices. Other voters will make their Assembly choice based on their presidential vote, because they use party labels as a cue when voting in a down-ticket race. " $[\mathrm{P}]$ arties are generally known by the presidential candidates they nominate, and candidates for state legislative races are a good deal less well known to voters than the congressional candidates who ride presidential coattails" (Campbell 1986, 46). Few voters, if any at all, will decide on an Assembly candidate first and
then vote for president on the basis of their Assembly vote preference. The causal arrow runs from the presidential vote to the Assembly vote, not from the Assembly vote to the presidential vote. This is why we speak of presidential coattails affecting legislative races, and not the other way around (Campbell 1986; Jacobson 2009).

The second reason why the presidential vote is exogenous to the Assembly vote is that it is not affected by local district-level conditions such as incumbency, spending, or candidate quality (Abramowitz, Alexander, and Gunning 2006, 87). The broader factors that influence the presidential vote, and the presidential candidates themselves, are the same in every Assembly district. The presidential vote is affected by underlying partisanship, national conditions and the characteristics of the presidential candidates, factors that are constant whether that vote is aggregated at the state, district, or ward levels.

To put it another way, a change in the statewide presidential vote is virtually certain to affect state legislative election results. Adding or subtracting hundreds of thousands of Democrats or Republicans will alter voting patterns at the district level. However, nobody would expect that the statewide presidential result will be affected by the configuration of legislative districts. The statewide presidential vote would be the same, no matter how the district lines are drawn. Consequently, we can consider the presidential vote as exogenous to, but a causal factor of, the state legislative vote.

## d. Independent Variables: Incumbency

The incumbency advantage is perhaps the most well-known feature of contemporary legislative elections (Jacobson 2009, 30-35). Legislative incumbents rarely lose, and usually win by large margins. All other things being equal, an incumbent will get more votes than a non-
incumbent. The causes of this advantage are less important in this context than its magnitude. ${ }^{13}$
The model takes into account the incumbency advantage by noting whether an incumbent is running in an Assembly district.

Incumbency effects are measured with a dummy variable equal to 1 when a candidate is an incumbent, and 0 otherwise, ${ }^{14}$ multiplied by the ward voting eligible population to create an interactive variable that accounts for differences in size from one ward to the next. Since the dependent variable is an actual vote count, the value of incumbency - in terms of how many additional votes incumbents receive - will vary with the number of voters who reside in a ward.

## e. Independent Variables: County Effects

The last set of variables estimate the effect that county geography has on the Assembly vote. Some counties in Wisconsin are heavily Republican (Ozaukee, Washington, Waukesha) and some heavily Democratic (Dane, Douglas, Milwaukee). It is possible that a voters' county of residence could have an effect on the vote choice, whether because of sorting, socialization or assimilation, or other unobserved effects. Including dummy variables for each county will capture these effects if they exist. There are 71 county variables (excluding Dunn County) set to 1 when a ward is located in that county, 0 otherwise.

[^13]
## f. Estimation and Results

Using Stata IC 11.2 I performed ordinary least squares regression, using 2012 ward data from contested districts where both Republican and Democratic candidates were on the ballot. ${ }^{15}$ Analyzing contested races solves the problem of trying to estimate partisan support in a district where voters have no opportunity to express their support for one side (Gelman and King 1994). The fact that Republicans registered 0 Assembly votes in the $78^{\text {th }}$ district (Madison), and Democrats 0 votes in the $58^{\text {th }}$ district (Washington County), does not mean there are no Republicans in the $78^{\text {th }}$ or Democrats in the $58^{\text {th }}$ districts, or that a Republican or Democratic candidate would receive zero votes if one were on the ballot. Using uncontested races in this initial analysis would produce inaccurate estimates of party strength in those districts.

The results for the Democratic and Republic regression models appear in Table 1. ${ }^{16}$ Most variables show the expected effects, particularly the very strong impact of the presidential vote. The $r^{2}$ values are extremely high, and the standard errors of the regression models (Root MSE) are low. The model is also extremely accurate: when compared to actual ward vote, the model's predictions of the Republican ward totals are within 16 votes, and the Democratic predictions are within 18 votes.

Figure 4 shows the overall accuracy of the model by plotting the predicted ward level vote totals by the actual vote totals in each ward. Predictions for both Democrats and

[^14]Republicans are grouped tightly around the 45-degree line where predicted and actual values would be equal.

Figure 5 shows the accuracy of the model at the district level, which is the more relevant quantity for real-world applicability. I calculated district level results by aggregating wards into the associated Assembly district, using LTSB assignments. The district-level estimates are very close to the actual vote totals, and the average absolute error is 356 votes for Democratic candidates and 344 votes for Republican candidates.

Table 1
Regression Results: 2012 Assembly Votes, Contested Districts
County fixed effect variables not shown,

| Dependent Variable | Independent Variable |  |
| :---: | :---: | :---: |
|  | Assembly Republican Votes | Assembly Democratic Votes |
| Total Voting <br> Eligible <br> Population | $\begin{gathered} 0.009 \\ (.0070) \end{gathered}$ | $\begin{gathered} -0.008 \\ (.0122) \end{gathered}$ |
| Black Voting <br> Eligible <br> Population | $\begin{aligned} & -0.026 \\ & (.0215) \end{aligned}$ | $\begin{aligned} & -0.021 \\ & (.044) \end{aligned}$ |
| Hispanic Voting eligible <br> Population | $\begin{aligned} & -0.0083 \\ & (.0321) \end{aligned}$ | $\begin{gathered} -0.149 * * \\ (.05) \end{gathered}$ |
| Democratic <br> Presidential Votes | $\begin{aligned} & 0.0072 \\ & (.0173) \end{aligned}$ | $\begin{gathered} 0.931^{* * *} \\ (.028) \end{gathered}$ |
| Republican <br> Presidential <br> Votes | $\begin{gathered} 0.946 * * * \\ (.0086) \end{gathered}$ | $\begin{aligned} & 0.013 \\ & (.013) \end{aligned}$ |
| Democratic Assembly Incumbent | $\begin{gathered} -0.021^{* * *} \\ (.006) \end{gathered}$ | $\begin{gathered} 0.028 * * * \\ (.007) \end{gathered}$ |
| Republican Assembly Incumbent | $\begin{gathered} 0.011^{* *} \\ (.0042) \end{gathered}$ | $\begin{gathered} -0.014^{* *} \\ (.005) \end{gathered}$ |
| Constant | $\begin{gathered} -0.92 \\ (7.52) \end{gathered}$ | $\begin{gathered} 9.8 \\ (5.4) \end{gathered}$ |
| $N$ | 5,282 | 5,282 |
| $r^{2}$ | . 9903 | . 9843 |
| Root MS Error | 15.8 | 17.7 |
| Robust standard errors clustered by Assembly District in parentheses.${ }^{*} \mathrm{p}<.05, * * \mathrm{p}<0.01, * * * \mathrm{p}<0.001$ |  |  |

Figure 4: Ward Level Predicted vs. Actual Assembly Vote - 2012


Figure 5: Predicted vs. Actual District Assembly Vote - 2012
Both Parties, Contested Districts


As important as the prediction of actual district vote totals is the model's ability to accurately identify the winner, as the efficiency gap calculation is sensitive to the party of the winners and losers. ${ }^{17}$ The accuracy of the model is shown in Table 2, which gives the actual and predicted vote percentages of the two-party vote for Republican candidates in contested districts. ${ }^{18}$

[^15]Table 2 - Predicted vs. Actual Vote Percentages,
Contested Districts

|  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  |  |
| District |  |  |  |  | | Actual |
| :---: |
| GOP Vote |
| $\%$ | | Predicted |
| :---: |
| GOP Vote |
| $\%$ | | Correct |
| :---: |
| Winner? |$\quad$| Error |
| :---: |
|  |
| 1 |


| 67 | $53.3 \%$ | $53.5 \%$ | $Y$ | $0.2 \%$ |
| :--- | ---: | :--- | :--- | ---: |
| 68 | $52.4 \%$ | $50.7 \%$ | $Y$ | $-1.8 \%$ |
| 69 | $61.2 \%$ | $58.5 \%$ | $Y$ | $-2.7 \%$ |
| 70 | $49.7 \%$ | $50.1 \%$ | N | $0.4 \%$ |
| 71 | $39.0 \%$ | $39.3 \%$ | $Y$ | $0.2 \%$ |
| 72 | $50.2 \%$ | $51.3 \%$ | $Y$ | $1.1 \%$ |
| 74 | $41.0 \%$ | $41.1 \%$ | $Y$ | $0.1 \%$ |
| 75 | $48.9 \%$ | $49.2 \%$ | $Y$ | $0.2 \%$ |
| 80 | $36.1 \%$ | $35.3 \%$ | $Y$ | $-0.8 \%$ |
| 81 | $38.1 \%$ | $39.6 \%$ | $Y$ | $1.4 \%$ |
| 82 | $60.3 \%$ | $61.6 \%$ | $Y$ | $1.4 \%$ |
| 83 | $69.8 \%$ | $71.6 \%$ | $Y$ | $1.9 \%$ |
| 84 | $62.8 \%$ | $61.8 \%$ | $Y$ | $-1.0 \%$ |
| 85 | $48.2 \%$ | $48.7 \%$ | $Y$ | $0.5 \%$ |
| 86 | $55.7 \%$ | $56.1 \%$ | $Y$ | $0.4 \%$ |
| 87 | $58.6 \%$ | $58.3 \%$ | $Y$ | $-0.3 \%$ |
| 88 | $52.5 \%$ | $54.1 \%$ | $Y$ | $1.7 \%$ |
| 89 | $59.1 \%$ | $59.2 \%$ | $Y$ | $0.1 \%$ |
| 90 | $39.6 \%$ | $37.7 \%$ | $Y$ | $-1.9 \%$ |
| 93 | $50.8 \%$ | $52.0 \%$ | $Y$ | $1.2 \%$ |
| 94 | $39.4 \%$ | $39.4 \%$ | $Y$ | $0.0 \%$ |
| 96 | $59.6 \%$ | $59.7 \%$ | $Y$ | $0.1 \%$ |
| 97 | $64.7 \%$ | $64.4 \%$ | $Y$ | $-0.3 \%$ |
| 98 | $70.5 \%$ | $70.0 \%$ | $Y$ | $-0.5 \%$ |
| 99 | $76.3 \%$ | $77.0 \%$ | $Y$ | $0.7 \%$ |

The regression model identifies the correct winner in 70 of 72 districts $(97.2 \%)$; that is, it accurately identifies the candidate who received the most votes. In the two misclassified races, the Republican candidates received $51.9 \%$ and $49.7 \%$ of the vote. The average absolute error in the vote margin is $1.49 \%$.

## g. Out of Sample Forecasting Accuracy

These results, which compare predicted election results to the actual election results, demonstrate that the model is very accurate. A harder test involves the accuracy of predictions using data not in the sample - that is, applying the model to data and election results that are different from the data used to estimate the model. To test the model's out of sample accuracy, I reran the model 72 times (once for every contested district) excluding every ward in one single
contested district each time, ${ }^{19}$ and then used the results of that estimation to predict the vote totals in wards in the excluded district using the independent variable values for those wards. For example, in the first run I excluded all wards in Assembly district 2 (see footnote 20), and estimated the model using data from the other seventy one contested districts. I then used the results to predict the vote totals in the 2 nd district, and compared the prediction to the actual vote totals. Since we know the actual election results in excluded districts, this exercise is a "hard test" of the model's general predictive ability.

Figure 6 and Table 3 show the results for the 60 contested districts in which the full model could be estimated. ${ }^{20}$ The average district forecast error of the Republican vote percentage increased slightly, to $2.1 \%$, but the out of sample forecasts identified the correct winner in 59 out of 60 races ( $98.3 \%$ ). In Figure 6, which plots the actual versus predicted vote totals, the points are not grouped as tightly around the 45-degree line as they are in the full model predictions (Figure 5), but still show a very high degree of accuracy.

Table 3 -Out of Sample Predicted vs. Actual Vote
Percentages, Contested Districts

| Assembly <br> District | Actual <br> GOP Vote <br> $\%$ | Predicted <br> GOP Vote <br> $\%$ | Correct <br> Winner? | Error |
| :---: | :---: | :---: | :---: | :---: |
| 2 | $58.7 \%$ | $59.0 \%$ | $Y$ | $0.3 \%$ |
| 3 | $60.4 \%$ | $57.5 \%$ | $Y$ | $-2.9 \%$ |
| 4 | $55.7 \%$ | $54.3 \%$ | $Y$ | $-1.3 \%$ |
| 5 | $55.9 \%$ | $58.9 \%$ | $Y$ | $2.9 \%$ |
| 13 | $60.6 \%$ | $60.4 \%$ | $Y$ | $-0.2 \%$ |

[^16]| 14 | 59.1\% | 61.0\% | Y | 1.8\% |
| :---: | :---: | :---: | :---: | :---: |
| 15 | 58.3\% | 56.7\% | Y | -1.6\% |
| 20 | 42.4\% | 39.9\% | Y | -2.5\% |
| 21 | 59.3\% | 56.3\% | Y | -3.1\% |
| 23 | 62.3\% | 61.4\% | Y | -0.9\% |
| 24 | 62.4\% | 60.2\% | Y | -2.3\% |
| 25 | 57.7\% | 55.7\% | $Y$ | -2.0\% |
| 26 | 51.3\% | 58.6\% | $Y$ | 7.3\% |
| 27 | 57.8\% | 50.3\% | Y | -7.5\% |
| 28 | 56.2\% | 55.1\% | $Y$ | -1.2\% |
| 29 | 55.9\% | 54.6\% | $Y$ | -1.3\% |
| 30 | 55.8\% | 57.2\% | $Y$ | 1.4\% |
| 31 | 56.5\% | 55.7\% | $Y$ | -0.9\% |
| 32 | 59.1\% | 60.2\% | $Y$ | 1.1\% |
| 33 | 64.9\% | 63.0\% | $Y$ | -1.9\% |
| 37 | 54.3\% | 56.3\% | $Y$ | 2.0\% |
| 38 | 60.0\% | 62.3\% | $Y$ | 2.3\% |
| 39 | 60.4\% | 59.0\% | $Y$ | -1.5\% |
| 41 | 58.0\% | 56.2\% | $Y$ | -1.7\% |
| 42 | 56.6\% | 51.8\% | $Y$ | -4.8\% |
| 43 | 42.3\% | 43.3\% | $Y$ | 1.1\% |
| 44 | 38.4\% | 40.8\% | $Y$ | 2.5\% |
| 45 | 36.1\% | 34.1\% | Y | -2.0\% |
| 46 | 35.2\% | 34.1\% | Y | -1.0\% |
| 47 | 29.0\% | 30.9\% | Y | 1.8\% |
| 50 | 51.7\% | 53.1\% | Y | 1.4\% |
| 51 | 51.9\% | 48.7\% | N | -3.2\% |
| 52 | 60.7\% | 59.4\% | Y | -1.3\% |
| 53 | 60.1\% | 64.4\% | Y | 4.4\% |
| 54 | 39.8\% | 43.8\% | Y | 4.0\% |
| 55 | 65.2\% | 56.0\% | $Y$ | -9.3\% |
| 56 | 58.3\% | 59.9\% | Y | 1.6\% |
| 60 | 71.2\% | 73.9\% | Y | 2.8\% |
| 61 | 55.7\% | 54.9\% | $Y$ | -0.8\% |
| 62 | 53.1\% | 54.5\% | Y | 1.4\% |
| 63 | 58.4\% | 57.1\% | Y | -1.3\% |
| 67 | 53.3\% | 54.7\% | Y | 1.4\% |
| 69 | 61.2\% | 57.2\% | Y | -4.0\% |
| 70 | 49.7\% | 49.7\% | Y | 0.0\% |
| 71 | 39.0\% | 40.1\% | $Y$ | 1.1\% |
| 72 | 50.2\% | 53.0\% | $Y$ | 2.8\% |
| 80 | 36.1\% | 35.1\% | $Y$ | -1.0\% |
| 81 | 38.1\% | 40.8\% | Y | 2.6\% |


| 82 | $60.3 \%$ | $62.0 \%$ | $Y$ | $1.8 \%$ |
| :---: | :---: | :---: | :---: | :---: |
| 83 | $69.8 \%$ | $71.8 \%$ | $Y$ | $2.0 \%$ |
| 84 | $62.8 \%$ | $61.7 \%$ | $Y$ | $-1.1 \%$ |
| 85 | $48.2 \%$ | $49.0 \%$ | $Y$ | $0.8 \%$ |
| 86 | $55.7 \%$ | $56.9 \%$ | $Y$ | $1.2 \%$ |
| 87 | $58.6 \%$ | $54.6 \%$ | $Y$ | $-3.9 \%$ |
| 88 | $52.5 \%$ | $54.6 \%$ | $Y$ | $2.1 \%$ |
| 89 | $59.1 \%$ | $59.0 \%$ | $Y$ | $-0.1 \%$ |
| 90 | $39.6 \%$ | $36.9 \%$ | $Y$ | $-2.7 \%$ |
| 97 | $64.7 \%$ | $64.2 \%$ | $Y$ | $-0.5 \%$ |
| 98 | $70.5 \%$ | $69.9 \%$ | $Y$ | $-0.5 \%$ |
| 99 | $76.3 \%$ | $77.3 \%$ | $Y$ | $1.0 \%$ |



The model does an excellent job accurately forecasting vote totals and election results, and provides a solid foundation for estimating hypothetical vote totals in an alternative district plan.

## h. Comparison to 2011 Republican Expert Baseline Partisanship Measure

The method I have outlined here is a standard technique in the analysis of redistricting plans: creating a baseline measure of partisanship that is independent of a particular district configuration, and applying those estimates to alternative hypothetical district plans.

Indeed, in preparing the district plan that would become Act 43, the state legislature went through the same analytical exercise, generating partisanship measures to forecast what the election results would be in the districts enacted in that plan. The expert that the legislative Republicans relied on to conduct that analysis, Dr. Ronald Keith Gaddie, described the process and method as "an effort to create a partisan normal vote measure or a partisan baselining measure to use to apply to different districts to ascertain their political tendency.,"21 The results of his regression analysis of the districts in Act 43 are in a spreadsheet used to evaluate the plan entitled "Final Map" which contains open seat baseline partisan estimates for existing and new Assembly districts.

Figure 7 compares Dr. Gaddie's open-seat baseline partisanship measure for the Act 43 districts with the equivalent results of my model, excluding the $8^{\text {th }}$ and $9^{\text {th }}$ Assembly districts which were redrawn by the Federal Court and are therefore not comparable. Gaddie's partisan baseline measure is plotted on the x -axis, and my measure on the y -axis. My measure is the expected partisan performance in actual Act 43 districts, with incumbency effects removed. ${ }^{22}$ The two measures are strongly related, indicating that both are capturing stable features of partisanship in Wisconsin. The line is a bivariate regression line produced by using Dr. Gaddie's partisanship estimate as the independent variable and my measure as the dependent variable.

[^17]The r -squared for this regression is 0.96 , indicating that the two measures are almost perfectly related, and are both capturing the same underlying partisanship.

The most important characteristics of Gaddie's measure is that it constitutes a true forecast of what was expected to occur in the 2012 elections, since the measure itself was generated in 2011 using data from the 2004-2010 elections. As I show below, this metric can be used to generate an efficiency gap measure of what was likely to happen (indeed, what did happen) in the 2012 election.


## 2. Step Two - Predicting Votes in a Demonstration District Plan

## a. Creating a Demonstration District Plan

With the model parameters in hand, I can estimate baseline partisanship and vote totals in every ward, including those uncontested by both parties (because I have independent variables in all wards, even when only one party is on the Assembly ballot). For uncontested districts, the predicted ward vote totals are what would be expected if both parties ran a candidate, based on the values of the independent variables in the wards. I then use these predicted ward level vote totals to generate vote estimates at the Census block level, and build a demonstration district using Census blocks as my basic unit. Because the variables used in the model are exogenous to district configuration and the out of sample predictions are accurate, the results of the analysis in Step one represent a valid measure of what the Assembly vote would have been in a different district configuration.

I calculated estimated "open seat" vote totals, by subtracting the incumbency advantage in every district in which an incumbent ran. This is a more accurate method of determining the baseline partisanship of a district, as it removes the effect of incumbents, who may or may not be running in an alternative plan. This baseline process is standard in the discipline, and was used by the expert retained by the state legislature, Dr. Ronald Keith Gaddie, to analyze the partisan effects of Act 43 during the redistricting process.

To obtain block level vote estimates, I disaggregated the ward level predicted values for the Democratic and Republican vote totals to individual blocks in that ward, based on each block's share of the ward vote eligible population. This technique is widely used and accepted in the discipline (McDonald 2014; Pavia. and López-Quílez 2013). Census blocks have a voting eligible population range between 0 and 2,988 , with an average of approximately 17 people. Wards contain an average of 40 blocks, although the range is substantial, with a minimum of 1
and a maximum of 740. At the end of this disaggregation process, I have a predicted Democratic and Republican Assembly vote total for each Census block in the state.

Table 4 shows an illustrative example, using Ward 23 in the city of Waukesha. This ward, located in the southeastern part of the city, had a 2010 Census population of 1,426, a voting age population of 1,089 , and a voting eligible population of 1,071 . The voting model generated estimates of 552 Republican and 318 Democratic votes in an open seat Assembly race in that ward. The ward contains twenty five Census blocks ranging in population from 0 to 127 , with a voting eligible population range of 0 to 115 .

The first column in Table 4 is the block's geographic identifier, a unique code. ${ }^{23}$ The next column is the block's voting eligible population (VEP) calculated as described in the previous section by removing noncitizens and institutionalized persons (although there are no prisons in this ward). The third column is the block's share of the ward's total VEP of 1,071 ; for the first block in the table it is $38 \div 1,071=.0352$, or $3.52 \%$. The next column is block level Republican vote estimate, calculated as $3.52 \%$ the ward Republican vote of 552, or 19.438. While the table rounds these vote totals, I use fractional values in the actual calculations.

[^18]Table 4 - Ward to Block Disaggregation
City of Waukesha Ward 23

| Ward Voting Eligible Population |  |  |  | 1,071 |
| :---: | :---: | :---: | :---: | :---: |
| Ward Estimated Republican Assembly Vote <br> Ward Estimated Democratic Assembly Vote |  |  |  | 552 |
|  |  |  |  | 318 |
| Block Geographic Identifier | Block VEP | Block Share of Ward VEP <br> (Block VEP $\div 1,071$ ) | Block Level Republican Vote Estimate <br> (Block Share * 522) | Block Level Democratic Vote Estimate <br> (Block Share * 318) |
| 551332024001002 | 38 | 3.52\% | 19 | 11 |
| 551332024001003 | 56 | 5.24\% | 29 | 17 |
| 551332024001004 | 65 | 6.06\% | 33 | 19 |
| 551332024001005 | 30 | 2.77\% | 15 | 9 |
| 551332024001007 | 47 | 4.37\% | 24 | 14 |
| 551332024001008 | 81 | 7.57\% | 42 | 24 |
| 551332024001009 | 12 | 1.11\% | 6 | 4 |
| 551332024001010 | 50 | 4.70\% | 26 | 15 |
| 551332024001011 | 26 | 2.46\% | 14 | 8 |
| 551332024001012 | 25 | 2.32\% | 13 | 7 |
| 551332024001013 | 44 | 4.14\% | 23 | 13 |
| 551332024001014 | 60 | 5.57\% | 31 | 18 |
| 551332024001015 | 30 | 2.77\% | 15 | 9 |
| 551332024001016 | 53 | 4.99\% | 28 | 16 |
| 551332024001017 | 0 | 0.00\% | 0 | 0 |
| 551332024002009 | 10 | 0.93\% | 5 | 3 |
| 551332024002010 | 50 | 4.68\% | 26 | 15 |
| 551332024002011 | 65 | 6.06\% | 33 | 19 |
| 551332024002012 | 37 | 3.44\% | 19 | 11 |
| 551332024002013 | 39 | 3.61\% | 20 | 12 |
| 551332024003036 | 41 | 3.78\% | 21 | 12 |
| 551332024003039 | 15 | 1.39\% | 8 | 4 |
| 551332024003040 | 62 | 5.76\% | 32 | 18 |
| 551332024003042 | 22 | 2.01\% | 11 | 6 |
| 551332025005011 | 115 | 10.73\% | 59 | 34 |

Next, I input this block level data into a commercial GIS software package used for redistricting (Maptitude for Redistricting 2013, Build 2060) matching each block in the database of estimated votes with the same block in the Maptitude data using the block identification code.

Finally, I drew a redistricting plan with the goal of minimizing the efficiency gap while adhering to the Wisconsin and federal Constitutional requirements of equal population, contiguity, compactness, and respect for political subdivisions. Beyond these criteria. the primary decision rule was creating competitive districts where possible, and balancing the number of districts with large Democratic and Republican majorities.

Figures 8 and 9 show the statewide map and the districts in the Milwaukee area.

Figure 8 - Demonstration Plan Statewide Map


Figure 9 - Demonstration Plan - Milwaukee Area


## b. Constitutional and Statutory Requirements

Table 5 shows the summary data for the Demonstration Plan (the full tables are in the annex to this report) and comparison data for the actual 2012 plan implemented in Act $43 .{ }^{24}$ The Demonstration Plan has a marginally larger population deviation, but is well below even the strictest standards applied to state legislative districts (a difference of $0.1 \%$ translates into 57 people). The population range in the Demonstration Plan is 57,191 to 57,686 , a difference of 495 people. Given the ideal Assembly district population of 57,444 , this is a deviation of $0.86 \%$. The Demonstration Plan is more compact on average than Act 43, and has fewer municipal splits (119 compared to 120 in Act 43). On all constitutional requirements, the Demonstration Plan is comparable to Act 43.

Table 5 - Plan Comparison to Act 43

|  | Demonstration <br> Plan | Act 43 |
| :---: | :---: | :---: |
| Population Deviation | $0.86 \%$ | $0.76 \%$ |
| Average Compactness (Reock) | 0.41 | 0.28 |
| Number of <br> Municipal Splits | County <br> City <br> Town <br> Village | 55 |

Act 43 created six majority-minority Black population districts (numbers 10-12 and 1618), ranging from $56.7 \%-67.6 \%$ Black population, and from 51.1\%-61.8\% Black voting age population. The Demonstration Plan retains six Majority Black Assembly districts, ranging from $60.0 \%$ to $63.4 \%$ Black population, and from $56.2 \%$ to $60.5 \%$ Black voting age population:

[^19]| Table 6 - Black Majority Districts in Demonstration Plan |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly <br> District | Population | Voting Age <br> Population | Black <br> Population | Black <br> Percentage <br> of Population | Black <br> Voting Age <br> Population | BVAP\% |
| 10 | 57,195 | 41,528 | 36,593 | $64.0 \%$ | 25,125 | $60.5 \%$ |
| 11 | 57,455 | 40,510 | 34,822 | $60.6 \%$ | 22,762 | $56.2 \%$ |
| 12 | 57,420 | 38,774 | 34,923 | $60.8 \%$ | 21,829 | $56.3 \%$ |
| 16 | 57,282 | 42,469 | 36,321 | $63.4 \%$ | 23,920 | $56.3 \%$ |
| 17 | 57,437 | 39,639 | 34,450 | $60.0 \%$ | 22,275 | $56.2 \%$ |
| 18 | 57,241 | 40,840 | 35,316 | $61.7 \%$ | 24,054 | $58.9 \%$ |

In Baldus et al. v. Brennan et al., a federal Court created a majority Latino district in Milwaukee (the $8^{\text {th }}$ Assembly District). The Demonstration Plan retains the boundaries of this district thereby insuring compliance with Section 2 of the Voting Rights Act.

## C. Efficiency Gap Calculations

With the model described in Step one above and the block-level partisanship baseline it generates, I can analyze any existing or hypothetical district configuration and generate predicted vote totals and efficiency gap measures for the Demonstration Plan.

## 1. Analysis of Act 43

Any discussion of Act 43 must begin with the basic fact that in 2012 Republicans achieved a 60-39 majority in the Assembly in an election in which the Democratic Party achieved $53.5 \%$ of the statewide two-party presidential vote. The imbalance between the Republican Party's statewide vote margin at the top of the ticket (46.5\%) and its Assembly majority $(60.6 \%)$ turns the very notion of partisan symmetry on its head. That standard, according to King and Grofman $(2007,8)$ "requires that the number of seats one party would
receive if it garnered a particular percentage of the vote be identical to the number of seats the other party would receive if it had received the same percentage of the vote" $(2007,8)$. Here, it means that Democrats would have had to obtain 60 Assembly seats with $46.5 \%$ of the vote, an absurd proposition that requires a party's legislative seat share to go $u p$ as its share of the vote goes down.

This result was achieved via the classic gerrymandering strategies of packing and cracking. Figure 10, a histogram of Republican two party vote percentages in 2012, shows the pattern. Here, the bars to the right of $50 \%$ indicate a Republican victory. Twenty three Democratic candidates were uncontested, indicating a significant level of packing (the bar at the far left side of the figure); uncontested races occur largely when one party sees zero probability of winning because the majority party has such overwhelming majorities in the district. By contrast, only four Republicans were uncontested. Act 43 also successfully cracked Democratic majorities in other districts, creating Republican majorities that were either marginal (twelve in the $50-55 \%$ range) or relatively safe (thirty nine in the $55-65 \%$ range). The 2012 results are consistent with what was forecast in 2011, as shown by Figure 11, a histogram of Dr. Gaddie's baseline partisanship measure for Act 43 districts. This measure forecast fifty one Assembly districts with between $50 \%$ and $65 \%$ Republican vote share. This is the same number that actually occurred, fifty one.

Figure 12 shows the baseline partisanship district forecasts for Act 43 , using the model outline in Step one, above. It is very similar to Dr. Gaddie's forecast and the actual results: it forecast fifty districts with between $50 \%$ and $65 \%$ Republican vote share.

Figure 10: Actual 2012 Republican Assembly Vote in Act 43 Districts


Figure 11: Republican Vote Forecast in Act 43 Districts- Gaddie Measure


Figure 12: Act 43 Baseline Partisan Measure


The treatment of the city of Sheboygan shows how this cracking was achieved.
Sheboygan is a city on the Lake Michigan shoreline with a population of 49,285. It is a strongly Democratic area, voting 58.7\%-41.3\% for Obama in 2012; my baseline partisanship estimate for the city is $58.2 \%$. The city is small enough to be contained in a single Assembly district in which it would constitute $86 \%$ of the ideal population, and it was entirely within the $26^{\text {th }}$ Assembly district in both the 1992 and 2001 redistricting rounds. The areas surrounding it the Village of Kohler and the Towns of Sheboygan and Wilson are all strongly Republican (with vote percentages for Romney of $62.8 \%, 56.3 \%$, and $59.4 \%$, respectively; together, these municipalities constitute an area that is $58.2 \%$ Republican, as measured by the presidential vote).

Keeping the city of Sheboygan together would have created a Democratic district, made up of the city itself (58.7\% Democratic) with the remaining 14\% of population drawn from one
of the Republican areas around it. The result would have been a District that was roughly $54 \%$ 56\% Democratic.

Act 43, however, split Sheboygan into separate Assembly districts, placing 32,640 residents of the city into the $26^{\text {th }}$ District, and 16,645 into the $27^{\text {th }}$. With the city split, these areas were combined into the Republican areas surrounding the city, producing two Republican districts: the $26^{\text {th }}$ ( $51.3 \%$ Republican in the 2012 Assembly race; baseline open seat partisanship measure of $53.3 \%$ ) and the $27^{\text {th }}$ ( $57.9 \%$ Republican in the 2012 Assembly race, baseline open seat partisanship measure of $52.3 \%$ ).

Figure 13, below, shows the split into Districts 26 and 27:


Figure 13-Act 43 Treatment of Sheboygan

## 2. Efficiency Gap Calculations for Act 43 and The Demonstration Plan

Recall that the efficiency gap is a measure of gerrymandering based on the difference in the number of "wasted votes." Votes cast for losing candidates are wasted, as are surplus votes for winning candidates above what is necessary to win. The gap is defined as the difference between the sum of wasted votes for the two parties, divided by the total number of votes cast in the election.

Comparing a hypothetical district plan (where vote totals are predicted) to an existing district plan (where vote totals are known) requires care, in large part because it can be difficult
to know with certainty what districts will have incumbents (or how incumbents might rearrange themselves after a redistricting cycle), and because not every district will be contested in an actual election (Stephanopoulos and McGhee 2015).

Handling uncontested races is a straightforward problem; the key is applying a consistent rule to all plans being compared. In the efficiency gap calculation for my plan, I measure underlying partisan strength in each district by estimating the number of votes that would be cast for each party in an open seat election each district, assuming that all races are contested. In the actual 2012 Assembly elections, only 72 of 99 seats were contested by both major parties, leaving 27 uncontested races. Uncontested races by themselves will not necessarily have a dramatic effect on efficiency gap calculations as long as the number of races is small, or if uncontested districts are evenly split between the parties (as a rule, one uncontested race with only a Democrat will cancel out one uncontested race with only a Republican, conditioned on the number of votes cast in each race). But a significant imbalance in uncontested races will have a material effect on the results. Of the 27 uncontested races in 2012, 23 were in Democratic districts and only 4 in Republican districts.

In the academic redistricting literature, uncontested seats are typically handled by imputing what the vote totals would have been if a race had been contested (Gelman and King 1990), or assigning each uncontested race a $75 \%-25 \%$ vote split in favor of the party whose candidate ran unopposed (Gelman and King 1994; Stephanopoulos and McGhee 2015). Because I have direct measures of partisanship and vote predictions, I am able to generate accurate estimates of what the vote totals would have been in Act 43's uncontested districts had both parties fielded candidates. In applying this method to the uncontested districts in the 2012 State Assembly elections, I create two directly equivalent sets of data: one for the Demonstration Plan,
with predicted values of open seat vote totals for all districts, and one for the districts created in Act 43, using open seat estimates for each district. Efficiency gap results for the two redistricting plans constructed this way can be compared directly.

Table 7 shows the full set of efficiency gap calculations for the Demonstration Plan, with incumbency effects removed. For each district I calculate an estimated Democratic and Republican vote total, and forecast a winner. The resulting columns show the number of "wasted votes," counting all votes cast for a losing candidates, and surplus votes for winning candidates (equal to $1 / 2$ of the margin of victory). Totals for each party are summed, and the efficiency gap calculated as the Net Wasted Votes (here, Democratic Wasted Votes - Republican Wasted Votes) divided by the total number of votes cast in the election.

The data in Table 7 (on page 48) show that the Demonstration Plan results in 741,984 wasted Democratic votes (column E), obtained by adding the number of lost Democratic votes cast for losing candidates $(566,634$, column $A$ ) and the number of surplus Democratic votes cast for winners above what was necessary to win $(175,350$, column C). The same calculation for Republicans (using columns B and D) results in 689,570 wasted Republican votes. The difference between these two numbers, 781,984-689,570 $=62,414$ net wasted Democratic votes. Dividing 62,414 by the predicted total number of votes $2,843,108$, produces the baseline efficiency gap for my plan, . 0220 , or $2.20 \%$.

Table 8 (on page 50) shows the same calculation for Act 43 districts, using estimated partisan vote totals with incumbent advantages removed. Act 43 resulted in a total of 332,552 net wasted Democratic votes. The efficiency gap of Act 43 is $11.69 \%$, more than five times larger than the Demonstration Plan.

Table 9 (on page 52) shows the efficiency gap calculation for the partisan baseline prediction used by Dr. Gaddie during the drawing of the Act 43 districts, applying his partisanship division to the total number of votes predicted from my model in each district. As described above in section $\operatorname{III}(\mathrm{B})(1)(\mathrm{h})$ above, this is the predicted baseline partisanship measure of Act 43. It produces a forecast Efficiency Gap for Act 43 of 12.36\%.

Table 10 summarizes these results:

|  | Table 10: Summary Statistics for Redistricting Plans |  |  |
| :---: | :---: | :---: | :---: |
|  | My Plan Baseline | Act 43 <br> Baseline | Act 43 - <br> Gaddie <br> Measure |
| party split (R-D) | 48-51 | 57-42 | 58-41 |
| Wasted Republican Votes | 679,570 | 544,893 | 535,057 |
| Wasted Democratic Votes | 741,984 | 877,445 | 886,403 |
| Gap | 62,414 | 332,552 | 351,346 |
| Total Democratic Votes | 1,454,117 | 1,454,717 | 1,394,018 |
| Total Republican Votes | 1,388,991 | 1,389,958 | 1,448,901 |
| Total Votes | 2,843,108 | 2,844,676 | 2,842,919 |
| Efficiency Gap (gap/total votes) | 2.20\% | 11.69\% | 12.36\% |

Three things are worth emphasizing. The first is that the predicted partisan effect of Act 43, represented by the Gaddie metric, produced an efficiency gap calculation (12.36\%) that was very close to the actual partisan effect of Act 43, as measured by the efficiency gap calculation for the actual 2012 partisan baseline (11.69\%). In brief, the architects of the Act 43 districts expected a partisan result that was almost identical to what actually occurred. The second is the large reduction in the efficiency gap that I am able to produce, which I have achieved without any departure from the core constitutional and statutory requirements of redistricting. The

Demonstration Plan is equivalent to Act 43 on all key criteria: population deviation, compactness, number of political subdivision splits, and compliance with the Voting Rights Act. At the same time, I have generated an efficiency gap score $82 \%$ smaller than the Act 43 gap. And third, I have reached this efficiency gap score with virtually identical numbers of Democratic and Republican voters as exist under Act 43. Given that my partisan estimates, once incumbency effects are removed, are entirely exogenous to any particular district configuration, these can be considered the same statewide set of voters. By placing the same voters as exist in Act 43 into a new set of districts designed to minimize the effects of gerrymandering while adhering to constitutional standards, I have generated a plan that is fair to both parties.

Figure 14 shows the distribution of baseline Republican vote predictions in the Demonstration Plan Assembly districts. The districts are far more balanced, with similar numbers of districts between $40 \%-50 \%$ (twenty seven) and between $50 \%-60 \%$ (twenty nine). There are also roughly equal numbers of districts above $65 \%$ (twelve) and below $35 \%$ (sixteen).

Figure 14: Predicted Vote - Demonstration Plan


Table 7-Efficiency Gap Calculation for Demonstration District Plan - No Incumbent Baseline

|  |  |  |  | A | B | C | D | E | F |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly District | Predicted Democratic Votes | Predicted Republican Votes | Predicted <br> Winning <br> Party | Lost Democratic Votes | Lost Republican Votes | Surplus Democratic Votes | Surplus Republican Votes | Wasted Democratic Votes ( $\mathrm{A}+\mathrm{C}$ ) | Wasted Republican Votes (B + D) | Net Wasted Votes (E - F) |
| 1 | 16,259 | 16,414 | Republican | 16259 | 0 | 0 | 78 | 16259 | 78 | 16181 |
| 2 | 11,805 | 10,025 | Democratic | 0 | 10025 | 890 | 0 | 890 | 10025 | -9136 |
| 3 | 11,243 | 17,807 | Republican | 11243 | 0 | 0 | 3282 | 11243 | 3282 | 7961 |
| 4 | 10,881 | 12,790 | Republican | 10881 | 0 | 0 | 955 | 10881 | 955 | 9926 |
| 5 | 13,497 | 13,845 | Republican | 13497 | 0 | 0 | 174 | 13497 | 174 | 13323 |
| 6 | 11,045 | 17,627 | Republican | 11045 | 0 | 0 | 3291 | 11045 | 3291 | 7753 |
| 7 | 22,822 | 10,214 | Democratic | 0 | 10214 | 6304 | 0 | 6304 | 10214 | -3910 |
| 8 | 7,192 | 1,695 | Democratic | 0 | 1695 | 2749 | 0 | 2749 | 1695 | 1054 |
| 9 | 10,497 | 5,635 | Democratic | 0 | 5635 | 2431 | 0 | 2431 | 5635 | -3205 |
| 10 | 25,348 | 3,270 | Democratic | 0 | 3270 | 11039 | 0 | 11039 | 3270 | 7769 |
| 11 | 22,374 | 4,855 | Democratic | 0 | 4855 | 8759 | 0 | 8759 | 4855 | 3904 |
| 12 | 20,041 | 4,039 | Democratic | 0 | 4039 | 8001 | 0 | 8001 | 4039 | 3962 |
| 13 | 15,950 | 16,510 | Republican | 15950 | 0 | 0 | 280 | 15950 | 280 | 15670 |
| 14 | 13,575 | 13,799 | Republican | 13575 | 0 | 0 | 112 | 13575 | 112 | 13464 |
| 15 | 13,412 | 14,901 | Republican | 13412 | 0 | 0 | 745 | 13412 | 745 | 12667 |
| 16 | 21,234 | 2,856 | Democratic | 0 | 2856 | 9189 | 0 | 9189 | 2856 | 6333 |
| 17 | 21,769 | 3,569 | Democratic | 0 | 3569 | 9100 | 0 | 9100 | 3569 | 5531 |
| 18 | 23,817 | 4,954 | Democratic | 0 | 4954 | 9431 | 0 | 9431 | 4954 | 4477 |
| 19 | 15,160 | 10,904 | Democratic | 0 | 10904 | 2128 | 0 | 2128 | 10904 | -8776 |
| 20 | 14,118 | 12,901 | Democratic | 0 | 12901 | 609 | 0 | 609 | 12901 | -12292 |
| 21 | 12,257 | 16,911 | Republican | 12257 | 0 | 0 | 2327 | 12257 | 2327 | 9930 |
| 22 | 18,335 | 14,831 | Democratic | 0 | 14831 | 1752 | 0 | 1752 | 14831 | -13079 |
| 23 | 10,922 | 25,459 | Republican | 10922 | 0 | 0 | 7268 | 10922 | 7268 | 3654 |
| 24 | 8,667 | 25,868 | Republican | 8667 | 0 | 0 | 8601 | 8667 | 8601 | 66 |
| 25 | 12,179 | 18,248 | Republican | 12179 | 0 | 0 | 3034 | 12179 | 3034 | 9145 |
| 26 | 13,251 | 14,527 | Republican | 13251 | 0 | 0 | 638 | 13251 | 638 | 12613 |
| 27 | 14,935 | 11,755 | Democratic | 0 | 11755 | 1590 | 0 | 1590 | 11755 | -10165 |
| 28 | 12,617 | 15,591 | Republican | 12617 | 0 | 0 | 1487 | 12617 | 1487 | 11131 |
| 29 | 14,180 | 12,954 | Democratic | 0 | 12954 | 613 | 0 | 613 | 12954 | -12341 |
| 30 | 11,308 | 15,165 | Republican | 11308 | 0 | 0 | 1929 | 11308 | 1929 | 9379 |
| 31 | 11,304 | 16,117 | Republican | 11304 | 0 | 0 | 2406 | 11304 | 2406 | 8898 |
| 32 | 12,685 | 13,787 | Republican | 12685 | 0 | 0 | 551 | 12685 | 551 | 12135 |
| 33 | 14,609 | 10,151 | Democratic | 0 | 10151 | 2229 | 0 | 2229 | 10151 | -7922 |
| 34 | 13,139 | 15,690 | Republican | 13139 | 0 | 0 | 1275 | 13139 | 1275 | 11864 |
| 35 | 11,288 | 16,503 | Republican | 11288 | 0 | 0 | 2607 | 11288 | 2607 | 8681 |
| 36 | 11,516 | 14,997 | Republican | 11516 | 0 | 0 | 1741 | 11516 | 1741 | 9775 |
| 37 | 9,222 | 22,240 | Republican | 9222 | 0 | 0 | 6509 | 9222 | 6509 | 2713 |
| 38 | 9,710 | 25,021 | Republican | 9710 | 0 | 0 | 7655 | 9710 | 7655 | 2055 |
| 39 | 10,747 | 17,526 | Republican | 10747 | 0 | 0 | 3390 | 10747 | 3390 | 7357 |
| 40 | 15,061 | 13,947 | Democratic | 0 | 13947 | 557 | 0 | 557 | 13947 | -13391 |
| 41 | 16,784 | 13,120 | Democratic | 0 | 13120 | 1832 | 0 | 1832 | 13120 | -11288 |
| 42 | 13,254 | 12,282 | Democratic | 0 | 12282 | 486 | 0 | 486 | 12282 | -11796 |
| 43 | 12,658 | 13,606 | Republican | 12658 | 0 | 0 | 474 | 12658 | 474 | 12184 |
| 44 | 16,477 | 10,886 | Democratic | 0 | 10886 | 2795 | 0 | 2795 | 10886 | -8091 |
| 45 | 16,352 | 13,589 | Democratic | 0 | 13589 | 1382 | 0 | 1382 | 13589 | -12207 |
| 46 | 20,583 | 11,418 | Democratic | 0 | 11418 | 4582 | 0 | 4582 | 11418 | -6835 |
| 47 | 20,208 | 9,888 | Democratic | 0 | 9888 | 5160 | 0 | 5160 | 9888 | -4728 |


| 48 | 24,457 | 8,840 | Democratic | 0 | 8840 | 7808 | 0 | 7808 | 8840 | -1032 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 49 | 13,625 | 13,477 | Democratic | 0 | 13477 | 74 | 0 | 74 | 13477 | -13403 |
| 50 | 12,289 | 13,709 | Republican | 12289 | 0 | 0 | 710 | 12289 | 710 | 11579 |
| 51 | 14,760 | 13,323 | Democratic | 0 | 13323 | 718 | 0 | 718 | 13323 | -12605 |
| 52 | 12,376 | 19,416 | Republican | 12376 | 0 | 0 | 3520 | 12376 | 3520 | 8857 |
| 53 | 12,388 | 13,362 | Republican | 12388 | 0 | 0 | 487 | 12388 | 487 | 11902 |
| 54 | 14,032 | 12,240 | Democratic | 0 | 12240 | 896 | 0 | 896 | 12240 | -11344 |
| 55 | 13,565 | 15,300 | Republican | 13565 | 0 | 0 | 868 | 13565 | 868 | 12697 |
| 56 | 12,553 | 14,518 | Republican | 12553 | 0 | 0 | 983 | 12553 | 983 | 11570 |
| 57 | 14,897 | 13,016 | Democratic | 0 | 13016 | 941 | 0 | 941 | 13016 | -12075 |
| 58 | 9,325 | 21,180 | Republican | 9325 | 0 | 0 | 5927 | 9325 | 5927 | 3398 |
| 59 | 11,565 | 21,984 | Republican | 11565 | 0 | 0 | 5209 | 11565 | 5209 | 6356 |
| 60 | 8,756 | 22,415 | Republican | 8756 | 0 | 0 | 6830 | 8756 | 6830 | 1926 |
| 61 | 12,933 | 16,576 | Republican | 12933 | 0 | 0 | 1822 | 12933 | 1822 | 11112 |
| 62 | 15,181 | 9,999 | Democratic | 0 | 9999 | 2591 | 0 | 2591 | 9999 | -7408 |
| 63 | 15,640 | 9,902 | Democratic | 0 | 9902 | 2869 | 0 | 2869 | 9902 | -7033 |
| 64 | 15,089 | 13,470 | Democratic | 0 | 13470 | 810 | 0 | 810 | 13470 | -12660 |
| 65 | 12,721 | 19,816 | Republican | 12721 | 0 | 0 | 3547 | 12721 | 3547 | 9173 |
| 66 | 16,286 | 6,362 | Democratic | 0 | 6362 | 4962 | 0 | 4962 | 6362 | -1401 |
| 67 | 15,321 | 14,226 | Democratic | 0 | 14226 | 547 | 0 | 547 | 14226 | -13678 |
| 68 | 11,958 | 12,124 | Republican | 11958 | 0 | 0 | 83 | 11958 | 83 | 11875 |
| 69 | 17,902 | 12,022 | Democratic | 0 | 12022 | 2940 | 0 | 2940 | 12022 | -9083 |
| 70 | 18,661 | 12,266 | Democratic | 0 | 12266 | 3197 | 0 | 3197 | 12266 | -9069 |
| 71 | 15,081 | 13,884 | Democratic | 0 | 13884 | 599 | 0 | 599 | 13884 | -13285 |
| 72 | 11,180 | 16,542 | Republican | 11180 | 0 | 0 | 2681 | 11180 | 2681 | 8500 |
| 73 | 17,137 | 10,785 | Democratic | 0 | 10785 | 3176 | 0 | 3176 | 10785 | -7609 |
| 74 | 17,712 | 14,219 | Democratic | 0 | 14219 | 1747 | 0 | 1747 | 14219 | -12472 |
| 75 | 13,902 | 17,700 | Republican | 13902 | 0 | 0 | 1899 | 13902 | 1899 | 12002 |
| 76 | 30,929 | 6,811 | Democratic | 0 | 6811 | 12059 | 0 | 12059 | 6811 | 5248 |
| 77 | 26,708 | 6,059 | Democratic | 0 | 6059 | 10325 | 0 | 10325 | 6059 | 4266 |
| 78 | 24,413 | 9,847 | Democratic | 0 | 9847 | 7283 | 0 | 7283 | 9847 | -2564 |
| 79 | 20,439 | 13,294 | Democratic | 0 | 13294 | 3572 | 0 | 3572 | 13294 | -9722 |
| 80 | 20,179 | 11,644 | Democratic | 0 | 11644 | 4267 | 0 | 4267 | 11644 | -7377 |
| 81 | 13,703 | 12,741 | Democratic | 0 | 12741 | 481 | 0 | 481 | 12741 | -12260 |
| 82 | 9,871 | 21,201 | Republican | 9871 | 0 | 0 | 5665 | 9871 | 5665 | 4206 |
| 83 | 9,241 | 23,075 | Republican | 9241 | 0 | 0 | 6917 | 9241 | 6917 | 2324 |
| 84 | 11,990 | 22,700 | Republican | 11990 | 0 | 0 | 5355 | 11990 | 5355 | 6634 |
| 85 | 10,028 | 13,190 | Republican | 10028 | 0 | 0 | 1581 | 10028 | 1581 | 8448 |
| 86 | 13,853 | 13,494 | Democratic | 0 | 13494 | 180 | 0 | 180 | 13494 | -13314 |
| 87 | 11,358 | 17,003 | Republican | 11358 | 0 | 0 | 2823 | 11358 | 2823 | 8535 |
| 88 | 14,209 | 11,142 | Democratic | 0 | 11142 | 1533 | 0 | 1533 | 11142 | -9609 |
| 89 | 13,374 | 15,771 | Republican | 13374 | 0 | 0 | 1199 | 13374 | 1199 | 12175 |
| 90 | 11,349 | 17,468 | Republican | 11349 | 0 | 0 | 3059 | 11349 | 3059 | 8290 |
| 91 | 14,807 | 13,845 | Democratic | 0 | 13845 | 481 | 0 | 481 | 13845 | -13364 |
| 92 | 14,907 | 14,594 | Democratic | 0 | 14594 | 157 | 0 | 157 | 14594 | -14437 |
| 93 | 12,441 | 18,057 | Republican | 12441 | 0 | 0 | 2808 | 12441 | 2808 | 9633 |
| 94 | 16,171 | 11,759 | Democratic | 0 | 11759 | 2206 | 0 | 2206 | 11759 | -9553 |
| 95 | 19,769 | 9,949 | Democratic | 0 | 9949 | 4910 | 0 | 4910 | 9949 | -5040 |
| 96 | 14,665 | 13,836 | Democratic | 0 | 13836 | 415 | 0 | 415 | 13836 | -13421 |
| 97 | 11,492 | 24,222 | Republican | 11492 | 0 | 0 | 6365 | 11492 | 6365 | 5128 |
| 98 | 9,864 | 24,773 | Republican | 9864 | 0 | 0 | 7454 | 9864 | 7454 | 2410 |
| 99 | 10,783 | 19,160 | Republican | 10783 | 0 | 0 | 4188 | 10783 | 4188 | 6594 |
| TOTALS | 1,454,117 | 1,388,991 |  | 566,634 | 536,783 | 175,350 | 142,787 | 741,984 | 679,570 | 62,414 |

Table 8 - Efficiency Gap Calculation for Act 43-No Incumbent Baseline

|  |  |  |  | A | B | C | D | E | F |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly District | Predicted Democratic Votes | Predicted Republican Votes | Predicted Winning Party | Lost Democratic Votes | Lost Republican Votes | Surplus Democra tic Votes | Surplus Republic an Votes | Wasted Democratic Votes $(A+C)$ | Wasted Republican Votes $(B+D)$ | Net <br> Wasted Votes (E - F) |
| 1 | 16,235 | 16,628 | Republican | 16235 | 0 | 0 | 197 | 16235 | 197 | 16038 |
| 2 | 12,398 | 16,357 | Republican | 12398 | 0 | 0 | 1980 | 12398 | 1980 | 10419 |
| 3 | 12,623 | 16,636 | Republican | 12623 | 0 | 0 | 2006 | 12623 | 2006 | 10617 |
| 4 | 13,926 | 15,576 | Republican | 13926 | 0 | 0 | 825 | 13926 | 825 | 13101 |
| 5 | 12,710 | 16,017 | Republican | 12710 | 0 | 0 | 1654 | 12710 | 1654 | 11056 |
| 6 | 10,929 | 14,938 | Republican | 10929 | 0 | 0 | 2005 | 10929 | 2005 | 8924 |
| 7 | 13,793 | 11,778 | Democratic | 0 | 11778 | 1007 | 0 | 1007 | 11778 | -10771 |
| 8 | 7,342 | 1,738 | Democratic | 0 | 1738 | 2802 | 0 | 2802 | 1738 | 1064 |
| 9 | 10,023 | 4,533 | Democratic | 0 | 4533 | 2745 | 0 | 2745 | 4533 | -1787 |
| 10 | 25,306 | 2,897 | Democratic | 0 | 2897 | 11205 | 0 | 11205 | 2897 | 8308 |
| 11 | 21,698 | 3,368 | Democratic | 0 | 3368 | 9165 | 0 | 9165 | 3368 | 5797 |
| 12 | 19,700 | 5,222 | Democratic | 0 | 5222 | 7239 | 0 | 7239 | 5222 | 2018 |
| 13 | 13,345 | 20,358 | Republican | 13345 | 0 | 0 | 3506 | 13345 | 3506 | 9839 |
| 14 | 14,499 | 21,025 | Republican | 14499 | 0 | 0 | 3263 | 14499 | 3263 | 11235 |
| 15 | 13,006 | 17,310 | Republican | 13006 | 0 | 0 | 2152 | 13006 | 2152 | 10853 |
| 16 | 22,293 | 2,342 | Democratic | 0 | 2342 | 9975 | 0 | 9975 | 2342 | 7633 |
| 17 | 24,088 | 4,047 | Democratic | 0 | 4047 | 10020 | 0 | 10020 | 4047 | 5973 |
| 18 | 22,204 | 2,692 | Democratic | 0 | 2692 | 9756 | 0 | 9756 | 2692 | 7064 |
| 19 | 22,759 | 10,364 | Democratic | 0 | 10364 | 6198 | 0 | 6198 | 10364 | -4166 |
| 20 | 16,066 | 12,856 | Democratic | 0 | 12856 | 1605 | 0 | 1605 | 12856 | -11252 |
| 21 | 12,566 | 15,324 | Republican | 12566 | 0 | 0 | 1379 | 12566 | 1379 | 11187 |
| 22 | 11,290 | 22,958 | Republican | 11290 | 0 | 0 | 5834 | 11290 | 5834 | 5456 |
| 23 | 14,260 | 21,633 | Republican | 14260 | 0 | 0 | 3687 | 14260 | 3687 | 10573 |
| 24 | 13,885 | 20,335 | Republican | 13885 | 0 | 0 | 3225 | 13885 | 3225 | 10659 |
| 25 | 12,032 | 15,933 | Republican | 12032 | 0 | 0 | 1950 | 12032 | 1950 | 10082 |
| 26 | 13,639 | 15,559 | Republican | 13639 | 0 | 0 | 960 | 13639 | 960 | 12679 |
| 27 | 14,709 | 16,360 | Republican | 14709 | 0 | 0 | 826 | 14709 | 826 | 13883 |
| 28 | 12,719 | 15,302 | Republican | 12719 | 0 | 0 | 1291 | 12719 | 1291 | 11428 |
| 29 | 12,909 | 14,662 | Republican | 12909 | 0 | 0 | 876 | 12909 | 876 | 12033 |
| 30 | 14,019 | 16,951 | Republican | 14019 | 0 | 0 | 1466 | 14019 | 1466 | 12553 |
| 31 | 13,273 | 15,615 | Republican | 13273 | 0 | 0 | 1171 | 13273 | 1171 | 12102 |
| 32 | 11,255 | 15,359 | Republican | 11255 | 0 | 0 | 2052 | 11255 | 2052 | 9203 |
| 33 | 11,226 | 18,298 | Republican | 11226 | 0 | 0 | 3536 | 11226 | 3536 | 7690 |
| 34 | 12,445 | 19,355 | Republican | 12445 | 0 | 0 | 3455 | 12445 | 3455 | 8991 |
| 35 | 12,270 | 15,525 | Republican | 12270 | 0 | 0 | 1628 | 12270 | 1628 | 10643 |
| 36 | 11,403 | 15,672 | Republican | 11403 | 0 | 0 | 2134 | 11403 | 2134 | 9269 |
| 37 | 12,707 | 16,202 | Republican | 12707 | 0 | 0 | 1747 | 12707 | 1747 | 10960 |
| 38 | 12,668 | 19,129 | Republican | 12668 | 0 | 0 | 3231 | 12668 | 3231 | 9437 |
| 39 | 11,491 | 17,211 | Republican | 11491 | 0 | 0 | 2860 | 11491 | 2860 | 8630 |
| 40 | 11,485 | 13,597 | Republican | 11485 | 0 | 0 | 1056 | 11485 | 1056 | 10429 |
| 41 | 11,719 | 14,492 | Republican | 11719 | 0 | 0 | 1387 | 11719 | 1387 | 10332 |
| 42 | 13,705 | 15,462 | Republican | 13705 | 0 | 0 | 879 | 13705 | 879 | 12826 |
| 43 | 17,380 | 13,075 | Democratic | 0 | 13075 | 2153 | 0 | 2153 | 13075 | -10923 |
| 44 | 16,680 | 10,304 | Democratic | 0 | 10304 | 3188 | 0 | 3188 | 10304 | -7116 |
| 45 | 15,153 | 9,691 | Democratic | 0 | 9691 | 2731 | 0 | 2731 | 9691 | -6959 |
| 46 | 19,173 | 11,534 | Democratic | 0 | 11534 | 3819 | 0 | 3819 | 11534 | -7714 |
| 47 | 21,609 | 9,340 | Democratic | 0 | 9340 | 6135 | 0 | 6135 | 9340 | -3205 |
| 48 | 24,517 | 7,635 | Democratic | 0 | 7635 | 8441 | 0 | 8441 | 7635 | 806 |
| 49 | 12,307 | 13,621 | Republican | 12307 | 0 | 0 | 657 | 12307 | 657 | 11650 |


| 50 | 12,467 | 12,326 | Democratic | 0 | 12326 | 71 | 0 | 71 | 12326 | -12256 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 51 | 14,173 | 13,048 | Democratic | 0 | 13048 | 563 | 0 | 563 | 13048 | -12485 |
| 52 | 11,294 | 15,656 | Republican | 11294 | 0 | 0 | 2181 | 11294 | 2181 | 9113 |
| 53 | 9,875 | 16,753 | Republican | 9875 | 0 | 0 | 3439 | 9875 | 3439 | 6437 |
| 54 | 15,180 | 12,882 | Democratic | 0 | 12882 | 1149 | 0 | 1149 | 12882 | -11733 |
| 55 | 12,634 | 16,971 | Republican | 12634 | 0 | 0 | 2169 | 12634 | 2169 | 10465 |
| 56 | 12,564 | 18,576 | Republican | 12564 | 0 | 0 | 3006 | 12564 | 3006 | 9559 |
| 57 | 14,387 | 11,676 | Democratic | 0 | 11676 | 1355 | 0 | 1355 | 11676 | -10321 |
| 58 | 8,843 | 22,417 | Republican | 8843 | 0 | 0 | 6787 | 8843 | 6787 | 2055 |
| 59 | 8,784 | 21,725 | Republican | 8784 | 0 | 0 | 6471 | 8784 | 6471 | 2313 |
| 60 | 9,848 | 23,989 | Republican | 9848 | 0 | 0 | 7071 | 9848 | 7071 | 2778 |
| 61 | 13,145 | 16,481 | Republican | 13145 | 0 | 0 | 1668 | 13145 | 1668 | 11477 |
| 62 | 14,828 | 17,309 | Republican | 14828 | 0 | 0 | 1240 | 14828 | 1240 | 13588 |
| 63 | 13,233 | 16,830 | Republican | 13233 | 0 | 0 | 1799 | 13233 | 1799 | 11434 |
| 64 | 15,702 | 11,307 | Democratic | 0 | 11307 | 2198 | 0 | 2198 | 11307 | -9109 |
| 65 | 15,105 | 7,929 | Democratic | 0 | 7929 | 3588 | 0 | 3588 | 7929 | -4341 |
| 66 | 16,162 | 5,472 | Democratic | 0 | 5472 | 5345 | 0 | 5345 | 5472 | -127 |
| 67 | 13,769 | 14,674 | Republican | 13769 | 0 | 0 | 453 | 13769 | 453 | 13316 |
| 68 | 13,663 | 13,005 | Democratic | 0 | 13005 | 329 | 0 | 329 | 13005 | -12676 |
| 69 | 11,083 | 14,347 | Republican | 11083 | 0 | 0 | 1632 | 11083 | 1632 | 9451 |
| 70 | 12,211 | 14,387 | Republican | 12211 | 0 | 0 | 1088 | 12211 | 1088 | 11123 |
| 71 | 17,614 | 11,383 | Democratic | 0 | 11383 | 3115 | 0 | 3115 | 11383 | -8267 |
| 72 | 14,294 | 13,895 | Democratic | 0 | 13895 | 199 | 0 | 199 | 13895 | -13696 |
| 73 | 17,353 | 10,784 | Democratic | 0 | 10784 | 3284 | 0 | 3284 | 10784 | -7500 |
| 74 | 17,095 | 13,772 | Democratic | 0 | 13772 | 1662 | 0 | 1662 | 13772 | -12110 |
| 75 | 15,000 | 13,418 | Democratic | 0 | 13418 | 791 | 0 | 791 | 13418 | -12627 |
| 76 | 30,939 | 6,805 | Democratic | 0 | 6805 | 12067 | 0 | 12067 | 6805 | 5262 |
| 77 | 26,925 | 6,041 | Democratic | 0 | 6041 | 10442 | 0 | 10442 | 6041 | 4402 |
| 78 | 24,163 | 9,857 | Democratic | 0 | 9857 | 7153 | 0 | 7153 | 9857 | -2704 |
| 79 | 20,753 | 13,975 | Democratic | 0 | 13975 | 3389 | 0 | 3389 | 13975 | -10586 |
| 80 | 20,369 | 12,604 | Democratic | 0 | 12604 | 3882 | 0 | 3882 | 12604 | -8722 |
| 81 | 16,310 | 12,356 | Democratic | 0 | 12356 | 1977 | 0 | 1977 | 12356 | -10379 |
| 82 | 12,168 | 18,085 | Republican | 12168 | 0 | 0 | 2959 | 12168 | 2959 | 9210 |
| 83 | 10,186 | 23,755 | Republican | 10186 | 0 | 0 | 6784 | 10186 | 6784 | 3401 |
| 84 | 12,503 | 18,765 | Republican | 12503 | 0 | 0 | 3131 | 12503 | 3131 | 9373 |
| 85 | 13,613 | 12,925 | Democratic | 0 | 12925 | 344 | 0 | 344 | 12925 | -12581 |
| 86 | 13,425 | 17,152 | Republican | 13425 | 0 | 0 | 1863 | 13425 | 1863 | 11561 |
| 87 | 11,780 | 15,118 | Republican | 11780 | 0 | 0 | 1669 | 11780 | 1669 | 10111 |
| 88 | 13,141 | 14,380 | Republican | 13141 | 0 | 0 | 620 | 13141 | 620 | 12521 |
| 89 | 11,610 | 15,516 | Republican | 11610 | 0 | 0 | 1953 | 11610 | 1953 | 9658 |
| 90 | 12,080 | 7,309 | Democratic | 0 | 7309 | 2385 | 0 | 2385 | 7309 | -4924 |
| 91 | 17,942 | 11,769 | Democratic | 0 | 11769 | 3086 | 0 | 3086 | 11769 | -8683 |
| 92 | 14,285 | 11,441 | Democratic | 0 | 11441 | 1422 | 0 | 1422 | 11441 | -10019 |
| 93 | 15,268 | 15,393 | Republican | 15268 | 0 | 0 | 62 | 15268 | 62 | 15206 |
| 94 | 17,408 | 12,954 | Democratic | 0 | 12954 | 2227 | 0 | 2227 | 12954 | -10727 |
| 95 | 19,804 | 9,627 | Democratic | 0 | 9627 | 5088 | 0 | 5088 | 9627 | -4539 |
| 96 | 10,950 | 14,873 | Republican | 10950 | 0 | 0 | 1962 | 10950 | 1962 | 8989 |
| 97 | 10,826 | 18,042 | Republican | 10826 | 0 | 0 | 3608 | 10826 | 3608 | 7219 |
| 98 | 10,182 | 21,855 | Republican | 10182 | 0 | 0 | 5837 | 10182 | 5837 | 4346 |
| 99 | 8,346 | 25,535 | Republican | 8346 | 0 | 0 | 8594 | 8346 | 8594 | -248 |
| TOTALS | 1,454,717 | 1,389,958 |  | 702,148 | 401,975 | 175,297 | 142,918 | 877,445 | 544,893 | 332,552 |

Table 9 - Efficiency Gap Calculation for
Act 432011 Gaddie Metric - No Incumbent Baseline

|  |  |  |  | A | B | C | D | E | F |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly District | Predicted Democratic Votes | Predicted Republican Votes | Predicted Winning Party | Lost Democratic Votes | Lost Republican Votes | Surplus Democratic Votes | Surplus Republican Votes | Wasted Democratic Votes ( $\mathrm{A}+\mathrm{C}$ ) | Wasted Republican Votes (B+D) | Net Wasted Votes (E - F) |
| 1 | 15,857 | 16,651 | Republican | 15857 | 0 | 0 | 397 | 15857 | 397 | 15461 |
| 2 | 12,983 | 15,766 | Republican | 12983 | 0 | 0 | 1391 | 12983 | 1391 | 11591 |
| 3 | 12,976 | 16,236 | Republican | 12976 | 0 | 0 | 1630 | 12976 | 1630 | 11346 |
| 4 | 13,742 | 15,791 | Republican | 13742 | 0 | 0 | 1025 | 13742 | 1025 | 12717 |
| 5 | 13,134 | 15,593 | Republican | 13134 | 0 | 0 | 1230 | 13134 | 1230 | 11904 |
| 6 | 10,779 | 15,088 | Republican | 10779 | 0 | 0 | 2155 | 10779 | 2155 | 8624 |
| 7 | 13,967 | 11,604 | Democratic | 0 | 11604 | 1181 | 0 | 1181 | 11604 | -10423 |
| 8 | 6,178 | 2,709 | Democratic | 0 | 2709 | 1735 | 0 | 1735 | 2709 | -974 |
| 9 | 10,173 | 4,184 | Democratic | 0 | 4184 | 2995 | 0 | 2995 | 4184 | -1189 |
| 10 | 24,623 | 3,547 | Democratic | 0 | 3547 | 10538 | 0 | 10538 | 3547 | 6992 |
| 11 | 20,235 | 4,927 | Democratic | 0 | 4927 | 7654 | 0 | 7654 | 4927 | 2728 |
| 12 | 18,066 | 6,856 | Democratic | 0 | 6856 | 5605 | 0 | 5605 | 6856 | -1251 |
| 13 | 13,929 | 19,774 | Republican | 13929 | 0 | 0 | 2922 | 13929 | 2922 | 11007 |
| 14 | 14,693 | 20,831 | Republican | 14693 | 0 | 0 | 3069 | 14693 | 3069 | 11624 |
| 15 | 13,497 | 16,819 | Republican | 13497 | 0 | 0 | 1661 | 13497 | 1661 | 11835 |
| 16 | 22,223 | 2,618 | Democratic | 0 | 2618 | 9803 | 0 | 9803 | 2618 | 7184 |
| 17 | 22,553 | 5,582 | Democratic | 0 | 5582 | 8486 | 0 | 8486 | 5582 | 2904 |
| 18 | 21,176 | 3,719 | Democratic | 0 | 3719 | 8728 | 0 | 8728 | 3719 | 5009 |
| 19 | 23,838 | 9,284 | Democratic | 0 | 9284 | 7277 | 0 | 7277 | 9284 | -2007 |
| 20 | 16,451 | 12,471 | Democratic | 0 | 12471 | 1990 | 0 | 1990 | 12471 | -10482 |
| 21 | 13,125 | 14,765 | Republican | 13125 | 0 | 0 | 820 | 13125 | 820 | 12305 |
| 22 | 11,364 | 22,885 | Republican | 11364 | 0 | 0 | 5761 | 11364 | 5761 | 5603 |
| 23 | 15,182 | 20,658 | Republican | 15182 | 0 | 0 | 2738 | 15182 | 2738 | 12444 |
| 24 | 14,205 | 20,015 | Republican | 14205 | 0 | 0 | 2905 | 14205 | 2905 | 11299 |
| 25 | 13,065 | 14,887 | Republican | 13065 | 0 | 0 | 911 | 13065 | 911 | 12154 |
| 26 | 12,853 | 16,338 | Republican | 12853 | 0 | 0 | 1743 | 12853 | 1743 | 11110 |
| 27 | 13,611 | 17,458 | Republican | 13611 | 0 | 0 | 1923 | 13611 | 1923 | 11688 |
| 28 | 12,609 | 15,412 | Republican | 12609 | 0 | 0 | 1401 | 12609 | 1401 | 11208 |
| 29 | 13,519 | 14,054 | Republican | 13519 | 0 | 0 | 267 | 13519 | 267 | 13251 |
| 30 | 14,267 | 16,601 | Republican | 14267 | 0 | 0 | 1167 | 14267 | 1167 | 13101 |
| 31 | 12,616 | 16,273 | Republican | 12616 | 0 | 0 | 1829 | 12616 | 1829 | 10787 |
| 32 | 10,038 | 16,566 | Republican | 10038 | 0 | 0 | 3264 | 10038 | 3264 | 6773 |
| 33 | 11,274 | 18,247 | Republican | 11274 | 0 | 0 | 3487 | 11274 | 3487 | 7788 |
| 34 | 14,239 | 17,558 | Republican | 14239 | 0 | 0 | 1660 | 14239 | 1660 | 12579 |
| 35 | 13,067 | 14,729 | Republican | 13067 | 0 | 0 | 831 | 13067 | 831 | 12236 |
| 36 | 12,227 | 14,848 | Republican | 12227 | 0 | 0 | 1310 | 12227 | 1310 | 10917 |
| 37 | 12,110 | 16,799 | Republican | 12110 | 0 | 0 | 2345 | 12110 | 2345 | 9766 |
| 38 | 12,574 | 19,218 | Republican | 12574 | 0 | 0 | 3322 | 12574 | 3322 | 9251 |
| 39 | 10,899 | 17,782 | Republican | 10899 | 0 | 0 | 3442 | 10899 | 3442 | 7457 |
| 40 | 10,514 | 14,561 | Republican | 10514 | 0 | 0 | 2024 | 10514 | 2024 | 8490 |
| 41 | 11,761 | 14,467 | Republican | 11761 | 0 | 0 | 1353 | 11761 | 1353 | 10407 |
| 42 | 13,152 | 16,036 | Republican | 13152 | 0 | 0 | 1442 | 13152 | 1442 | 11710 |
| 43 | 17,339 | 13,113 | Democratic | 0 | 13113 | 2113 | 0 | 2113 | 13113 | -10999 |
| 44 | 16,941 | 10,043 | Democratic | 0 | 10043 | 3449 | 0 | 3449 | 10043 | -6595 |
| 45 | 14,886 | 9,957 | Democratic | 0 | 9957 | 2464 | 0 | 2464 | 9957 | -7493 |
| 46 | 17,681 | 13,010 | Democratic | 0 | 13010 | 2336 | 0 | 2336 | 13010 | -10674 |


| 47 | 20,628 | 10,322 | Democratic | 0 | 10322 | 5153 | 0 | 5153 | 10322 | -5169 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 48 | 23,290 | 8,861 | Democratic | 0 | 8861 | 7215 | 0 | 7215 | 8861 | -1646 |
| 49 | 13,071 | 12,859 | Democratic | 0 | 12859 | 106 | 0 | 106 | 12859 | -12752 |
| 50 | 11,887 | 12,908 | Republican | 11887 | 0 | 0 | 511 | 11887 | 511 | 11376 |
| 51 | 14,637 | 12,584 | Democratic | 0 | 12584 | 1026 | 0 | 1026 | 12584 | -11558 |
| 52 | 11,034 | 15,918 | Republican | 11034 | 0 | 0 | 2442 | 11034 | 2442 | 8592 |
| 53 | 9,930 | 16,099 | Republican | 9930 | 0 | 0 | 3084 | 9930 | 3084 | 6846 |
| 54 | 15,372 | 12,690 | Democratic | 0 | 12690 | 1341 | 0 | 1341 | 12690 | -11348 |
| 55 | 13,302 | 16,297 | Republican | 13302 | 0 | 0 | 1498 | 13302 | 1498 | 11804 |
| 56 | 12,809 | 18,326 | Republican | 12809 | 0 | 0 | 2759 | 12809 | 2759 | 10050 |
| 57 | 14,436 | 11,575 | Democratic | 0 | 11575 | 1431 | 0 | 1431 | 11575 | -10145 |
| 58 | 9,211 | 22,056 | Republican | 9211 | 0 | 0 | 6422 | 9211 | 6422 | 2789 |
| 59 | 9,669 | 20,843 | Republican | 9669 | 0 | 0 | 5587 | 9669 | 5587 | 4083 |
| 60 | 10,307 | 23,508 | Republican | 10307 | 0 | 0 | 6601 | 10307 | 6601 | 3706 |
| 61 | 12,661 | 16,935 | Republican | 12661 | 0 | 0 | 2137 | 12661 | 2137 | 10524 |
| 62 | 13,959 | 18,175 | Republican | 13959 | 0 | 0 | 2108 | 13959 | 2108 | 11851 |
| 63 | 11,973 | 17,692 | Republican | 11973 | 0 | 0 | 2860 | 11973 | 2860 | 9113 |
| 64 | 15,452 | 11,524 | Democratic | 0 | 11524 | 1964 | 0 | 1964 | 11524 | -9560 |
| 65 | 14,760 | 8,274 | Democratic | 0 | 8274 | 3243 | 0 | 3243 | 8274 | -5031 |
| 66 | 14,776 | 6,861 | Democratic | 0 | 6861 | 3957 | 0 | 3957 | 6861 | -2904 |
| 67 | 13,748 | 14,698 | Republican | 13748 | 0 | 0 | 475 | 13748 | 475 | 13273 |
| 68 | 13,508 | 13,177 | Democratic | 0 | 13177 | 165 | 0 | 165 | 13177 | -13011 |
| 69 | 11,657 | 13,773 | Republican | 11657 | 0 | 0 | 1058 | 11657 | 1058 | 10599 |
| 70 | 13,105 | 13,493 | Republican | 13105 | 0 | 0 | 194 | 13105 | 194 | 12911 |
| 71 | 17,189 | 11,807 | Democratic | 0 | 11807 | 2691 | 0 | 2691 | 11807 | -9116 |
| 72 | 13,674 | 14,514 | Republican | 13674 | 0 | 0 | 420 | 13674 | 420 | 13254 |
| 73 | 16,837 | 11,300 | Democratic | 0 | 11300 | 2769 | 0 | 2769 | 11300 | -8531 |
| 74 | 17,628 | 13,239 | Democratic | 0 | 13239 | 2195 | 0 | 2195 | 13239 | -11044 |
| 75 | 13,590 | 14,829 | Republican | 13590 | 0 | 0 | 620 | 13590 | 620 | 12970 |
| 76 | 32,275 | 5,469 | Democratic | 0 | 5469 | 13403 | 0 | 13403 | 5469 | 7934 |
| 77 | 26,627 | 6,339 | Democratic | 0 | 6339 | 10144 | 0 | 10144 | 6339 | 3804 |
| 78 | 23,528 | 10,492 | Democratic | 0 | 10492 | 6518 | 0 | 6518 | 10492 | -3974 |
| 79 | 20,211 | 14,516 | Democratic | 0 | 14516 | 2848 | 0 | 2848 | 14516 | -11668 |
| 80 | 20,251 | 12,704 | Democratic | 0 | 12704 | 3773 | 0 | 3773 | 12704 | -8931 |
| 81 | 15,887 | 12,770 | Democratic | 0 | 12770 | 1559 | 0 | 1559 | 12770 | -11211 |
| 82 | 12,985 | 17,269 | Republican | 12985 | 0 | 0 | 2142 | 12985 | 2142 | 10843 |
| 83 | 10,756 | 23,185 | Republican | 10756 | 0 | 0 | 6215 | 10756 | 6215 | 4541 |
| 84 | 13,414 | 17,854 | Republican | 13414 | 0 | 0 | 2220 | 13414 | 2220 | 11194 |
| 85 | 13,703 | 12,843 | Democratic | 0 | 12843 | 430 | 0 | 430 | 12843 | -12413 |
| 86 | 15,780 | 14,789 | Democratic | 0 | 14789 | 495 | 0 | 495 | 14789 | -14294 |
| 87 | 12,413 | 14,420 | Republican | 12413 | 0 | 0 | 1004 | 12413 | 1004 | 11409 |
| 88 | 12,882 | 14,638 | Republican | 12882 | 0 | 0 | 878 | 12882 | 878 | 12004 |
| 89 | 12,009 | 15,118 | Republican | 12009 | 0 | 0 | 1554 | 12009 | 1554 | 10455 |
| 90 | 11,556 | 7,833 | Democratic | 0 | 7833 | 1861 | 0 | 1861 | 7833 | -5972 |
| 91 | 18,044 | 11,816 | Democratic | 0 | 11816 | 3114 | 0 | 3114 | 11816 | -8701 |
| 92 | 14,313 | 11,383 | Democratic | 0 | 11383 | 1465 | 0 | 1465 | 11383 | -9919 |
| 93 | 15,014 | 15,690 | Republican | 15014 | 0 | 0 | 338 | 15014 | 338 | 14676 |
| 94 | 14,601 | 15,761 | Republican | 14601 | 0 | 0 | 580 | 14601 | 580 | 14022 |
| 95 | 18,730 | 10,701 | Democratic | 0 | 10701 | 4014 | 0 | 4014 | 10701 | -6687 |
| 96 | 13,841 | 11,982 | Democratic | 0 | 11982 | 930 | 0 | 930 | 11982 | -11052 |
| 97 | 10,706 | 18,158 | Republican | 10706 | 0 | 0 | 3726 | 10706 | 3726 | 6979 |
| 98 | 10,566 | 21,472 | Republican | 10566 | 0 | 0 | 5453 | 10566 | 5453 | 5113 |
| 99 | 8,517 | 25,349 | Republican | 8517 | 0 | 0 | 8416 | 8517 | 8416 | 102 |
| TOTALS | 1,448,901 | 1,394,018 |  | 726,238 | 402,334 | 160,165 | 132,723 | 886,403 | 535,057 | 351,346 |

## D. Conclusions

In this report, I have outlined a method that generates accurate estimates of underlying partisanship using the 2012 presidential election vote, demographics, incumbency, and geographic features to explain patterns of voting in Assembly elections. This method is accurate, as demonstrated by its ability to forecast vote totals at both the individual ward and district levels, and I demonstrate that it generates valid out of sample estimates. It produces results that are very similar to those derived by the expert witness retained by the state legislature during its development of the redistricting map implemented in Act 43.

The results demonstrate that Act 43 was an egregious gerrymander, packing Democratic voters into a small number of districts and distributing Republican voters efficiently in a large number of districts in which they constituted safe majorities. As I demonstrated with the treatment of the city of Sheboygan in Act 43, areas of Democratic strength large enough to constitute majorities in single districts were unnecessarily split and then combined with larger Republican populations to create additional Republican districts and eliminate Democratic districts. The city, which had been in a single Democratic Assembly district since 1992, was split into two Republican districts. This packing and cracking was so successful that Republicans won $61 \%$ of Assembly seats in 2012, while obtaining only $46.5 \%$ of the statewide presidential vote.

The scope of the gerrymander is demonstrated by the efficiency gap calculation for Act 43: $11.69 \%$. Based on the baseline partisanship estimates produced by Dr. Ronald Keith Gaddie during the drawing of the Act 43 plan, this was the intended outcome: using Gaddie's baseline estimates, Act 43 had an expected efficiency gap of $12.36 \%$.

However, I drew a demonstration districting plan that was equivalent to Act 43 on population deviation, municipal splits, and compliance with the Voting Rights Act, and better on compactness, with a dramatically lower efficiency gap score of $2.20 \%$. This proves that Act 43 's extreme partisan effects were not required by these constitutional or statutory mandates.

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## I. Data Issues

The largest errors in the Legislative Technology Services Bureau (LTSB) data occurred because the two data sets used to create this data do not precisely overlap. In GIS argot, the two sets of data are not reported in the same geography. The LTSB files contained data at the individual ward level, while the official election data is aggregated by reporting unit. Wisconsin elections are administered at the ward level, but are often tabulated and released in reporting units consisting of multiple wards. ${ }^{1}$ Of Wisconsin's roughly 6,530 populated wards, only about a third report election results at the individual ward level; the rest report results by combining wards into reporting units. As one example, the city of Manitowoc (2010 population 33,736) has 25 wards, but reports election results in 10 reporting units of between 2 and 6 wards each. ${ }^{2}$

In order to generate data at the ward level, my understanding is that the LTSB disaggregated reporting unit results to individual wards based on the fraction of Voting Age Population in each ward comprising the reporting unit. In the process a number of anomalies crept into the data. The LTSB file for 2012 contains wards where the number of votes cast exceeds the voting age population; wards with large voting age populations and an unusually low number of votes, often zero, recorded; wards, municipalities, and districts with vote totals that differ substantially from what the Government Accountability Board (GAB) reports; votes allocated to the wrong district; incorrectly numbered and duplicated wards; and wards in uncontested Assembly districts with votes recorded for both political parties.

[^20]In most cases, correcting the errors in the LTSB data involved manually changing the incorrect ward totals to reflect GAB results. When the GAB data were combined into reporting units, I allocated votes to each ward in the unit based on the ward's share of the voting eligible population, removing noncitizen and prison populations. ${ }^{3}$ This process generated more accurate ward level data, and is a standard technique when allocating votes into different geographic levels (McDonald 2014; Pavia and López-Quílez 2013). At times, however, the LTSB and GAB data could not be reconciled, because of wards that appeared in one file but not in the other, or discrepancies in ward geography. The votes I was not able to allocate constituted only $0.21 \%$ of the total votes cast in the 2012 Assembly election, and have no effect on any subsequent analysis or my conclusions.

The following table shows some of the problems with the data recorded by the LTSB. It displays the errors in the LTSB 2012 presidential vote totals for the city of Mequon. The GAB Reports columns show the vote totals for each of the city's reporting units taken from the 2014 Wisconsin Blue Book, which I take to be authoritative. ${ }^{4}$ The LTSB Data columns show the results of combining the individual ward data in the LTSB ward file into the GAB reporting units. The Difference columns show the errors in the LTSB data. While the vote totals for the municipality are the same in both data sets, every ward total is different.

[^21]Differences Between GAB Reports and LTSB Data 2012 Presidential Election Results for Mequon, WI (Ozaukee County)

| GAB Reports |  |  |  | LTSB Data |  |  | Difference |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Reporting Unit (wards) | Obama Votes | Romney Votes | Total <br> Votes | Obama Votes | Romney Votes | Total Votes | Obama <br> Votes | Romney Votes | Total Votes |
| 1 | 534 | 890 | 1424 | 849 | 1,522 | 2,371 | 315 | 632 | 947 |
| 2 | 120 | 391 | 511 | 240 | 633 | 873 | 120 | 242 | 362 |
| 3,4 | 637 | 1,249 | 1886 | 415 | 833 | 1,248 | (222) | (416) | (638) |
| 5, 7B | 205 | 603 | 808 | 155 | 311 | 466 | (50) | (292) | (342) |
| 6, 7A | 392 | 909 | 1301 | 292 | 589 | 881 | (100) | (320) | (420) |
| 8,9,10 | 737 | 1,245 | 1982 | 477 | 956 | 1,433 | (260) | (289) | (549) |
| 11, 12 | 635 | 1,126 | 1761 | 527 | 1,057 | 1,584 | (108) | (69) | (177) |
| 13, 14 | 353 | 770 | 1123 | 253 | 506 | 759 | (100) | (264) | (364) |
| 15 | 380 | 494 | 874 | 579 | 896 | 1,475 | 199 | 402 | 601 |
| 16 | 221 | 491 | 712 | 357 | 766 | 1,123 | 136 | 275 | 411 |
| 17 | 336 | 459 | 795 | 517 | 824 | 1,341 | 181 | 365 | 546 |
| 18 | 204 | 368 | 572 | 322 | 607 | 929 | 118 | 239 | 357 |
| 19,20,21 | 639 | 1,331 | 1970 | 410 | 826 | 1,236 | (229) | (505) | (734) |
| Totals | 5,393 | 10,326 | 15,719 | 5,393 | 10,326 | 15,719 | 0 | 0 | 0 |

Correcting these totals required manually changing the single-ward vote counts to match the GAB data, and allocating votes in reporting units to the individual wards based on the votingeligible population in each ward in the unit (in the following table, wards in a reporting unit are framed together):

Allocation of Reporting Unit Data to Ward Data
City of Mequon, 2012 Presidential Vote

| GAB Data |  |  |  |  | Data Used in Voting Model |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ward | Obama Votes | Romney Votes | Ward Voting Eligible Population | Ward Share of Reporting Unit VEP | Obama Votes | Romney <br> Votes | Total Votes |
| 1 | 534 | 890 | - | - | 534 | 890 | 1,424 |
| 2 | 120 | 391 | - | - | 120 | 391 | 511 |
| 3 | 637 | 1249 | 1063 | 53\% | 336 | 658 | 994 |
| 4 |  |  | 954 | 47\% | 301 | 591 | 892 |
| 5 | 205 | 603 | 501 | 67\% | 137 | 402 | 539 |
| 7B |  |  | 250 | 33\% | 68 | 201 | 269 |
| 6 | 392 | 909 | 1240 | 87\% | 343 | 794 | 1,137 |
| 7A |  |  | 179 | 13\% | 49 | 115 | 164 |
| 8 | 737 | 1245 | 599 | 26\% | 192 | 324 | 516 |
| 9 |  |  | 457 | 20\% | 146 | 247 | 393 |
| 10 |  |  | 1247 | 54\% | 399 | 674 | 1,073 |
| 11 | 635 | 1126 | 1530 | 60\% | 380 | 673 | 1,053 |
| 12 |  |  | 1029 | 40\% | 255 | 453 | 708 |
| 13 | 353 | 770 | 761 | 63\% | 221 | 482 | 703 |
| 14 |  |  | 455 | 37\% | 132 | 288 | 420 |
| 15 | 380 | 494 | - | - | 380 | 494 | 874 |
| 16 | 221 | 491 | - | - | 221 | 491 | 712 |
| 17 | 336 | 459 | - | - | 336 | 459 | 795 |
| 18 | 204 | 368 | - | - | 204 | 368 | 572 |
| 19 |  |  | 908 | 46\% | 291 | 606 | 897 |
| 20 | 639 | 1331 | 776 | 39\% | 249 | 518 | 767 |
| 21 |  |  | 310 | 16\% | 99 | 207 | 306 |
| Totals | 5,393 | 10,326 |  |  | 5,393 | 10,326 | 15,719 |

I repeated this process for every instance of inaccurate vote totals in the LTSB, using
GAB data as the reference.

## II. Full Regression Results

Republican vote totals (bold variables have $\mathrm{p}<.05$ )

Independent Variable: Assembly Republican Votes
$\left.\begin{array}{|r|c|c|c|c|}\hline \begin{array}{l}\text { Dependent } \\ \text { Variable }\end{array} & \text { Coefficient } & \begin{array}{c}\text { Robust } \\ \text { Std. Error }\end{array} & \text { t-statistic } & \text { P-value } \\ \hline \begin{array}{r}\text { Total Voting } \\ \text { Eligible }\end{array} & 0.01 & & & \\ \text { Population }\end{array}\right)$

| Columbia | 15.01 | 10.08 | 1.49 | 0.141 |
| :---: | :---: | :---: | :---: | :---: |
| Crawford | 28.20 | 7.24 | 3.90 | 0 |
| Dane | 1.55 | 8.53 | 0.18 | 0.857 |
| Dodge | 8.54 | 7.88 | 1.08 | 0.282 |
| Door | 16.98 | 7.23 | 2.35 | 0.022 |
| Douglas | -3.14 | 7.65 | -0.41 | 0.682 |
| EauClaire | 0.47 | 7.83 | 0.06 | 0.953 |
| Florence | -7.34 | 7.52 | -0.98 | 0.332 |
| FondduLac | 4.74 | 8.07 | 0.59 | 0.559 |
| Forest | -1.91 | 7.39 | -0.26 | 0.796 |
| Grant | 24.64 | 7.23 | 3.41 | 0.001 |
| Green | 14.41 | 9.95 | 1.45 | 0.152 |
| GreenLake | 11.96 | 7.36 | 1.62 | 0.109 |
| Iowa | 15.04 | 8.08 | 1.86 | 0.067 |
| Iron | 20.54 | 7.68 | 2.67 | 0.009 |
| Jackson | 5.74 | 7.53 | 0.76 | 0.449 |
| Jefferson | 2.37 | 8.41 | 0.28 | 0.779 |
| Juneau | -4.31 | 7.29 | -0.59 | 0.556 |
| Kenosha | 3.73 | 7.99 | 0.47 | 0.642 |
| Kewaunee | -14.13 | 7.24 | -1.95 | 0.055 |
| LaCrosse | -26.58 | 8.43 | -3.15 | 0.002 |
| Lafayette | 18.18 | 7.29 | 2.49 | 0.015 |
| Langlade | 4.35 | 8.30 | 0.52 | 0.602 |
| Lincoln | -0.38 | 7.53 | -0.05 | 0.96 |
| Manitowoc | 19.35 | 9.36 | 2.07 | 0.042 |
| Marathon | 2.01 | 8.56 | 0.24 | 0.815 |
| Marinette | 19.89 | 8.04 | 2.48 | 0.016 |
| Marquette | 6.91 | 7.26 | 0.95 | 0.344 |
| Menominee | -3.08 | 7.32 | -0.42 | 0.675 |
| Milwaukee | 1.96 | 11.98 | 0.16 | 0.871 |
| Monroe | 19.47 | 7.72 | 2.52 | 0.014 |
| Oconto | 3.21 | 7.95 | 0.40 | 0.687 |
| Oneida | 12.01 | 7.95 | 1.51 | 0.136 |
| Outagamie | 1.90 | 8.02 | 0.24 | 0.814 |
| Ozaukee | 13.71 | 8.82 | 1.55 | 0.125 |
| Pepin | -9.83 | 7.27 | -1.35 | 0.181 |
| Pierce | -9.31 | 7.18 | -1.30 | 0.199 |
| Polk | -3.47 | 7.24 | -0.48 | 0.633 |
| Portage | -20.74 | 7.71 | -2.69 | 0.009 |
| Price | 5.25 | 7.75 | 0.68 | 0.501 |
| Racine | -6.90 | 8.23 | -0.84 | 0.404 |
| Richland | 16.24 | 8.55 | 1.90 | 0.062 |
| Rock | 9.24 | 8.32 | 1.11 | 0.27 |


| Rusk | 3.71 | 7.37 | 0.50 | 0.616 |
| :---: | :---: | :---: | :---: | :---: |
| SaintCroix | 13.80 | 9.31 | 1.48 | 0.143 |
| Sauk | $\mathbf{1 6 . 6 8}$ | $\mathbf{8 . 2 7}$ | $\mathbf{2 . 0 2}$ | $\mathbf{0 . 0 4 8}$ |
| Sawyer | -0.90 | 7.40 | -0.12 | 0.903 |
| Shawano | 2.70 | 7.86 | 0.34 | 0.733 |
| Sheboygan | -6.50 | 15.54 | -0.42 | 0.677 |
| Taylor | 9.96 | 7.30 | 1.37 | 0.176 |
| Trempealeau | 1.29 | 7.21 | 0.18 | 0.859 |
| Vernon | $\mathbf{3 1 . 5 4}$ | $\mathbf{7 . 2 9}$ | $\mathbf{4 . 3 3}$ | $\mathbf{0}$ |
| Vilas | 3.61 | 7.64 | 0.47 | 0.638 |
| Walworth | -2.00 | 8.17 | -0.24 | 0.807 |
| Washburn | -10.80 | 7.31 | -1.48 | 0.144 |
| Washington | 14.16 | 12.70 | 1.12 | 0.269 |
| Waukesha | 1.18 | 7.93 | 0.15 | 0.882 |
| Waupaca | -8.08 | 7.26 | -1.11 | 0.27 |
| Waushara | -3.47 | 7.30 | -0.48 | 0.636 |
| Winnebago | 30.00 | 17.09 | 1.76 | 0.084 |
| Wood | -7.60 | 8.96 | -0.85 | 0.399 |
| Constant | -0.92 | 7.52 | -0.12 | 0.903 |


| N | 5282.00 |
| :---: | :---: |
| R-squared | 0.9903 |
| Root MSE | 15.823 |

Democratic vote totals

Independent Variable: Assembly Democratic Votes

| Dependent <br> Variable | Coefficient | Robust <br> Std. Error | t-statistic | P-value |
| :---: | :---: | :---: | :---: | :---: |
| Total Voting Eligible Population | -0.01 | 0.01 | -0.65 | 0.52 |
| Black Voting <br> Eligible <br> Population | -0.02 | 0.04 | -0.49 | 0.63 |
| Hispanic <br> Voting Eligible <br> Population | -0.15 | 0.05 | -3.01 | 0.00 |
| Democratic <br> Presidential <br> Votes | 0.93 | 0.03 | 33.33 | 0.00 |
| Republican <br> Presidential <br> Votes | 0.01 | 0.01 | 0.98 | 0.33 |
| Democratic Assembly Incumbent | 0.03 | 0.01 | 3.85 | 0.00 |
| Republican Assembly Incumbent | -0.01 | 0.01 | -2.77 | 0.01 |
| Adams | -14.45 | 6.73 | -2.15 | 0.04 |
| Ashland | -4.78 | 5.58 | -0.86 | 0.40 |
| Barron | 14.57 | 4.04 | 3.60 | 0.00 |
| Bayfield | -2.82 | 5.58 | -0.50 | 0.62 |
| Brown | -21.57 | 7.80 | -2.77 | 0.01 |
| Buffalo | 5.10 | 4.86 | 1.05 | 0.30 |
| Burnett | -3.84 | 4.69 | -0.82 | 0.42 |
| Calumet | -26.32 | 5.81 | -4.53 | 0.00 |
| Chippewa | 0.98 | 9.53 | 0.10 | 0.92 |
| Clark | -6.83 | 4.80 | -1.42 | 0.16 |
| Columbia | -19.51 | 8.15 | -2.39 | 0.02 |
| Crawford | -32.57 | 4.33 | -7.51 | 0.00 |
| Dane | -9.39 | 7.20 | -1.31 | 0.20 |
| Dodge | -8.49 | 5.27 | -1.61 | 0.11 |
| Door | -11.92 | 4.51 | -2.64 | 0.01 |
| Douglas | -7.18 | 5.40 | -1.33 | 0.19 |
| EauClaire | 1.05 | 7.22 | 0.14 | 0.89 |
| Florence | -13.53 | 5.33 | -2.54 | 0.01 |
| FondduLac | -25.18 | 4.92 | -5.12 | 0.00 |
| Forest | -10.83 | 6.06 | -1.79 | 0.08 |


| Grant | -23.14 | 4.26 | -5.43 | 0.00 |
| :---: | :---: | :---: | :---: | :---: |
| Green | -15.68 | 6.63 | -2.36 | 0.02 |
| GreenLake | -17.01 | 4.65 | -3.66 | 0.00 |
| Iowa | -19.48 | 4.91 | -3.96 | 0.00 |
| Iron | -30.91 | 5.54 | -5.58 | 0.00 |
| Jackson | -12.37 | 6.44 | -1.92 | 0.06 |
| Jefferson | -17.18 | 7.09 | -2.42 | 0.02 |
| Juneau | -5.78 | 4.55 | -1.27 | 0.21 |
| Kenosha | 1.78 | 5.33 | 0.33 | 0.74 |
| Kewaunee | 17.69 | 4.41 | 4.01 | 0.00 |
| LaCrosse | 25.17 | 6.69 | 3.76 | 0.00 |
| Lafayette | -22.66 | 4.58 | -4.95 | 0.00 |
| Langlade | -22.20 | 6.05 | -3.67 | 0.00 |
| Lincoln | -13.42 | 5.15 | -2.61 | 0.01 |
| Manitowoc | -15.90 | 5.49 | -2.90 | 0.01 |
| Marathon | -5.64 | 6.20 | -0.91 | 0.37 |
| Marinette | -26.28 | 4.22 | -6.23 | 0.00 |
| Marquette | -15.87 | 4.48 | -3.54 | 0.00 |
| Menominee | -61.44 | 4.41 | -13.95 | 0.00 |
| Milwaukee | -29.20 | 6.47 | -4.51 | 0.00 |
| Monroe | -26.83 | 5.44 | -4.93 | 0.00 |
| Oconto | -12.99 | 4.42 | -2.94 | 0.00 |
| Oneida | -35.94 | 5.19 | -6.92 | 0.00 |
| Outagamie | -14.60 | 6.94 | -2.10 | 0.04 |
| Ozaukee | -17.19 | 5.83 | -2.95 | 0.00 |
| Pepin | 6.62 | 4.52 | 1.46 | 0.15 |
| Pierce | 12.49 | 4.00 | 3.12 | 0.00 |
| Polk | 5.81 | 4.32 | 1.35 | 0.18 |
| Portage | -0.04 | 5.13 | -0.01 | 0.99 |
| Price | -14.62 | 5.64 | -2.59 | 0.01 |
| Racine | 4.42 | 5.29 | 0.83 | 0.41 |
| Richland | -26.22 | 5.30 | -4.95 | 0.00 |
| Rock | -4.48 | 8.87 | -0.50 | 0.62 |
| Rusk | -8.01 | 4.90 | -1.64 | 0.11 |
| SaintCroix | -6.89 | 6.67 | -1.03 | 0.31 |
| Sauk | -19.42 | 6.51 | -2.98 | 0.00 |
| Sawyer | -6.06 | 4.64 | -1.30 | 0.20 |
| Shawano | -14.93 | 4.58 | -3.26 | 0.00 |
| Sheboygan | 15.96 | 17.17 | 0.93 | 0.36 |
| Taylor | -6.81 | 4.56 | -1.49 | 0.14 |
| Trempealeau | -3.89 | 4.29 | -0.91 | 0.37 |
| Vernon | -32.42 | 4.52 | -7.18 | 0.00 |
| Vilas | -27.14 | 5.48 | -4.95 | 0.00 |


| Walworth | 0.34 | 5.26 | 0.07 | 0.95 |
| :--- | :---: | :---: | :---: | :---: |
| Washburn | 6.43 | 4.74 | 1.36 | 0.18 |
| Washington | $\mathbf{- 1 9 . 2 3}$ | $\mathbf{9 . 7 5}$ | $\mathbf{- 1 . 9 7}$ | $\mathbf{0 . 0 5}$ |
| Waukesha | $\mathbf{- 1 7 . 6 3}$ | $\mathbf{5 . 5 5}$ | $\mathbf{- 3 . 1 8}$ | $\mathbf{0 . 0 0}$ |
| Waupaca | $\mathbf{- 1 0 . 4 8}$ | $\mathbf{4 . 3 7}$ | $\mathbf{- 2 . 4 0}$ | $\mathbf{0 . 0 2}$ |
| Waushara | 0.21 | 4.64 | 0.04 | 0.97 |
| Winnebago | $\mathbf{- 3 2 . 1 2}$ | $\mathbf{1 5 . 9 4}$ | $\mathbf{- 2 . 0 2}$ | $\mathbf{0 . 0 5}$ |
| Wood | 8.14 | 6.01 | 1.35 | 0.18 |
| Constant | 9.80 | 5.39 | 1.82 | 0.07 |


| N | 5282.00 |
| :---: | :---: |
| R-squared | 0.9843 |
| Root MSE | 17.675 |

III. Plan characteristics
A. Population deviation

| Assembly <br> District | Population | Deviation <br> from <br> Ideal | $\%$ <br> Deviation |
| :---: | ---: | ---: | ---: |
| 1 | 57,487 | 43 | $0.07 \%$ |
| 2 | 57,590 | 146 | $0.25 \%$ |
| 3 | 57,686 | 242 | $0.42 \%$ |
| 4 | 57,406 | -38 | $-0.07 \%$ |
| 5 | 57,633 | 189 | $0.33 \%$ |
| 6 | 57,480 | 36 | $0.06 \%$ |
| 7 | 57,208 | -236 | $-0.41 \%$ |
| 8 | 57,196 | -248 | $-0.43 \%$ |
| 9 | 57,420 | -24 | $-0.04 \%$ |
| 10 | 57,195 | -249 | $-0.43 \%$ |
| 11 | 57,455 | 11 | $0.02 \%$ |
| 12 | 57,420 | -24 | $-0.04 \%$ |
| 13 | 57,248 | -196 | $-0.34 \%$ |
| 14 | 57,333 | -111 | $-0.19 \%$ |
| 15 | 57,514 | 70 | $0.12 \%$ |
| 16 | 57,282 | -162 | $-0.28 \%$ |
| 17 | 57,437 | -7 | $-0.01 \%$ |
| 34 | 57,241 | -203 | $-0.35 \%$ |
| 36 | 57,377 | -67 | $-0.12 \%$ |


| 37 | 57,671 | 227 | 0.40\% |
| :---: | :---: | :---: | :---: |
| 38 | 57,572 | 128 | 0.22\% |
| 39 | 57,457 | 13 | 0.02\% |
| 40 | 57,495 | 51 | 0.09\% |
| 41 | 57,671 | 227 | 0.40\% |
| 42 | 57,559 | 115 | 0.20\% |
| 43 | 57,444 | 0 | 0.00\% |
| 44 | 57,434 | -10 | -0.02\% |
| 45 | 57,242 | -202 | -0.35\% |
| 46 | 57,463 | 19 | 0.03\% |
| 47 | 57,494 | 50 | 0.09\% |
| 48 | 57,568 | 124 | 0.22\% |
| 49 | 57,389 | -55 | -0.10\% |
| 50 | 57,465 | 21 | 0.04\% |
| 51 | 57,247 | -197 | -0.34\% |
| 52 | 57,384 | -60 | -0.10\% |
| 53 | 57,444 | 0 | 0.00\% |
| 54 | 57,443 | -1 | 0.00\% |
| 55 | 57,446 | 2 | 0.00\% |
| 56 | 57,342 | -102 | -0.18\% |
| 57 | 57,404 | -40 | -0.07\% |
| 58 | 57,436 | -8 | -0.01\% |
| 59 | 57,554 | 110 | 0.19\% |
| 60 | 57,547 | 103 | 0.18\% |
| 61 | 57,605 | 161 | 0.28\% |
| 62 | 57,632 | 188 | 0.33\% |
| 63 | 57,299 | -145 | -0.25\% |
| 64 | 57,266 | -178 | -0.31\% |
| 65 | 57,601 | 157 | 0.27\% |
| 66 | 57,459 | 15 | 0.03\% |
| 67 | 57,378 | -66 | -0.11\% |
| 68 | 57,254 | -190 | -0.33\% |
| 69 | 57,424 | -20 | -0.03\% |
| 70 | 57,415 | -29 | -0.05\% |
| 71 | 57,228 | -216 | -0.38\% |
| 72 | 57,654 | 210 | 0.37\% |
| 73 | 57,491 | 47 | 0.08\% |
| 74 | 57,320 | -124 | -0.22\% |
| 75 | 57,255 | -189 | -0.33\% |
| 76 | 57,586 | 142 | 0.25\% |
| 77 | 57,398 | -46 | -0.08\% |
| 78 | 57,579 | 135 | 0.24\% |
| 79 | 57,341 | -103 | -0.18\% |


| 80 | 57,385 | -59 | $-0.10 \%$ |
| ---: | ---: | ---: | ---: |
| 81 | 57,266 | -178 | $-0.31 \%$ |
| 82 | 57,641 | 197 | $0.34 \%$ |
| 83 | 57,612 | 168 | $0.29 \%$ |
| 84 | 57,375 | -69 | $-0.12 \%$ |
| 85 | 57,529 | 85 | $0.15 \%$ |
| 86 | 57,477 | 33 | $0.06 \%$ |
| 87 | 57,661 | 217 | $0.38 \%$ |
| 88 | 57,533 | 89 | $0.15 \%$ |
| 89 | 57,490 | 46 | $0.08 \%$ |
| 90 | 57,617 | 173 | $0.30 \%$ |
| 91 | 57,374 | -70 | $-0.12 \%$ |
| 92 | 57,421 | -23 | $-0.04 \%$ |
| 93 | 57,280 | -164 | $-0.29 \%$ |
| 94 | 57,509 | 65 | $0.11 \%$ |
| 95 | 57,496 | 52 | $0.09 \%$ |
| 96 | 57,406 | -38 | $-0.07 \%$ |
| 97 | 57,487 | 43 | $0.07 \%$ |
| 98 | 57,485 | 41 | $0.07 \%$ |
| 99 | 57,657 | 213 | $0.37 \%$ |

B. Compactness (Reock or smallest circle measure)

| Assembly <br> District | Smallest <br> Circle <br> Measure |
| :---: | :---: |
| 1 | 0.44 |
| 2 | 0.46 |
| 3 | 0.42 |
| 4 | 0.55 |
| 5 | 0.39 |
| 6 | 0.35 |
| 7 | 0.52 |
| 8 | 0.66 |
| 9 | 0.39 |
| 10 | 0.45 |
| 11 | 0.39 |
| 12 | 0.36 |
| 13 | 0.28 |
| 14 | 0.44 |
| 15 | 0.49 |
| 16 | 0.52 |
| 17 | 0.52 |
|  |  |


| 18 | 0.30 |
| :--- | :--- |
| 19 | 0.30 |
| 20 | 0.44 |
| 21 | 0.40 |
| 22 | 0.34 |
| 23 | 0.42 |
| 24 | 0.42 |
| 25 | 0.57 |
| 26 | 0.49 |
| 27 | 0.53 |
| 28 | 0.31 |
| 29 | 0.49 |
| 30 | 0.50 |
| 31 | 0.60 |
| 32 | 0.45 |
| 33 | 0.30 |
| 34 | 0.42 |
| 35 | 0.49 |
| 36 | 0.43 |
| 37 | 0.34 |
| 38 | 0.24 |
| 39 | 0.30 |
| 40 | 0.51 |
| 41 | 0.39 |
| 42 | 0.33 |
| 43 | 0.29 |
| 44 | 0.43 |
| 45 | 0.37 |
| 46 | 0.35 |
| 47 | 0.26 |
| 48 | 0.43 |
| 49 | 0.35 |
| 50 | 0.44 |
| 51 | 0.53 |
| 52 | 0.56 |
| 53 | 0.27 |
| 54 | 0.28 |
| 59 | 0.37 |
| 55 | 0.37 |
| 56 | 0.57 |
| 57 | 0.26 |
| 59 | 0.40 |
|  | 0.37 |


| 61 | 0.39 |
| :--- | :--- |
| 62 | 0.25 |
| 63 | 0.43 |
| 64 | 0.27 |
| 65 | 0.32 |
| 66 | 0.32 |
| 67 | 0.56 |
| 68 | 0.52 |
| 69 | 0.31 |
| 70 | 0.28 |
| 71 | 0.34 |
| 72 | 0.35 |
| 73 | 0.28 |
| 74 | 0.37 |
| 75 | 0.36 |
| 76 | 0.23 |
| 77 | 0.39 |
| 78 | 0.51 |
| 79 | 0.59 |
| 80 | 0.33 |
| 81 | 0.55 |
| 82 | 0.37 |
| 83 | 0.26 |
| 84 | 0.28 |
| 85 | 0.58 |
| 86 | 0.36 |
| 87 | 0.35 |
| 88 | 0.35 |
| 89 | 0.56 |
| 90 | 0.52 |
| 91 | 0.49 |
| 92 | 0.49 |
| 93 | 0.42 |
| 94 | 0.44 |
| 95 | 0.42 |
| 96 | 0.39 |
| 97 | 0.32 |
| 98 | 0.41 |
| 99 | 0.30 |
|  |  |
|  |  |

IN THE UNITED STATES DISTRICT COURT FOR THE WESTERN DISTRICT OF WISCONSIN

WILLIAM WHITFORD, et al.,
Plaintiffs,
v.

Case No. 15-CV-421-bbc
GERALD NICHOL, et al.,
Defendants.

## DECLARATION OF SEAN P. TRENDE

Sean Patrick Trende, under penalty of perjury, makes the following declaration:

1. I am over 18 years of age and am competent to testify regarding the matters discussed in this declaration.
2. I have been retained in this matter to provide expert testimony. I am compensated at a rate of $\$ 300$ per hour, excluding travel time.
3. My curriculum vitae is attached to this declaration as Exhibit 1.
4. A list of materials upon which I relied in the preparation of this declaration are attached as Exhibit 2.

## INTRODUCTION

5. Plaintiffs in this case attempt to solve the decades-old problem of identifying partisan gerrymanders that are severe enough to violate the federal constitution by introducing a novel measure of partisan gerrymandering, based upon the concept of "wasted votes." The basis for this theory is that a party gerrymanders when members of the opposing party are "packed" into single districts. This allows the gerrymandering party to spread their remaining members over a large number of districts, creating just enough partisan density to win. Because members of the opposing party are packed into districts far in excess of what is needed to win those districts, this should manifest in the opposing party having a disproportionate number of "wasted votes," that is, votes in excess of what are needed to win in given districts.
6. I have a tremendous amount of respect for Dr. Jackman's work (I'm not personally familiar with Dr. Mayer), as well as Dr. McGhee, upon whose work the reports here are based. Nevertheless, there are multiple problems with utilizing this approach to identify unconstitutional partisan gerrymanders.
7. First, plaintiffs' experts do not provide a single measurement for the efficiency gap ("EG") for courts to use. Their methods are based upon the same approach, but utilize differing assumptions without providing a basis for the Court to choose among those assumptions. Their two equations lead to different results, which are large enough that they could represent the difference between a plan inviting Court scrutiny and a plan being presumed constitutional.
8. Second, the metric fails to account for the "natural" packing that can occur if party members are disproportionately clustered in certain types of areas, or if a law such as the Voting Rights Act forces packing of partisans of one party, but not of the other. This is important because if efficiency gaps are not accounting for "natural" clustering, then at least some of the asymmetry they are remedying is not a result of state action. If significant geographic clustering occurs, and is not accounted for, then the EG is really acting as a sort of "make up call" for natural effects and for the effects of the Voting Rights Act. This is true even if a mapmaker can draw a map with a smaller efficiency gap.
9. To better understand the issue of geographic clustering, and why it is so crucial to understanding the limitations of the wasted votes metric, consider the following examples.
10. The following maps depict a hypothetical state "X." It has 100 individual voters, who live conveniently on a ten-by-ten grid. Voters who always vote for the Republican candidate are color coded red, while voters color who always vote for the Democratic candidate are color coded blue. The state has four legislative districts.
11. We start with an example where the voters are proportionally clustered, with Republicans living in the eastern half of the state and Democrats living in the western half:

| - | - | - | - | - | - | - | - | - | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| - | - | - | - | - | - | - | - | - | - |
| - | - | - | - | - | - | - | - | - | - |
| - | - | - | - | - | - | - | - | $\bullet$ | - |
| - | - | - | - | - | - | - | - | $\bullet$ | - |
| - | - | - | - | - | - | - | - | - | - |
| - | - | - | - | - | - | - | - | - | - |
| - | - | - | - | - | - | - | - | - | - |
| - | - | - | - | - | - | - | - | - | - |
| - | - | - | - | - | - | - | - | - | - |

12. We can further imagine a scenario where mapmakers attempt to draw compact districts under neutral principles, and so simply divide the state into evenly matched quadrants:

13. In this scenario, the parties are evenly matched, and the EG is zero.
14. Note that a similar effect would occur if there were zero clustering, and every red voter lived "next door" only to blue voters. In fact, it would be very difficult to draw districts that were not evenly matched under that scenario.
15. Of course, it is still possible to draw maps to partisan advantage in this scenario. For example, the following lines would result in one district that would have 25 Democratic voters and zero Republicans, one that would have 8 Democrats and 17 Republicans, one that would have 7 Democrats and 18 Republicans, and one that would have 10 Democrats and 15 Republicans. Under this, the EG is equal to -.25 , inviting court scrutiny under plaintiffs’ standard.

16. At the same time, if you flipped the lines around a vertical line in the middle of the state, creating a mirror image of the above map, you would have a map with an identical Democratic advantage. In other words, in this scenario the Republicans and Democrats have equal abilities to draw lines to their advantage.
17. If this were how partisans were actually dispersed, there might be merit to plaintiffs' approach, as we would have a baseline for what efficiency gaps should be under neutral principles. But the world is not so tidy. Imagine a slightly different scenario, where a state's Democratic voters are moderately clustered toward the southern edge of the state. The remaining voters are evenly dispersed throughout the state.

18. In this scenario, our even division of the state into quadrants results in two Republican and two Democratic districts, but it is a closer call. Beginning in the top left quadrant, and proceeding clockwise, the districts have: 17 Republicans and 8 Democrats, 17 Republicans and 8 Democrats, 10 Republicans and 15 Democrats, and 6 Republicans and 19 Democrats. Under this scenario, the EG is zero.
19. But note how sensitive this scenario is to slight shifts in partisanship. If three Democrats in the southeastern portion of the state vote differently, we have three Republican districts in a state that would still be evenly split. An even efficiency gap would be transformed
into an efficiency gap of -.25 under the Jackman approach, and of -.19 under the Mayer approach. Court scrutiny would be invited as a result of just three percent of voters changing their underlying voting pattern.
20. The northeastern and southeastern districts can be tweaked to draw three Republican districts with relative ease, while maintaining true compactness in the western portion of the state; drawing three Democratic maps is difficult:

State ' $X$ ' with Moderate Clustering, 3 R Districts

21. Let's imagine one final scenario to bring the point home. In this scenario the voters in the state are heavily clustered in the southwestern corner of the state, while the remaining partisans are more evenly dispersed. We again draw our familiar "grid" districts:

State ' $X$ ' with Heavy Clustering

22. Under this scenario, utilizing our original "neutral" map drawing techniques actually results in three reliably Republican districts. Beginning in the northwestern corner and proceeding clockwise, the districts contain: 17 Republicans and 8 Democrats, 17 Republicans and 8 Democrats, 16 Republicans and 9 Democrats, and 25 Democrats. Under this scenario, the efficiency gap is -.25 . Court scrutiny is invited as a result of applying neutral principles.
23. Of course, you can still draw two, or even three Democratic districts under our "clustered" scenario using relatively compact districts. But this misses the point. The point is that if significant partisan clustering occurs in a state, application of undeniably neutral redistricting principles would nevertheless result in a disproportionate number of wasted Democratic votes, and could invite court scrutiny. Moreover, it is easy to draw Republicanleaning districts - it takes a few minutes of effort - while drawing Democratic leaning districts requires some ingenuity.
24. In short, under a scenario where significant clustering occurs, you actually have to engage in what would traditionally be called gerrymandering in order to draw a neutral map.
25. As the report shows, this is exactly what occurred in Wisconsin. This is obvious from a simple visual inspection of maps of Wisconsin precincts and counties over time. The Democratic vote begins dispersed across the state, but becomes increasingly clustered geographically over time. The gradual consolidation of the Democratic vote into a few key Wisconsin counties coincides with the growth of wasted Democratic votes.
26. This report also measures the consolidation of the Democratic vote quantitatively, finding that heavily Democratic precincts tend to be clustered closer to other heavily Democratic precincts in Wisconsin than Republican precincts are to other Republican precincts, and that this trend has accelerated over the course of the past decade.
27. A failure to account for this and the "natural" wasted votes that occur as a result of clustering calls into question the usefulness of the wasted votes metric as a measurement of gerrymandering - at least as gerrymandering is commonly understood. When significant clustering occurs, a party can "gerrymander" while drawing lines without partisan intent.
28. This leads to the third problem with plaintiffs' approach: It is both underinclusive and overinclusive. The report examines those states that would invite court scrutiny under the metric, and finds an odd mixture of maps that were drawn with obvious partisan intent, as well as maps that could not reasonably qualify as partisan gerrymanders.
29. For example, the EG metric finds that New York and Wisconsin in the 2000s would qualify as partisan Republican gerrymanders. But Democrats drew Assembly districts in New York, while Wisconsin's map in the 2000s was drawn by a Court. Both are examples of states where there is a high degree of partisan clustering: in New York City and in Dane/Milwaukee/Rock counties respectively.
30. At the same time, almost all observers agree that Democrats gerrymandered aggressively in Illinois, at least as commonly understood, in a bid to shore up their majorities in the state. Yet those maps would not invite scrutiny under the proffered standard.
31. Because the standard does not account for the naturally occurring clustering of partisans that has grown in Wisconsin recently, and because the metric brings under its ambit maps that are clearly not partisan gerrymanders, as commonly understood, while excluding maps that were clearly drawn with heavy partisan intent, it is not a solution to the problem of identifying unconstitutional partisan gerrymanders that has flummoxed federal courts for decades.
32. Fourth, the imputation strategy employed to solve the problem of uncontested districts results in a skewing of efficiency gaps in Wisconsin.
33. Fifth, the EG metric fails to account for important effects, such as incumbency and campaign spending.
34. Sixth, the EG metric is overly sensitive to slight changes in votes.
35. Seventh, EGs do not mean that parties are effectively locked out of the political process.

## EXPERT CREDENTIALS

36. I have studied and followed United States elections on both a part-time and fulltime basis for almost two decades.
37. I received a B.A. from Yale University in 1995, with a double major in history and political science.
38. I received a J.D. from Duke University in 2001.
39. I also received an M.A. from Duke University in 2001, in political science.
40. I joined RealClearPolitics in January of 2009 as their Senior Elections Analyst. I assumed a fulltime position with RealClearPolitics in March of 2010.
41. RealClearPolitics is one of the most heavily trafficked political websites in the world. It serves as a one-stop shop for political analysis from all sides of the political spectrum and is recognized as a pioneer in the field of poll aggregation. It is routinely cited by the most influential voices in politics, including David Brooks of The New York Times, Brit Hume of Fox News, Michael Barone of The Almanac of American Politics, Paul Gigot of The Wall Street Journal, and Peter Beinart of The Atlantic.
42. My main responsibilities with RealClearPolitics consist of tracking, analyzing, and writing about elections. I also am in charge of rating the competitiveness of House of Representatives races, and collaborate in rating the competitiveness of Presidential, Senate and gubernatorial races. As a part of carrying out these responsibilities, I have studied and written extensively about demographic trends in the country, as well as the approaches that parties use to draw lines.
43. As part of familiarizing myself with how parties have drawn lines over the decades, as well as learning the political geography of the United States, I drew, using Adobe Illustrator, complete maps of every congressional district ever drawn, dating back to 1789 . Examples of these maps are attached as Exhibits 3-12.
44. I am also a Senior Columnist for Dr. Larry Sabato's "Crystal Ball." I began writing for the Crystal Ball in January of 2014.
45. The overarching purpose of my writings, both at RealClearPolitics and the Crystal Ball, is to try to convey more rigorous statistical understandings of elections than are typically found in journalistic coverage of elections to a lay audience.
46. I am the author of The Lost Majority: Why the Future of Government is up For Grabs and Who Will Take It. The book offers a revisionist take on realignment theory. It argues that realignments are a poor concept that should be abandoned. As part of this analysis, it conducts a thorough analysis of demographic and political trends beginning around 1920 and continuing through the modern times. It was one of the first examples of the dangers the Democratic Party faced from the increased geographic concentration of its coalition.
47. I also authored a chapter in Dr. Larry Sabato's Barack Obama and the New America: The 2012 Election and the Changing Face of Politics, which discussed the demographic shifts accompanying the 2012 elections. I also authored a chapter in Sabato's The Surge: 2014's Big GOP Win and What It Means for the Next Presidential Election, which discusses demographics and Electoral College shifts.
48. I co-authored the 2014 Almanac of American Politics. The Almanac is considered the foundational text for understanding congressional districts and the representatives of those districts, as well as the dynamics in play behind those elections. PBS's Judy Woodruff described the book as "the oxygen of the political world," while NBC's Chuck Todd noted that "[r]eal political junkies get two Almanacs: one for the home and one for the office." My focus was researching the history of and writing descriptions for many of the newly-drawn districts.
49. I have spoken on these subjects before audiences from across the political spectrum, including at the Heritage Foundation, the American Enterprise Institute, the CATO Institute, the Bipartisan Policy Center, and the Brookings Institution. In 2012, I was invited to Brussels to speak about American elections to the European External Action Service, which is the European Union's diplomatic corps.
50. It is my policy to appear on any news outlet that invites me, barring scheduling conflicts, and I have appeared on both Fox News and MSNBC to discuss electoral and demographic trends. I have spoken on a diverse array of radio shows such as First Edition with Sean Yoes, the Diane Rehm Show, the Brian Lehrer Show, the John Batchelor Show, the Bill Bennett Show, and Fox News Radio. I have been cited in major news publications, including The New York Times, The Washington Post, The Los Angeles Times, The Wall Street Journal, and USA Today.
51. I sit on the advisory panel for the "States of Change: Demographics and Democracy" project. This three-year project is sponsored by the Hewlett Foundation and involves three premier think tanks: The Brookings Institution, the American Enterprise Institute, and the Center for American Progress. The group takes a detailed look at trends among eligible voters and the overall population, both nationally and in key states, in an attempt to explain the impact of these changes on American politics, and to create population projections, which the Census Bureau abandoned in 1995.
52. I previously authored an expert report in Dickson v. Rucho, No. 11-CVS-16896 (N.C. Super Ct., Wake County), in which I was asked to identify the partisanship of various districts and opine as to whether they were drawn with partisan intent. It is my understanding that my report was accepted without objection. I have also authored an expert report in a nearly identical version of this litigation, brought in federal court.
53. I also previously authored two expert reports in NAACP v. McCrory, No. 1:13CV658 (M.D.N.C.), which involves challenges to multiple changes to North Carolina's voter laws, including a reduction in early voting days and elimination of same-day registration. I testified at the trial phase of that litigation.
54. I also previously authored an expert report in NAACP v. Husted, No. 2:14-cv-404
(S.D. Ohio). There was no live testimony at the preliminary injunction phase of that litigation, but it is my understanding that my expert report was accepted by and cited to by the Court without objection. I have also authored an expert report in a later iteration of that litigation, Ohio Democratic Party v. Husted, No. 2:15-CV-1802 (S.D. Ohio), and testified at trial.

## OPINIONS

## I. Plaintiffs' Experts Do Not Offer A Unified Definition of the Efficiency Gap

55. It is at times difficult to critique plaintiffs' conception of the efficiency gap, because their experts offer two different formulas for measuring that gap. This difference can be consequential.
56. Dr. Jackman calculates the EG with respect to the votes-to-seats curve. For him, the EG is generated from the equation "EG=S-.5-2(V-.5)," where " S " is the share of seats a party wins in a given jurisdiction and "V" is the share of votes that a party wins. Jackman at 16.
57. Dr. Mayer, by contrast, defines the efficiency gap as "the difference between the sum of wasted votes for the two parties, divided by the total number of votes cast in the election." Mayer at 43. Dr. Mayer also expresses his metric in terms of percentages, while Dr. Jackman expresses his metric in decimal form, although in mathematical terms the scale is identical. For purposes of this report, I will express both in decimal form.
58. To see how these values can vary, consider two examples provided in Dr. Mayer's report. On page 50, Dr. Mayer estimates the results Act 43 would have produced had all seats been open. On page 48, he estimates the results from his sample plan.
59. According to Dr. Mayer's calculations, the EG for Act 43 is -.1169 . But employing Dr. Jackman's formula, the EG is -.0985 .
60. Similarly, according to Dr. Mayer's calculations, the EG for his demonstration plan is -.219. Under Dr. Jackman's formula, the EG is -. 0077 .
61. The difference in measurement with respect to Dr. Mayer's estimated Act 43 result is .0141 points. The difference in measurement with respect to Dr. Mayer's estimated demonstration plan is .0184 points. When one considers that Dr. Jackman's measurements of historic efficiency gaps stretch only from -. 18 to .2 , this is a substantial, meaningful amount of uncertainty. If a court adopts Dr. Jackman's approach to the efficiency gap, it will likely result in a somewhat different universe of states found presumptively unconstitutional than if it adopts Dr. Mayer's approach.

## II. The Clustering of the Democratic Coalition creates "natural" packing, which the Efficiency Gap metric does not account for.

62. In 2002, John Judis and Ruy Teixeira wrote a book entitled "The Emerging Democratic Majority." In their telling, the Democratic Party of the 1990s was undergoing a transformation, and would emerge as a dominant party as a result of its coalition of minorities, women, creative class professionals and working class voters. This, they surmised, would enable Democrats to control the House, Senate and presidency into the future.
63. In 2011, I wrote a book called "The Lost Majority: Why The Future of Government is Up for Grabs, and Who Will Take It." It observed that Judis and Teixeira had been correct about a great many things, but had also overlooked the degree to which the new coalition would alienate older members of the Democratic coalition (as well as relying upon a faulty political science concept known as realignment theory). In particular, the increasingly liberal Democratic coalition alienated more conservative working class and rural voters, which Judis and Teixeira assumed would form the fourth portion of the Democrats' coalition.
64. My book argued that this trend among white working class voters and rural voters would help keep Republicans competitive at the presidential level for the foreseeable future. It also concluded that this should not have been surprising, as the story of American politics is one of ever changing coalitions, as the growth of one group pushes a group without countervailing interests into the arms of the other party.
65. But I noted that the Democrats' new coalition was uniquely problematic at the state legislative and congressional level. Because liberals, young voters, minorities, and other members of the Democrats' coalition tend to be concentrated in cities and/or placed into minority majority districts, this damaged their ability to win congressional districts, which reward parties with a wide geographic reaches (as illustrated in the introduction to this report).
66. Consider the West South Central region of the country. The following maps show the counties won by Republican and Democratic presidential candidates, utilizing the familiar red/blue color scheme.

67. When Bill Clinton ran for re-election in 1996, he won nationally by about eight points. As we can see, his support in the region was geographically dispersed, which allowed him to carry around 54 percent of the Congressional districts in the region. This, in turn, helped Democrats win around 50 percent of these districts.
68. Barack Obama won nationally in 2008 by about seven points, yet this did not translate into success in the region. He ran about eight points behind Clinton's 1996 showing here. Interestingly, he actually ran about three points ahead of John Kerry in this region, yet carried fewer counties. The difference is that he carried several urban counties that neither Kerry
nor Clinton carried, such as Harris County, Texas (Houston), Jefferson County, Alabama (Birmingham) and Dallas County, Texas. But because his coalition shrank geographically, the net result was disadvantageous to Congressional Democrats; then-Senator Obama carried only 23 percent of the Congressional Districts in the region, with Democrats winning 39 percent of the seats. The latter number fell to 26 percent in 2010.
69. You can see the effects of geographic clustering in sharpest relief in a state like Virginia. Here, Barack Obama won by six points in 2008, while Bill Clinton had lost by two (despite the fact that they had won by similar margins nationally). Yet, from a geographic perspective, Obama's coalition was quite a bit narrower.
70. Obama shed voters, even from Kerry's losing coalition, in the western portion of the state, carrying only Montgomery County (Virginia Tech). He and Kerry added Albemarle County outside fast-growing Charlottesville (University of Virginia), and he performed well in the African American rural counties. He also added suburban Henrico County near Richmond, and carried some counties in the Hampton Roads area that Kerry and Clinton failed to carry. But the biggest gains are obvious, coming in northern Virginia. Obama became the first Democrat since LBJ to carry Loudoun and Prince William counties, and the second to carry Fairfax (Kerry was the first).

71. There is little doubt that the Democratic vote in Wisconsin is also increasingly concentrated in fewer counties. To understand the following analysis, we must first understand the concept of a state's Partisan Index.
72. A state's Partisan Index is computed by subtracting the share of the state that voted for the Republican presidential candidate from the share of the nation that voted for the Republican presidential candidate. For purposes of these calculations, third parties and independent candidates are excluded (i.e., we use what political scientists call the "two-party vote").
73. To illustrate the utility of the Partisan Index, consider the following example. In 1984, Ronald Reagan won 51.4 percent of the two-party vote in Massachusetts. In absolute terms, one could consider Massachusetts a swing state. But no one would have considered Massachusetts a swing state, because it had two Democratic senators, a Democratic governor,
and an overwhelmingly Democratic legislature. Ten of the state's eleven congressional districts elected Democrats, and the one Republican, Silvio Conte, was very liberal Republican.
74. Moreover, one would conclude that, using absolute terms, the state has swung wildly toward Democrats in the interim, since Barack Obama won 61.8 percent of the two-party vote in the state in 2012.
75. But Reagan's 51.4 percent win in Massachusetts has to be viewed in the context of his winning 59.2 percent of the two-party vote nationally. Compared to the country as a whole, Massachusetts actually had a Democratic lean of 7.8 points in 1984.
76. Likewise, Obama's 61.7 percent win in Massachusetts has to be viewed in the context of his winning 52 percent of the two-party vote nationally. Compared to the country as a whole, Massachusetts actually had a Democratic lean of 9.8 points in 2012. Viewed in this light, Massachusetts has actually had relatively stable politics since 1984, with only a slight shift toward Democrats.
77. In short, Partisan Index allows us to control for national effects, and compare results across elections.
78. Wisconsin's PI has been mostly stable since the 1980s. After dipping to nearneutrality, during the 1990s, it shifted modestly leftward in the 00 's.

79. In this report, we begin by looking at Partisan Index on the county level across Wisconsin in a series of maps, with particular attention paid to 1996, 2004 and 2012, which represent years where the PIs of the state were similar (1.19, 1.43, and 1.54, respectively).

## Wisconsin County PI 1988


80. In 1988, the Democratic Party in Wisconsin had a broad geographic reach. It was strongest on the Menominee Indian Reservation ( $\mathrm{PI}=26.86$ ), as is the case today. The other four most Democratic counties include Douglas (22.47), Milwaukee (15.34), Ashland (14.63) and Dane (14.3). Seventy-one percent of counties had Democratic leans, and the Democratic Party covered the entire western portion of the state, particularly in the northwest. Republicans were relegated to the German-settled counties in the southeast and east-central portions of the state (note: The map caps the color-coding at PIs of -.1 and .1, in order to minimize the effect of outliers on the overall color-coding scheme).

## Wisconsin County PI 1996


81. By 1996, the state as a whole had become modestly more Republican compared to the country as a whole, so it is unsurprising that the number of Republican counties increased; 45 counties ( 62.5 percent) had Democratic leans. But the shift was uneven. Democratic performance fell by just 4.5 points and 4.2 points in Milwaukee and Dane Counties, respectively. It fell by nine points in Douglas County, however, as the northwest became noticeably less Democratic.

## Wisconsin County PI 2004


82. In 2004, Wisconsin was once again marginally more Democratic than the country as a whole, but the political divisions looked quite different than they had in 1996. Democrats maintained their strength in the three industrial counties on the Lake Superior shoreline, as well as in the southwestern portion of the state. Milwaukee and Menominee Counties were Democratic as well. Ashland, Bayfield, and Douglas counties were 2.5 percent, 3.5 percent, and 4.2 percent more Republican than the country as a whole, respectively, than they had been in 1996. Milwaukee was 3.8 percent more Democratic. Menominee and Dane counties were both 7.9 percent more Democratic than they had been in 1996.
83. It was a different story in less populated counties. Forest County swung 9.2 points toward Republicans, Crawford County swung 1.2 points toward Republicans, and Adams County swung four points toward Republicans. The total number of Democratic-leaning counties dropped to 33 , or just 46 percent of the counties in the state. Overall, the bluest counties tended to become bluer, while the rest of the state shifted rightward.

## Wisconsin County PVI Change, 1996-2004


84. In 2012, the state was in roughly the same position relative to the country as a whole as it had been in 2004. But the stable orientation of the state overall masked significant internal movement. Dane and Milwaukee Counties swung a couple of points toward Democrats, along with some of the southwestern counties. Douglas and Ashland counties, along with most of the northwestern portion of the state, actually moved a touch toward Republicans. Overall, although the state was almost identically as Democratic in 2012 as it was in 1996, only 27 counties retained a Democratic lean in the latter year, or just 37.5 percent of the state. Moreover, these counties were geographically concentrated, in the southwestern portion of the state, in the far northwest, and in Milwaukee.

## Wisconsin County PI 2012



Wisconsin County PVI Change, 2004-2012

85. Overall, from 1996 to 2012, the Democratic Party became substantially less competitive in the northwestern portion of the state, as well as in the rural portions of the state outside of the southwestern corner. Its reach was limited to fewer counties, and those counties were clustered in geographically compact regions. You can see this in the map of changes
occurring across the entire time period; Democrats gained primarily in counties that already leaned Democratic at the beginning of the time period, while Republicans gained in places where they had been weak. The state didn't budge politically, but the internal movement was unmistakable. As was the case with the country as a whole, the Democrats' coalition became deeper, but narrower.

## Wisconsin County PVI Change, 1996-2012


86. To put this into further perspective, Dane, Milwaukee, and Rock counties have provided Democrats with their three largest vote margins in every election since 1992 (inclusive). In 1996, Bill Clinton carried these three counties with 64 percent of the two-party vote. He also, however, carried the rest of the state with 52 percent of the vote, for a difference of twelve percent. In 2012, by contrast, even though Barack Obama was winning with a lower vote share (both in Wisconsin and nationally) than Clinton had in 1996, he carried Dane, Milwaukee and Rock counties with 69 percent of the vote. He lost the rest of the state, however, to Mitt Romney, 47 percent to 53 percent. The gap between those three counties and the rest of the state was 22 points. If we look in terms of Partisan Index, we see a similar trend; the gap
between the three counties above and the rest of the state was 12 points in 1996, 18 points in 2004, and 22 points in 2012.
87. We can also take a more rigorous approach to this. Consider the following map of Wisconsin wards in 2012, using Dr. Mayer's modified ward values.

88. This allows us to see that the clustering that is apparent at the county level filters down to the ward level, with Democrats concentrated in the northwest, southwest, and in Milwaukee County.
89. We can see this further in the following chart, reproduced from Jowei Chen and Jonathan Rodden, "Unintentional Gerrymandering: Political Geography and Electoral Bias in Legislatures," 57 Quarterly Journal of Poli. Sci. 200 (2013):

90. Each of these dots represents an estimate for voting units in each state, generated from vote files and the U.S. census. The figure charts them by partisanship (e.g., how heavily each unit voted for the Republican presidential ticket in 2000) on the vertical axis, and by population density on the horizontal axis. As you can see, in Wisconsin (as in many other states), as the units become more heavily Democratic, they also become more densely populated. This suggests that the Democratic vote is heavily concentrated in cities. Even as of 2000, as population density increased in Wisconsin, the Republican share of the vote dropped.
91. We can validate our assumption numerically through a two-step process. First, we want to see whether Wisconsin's wards have become increasingly polarized. That is, are there more heavily Democratic wards today than there were a decade ago? Second, we want to know whether the heavily Democratic wards are located more closely together than heavily Republican wards.
92. From 2002 to 2014, I looked at the top of the ticket race in the state (note: I tested both the "raw" LTSB data and the data recalculated under Dr. Mayer's metric for 2004, 2008 and 2012, and determined that, in this context, utilizing the raw data did not alter any conclusions).
93. To accomplish the first goal, I calculated the average Democratic lean of wards that leaned toward Democrats over the course of the past decade:

94. As you can see, the mean Democratic ward in Wisconsin has moved leftward over the course of the past decade. That is to say, the average Democratic ward in 2014 was 2.5 percent more Democratic than in 2002.
95. At the same time, we do not see any similar effect for Republican wards:

96. This answers the question of whether the Democratic-leaning wards in Wisconsin have become more heavily Democratic over time. To answer our second question, we first need
to sort the wards for each cycle into partisan-filtered maps, using the partisan index as a guide to the state's overall partisanship. That is a complicated way of saying that I took the $\mathrm{D}+1$ wards as a group, the $\mathrm{D}+2$ wards as a group, and so forth.
97. Next, the distance to the nearest neighbor for each ward was calculated, for each subset of partisan indices. To visualize this, imagine creating a grid with all of the $\mathrm{D}+1$ wards listed both horizontally and vertically (if you prefer, an ix j matrix where both dimensions are defined as including the number of wards). The distance from the first ward to every other ward is calculated, filling in the first row of our grid. The smallest value is noted, which represents the distance from ward 1 to the nearest other ward of similar partisan index. The process then repeats for ward 2 , ward 3, and so forth. At the end, the median of the smallest distances is calculated, which gives us an idea how close the $\mathrm{D}+1$ wards are to each other (I utilized the median rather than the mean here because outlying wards, such as Menominee County, exert an undue amount of leverage on averages). The process is then repeated for $\mathrm{D}+2, \mathrm{D}+3$ and so forth. If Wisconsin has, in fact, become more clustered over time, then we should see the median distance decline as the partisanship of wards increases.
98. In fact, this is exactly what we see. The following charts show the wards grouped and labeled from most Democratic to least Democratic, and most Republican to least Republican. It shows the median distance for each grouping from every ward to its closest neighbor of similar partisanship. The quantiles from .45 to .55 are excluded, since they are effectively neutral.


99. As the wards become more Democratic, the distances between them shrinks. By contrast, the Republican ward distances tend to be fairly stable, until we get to the most heavily Republican wards, which are actually more spread out than the more neutral wards.
100. Taken together, these analyses demonstrate that, over the course of the past two decades, Wisconsin's Democratic vote has increasingly found itself relegated to Milwaukee County, the southwestern portion of the state, and a few counties in the northwestern portion of the state. This, in turn, shifts Wisconsin the baseline of Wisconsin maps rightward.
101. We see an example of how this plays out in Dr. Mayer's analysis. He proceeded with a mandate to "draw[] a legislative plan that has an efficiency gap as close to zero as possible while complying with federal and state requirements at least as well as the plan enacted by the Wisconsin legislature in Act 43." Mayer Report at 2. Yet after several days of mapmaking, Dr. Mayer ultimately failed to draw a map with a zero efficiency gap; the efficiency gap was actually -.022 . That is almost $1 / 3$ of the way to being a gerrymander under the standard that plaintiffs urge.
102. Plus we must remember what it means that Dr. Mayer sought to "comply[] with federal and state requirements at least as well as the plan enacted by the Wisconsin legislature." First, it is not clear that he succeeded; his districts have larger population deviations and split more localities (though they split fewer counties) than the Act 43 districts. Id. at 37.
103. But second, and more importantly, plaintiffs' theory is that Act 43 represents an egregious, unconstitutional gerrymander. There is something of a Hobson's choice at work here. Either Act 43 complies with traditional redistricting criteria well, which would divorce plaintiffs' metric from most understandings of gerrymandering even further, or it does not comply with traditional criteria well, in which case it is unclear that even a gerrymander (under most understandings of the term) pointing the other direction would be able to eliminate the efficiency gap entirely.
104. This is important because the efficiency gap metric assumes there is a baseline of zero - that is, if maps were drawn under neutral criteria with neutral intent, there would be no efficiency gap. But as the drawings in our introduction demonstrate, this is not necessarily the case. When natural clustering of Democrats occurs, the efficiency gap created by neutral processes drifts rightward; efficiency gaps increasingly present as a result of factors other than action by the state. This is likely one reason why, as plaintiffs' experts observe, the national trend has been toward increasingly Republican-leaning efficiency gaps, while the larger proDemocratic efficiency gaps tend to occur in earlier decades.
105. What plaintiffs' standard does, at least in part, is force legislatures to enact "make up calls" for natural clustering of Democrats and for the clustering of Democratic-leaning groups required by the Voting Rights Act. In an odd way, by failing to account for the natural distribution of partisans, plaintiffs force legislatures to draw lines with partisan intent.

## III. Plaintiffs' Standard is Both Underinclusive and Overinclusive.

106. This "natural gerrymandering" leads to an additional problem: The efficiency gap invites court scrutiny of maps that are clearly not partisan gerrymanders, while absolving maps where legislators clearly acted overwhelmingly with partisan intent.
107. While the Supreme Court has dismissed partisan intent or proportionality as a workable standard for gerrymandering, it has never intimated that gerrymanders could exist without partisan intent or disproportionate outcomes. The problem lies in creating workable limits determining how much partisan intent is too much partisan intent, or in constructing the counterfactual to predict disproportionate outcomes. At the same time, almost everyone's conception of gerrymandering involves intent to disadvantage a party, and to create disproportionate outcomes. If a proposed standard ignores a large number of maps drawn with
clear, overwhelming partisan intent, or includes a large number of maps that could not reasonably be argued to be gerrymanders, there is a good chance that the metric so radically alters the understanding of gerrymandering that it, in fact, is capturing something entirely different than gerrymandering.
108. Dr. Jackman identifies 17 maps with an "unambiguous history" of having a consistent efficiency gap sign over the lifespan of the plan. Jackman at 55.
109. But many of the states that would be included in the definition of a gerrymander here are poor candidates for the label, at least as most people would understand it. Table 1 shows the states on the list, as well as the party that controlled the governorship, state senate, and state house in the year prior to reapportionment.

Table 1: Partisan Control of Redistricting, Maps Id'd With "Unambiguous EGs"

| State | Year | Gov | House | Sen |
| :--- | :---: | :---: | :---: | :---: |
| FL | 2002 | R | R | R |
| CA | 1992 | R | D | D |
| CO | 1982 | D | R | R |
| CO | 1972 | R | R | R |
| IL | 1992 | R | D | D |
| MI | 2002 | R | R | R |
| MI | 1992 | R | R | D |
| MO | 2002 | D | D | R |
| NY | 2002 | R | R | D |
| NY | 1992 | D | R | D |
| NY | 1972 | R | R | R |
| NY | 1982 | D | R | D |
| OH | 2002 | R | R | R |
| OH | 1994 | R | R | D |
| PA | 1982 | R | R | R |
| WI | 2002 | R | D | R |
| FL | 1972 | D | D | D |

110. Only seven of the seventeen states included in the list of gerrymandered states feature unified partisan control of redistricting in the year where reapportionment was conducted (Ohio in 1992 drew its district lines through a Republican-controlled apportionment board). In five of those seven instances (the two Florida maps being the exception), control of at least one of the maps that produced unambiguous histories of consistent efficiency gaps switched partisan
hands at least once. The results of New York's 1972 map were particularly dramatic; by the end of the decade an 83-66 Republican lead in the state Assembly had transformed to an 85-64 Democratic lead; the 1972 elections actually marked the last election where Republicans would control the Assembly. This suggests that even enduring efficiency gaps do not necessarily translate into one side or the other being locked out of the legislative process (see below).
111. The remaining maps are poor candidates for gerrymanders, at least as the term is commonly understood. The Almanac of American Politics 1994 described the 1992 California plan (to simplify things, I refer to the year the plan was implemented, rather than the year it was actually adopted): "The key decisions for the 1990s California maps were made by the voters in 1990 and 1986. In 1990 they elected Republican Governor Pete Wilson, thus depriving Democrats of the untrammeled control they had over redistricting in 1982 and $1962 \ldots$. Wilson held solid to his plan to appoint a redistricting commission to draw up plans for Congress and the legislature, and then handed them over to the state Supreme Court, which in January 1992 adopted them. In fact, the plan is more evenhanded than a Republican redistricter of, say, Phil Burton's abilities would have concocted. The lines are far more regular than in the ultrapartisan plan passed in Texas by the Democrats (this decade's winner of the Burton award)." Almanac of American Politics 1994 at 86.
112. In 1992, the Michigan state legislature failed to pass a reapportionment plan. The state Supreme Court appointed a panel of three special masters, which rejected the plans submitted by the state parties as excessively political. It instead implemented its own plan, which the state Supreme Court approved. See NAACP v. Austin, 857 F. Supp. 560 (E.D. Mich. 1994).
113. In 2002, the Missouri legislature deadlocked, and failed to pass any redistricting plan. The map was drawn by a committee of court of appeals judges.
114. The inclusion of New York's maps as potential gerrymanders is particularly perplexing. Control of redistricting has been split since the 1970 maps were drawn, and the tradition that has emerged is that the Republican-controlled senate draws the 63 -member senate map, while the Democratic-controlled assembly draws the 150 -member assembly map. The reason New York consistently presents as a Republican gerrymander has little to do with the lines drawn, but rather derives from the concentration of the Democratic vote. In 2012, Barack Obama carried New York state by two million votes, but carried the area outside of New York City by just 441,000 votes. These votes are also concentrated (in places like Hempstead and Islip on Long Island), which means that, even with Democrats drawing the Assembly lines and a 441,000 presidential vote advantage to work with, they are able only to split the Long Island and upstate districts evenly with Republicans.
115. The standard also overlooks some of the more obvious examples of redistricting with partisan intention. For example, at the congressional level, the 2004 Almanac of American Politics describes the 2002 redistricting process in Alabama as follows: "[t]he Democrats in control of redistricting in Alabama in 2002 did a pretty good job of helping their party in drawing the boundaries of the state's seven congressional districts, but not quite good enough of a job to add to the two seats they have held since 1994." Id. at 54. The map the Democrats produced in a bid to shore up their majorities produces an efficiency gap of -.125 , which would invite court scrutiny as a Republican gerrymander.
116. In Colorado in 2002, a court selected a Democratic-drawn map for Congress and state House; Republicans were so infuriated by this that they attempted a mid-decade
redistricting when they next controlled the legislature. Id. at 303-04. But the Democratic plan actually produces an efficiency gap of -.09 , which would invite court scrutiny as a Republican gerrymander.
117. On the other hand, Georgia in 2002 was considered a strongly Democratic gerrymander. The Almanac describes the process: "[a]fter the 1990 and 2000 Censuses, Georgia Democrats, led by Speaker Thomas Murphy, pushed through convoluted redistricting plans arguably the most convoluted in the nation each time - to guarantee majorities for their party in the state's House delegation." Id. at 454. To do saw, the Georgia legislature drew highly convoluted lines, including the new 13th, which has been liked to a "sick chicken." But the map actually had a slight Republican efficiency gap of -.01 .

118. Illinois's congressional districts in 2002 represented a negotiated, bipartisan plan that was broadly acceptable to members of both parties. Id. at 528-29. Yet it presents with an efficiency gap of -.09 , which would invite Court scrutiny as a Republican gerrymander.
119. Iowa's Legislative Services Bureau is often held up as an exemplar of how nonpartisan redistricting ought to work. Yet in 2002, it presents with an efficiency gap of -.2, which would invite court scrutiny as a Republican gerrymander.
120. North Carolina's 2002 redistricting was likewise controlled by Democrats, who sought to weaken Republican Robin Hayes in the 8th District while shoring up Democrat Mike McIntyre in the 7th District. It was successful in doing just that later in the decade. But in 2002, it presented a marginal Republican lean, with an EG of -.026. It is not a gerrymander under the efficiency gap metric, despite plain partisan intent and convoluted districts, including the second district, which resembles a dragon in flight:

## North Carolina 2002


121. In 2012, the Arizona congressional lines were drawn by an independent redistricting commission. In 2012, it nevertheless presented with a .16 EG , which would invite court scrutiny as a Democratic gerrymander.
122. In 2012, a Colorado district court judge selected a Democratic redistricting plan for Congress. See Almanac of American Politics 2014 at 290-91. In 2012, it nevertheless presented with a -. 099 EG , which would invite court scrutiny as a Republican gerrymander.
123. In 2011, Illinois instituted some of the most aggressive redistricting in the country. As the Almanac reported "[u]nder heavy pressure from party leaders desperate to offset Republican gains in other states, Democrats in May 2011 released a map designed to eliminate up to six Republican seats. . . . The state's Republican delegation immediately put out a joint
statement calling it 'little more than an attempt to undo the results of the elections held just six months ago' and they were largely right." Id. at 541 . Yet the map only presented with an efficiency gap of .058 , which would not trigger court scrutiny.
124. Perhaps most strikingly, the Supreme Court conceded in Veith v. Jubelirer that the Pennsylvania map for the 2000s was a "partisan redistricting plan;" the case failed because of the lack of a manageable standard. While it presented as a partisan redistricting plan in 2002, in 2006 the efficiency gap was only -.04 , while in 2008 it was actually a positive .033 . In other words, had the national environment been worse for Republicans in 2002, the efficiency gap might conclude that the Veith map was actually a modest Democratic gerrymander.
125. In Figure 36, which examines the current legislative maps, Dr. Jackman finds actionable EGs for Rhode Island and Vermont on the Democratic side, and for Florida, Michigan, Virginia, North Carolina, Kansas, Indiana, New York, and Wyoming. A majority of states overall appear to have at least one year of 2012 or 2014 outside of the actionable .07 threshold identified by Dr. Jackman.
126. But as seen in the chart reproduced from Chen \& Rodden above, there simply are not many precincts in Wyoming that lean Democratic; the same is likely true in Vermont and Rhode Island (oddly, efficiency gaps can present when the opposite of clustering occurs: When one party is politically dominant and partisans for the other party are so spread out that it is impossible to cluster them into districts). Democrats drew the Assembly in New York, while Kansas is a judge-drawn map. At the same time, maps that are generally thought to represent aggressive partisan maps, such as Arkansas and Illinois, appear as neutral maps under plaintiffs' standard.
127. Finally, the EG narrative is problematic for Wisconsin in particular. If the EG were a good measure of gerrymandering, we'd expect some sort of measurable difference between gaps to occur in redistricting years. In other words, we would expect that there would be natural variations over time, but overall we should see a "stepped" pattern to the chart of efficiency gaps over time, with the steps corresponding to redistricting years.
128. But this is not what we see in Wisconsin. As Dr. Jackman notes in Fig. 35, the time period from 1970 to 1996 shows relatively stable EGs in Wisconsin, regardless of who controls redistricting. But 1996 is the last year for which we see such balance. A substantial, fairly steady dropoff begins in 1998. Six of the nine post-1996 EGs appear to be large enough to be actionable under plaintiffs' theory. Worse, three of those six cases occurred under the 2001 redistricting, which resulted in a court-drawn legislative map. Indeed, it is not at all clear that the current map is appreciably different in terms of gaps from the map that was drawn by the court for the 2000s.
129. Instead, what we observe appears to ape national trends. The following two charts are taken from page 873 of the article from Drs. Eric McGhee and Nicholas Stephanopolous (which underlies this litigation). They show the average net and absolute efficiency gaps from 1972 to 2012 for Congressional and state legislative seats:

130. In both instances, we see the same thing: A clear pro-Republican trend in the overall net efficiency gap, but one that is not keyed off of redistricting years. Instead, the congressional chart begins a steady downward trajectory beginning with the 1994 elections (with the largest drop occurring in 1996), while the state house chart shows a dropoff beginning in 1990 (with a similar acceleration occurring in 1996). The EGs demonstrated in 2012 in both maps are similar to EGs that manifested in 2006, and the large drop-offs tend not to occur in
redistricting years. This suggests that the efficiency gaps we see are in large part due to exogenous forces, such as natural partisan clustering, rather than gerrymandering.
131. Efficiency gaps are growing in ways that gerrymandering has difficulty explaining, and are present in maps drawn by courts, by independent commissions, and by members of the opposing party. Given this, it is unclear why the existence of an efficiency gap would provide prima facie evidence that members of a party have had their right to vote diminished by state action.

## IV. Dr. Jackman's Imputation Strategy is Problematic.

132. One of the great challenges of utilizing the efficiency gap is dealing with the problem of uncontested districts. Unopposed candidates will artificially inflate a party's popular vote total, and can skew the efficiency gap if they are disproportionally allocated to one party or the other.
133. Dr. Jackman's solution, when the data are available, is to use presidential vote share in the district (he has a different solution when presidential votes are not available). He notes that there is a tight correlation between the presidential vote share and state house vote share. Therefore, when state house vote shares are missing because of an uncontested election, Dr. Jackman substitutes presidential vote share from a similar district.
134. But there are two interrelated problems with this. First, we are not simply concerned with the r-square here (which, in lay terms, tells us how well knowing the value of variable A helps us to predict the value of variable B). We are also concerned with the coefficient, or the slope of the best fit line. If every percent increase in presidential vote share yielded a .5 percent increase in state house vote share, we would have a very high r-square, but we would not want to use this as a substitute.
135. Second, plaintiffs' own experts provide some good evidence suggesting that there may, in fact, be a systemic bias involved in imputing presidential results to state House results. Dr. Mayer demonstrates that there were many fewer unconstested Republican districts in 2012 than uncontested Democratic districts. Mayer at 40. Therefore, Dr. Jackman is imputing votes for far more Democratic districts than Republican districts.
136. In and of itself, this is not a problem if the imputation strategy is correct. But on page 15, Dr. Mayer plots a line that represents a 1:1 ratio between presidential and assembly votes for Republicans and Democrats. That is, if every ward showed the same number of votes for president and assembly, every dot would fall on the line.
137. Figure 2 demonstrates that imputation is acceptable for Republican wards in Wisconsin, since the dots appear to fall more-or-less on the line.
138. For Democrats, however, the dots systematically fall below the line, often creating differences on the order of 10 percent.
139. The net effect of this will be to skew the imputation. It suggests that too many votes are being imputed in wards reporting a high number of Democratic votes, which will skew popular vote totals. In other words, a ward with 100 votes for Romney and 900 votes for Obama probably should not be reported as a 90 percent Democratic ward with 1,000 votes cast. It should probably be reported as an 89 percent Democratic ward with 900 votes cast. The impact of this will be particularly pronounced, given that there are more imputations being performed for Democratic districts than Republican districts.

## V. The Efficiency Gap Metric Ignores Important Factors, Such as Incumbency, Candidate Quality, Campaign Spending, and Recruiting Advantages.

140. When Dr. Mayer models his efficiency gaps, he notes that incumbency has a statistically significant impact on vote totals (this is one reason why he ultimately models results
without any incumbents). Other factors, such as candidate quality, campaign spending, and recruiting advantages are acknowledged as having positive effects on turnout. E.g., Eric McGhee \& John Sides, "Do Campaigns Drive Partisan Turnout?" 33 Polit. Behav. 313-333 (2010).
141. In other words, if one party has a disproportionately strong get-out-the-vote effort in place, or better candidates, or fewer incumbents, it can alter the popular vote totals and alter the efficiency gap.
142. In other words, there are important factors in addition to clustering that can alter the efficiency gap, and which the presented EG metric does not account for.

## VI. Efficiency Gaps Are Sensitive To Slight Changes.

143. This might not be a problem if the Efficiency Gap was not sensitive to slight changes in turnout or voting behavior. But it is. Consider the following scenario: A Republican legislature redistricts a Democratic-leaning state. It creates five 90\% Democratic districts, a 60\% Democratic district, four 90\% Republican districts, six 55\% Republican districts, a 53\% Republican district, and three 49\% Democratic districts.

Table 2: Redistricting in Hypothetical State

| District | $\mathrm{D} \%$ | $\mathrm{R} \%$ |
| :---: | :---: | :---: |
| 1 | 10 | 90 |
| 2 | 10 | 90 |
| 3 | 10 | 90 |
| 4 | 10 | 90 |
| 5 | 45 | 55 |
| 6 | 45 | 55 |
| 7 | 45 | 55 |
| 8 | 45 | 55 |
| 9 | 45 | 55 |
| 10 | 45 | 55 |
| 11 | 47 | 53 |
| 12 | 51 | 49 |
| 13 | 51 | 49 |
| 14 | 51 | 49 |
| 15 | 60 | 40 |
| 16 | 90 | 10 |
| 17 | 90 | 10 |
| 18 | 90 | 10 |
| 19 | 90 | 10 |
| 20 | 90 | 10 |

144. In the first year after redistricting, if everyone votes as expected, we would see a -. 06 efficiency gap, suggesting that the map was not a Republican gerrymander, under the plaintiffs' proposed standard.
145. But assume that we saw a national Republican wave in the first year, and Republicans fared two points better across-the-board. The map would result in a -.19 efficiency gap, which would constitute a gross "gerrymander."
146. The result would not have to be that dramatic, however. Assume instead that Republicans ran a slightly stronger candidate in district 12, and carried it. The efficiency gap would be -.109 , and the map would be presumed unconstitutional.
147. This is not a wholly hypothetical concern. As discussed above, Dr. Mayer measures Act 43, sans incumbents, of having an EG of 11.69. But assume that through a modestly better GOTV effort, Democrats win 400 more votes in District 1, and 200 more votes in District 94 in the 2012 election. The EG falls by more than two points off these modest shifts, to 9.466 .
148. In other word, the EG metric is sensitive enough that relatively small differences in the electoral outcome can make a difference between whether a map is presumptively unconstitutional or not. While this shift would not make a difference in terms of whether the Wisconsin map invited Court scrutiny, as a national standard, it almost certainly would in other states.

## VI. Efficiency gaps do not mean that stability is created or that parties are locked out of the process.

149. Finally, it is worth noting that EGs do not correlate to partisan outcomes. That is to say, to the extent an equal protection violation derives from foreclosing a party from adequately participating in the political process, the EG does not reveal such a pattern.
150. For example, as noted above, even though New York has consistently had a proRepublican efficiency gap, Republicans have never claimed control of the Assembly. The most severe Republican gerrymander, under the EG standard, came in 2002. Yet despite the fact that the EG never rises above -. 078 under that map -every election results in an actionable Republican gerrymander-Democrats always controlled the Assembly by a large margin.
151. The Michigan 2002 map is counted as a Republican gerrymander, yet Democrats won the state House in 2006 and 2008. Likewise, the Michigan 1992 map is counted as a Republican gerrymander, yet Democrats controlled the state House throughout the decade.
152. The Colorado 1972 map is counted as a Republican gerrymander, yet Democrats won the state House twice under the map (in what was then considered a Republican state).
153. Likewise, even though California's 1992 map is counted as a Republican gerrymander, Democrats managed to win unified control of the legislature in 1996, 1998 and 2000.
154. This is not to say that partisan outcome provides a workable legal standard for analyzing gerrymanders. If anything, the foregoing merely proves the point that forecasting actual partisan outcomes over the course of a decade can be difficult. But when a standard for gerrymandering does not align with outcomes in a backward-looking analysis, it calls into question the utility of the metric as a standard overall.

## CONCLUSION

155. The EG is a clever metric, propounded by some of the political scientists I hold in the highest regard. But as a legal standard, it is highly problematic. For a variety of reasons described above, it casts its net both too widely and not widely enough. Moreover, it effectively forces mapmakers to gerrymander to "fix" things that do not result from state action.

This the 2nd day of December, 2015.


Sean P. Trende

Rebuttal Report: Response to Expert Reports of Sean Trende and Nicholas Goedert

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December 21, 2015

This report presents my responses to the criticisms that Sean Trende and Professor Nicholas Goedert make of my report. ${ }^{1}$

## I. Summary

A. Both Trende and Goedert erroneously argue that Democrats are more geographically concentrated than Republicans in Wisconsin, which creates a natural pro-Republican bias even under a neutrally-drawn district plan. Both arguments are based on unreliable methodologies, flawed measures, and lead to inaccurate conclusions. Trende's methodology for measuring partisan concentration relies on an unorthodox method (the PVI) far more common among political commentators than academics who study spatial patterns of concentration and isolation. Moreover, as he applies it here, Trende relies on fundamentally inaccurate measures of geography that are guaranteed to demonstrate that Democratic wards are closer to one another than are Republican wards.

Goedert's arguments about geographic concentration are analogous to Trende's, and suffer from the same flaws in that they are based on superficial claims that do not rely on actual measures of spatial concentration or isolation. Moreover, Goedert's claims here contradict his own research, in which he finds that even after controlling for urbanization (a proxy for concentration), Republican control of the redistricting process has a large and statistically significant impact on a plan's bias. A model in one of his papers (Goedert 2015) also shows that a court-drawn or bipartisan map in Wisconsin would be expected to produce a pro-Democratic bias. The model generates the same expectation for a court-drawn or bipartisan map in a state that resembles the country as a whole. Accordingly, based on Goedert's own analysis, there is no natural pro-Republican tilt in either Wisconsin or the typical U.S. state.

In contrast to Trende's and Goedert's unorthodox techniques, widely (even universally) accepted measures of spatial distributions, such as Global Moran's I (Cho 2003) and the Isolation Index (Reardon 2004), show that Wisconsin's Republicans and Democrats are equally spatially concentrated and equally spatially isolated from each other, and that in some election years Republicans are more concentrated than Democrats.
B. Trende criticizes my method of estimating the partisanship of uncontested Assembly districts as biased. But his criticism stems from a superficial and erroneous discussion of a single figure in my report (Figure 2), and he erroneously believes that I set the Assembly votes in uncontested districts to the presidential vote in those districts. He does not take notice of the fact that my analysis was based on a comprehensive multiple regression model that controlled for the very factors that he claims create bias, nor that my model produces extraordinarily accurate forecasts of the actual data, using multiple methods.

[^22]C. Trende criticizes my baseline measure of partisanship for not taking into account factors such as incumbency, candidate quality, and spending. This is an inaccurate criticism, because estimating baseline partisanship is designed to control for incumbency, campaign spending, and candidate quality. This is the method preferred in the academic literature on redistricting, which seeks to understand the consequences of hypothetical plans (in which candidate quality, spending, and incumbency are unknown). My approach is identical to the method used by Professor Gaddie, who produced the baseline partisan estimates used by Wisconsin's map drawers in 2011.
D. Goedert challenges my model for estimating baseline partisanship in 2012, contending that I took into account information that the authors of Act 43 did not have (the 2012 election results). However, my baseline estimates of partisanship are nearly identical to those generated by Gaddie in 2011, indicating the same conclusions follow whether 2012 or pre-2012 data are used in the analysis. In addition, pre-2012 election results are highly correlated with 2012 election results, indicating that it would make no difference if I had used earlier election results. Goedert dismisses the convergence between my estimates and Gaddie's estimates as "mostly coincidental," but offers no evidence or data to support his assertion.
E. Geodert also challenges my efficiency gap calculations for ignoring the effects of incumbency, which he asserts that any author of a redistricting plan would incorporate. His criticism fails to acknowledge that controlling for incumbency is the standard methodology for estimating the partisan consequences of a hypothetical district plan. Nevertheless, I recalculated efficiency gap estimates for both Act 43 and my Demonstration Plan, taking incumbency into account. The substantive conclusions are identical: the efficiency gap for my plan increases slightly (but is still well within acceptable limits), as does the efficiency gap for Act 43. The difference between the two plans' efficiency gaps remains enormous.
F. Goedert criticizes my efficiency gap calculations for not including any sensitivity testing to determine whether my results are robust to changes in the statewide electoral environment. I conducted a uniform swing analysis over the range of plausible election results, based on the maximum and minimum statewide Democratic Assembly vote since 1992. This analysis shows that the efficiency gaps of both Act 43 and the Demonstration Plan are robust: Act 43's efficiency gap remains very high across this range, always significantly above the plaintiffs' suggested 7\% threshold, and the Demonstration Plan's efficiency gap remains very low, and is always well below the threshold. Goedert is simply incorrect in asserting that the plans' respective efficiency gaps are not robust, and, again, offers no data or evidence to support his claim.
G. Throughout their reports, neither Trende nor Goedert has actually done any analysis that identifies problems with my analysis, or that specifically shows where my analysis is
incorrect. Trende and Goedert merely offer speculative and unsubstantiated criticism, but never offer any substantive data or evidence that supports their arguments. And, as I will show, when they attempt to analyze Wisconsin's political geography, their conclusions are utterly wrong.

## II. The Claim that Wisconsin's Political Geography Has a Pro Republican Bias

While I will go into more detail on the specific points each report makes, I focus first on a central argument both Trende and Goedert make: that Wisconsin has a natural distribution of Republicans and Democrats that produces an intrinsic pro-Republican bias in a neutrally-drawn redistricting plan. They claim that because Democrats in Wisconsin happen to be (allegedly) naturally concentrated in small pockets of overwhelming Democratic strength, even a neutrallydrawn map would produce a large pro-Republican efficiency gap. As a result, they conclude, it is not possible to consider a large pro-Republican efficiency gap as evidence of gerrymandering.

I begin by noting that both Trende and Goedert ignore the role that political geography already plays in plaintiffs' proposed test. Under the test's first prong, if the state's motive in enacting its plan was simply to follow the contours of the state's geography, then partisan intent would not be present and plaintiffs would proceed no further in their claim. Similarly, under the test's third prong, if the state can show that its plan's large efficiency gap was necessitated by the geographic distribution of the state's voters, then the plan would be upheld. These points mean that geography is already properly incorporated into plaintiffs' proposal.

There are, additionally, two points that fundamentally negate the utility of this line of attack. First, the geographic concentration argument is predicated on the foundational assumption that a neutrally-drawn map would have produced a pro-Republican bias. Even if Trende and Goedert are correct in this assumption (which they are not), they take no position on whether the process in Wisconsin was, in fact, neutral. The record of the federal redistricting trial clearly shows that Act 43 was designed with the predominant purpose of benefiting Republicans and disadvantaging Democrats, and neither Trende nor Goedert contradicts the findings in my report of examples of blatant packing and cracking that are the very DNA of a partisan gerrymander.

And second, even if the state's experts are correct that political geography has produced the pro-Republican bias in Wisconsin's state legislative district plan (which they are not), it is impossible for them to quantify how much of an effect geography has had: is it $5 \%$ ? $10 \%$ ? $90 \%$ ? $100 \%$ ? Neither Trende nor Goedert have actually done any analysis that demonstrates that the alleged concentration of Democrats in Wisconsin will produce a pro-Republican efficiency gap, or any work that quantifies how concentration is related to efficiency gap calculations. They simply assert (incorrectly) that Democrats are more concentrated than Republicans, and therefore that even a neutral map will produce a pro-Republican bias.

But they are also wrong on the facts. Their argument about geographic concentration is based on flawed data and measures, and has no basis in accepted methods of measuring geographic concentration and isolation. Trende, in particular, uses an unorthodox method with
no support in the peer-reviewed literature, and one that is guaranteed to produce a biased result that shows Democrats far more concentrated than they actually are. Goedert's argument contradicts his own published work, which shows that partisan control of redistricting generates a substantial bias even after partisan concentration is taken into account. His argument, further, falls victim to the Modified Areal Unit Problem, in that it is based entirely on the analysis of wards, ignoring the fact that wards are aggregated into districts. As I demonstrate, this aggregation process completely changes the applicability of Goedert's conclusions.

When I analyze the geographic distribution of Wisconsin's Democrats and Republicans using widely accepted measures of spatial concentration and isolation (Global Moran's I and the Isolation Index), I find that there is very little evidence of significant disparities in how the parties' voters have been distributed in recent election cycles. Republicans are in fact more concentrated than Democrats when measured by the 2012 Assembly vote.

## A. Trende

Trende spends nearly half of his report (paragraphs 62-105) arguing that Democrats are naturally more concentrated ("clustered") than Republicans in Wisconsin, which creates a natural packing effect. Much of this discussion is entirely irrelevant to Wisconsin (Trende's discussion of patterns in the southern United States, Virginia, and differences between the 1996 and 2008 Democratic coalitions; see paragraphs 62-77). Trende also simply asserts that "there is little doubt that the Democratic vote in Wisconsin is also increasingly concentrated in fewer counties" (paragraph 71). He neither explains the relevance of the county vote to the issue of geographic distribution and legislative redistricting, nor why the county vote pattern in 1988 or 1996 is germane to the environment in 2012.

## 1. The PVI (partisan vote index) is the wrong quantity of interest

As applied to Wisconsin, Trende attempts to demonstrate that over the last 20 years Democrats have become more concentrated. His method relies on a quantity he calls the Partisan Lean Index, which is the party's county or ward vote share minus the party's statewide vote share, and appears to be analogous to the Cook PVI, which is the same quantity calculated using the congressional district vote and the national presidential vote. Trende argues that Democratic wards are closer together than Republican wards, which to him is evidence of geographic clustering that produces a natural pro-Republican redistricting bias.

The PVI (which is how Trende abbreviates the measure) is a quantity that is not commonly used in the academic literature, and when it is, it is used largely as a simple descriptive statistic. What this index does is simply redistribute the ward vote around the statewide average, and thus tells us which areas are more Democratic (or Republican) than the state as a whole, and which areas are less so. ${ }^{2}$ It tells us little about overall partisan strength, and

[^23]is useful only in comparing elections at one level (here, counties or wards) to elections at another (the state).

The PVI is used almost exclusively by political commentators to describe congressional districts (the most widely known is the Cook PVI, which compares the average congressional district vote split over two consecutive elections to the average national presidential vote over those same elections). It is used less frequently in academic research, and then largely as a basic descriptive statistic used to classify districts as competitive or not. It is not used in the context of state legislative redistricting (Trende did not cite any studies that support the use of his measure, and could not identify any in his deposition).

Moreover, Trende appears to have made two errors in his calculation of the PVI. ${ }^{3}$ First, while he states that his PVI is based on the top-of-the-ticket race in each year, he uses the gubernatorial elections as his top-of-the-ticket race in 2002, 2010, and 2014, but the U.S. Senate race in 2006, even though there was a gubernatorial race that year. While scholars may differ on whether a gubernatorial or U.S. Senate election is the correct top-ticket race, there is no justification whatsoever for being inconsistent. ${ }^{4}$

Second, in calculating his 2014 PVI, Trende mistakenly subtracted the 2014 statewide percentages from the 2012 ward totals (this is the code he used to generate the PVI for 2014; the error is highlighted, and "map_2012\$r_share" is the ward vote for 2012):

```
map_2014=readOGR("Wards_Final_Geo_111312_2014_ED.shp",
"Wards_Final_Geo_111312_2014_ED")
map_2014=spTransform(map_2014, CRS("+proj=longlat +datum=WGS84"))
map_2014$r_share=map_2014$GOVREP14/(map_2014$GOVREP14 + map_2014$GOVDEM14)
map_2014$pvi=map_2012$r_share -
sum(map_2014$GOVREP14)/(sum(map_2014$GOVREP14) + sum(map_2014$GOVDEM14))
map_2014$pvi[which(is.nan(map_2014$pvi))]=0
```

Instead of the PVI, the actual ward level vote (or party vote share) is a much more direct measure of ward partisanship. I used LTSB ward level data from 2002 to 2014 to calculate the average Democratic percentage of the vote in a Democratic ward (all wards that were more than $50 \%$ Democratic in the top-ticket race), and the average Republican vote in wards where Republicans won more than $50 \%$ of the top-ticket vote. A graph of this data shows a very different pattern from what Trende claims (Republicans are in red; Democrats in blue):

[^24]

Here, we see that Democrats and Republicans have moved in almost identical fashion between 2002 and 2014. In 2002, Democrat wards were about $60.8 \%$ Democratic, and Republican wards were about $60.5 \%$ Republican in the top-ticket races. In 2014, similarly, both Democratic and Republican wards became more partisan: Democratic wards were $63.3 \%$ Democratic, and Republican wards 63.6\% Republican.

Trende's claim that Democratic wards have become more Democratic, while Republican wards have not become more Republican (paragraphs 91-95), is simply false.

Trende offers no justification or support for why he is relying on the PVI measure rather than more direct indicators of ward partisanship; he merely asserts that it is a relevant quantity. Given that there are far more widely used and relevant measures of district level partisanship, his reliance on it in this context is unsupportable.

## 2. Trende's "Nearest Neighbor" Method is Inappropriate and Inaccurate

After introducing the PVI, Trende attempts to use it to demonstrate that Democrats have become more closely packed than Republicans (which, he asserts, produces a natural proRepublican gerrymander). Apart from the irrelevance of the PVI, Trende's analysis uses a fundamentally flawed measure that is guaranteed to exaggerate the extent of Democratic concentrations. Instead of his measure, widely used and academically accepted metrics of concentration and isolation show that Democrats and Republicans are both highly segregated, and to about the same extent. Just as there are core areas of high Democratic strength in Milwaukee and Madison, there are similar Republican core areas in the "collar counties" of Waukesha, Ozaukee, and Washington.

The premise of Trende's argument is that pro-Democratic wards are closer to other proDemocratic wards than are pro-Republican wards to other pro-Republican wards. His method, which I infer from his description, is to identify a pro-Democratic or pro-Republican ward of a certain percentage lean, and then to find the distance to the nearest ward with the same partisan lean. He determines the median distance between similar wards, and presents two graphs (about paragraph 98 in his report) showing that the median distance between similar Democratic wards is smaller than for Republican wards, and that as Democratic wards become more Democratic, they become closer to one another.

This is reminiscent of the nearest neighbor method used in the study of populations, but it bears little resemblance to how the concept is actually used in the literature, even in its earliest form (Clark and Evans (1954) used it to study the distribution of plant and animal populations). ${ }^{5}$ His application of this method is highly unorthodox, unsuited to the study of redistricting, and not based on any accepted peer-reviewed academic work (he does not cite a single study in support of his method).

Trende's method is to start with a ward (call it $i$ ), calculate its PVI and assign it to a quantile, and then locate the closest ward that shares this PVI quantile (call it $j$ ). The geographic distance between wards $i$ and $j$ (presumably calculated using the ward centroids, although Trende fails to specify this key detail) is then recorded (paragraph 97). The process is repeated for every ward over every election from 2002 to 2014 , producing for each election a matrix consisting of every ward and the distance to the nearest ward with the same PVI quantile. He then calculates median distances between wards of the same PVI quantiles, which he claims shows that Democratic wards are, and have been continuing to move, closer together than Republican wards.

There are several problems with this approach. First, and most fundamentally, the proximity of similar wards is simply not a measure of geographic concentration or clustering. Trende's method tells us nothing about which wards are actually adjacent to wards of a certain PVI. It only tells us how far these wards tend to be from other wards of the same partisan lean. It is entirely possible for wards of the same partisan makeup to be far apart but still easy to join in the same district (think of a sparsely populated but uniformly partisan area). Likewise, it is entirely possible that wards of the same partisan makeup are close together but quite difficult to combine in the same district (think of a densely populated but politically heterogeneous area). Trende's method cannot distinguish between these scenarios, and as a result it cannot tell us anything about the geographic patterns that actually matter for redistricting.

Second, Trende does not explicitly define in his report what a "similar partisan index" (paragraph 97) means. Clearly, Trende is classifying them in some way, defining "similar" as within some range, as his vague discussion of quantiles indicates (paragraph 98). But without specifying the range, it is impossible to know whether his measure has any meaning. Different

[^25]classification methods -- requiring a match of, say, within 0.1 percentage points, or classifying according to deciles or some other method -- are likely to yield very different results than requiring a match of within 0.5 or 1.0 percentage points or using a larger number of categories. His graphs suggest he is using some type of percentile distribution (the x axis label refers to "(.05\% is the most Democratic [or Republican] Ward)," but he does not explicitly define why he chose this particular scheme or how he calculated the quantiles. On this point alone, his method lacks validity or replicability.

But there are two additional serious - fatal, in fact - flaws in this method. First, in treating the geographic distances between wards as his quantity of interest, Trende does not take into account the fact that wards in Wisconsin are not uniform in area. Ward areas actually vary widely: some are very small, others are moderate in size, and still others are very large (wards are drawn within specified population limits, but their geographic areas are not similarly constrained).

Table A shows the mean and median areas (in square miles) of Wisconsin wards. The average is $8.41 \mathrm{mi}^{2}$, but the range is huge: the smallest ward with a nontrivial population is in the City of Middleton: ward 19, with 690 people in an area of $0.0071 \mathrm{mi}^{2}$. The largest ward in the state is in the Town of Winter: ward 2 (in Sawyer County), with 565 people in an area of 227.7 $\mathrm{mi}^{2}$.

Geographic distances between ward centroids will, obviously, depend on how large the wards are. Although centroid-to-centroid distances will not map perfectly onto area differences (because the distances will vary with the shape and orientation of wards), two large wards - even if they are adjacent - will show up as much farther apart than two smaller wards that might be separated by numerous other wards and municipal boundaries.

The problem is magnified when we observe that ward sizes are correlated with other relevant variables, particularly whether a ward is in a city, and most crucially, whether it is a Democratic or Republican ward:

| Table A 2012 Ward Sizes (square miles) ${ }^{6}$ |  |  |
| :---: | :---: | :---: |
|  | Mean | Median |
| Statewide Average | 8.41 | 1.12 |
| City of Milwaukee | 0.29 | 0.20 |
| Rest of State | 8.83 | 1.27 |
| Democratic Wards | 5.91 | 0.56 |
| Republican Wards | 10.96 | 3.45 |

Wards in the city of Milwaukee have a mean area of only $0.29 \mathrm{mi}^{2}$, which is $3 \%$ of the size of the mean area statewide. Democratic wards (measured by whether the 2012 Democratic presidential vote was above $50 \%$ ) are, on average, only about half the size of Republican wards ( $5.91 \mathrm{mi}^{2}$ vs. $10.96 \mathrm{mi}^{2}$ ).

In relying on the distance between wards, Trende is thus putting his thumb on the scale; all other things equal, this method will always show Democratic wards to be much closer than Republican wards, irrespective of whether this concentration is real or merely an artifact of ward area. To put it most simply, smaller Democratic wards will always appear closer than larger Republican wards.

But a second and equally serious problem lurks. Trende does not use the mean distance between wards as his quantity of interest, but rather the median. He justifies this choice "because outlying wards, such as Menominee County, exert an undue amount of leverage on averages" (paragraph 97).

This is the wrong measure, because the "nearest neighbor" approach is unlikely to pair, say, a ward in Milwaukee with a ward in northwest Wisconsin. Menominee County will not exercise "an undue amount of leverage" because it is an outlying ward. It will exercise an undue amount of leverage because it has a very large area $\left(222.8 \mathrm{mi}^{2}\right)$, which is something Trende should, but does not, correct for.

His use of the median rather than the mean further exaggerates the difference between Republican ward distances and Democratic ward distances. The average Republican ward area is 1.9 times larger than the average Democratic ward area ( 10.96 vs. $5.91 \mathrm{mi}^{2}$ ). But the median Republican ward is 6.2 times larger than the median Democratic ward ( $3.45 \mathrm{mi}^{2}$ vs. $0.56 \mathrm{mi}^{2}$ ).

[^26]Because the disparity is three times larger for the median versus the mean area, Trende is further stacking the deck in favor of his preferred hypothesis.

I was able to replicate Trende's analysis, using LTSB data and the R code he disclosed. When the mean distances between similar wards are included, Figure B is the result for the 2012 Election: ${ }^{7}$

Figure B
Nearest Neighbor Analysis: Mean vs. Median Distance


In this graph, the dotted lines are the median nearest neighbor distances for Democratic (blue) and Republican (red) wards, replicating what Trende did in his median distance graphs around paragraph 98 in his report. Wards become more partisan as we move from right to left.

The mean distances are shown with solid lines. While Republican wards remain farther apart than Democratic wards, the mean distances for both parties are much larger than the median distances. Proportionally, Republican and Democratic wards are much closer together in mean than in median distances (which is what one would expect, given the exaggerated difference between median Democratic and Republican ward sizes). Specifically, the mean distance between Republican wards is only about 70\% larger than the mean distance between Democratic wards, compared to a $180 \%$ difference between the median Republican and Democratic distance.

[^27]More relevant is the shape of the mean distance lines. They show that Republican and Democratic distances move precisely in parallel, and that strongly Democratic wards are significantly farther apart than weaker Democratic wards (as are strongly Republican wards). This is the complete opposite of Trende's claim that stronger Democratic wards are closer together than weaker Democratic wards, and it obliterates the core of Trende's report: the assertion that the pro-Republican bias evident in Act 43 is the natural result of Democrats being more geographically concentrated.

To conclude, Trende's argument about Democratic concentration is based on an irrelevant measure of partisanship (PVI) that is incorrectly calculated, applies a methodology that bears no relationship to any scholarship or actual research on spatial distribution, ignores a key feature of Wisconsin's actual political geography (ward area), relies on an improper distance measure that is enormously biased in favor of his hypothesis, and produces a result that fundamentally misrepresents what the data actually shows. Because of his use of a questionable method and fundamentally flawed measures, Trende's opinions should be regarded as uninformative.

## B. Goedert

Goedert, like Trende, asserts that Wisconsin's natural geography creates an intrinsic proRepublican bias in redistricting (p. 17). He cites his own research that geography produced a pro-Republican bias in the 2012 congressional election (p. 19).

The only analysis Goedert conducts as to Wisconsin is an examination of wards, which he claims shows "the bias inherent in Wisconsin's geography" (p. 21). His analysis is a simple "uniform swing" study of wards in 2012, adjusting the Democratic presidential vote in each ward downward by $3.5 \%$ to determine the overall ward distribution in the event of a tied election (Figure 1, p. 22). He asserts that based on this analysis, "Republicans would win $60.2 \%$ of wards, comprising $54.4 \%$ of the voting population" in a tied election (p.22). This is the extent of his analysis.

This analysis, however, is a non sequitur, because it fails to aggregate wards to the relevant geographic level, which is districts. Goedert's failure to take this into account is an example of the Modified Areal Unit Problem, in which inferences at one level of geography frequently do not hold at other levels of aggregation; see King (1996). In this example, the ward level vote is far less relevant than the district level vote, because it is entirely possible that wards will be aggregated in such a way that the pattern he observes either disappears (or even reverses).

When we examine the distribution of districts, which have a population deviation small enough that we can consider them equal (the deviation under Act 43 is $0.76 \%$ ), we in fact see almost the reverse pattern. The following graph (Figure C) displays Goedert's adjusted ward level presidential vote in a simulated 50-50 election, along with an adjusted baseline forecast for Act 43 districts, using my baseline open seat model, in a simulated tied election. Both wards and districts are weighted based on the number of votes cast in each unit. This allows me to directly compare ward level results to district level results:

Figure C: Distribution of Wards vs. Districts


What this figure demonstrates is that as wards are aggregated into districts, the distribution substantially changes. The red line is a kernel density plot of the ward Democratic vote percentage in a simulated tied election; it is a continuous version of the histogram Goedert presents in his Figure 1. The dotted blue line shows the predicted Democratic vote in Act 43 districts in a simulated tied election - or, what occurs after the wards are aggregated into Assembly districts. The overall shape of the curves, the mode of each distribution, and even the mean vote percentage vary as we aggregate from wards up to districts. Knowing the ward distribution ultimately does not tell us much about what the distribution of districts will look like; the process of aggregation is crucial.

More significantly, the district distribution is much more tilted in a Republican direction than is the ward distribution. The ward distribution is nearly normal in shape, and has a peak very close to $50 \%$ Democratic. In contrast, the district distribution is skewed to the right, and has a much higher peak around $42 \%$ Democratic, meaning that there are many more districts that Republicans win by relatively small margins (indicating that Democrats are cracked), and many more districts where Democrats win by much larger margins (indicating packing). Accordingly, the district distribution does not mirror the underlying distribution of wards. Rather, it reveals that Act 43's designers were able to distort a fairly neutral ward distribution into a far more advantageous district distribution, through gerrymandering.

## 1. Goedert's Published Work Contradicts His Report

Goedert's own prior work indicates that unified party control of state government has an independent and significant effect on the bias of redistricting plans, even after controlling for
population concentration. This work also indicates that if Wisconsin, or a state resembling the country as a whole, had a court-drawn or bipartisan map in 2012, this map would have had a slight pro-Democratic bias. These findings further obliterate the claim that Act 43's extreme partisan tilt resulted from Wisconsin's natural political geography.

In a 2014 article, Goedert analyzes the consequences of different redistricting processes, looking for evidence that partisanship and geography each have an independent effect on the partisan bias of redistricting plans. ${ }^{8}$ Using an unorthodox definition of gerrymandering Goedert defines any redistricting plan created in a state with unified party control of state government as a partisan gerrymander - he finds that in states with more than six congressional districts, both urbanization (a proxy for Democratic concentration) and unified party control have a strong and statistically significant effect on the bias of a district plan $(2014,6)$. Goedert interprets his results as indicating that geography matters, and that higher urban concentration leads to more bias against Democrats $(2014,6)$. But what his results also show is that even after taking urbanization into account, the partisanship of the map drawers introduces a separate and significant bias: Republican-drawn maps are associated with an additional $13.6 \%$ proRepublican bias.

Geodert updated his 2014 article in a more recent manuscript, which incorporated the results of the 2014 midterm elections. Here, he finds that urbanization no longer has a statistically significant effect on the bias of district plans $(2015,6)$. Yet he stills finds evidence that the partisanship of map-drawers has a significant effect on district plans' bias (in 2014, a Republican-drawn plan adds $12.4 \%$ bias, or roughly the same as the $13.6 \%$ estimate for 2012).

So, on the one hand, Goedert's own work comes to different conclusions about the impact of urbanization (or Democratic concentration): sometimes it matters, other times it does not. But his work is consistent about the effect of partisan control: when partisans draw maps, they always do so in ways that dramatically bias plans in their favor. The clear inference is that geography matters much less than partisan control in explaining plans' electoral consequences.

Furthermore, we can use Goedert's regression model to generate a forecast of what would have occurred in 2012 in Wisconsin - as well as in a state resembling the country as a whole under a neutral process (i.e., a court-drawn or bipartisan plan). His regression model includes the following variables $(2015,11)$ :

1. Whether a district plan was drawn by Democrats or Republicans (court-drawn and bipartisan plans are the excluded category)
2. A state's African American population percentage
3. A state's Hispanic population percentage

[^28]4. The percentage of a state that is urbanized (according to the Census)
5. The statewide Democratic vote
6. The number of congressional seats.

With the coefficients of this model, and the appropriate data for Wisconsin (or any other state), we can calculate what the expected bias would be for a plan in $2012 .{ }^{9}$ The dependent variable here is a measure of bias almost identical to the efficiency gap, with positive values indicating a pro-Democratic bias, and negative values a pro-Republican bias. Because this is a linear regression, we can multiply each coefficient by the value of the independent variable, and then sum the results to generate a forecast from any set of data values. In Table B, I set both Democratic and Republic Gerrymanders to 0, simulating a neutrally-drawn plan:

[^29]| Table B <br> Goedert's Regression Model for 2012 <br> Dependent Variable: <br> Pro-Democratic Bias in a District Plan |  |  |  |
| :---: | :---: | :---: | :---: |
| Variable <br> Name | (a) <br> Coefficient Value | (b) <br> Variable value for Wisconsin | Value $\text { (a) } \times(b)$ |
| Democratic Gerrymander | 16.6 | 0 | 0 |
| Republican Gerrymander | -13.6 | 0 | 0 |
| \% Black | -0..29 | 6.6 | -1.914 |
| \% Hispanic | 0.77 | 6.5 | 5.005 |
| \% Urbanized | -0.72 | 70.2 | -50.544 |
| Statewide <br> Democratic Congressional Vote | 0.11 | $\begin{gathered} 50.8 \\ (2012) \end{gathered}$ | 5.588 |
| Number of Seats | -0.16 | 8 | -1.28 |
| Constant | 45.0 | 1 | 45 |
| Total | (sum of | values) | 1.855 |

Goedert's regression model thus predicts that if Wisconsin had a neutrally drawn plan in 2012, the resulting map would have had a pro-Democratic bias of $1.855 \%$. In other words, in the absence of unified Republican control over the redistricting process, Wisconsin's demographic, geographic, and political characteristics would have resulted in a small natural Democratic advantage. And this is no fluke of the state or the election year. We can also use Goedert's model to predict what would happen in a state resembling the United States as a whole (i.e., a state that is $13.2 \%$ black, $17.4 \%$ Hispanic, $80.7 \%$ urbanized, $51 \%$ Democratic, and with
8.7 congressional seats ${ }^{10}$ ). Substituting these values into the regression model shows that in an "average" state, a neutrally-drawn map would have had a pro-Democratic bias of $0.684 \%$ in 2012.

Goedert's 2014 variant of the model $(2015,13)$ further predicts that Wisconsin would have had a pro-Democratic bias of $4.392 \%$ in 2014, and that the average state would have had a pro-Democratic bias of $1.589 \%$. At this point, it is hard to see what is left of the thesis that political geography inherently favors Republicans. If anything, Goedert's own published analysis shows that Wisconsin's political geography slightly favors Democrats.

## C. Accepted Measures of Geographic Concentration and Isolation Show that Democrats and Republicans are Equally Dispersed

In arguing that Republicans in Wisconsin enjoy a natural geographic advantage, both Trende and Geodert use ad hoc, unorthodox measures of concentration that are neither relevant nor accepted by the academic literature. In fact, there exist widely accepted metrics of geographic concentration and dispersion, used by geographers and demographers to study spatial patterns. Two of the most common are Global Moran's I (Anseln 1995; Cho 2003), and the Isolation Index (Glaeser and Vigdor 2012; Reardon 2004). I use these metrics to determine how Democrats and Republicans in Wisconsin are actually distributed.

Moran's I is a measure of spatial autocorrelation, or how values of a variable in space correlate with values in nearby space. It can be calculated for an entire geographic system (Global Moran's I), or for any specific point in space (Local Moran's I). The Isolation Index indicates, for the average member of a group residing in a certain geographic unit (such as a ward), what share of the member's neighbors in the unit belong to the same group (Iceland and Weinberg 2002, 120). It measures how geographically isolated a group is (Reardon 2004, 153), and it can easily be adjusted, by deducting a group's share of the statewide population, to show how much more isolated a group is than we would expect given its statewide size (Glaeser and Vigdor 2012, 2). Both Moran's I and the Isolation Index are widely used in studies of residential segregation and sorting (Chung and Brown 2007; Massey and Denton 1989; Glaeser and Vigdor 2012; Dawkins 2007; Reardon 2004; Iceland and Weinberg 2002), epidemiology (Moore and Carpenter 1999), network effects (Cho 2003), and political geography (Glaeser and Ward 2005). The measures are also used by the U.S. Census Bureau itself (Iceland and Weinberg 2002).

Both Moran's I and the Isolation Index are directly applicable to the issue of measuring the geographic distribution of Democrats and Republicans in Wisconsin. In this context, Global Moran's I tells us how likely Democrats are to live clustered next to other Democrats (and Republicans to Republicans), and the Isolation Index, adjusted as noted above, tells us to what extent the average Democrat (or Republican) lives in a ward that is more heavily Democratic (or Republican) than the state as a whole. I use these indices to directly assess the geographic distribution of Democrats, and, more importantly, to compare it to the geographic distribution of Republicans.

[^30]Global Moran's I is analogous to a correlation coefficient, and ranges from -1 to 1 ; scores close to 1 indicate a very high spatial correlation (i.e., clustering) of Democrats (or Republicans). The Isolation Index ranges from 0 to 1 , and, adjusted as noted above, indicates to what extent the average Democrat or Republican lives in a ward that is more heavily Democratic or Republican than Wisconsin as a whole. In calculating both measures, I use the ward as the basic unit of geography and actual Assembly votes. ${ }^{11}$ Because I only have geodata for the current wards, I only estimate Global Moran's I for 2012 and 2014. For the Isolation Index, I compute scores dating back to 2004. Both Global Moran's I and the Isolation Index are asymmetrical, and so must be calculated separately for Democrats and Republicans.

Table C shows the values of the Isolation Index, adjusted as noted above, for Democrats and Republicans in Wisconsin from 2004 to 2014:

|  | Table C <br> Isolation Index |  |
| :---: | :---: | :---: |
|  | Dem- <br> Rep | Rep- <br> Dem |
| 2014 | 0.23 | 0.20 |
| 2012 | 0.14 | 0.12 |
| 2010 | 0.15 | 0.17 |
| 2008 | 0.15 | 0.14 |
| 2006 | 0.16 | 0.17 |
| 2004 | 0.20 | 0.21 |

As is evident from Table C, Democrats were slightly less isolated than Republicans in 2004, 2006, and 2010, and slightly more so in 2008, 2012, and 2014. In all cases, the differences in isolation were very small, amounting to only one to three percentage points (out of a scale extending from $0 \%$ to $100 \%$ ). In the 2012 election, for instance, the average Democrat lived in a ward whose Democratic vote share was $14 \%$ more Democratic than the state as a whole; analogously, the average Republican lived in a ward whose Republican vote share was $12 \%$ more Republican than the entire state. In the previous election, it was Republican voters who were more isolated than Democratic voters ( $17 \%$ versus $15 \%$ ). This analysis in no way supports the claim that Republicans are more advantageously distributed than Democrats; on the contrary, both parties' supporters are almost identical in their geographic isolation over the last decade, and there is no clear temporal pattern. In some years, Democrats are marginally more isolated than Republicans, and in other years Republicans are marginally more isolated than Democrats.

[^31]The results are very similar with the Global Moran's I, again calculated for Democrats and Republicans in Wisconsin, although only for the two elections (2012 and 2014) for which the geodata is readily available:

|  | Table D <br> Global Moran's I |  |
| :---: | :---: | :---: |
|  | Democrats | Republicans |
| 2014 | 0.75 | 0.68 |
| 2012 | 0.68 | 0.69 |

Here, we see that Democrats were slightly less spatially concentrated than Republicans in 2012, but slightly more spatially concentrated in 2014. The differences in both cases are tiny: 0.01 in 2012 and 0.07 in 2014, on a scale that stretches from -1 to 1 . The message is quite clear: both Democrats and Republicans in Wisconsin tend to live near one another in distinct clusters, but there is no evidence that Democrats are more geographically clustered than Republicans.

Accordingly, two widely used and accepted measures of geographic distribution show no consistent pattern, and no material difference in how Wisconsin's Democrats and Republicans are dispersed spatially. In no sense, therefore, is it an accurate statement that Democrats are much more concentrated than Republicans - the unsubstantiated claim that comprised the core of both Trende's and Geodert's arguments about natural gerrymanders.

## III. Trende's Claim That My Vote Model Is Biased Is Incorrect

Trende claims that there may be "a systematic bias involved in imputing presidential results to state House results" (paragraph 135). As evidence he points to Figures 2 and 3 in my original report, which display the relationship between the ward level presidential vote and the ward level Assembly vote. Trende notes that Figure 2 shows that there is close to a $1: 1$ relationship between Republican presidential and Assembly votes, as the dots on the graph are distributed around the 45-degree line:


However, Trende claims that the relationship is different for Democratic votes (Figure 3 in my original report):


Here, Trende argues, the "dots systematically fall below the line, often creating differences on the order of 10 percent" (paragraph 138). This pattern, he asserts, will "skew the imputation" of votes, resulting in "too many votes [being] imputed in wards reporting a high number of Democratic votes" (paragraph 139).

Trende is completely and unambiguously wrong in this claim, which belies a fundamental lack of understanding of multiple regression and the causes of bias in statistical models. Trende appears to believe that I simply assumed that ward level Democratic Assembly votes are actually equal to ward level Democratic presidential votes, or that in estimating the Assembly vote in uncontested wards I merely used the value of the presidential vote (presumably because that is how he imputes the vote in uncontested districts in his own analysis; deposition page 83).

That is wrong. I displayed this graph merely to show that there is in fact a strong relationship between the two variables. The fact that the Democratic Assembly vote tends to fall below the presidential vote is completely irrelevant to any possible bias. In fact, regression analysis estimates the relationship between the two quantities by identifying the slope of the line that relates them, not how the relationship varies across a 45-degree line.

Below (Figure D) is a graph that plots the data in Figure 3 of my original report along with a fitted line of predicted values from a bivariate regression of the Democratic Assembly vote on the Democratic presidential vote. The red line consists of the predicted values of the Democratic Assembly vote in each ward:


Here, we see that the fitted line runs exactly down the middle of the plotted points. My regression analysis of the Democratic Assembly vote (Table 1 in my original report) shows that the coefficient for the Democratic presidential vote is 0.931 ( $\mathrm{p}<0.0001$ ), which is precisely the pattern than we see in the bivariate relationship above. In a linear model, this coefficient is the
slope of the line that relates the presidential vote to the assembly vote. It is less than 1 (a 45degree line), indicating that the Assembly vote rises more slowly than the presidential vote; i.e., the predicted Assembly vote will lie below the 45-degree line in Figure 2.

And, as is immediately apparent from the actual results of my regression (Figure 4 in my original report, which plots the actual vs. predicted ward level votes), there is no bias in the results. In this graph, the 45 -degree line is where the predicted Assembly vote would fall if it were exactly equal to the actual Assembly vote:

Figure 4: Ward Level Predicted vs. Actual Assembly Vote - 2012


Trende's criticism on this point is utterly misinformed. No one with a solid understanding of quantitative methods or regression analysis would have made it.

## IV. Trende's Claim That My Efficiency Gap Calculations Ignore Incumbency, Candidate Quality, and Campaign Spending

In paragraphs 140-143, Trende criticizes my efficiency gap calculations for failing to take into account factors that can affect election results, such as get-out-the vote drives, candidate quality, recruitment, and campaign spending.

Trende offers no evidence that these factors would actually have a material effect on my estimates if I had more directly taken them into account. And he ignores the fact that any
estimation of the results of a hypothetical district plan utilizes baseline estimates that, in effect, average out the effects of these factors (Gelman and King 1990; 1994). That is to say, my regression model does implicitly incorporate these factors, in its analysis of the relationship between the presidential vote (where none of these variables will affect the vote) and the Assembly vote (where they are all incorporated into the estimates).

Moreover, Trende's criticism overlooks the point that my model is based on precisely the same information that the authors of Act 43 considered in estimating the likely partisan effects of the new districts. In particular, Gaddie's analysis of the partisan effects in the new Act 43 districts was functionally equivalent to mine and based on exactly the same considerations.

Like his complaints about alleged bias in the regression analysis that I discuss above, Trende's criticism is uninformed and betrays a lack of knowledge of how hypothetical district plans are evaluated.

## V. Goedert's Claim That My Efficiency Gap Calculations Incorporate Information Not Available to Act 43's Designers, and Ignore the Effects of Incumbency

Goedert criticizes my analysis for incorporating information that map drawers did not have (2012 election results), and for ignoring information that map drawers would have taken into account (incumbency in particular).

The first criticism is incorrect, as Act 43 's designers in fact had information functionally equivalent to the 2012 election results in their possession, in the form of Gaddie's Act 43 district level estimates. These estimates, like my own, are baseline measures of partisanship, and they correlate almost perfectly with my results $\left(\mathrm{r}^{2}=0.96\right)$. In his deposition, Gaddie described in detail his method, which like mine assumed that all seats would be contested and that no incumbents would run (Gaddie Deposition, pp. 197, 198, 201, 202, 204):

Let's suppose we have a seat with an incumbent and a seat without an incumbent and each one has an Assembly election. The party of the incumbent is presumably going to do a little stronger in the district where they have an incumbent than in an open seat. So I can't really take -- Let's suppose I move precincts from the open seat into that incumbent seat. I can't really take those open seat Assembly votes, add them, compare them to the percentage for the incumbent running for the same party, get an accurate estimation of the partisanship and the competitiveness of the district. So we attempt to create a substitute measure. Statewide elections are held in all precincts, they're held in all constituencies, so one thing that we often do is we do what we call reconstituted elections, or proxy elections, where we'll take one election or a composite of elections, like I described previously, and attempt to create some measure of partisan competitiveness, an expected vote or what we call a normal vote, what the vote would usually do without an incumbent in the district." (Gaddie Deposition, pp. 204-5)

To highlight the similarity between Gaddie's pre-2012 estimates and my own estimates using 2012 election results, below is a graph plotting the two sets of data (Figure 7 in my original report, p. 30):


This graph shows that the information the Act 43 authors relied on when drawing their map (the Gaddie estimates) and my estimates, are nearly identical. This is largely because they are both estimates of the same underlying quantity - the baseline partisanship of a hypothetical Assembly district. Goedert dismisses the nearly perfect correlation as "mostly coincidental" (p. 17), but offers no analysis or data to support this conclusion. It is simply an assertion offered without evidence.

And it is an entirely unpersuasive assertion for the additional reason that election results in Wisconsin (and in most states) are extremely highly correlated from one election to the next. For example, Wisconsin's counties remained geographically constant between 2008 and 2012, and Trende supplied information about the presidential vote in each county in each of these years. The 2008 county level presidential vote and the 2012 county level presidential vote are almost perfectly correlated $\left(r^{2}=0.96\right)$, indicating that it would make no difference whether Act 43 was assessed using the former or the latter. ${ }^{12}$ Either way, the same conclusion would follow: that

[^32]the map is an extreme Republican gerrymander, and that the authors of Act 43 had information in their possession that predicted it.

Second, Goedert claims that map drawers do not ignore incumbency when drawing maps. That will generally be true when map drawers are trying to figure out which incumbent should be included in which district. But when it comes to estimating the likely partisanship of the new districts, ignoring incumbency (that is, controlling for it) is precisely what the drawers of Act 43 did, as Gaddie noted in his description of his methods. This approach is sensible since incumbents can be defeated, retire, run for higher office, or switch parties over a plan's decadelong lifespan. A map's authors will typically want to ensure that their projections do not depend on particular incumbents continuing to run in particular districts.

In any event, including incumbency in no way changes my substantive conclusions about Act 43 or the Demonstration Plan. I recalculated the efficiency gap for both maps, using my baseline partisan estimate and then incorporating incumbency into the model. For Act 43, I used the actual incumbents who ran in the plan's districts, with the adjustments noted in my report to account for paired incumbents and those who lost in primaries (p. 18, footnote 14). ${ }^{13}$ For my plan, I geocoded incumbents' home addresses ${ }^{14}$ and then identified which districts had incumbents residing in them using Maptitude for Redistricting. Table E shows the resulting efficiency gap calculations, and compares them to the open seat baseline I generated in my report:

| Table E |  |
| :---: | :---: | :---: |
| Efficiency Gap Calculations |  |
| with Incumbents |  |

The efficiency gap increases marginally for both plans (by $1.5 \%$ for the Demonstration Plan and $1.4 \%$ for Act 43), in large part because there were more Republican (50) than

[^33]Democratic (24) incumbents running in 2012. With twice as many incumbents, Republicans will win more seats than in the open seat baseline even though the Republican vote percentage remains below $50 \%$ in both cases. It is thus apparent that taking incumbency into account has no effect on my conclusion that Act 43 was an egregious partisan gerrymander; the substantive inferences are identical, with or without incumbency. ${ }^{15}$

## VI. Goedert's Claim That I Did Not Perform Sensitivity Testing for Act 43's or the Demonstration Plan's Efficiency Gaps

Goedert criticizes the efficiency gap calculations for both Act 43 and the Demonstration Plan, arguing that I "provide no estimates for the efficiency gap of the demonstration plan under the range of plausible election outcomes facing legislators at the time they were drawing the map" (p. 16), and that I conduct no "sensitivity testing" of my calculations of Act 43's efficiency gap.

I note that Goedert has not provided any actual analysis showing that this sensitivity testing would have materially altered my conclusions, or even any citations showing that such testing is necessary to evaluate the adequacy of my calculations.

Still, it is possible to show that my calculations are robust to significant changes in the electoral environment. Using Jackman's historical estimates of the statewide Assembly vote in Wisconsin, I can determine the plausible variation of the overall vote over the course of a decade. Since 1992, the statewide Democratic percentage of the Assembly vote has ranged from a high of $54.6 \%$ (in 2006) to a low of $46.4 \%$ (in 2010). The Democratic share of the statewide vote in 2012 was $51.2 \%$ in my baseline calculations, which suggests a plausible range of $-5 \%$ to $+3 \%$ in conducting a sensitivity analysis. In effect, this approach asks whether Act 43 's and the Demonstration Plan's efficiency gaps would be durable in the face of massive Democratic or Republican waves - an extremely rigorous test that exceeds what is normally found in the literature.

Following Goedert's method of applying a uniform swing (p.21), I can estimate the effects that these swings will have on the efficiency gap, both for Act 43 and for the Demonstration Plan. To maintain consistency and to address his concern that I did not incorporate incumbency in my baseline, I estimate the effects using the incumbent baseline (that is, including the incumbents who ran in 2012).

[^34]The results are shown in the following two tables, the first for the Demonstration Plan (Table F), and the second for Act 43 (Table G). For the Demonstration Plan, the efficiency gap remains well below the plaintiffs' suggested 7\% threshold, even when the statewide vote reaches the most extreme values either party has seen over the last three decades. Specifically, the efficiency gap goes to $3.9 \%$ in the event of a Democratic wave akin to that of 2006, and to $-2.0 \%$ if a Republican wave like that of 2010 occurs. For Act 43, however, the efficiency gap remains extremely large and above the threshold at all times, ranging from $10.7 \%$ in a Democratic wave to $8.8 \%$ in a Republican wave. Moreover, the sensitivity testing shows that even if the Democrats obtained over $54 \%$ of the statewide Assembly vote - equal to their best performance in a generation - they still would not capture a majority of the Assembly, gaining only 48 seats. Act 43's gerrymandering thus effectively insulates the Republican Assembly majority from all plausible shifts in voter sentiment.

|  | Table F <br>  <br> Efficiency Gap Estimates, Uniform <br> Swing |  |  |
| :---: | :---: | :---: | :---: |
| Demonstration Plan |  |  |  |
| party split (R-D) | D Minus 5 | My Plan <br> Incumbent <br> Baseline | D Plus 3 |
| Rep share of <br> Seats | $52 \%$ | $48 \%$ | $43-56$ |
| Wasted <br> Republican Votes | 737,557 | 659,821 | 659,390 |
| Wasted <br> Democratic Votes | 681,900 | 765,561 | 769,546 |
| Gap | $(55,657)$ | 105,740 | 110,156 |
| Total Democratic <br> Votes | $1,336,168$ | $1,484,631$ | $1,573,709$ |
| Total Republican <br> Votes | $1,502,745$ | $1,366,132$ | $1,284,164$ |
| Total Votes | $2,838,913$ | $2,850,763$ | $2,857,873$ |
| Efficiency Gap <br> (gap/total votes) | $\mathbf{- 1 . 9 6 \%}$ | $\mathbf{3 . 7 1 \%}$ | $\mathbf{3 . 8 5 \%}$ |


|  | Table G <br> Efficiency <br>  <br> Gap Estimates, Uniform <br> Swing <br> Act 43 Districts |  |  |
| :---: | :---: | :---: | :---: |
|  | D Minus 5 | Act 43 <br> Actual | D Plus 3 |
| Party Split (R-D) | $64-35$ | $60-39$ | $51-48$ |
| Rep share of <br> Seats | $65 \%$ | $61 \%$ | $52 \%$ |
| Wasted <br> Republican <br> Votes | 585,668 | 504,553 | 560,840 |
| Wasted <br> Democratic <br> Votes | 835,968 | 876,153 | 866,725 |
| Gap | 250,300 | 371,600 | 305,885 |
| Total <br> Democratic <br> Votes | $1,316,158$ | $1,462,397$ | $1,550,141$ |
| Total <br> Republican <br> Votes | $1,527,115$ | $1,388,286$ | $1,304,989$ |
| Total Votes | $2,843,273$ | $2,850,684$ | $2,855,130$ |
| Efficiency Gap <br> (gap/total votes) | $\mathbf{8 . 8 0 \%}$ | $\mathbf{1 3 . 0 4 \%}$ | $\mathbf{1 0 . 7 1 \%}$ |

Figure E below shows these results graphically: the red x's are the efficiency gap estimates for the Demonstration Plan, and the blue diamonds the estimates for Act 43. The dotted line is at plaintiffs' suggested threshold of $7 \%$. The figure clearly demonstrates that even across huge partisan swings, the efficiency gap under Act 43 remains very large, and the efficiency gap for the Demonstration Plan remains very small. This is further powerful confirmation of the durability of Act 43's bias - and the durable lack of bias of the Demonstration Plan.

# Figure E: Sensitivity Analysis -5 to +3 Democratic Swing 

15\%-1

- 0
- +3
- -5
$\qquad$
$-$
+3 关
* -5
* Demonstration Plan
- Act 43 Districts


## VII. Conclusion

In their criticism of my report, both Trende and Goedert offer nothing but supposition, speculation, irrelevant discourse about Wisconsin political history, extraneous discussion of congressional redistricting in other parts of the United States, wildly inapposite and inaccurate conjecture about the geographic concentration of Democrats as a possible source of the proRepublican bias of Act 43, unreliable methodologies, and minor quibbles that have no consequences for my conclusions. Neither Trende nor Goedert has conducted any valid analysis of either Act 43 or the Demonstration Plan - in fact, they make no mention at all of the specifics of the Demonstration plan.

Most significantly, nothing in their reports undercuts my fundamental conclusion that Act 43 constituted an egregious and durable gerrymander, and that it was entirely possible to draw a neutral map that met or exceeded Act 43 on all legal dimensions. If anything, the sensitivity
testing substantially bolsters this conclusion, since it shows that Act 43's large efficiency gap and the Demonstration Plan's small one are durable in the face of enormous changes in Wisconsin's electoral environment.

Dated: December 21, 2015
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# Assessing the Current Wisconsin State Legislative Districting Plan 

Simon Jackman

July 7, 2015

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## 1 Introduction

My name is Simon Jackman. I am currently a Professor of Political Science at Stanford University, and, by courtesy, a Professor of Statistics. I joined the Stanford faculty in 1996. I teach classes on American politics and statistical methods in the social sciences.

I have been asked by counsel representing the plaintiffs in this lawsuit (the "Plaintiffs") to analyze relevant data and provide expert opinions in the case titled above. More specifically, I have been asked

- to determine if the current Wisconsin legislative districting plan constitutes a partisan gerrymander;
- to explain a summary measure of a districting plan known as "the efficiency gap" (Stephanopolous and McGhee, 2015), what it measures, how it is calculated, and to assess how well it measures partisan gerrymandering;
- to compare the efficiency gap to extant summary measures of districting plans such as partisan bias;
- to analyze data from state legislative elections in recent decades, so as to assess the properties of the efficiency gap and to identify plans with high values of the efficiency gap;
- to suggest a threshold or other measure that can be used to determine if a districting plan is an extreme partisan gerrymander;
- to describe how the efficiency gap for the Wisconsin districting plan compares to the values of the efficiency gap observed in recent decades elsewhere in the United States;
- to describe where the efficiency gap for the current Wisconsin districting plan lies in comparison with the threshold for determining if a districting plan constitutes an extreme partisan gerrymander.

My opinions are based on the knowledge I have amassed over my education, training and experience, and follow from statistical analysis of the following data:

- a large, canonical data set on candidacies and results in state legislative elections, 1967 to the present available from the Inter-University Consortium for Political and Social Research (ICPSR study number 34297); I use a release of the data updated through 2014, maintained by Karl Klarner (Indiana State University and Harvard University).
- presidential election returns, 2000-2012, aggregated to state legislative districts.


## 2 Qualifications, Publications and Compensation

My Ph.D. is in Political Science, from the University of Rochester, where my graduate training included courses in econometrics and statistics. My curriculum vitae is attached to this report.

All publications that I have authored and published in the past ten years appear in my curriculum vitae. Those publications include peer-reviewed journals such as: The Journal of Politics, Electoral Studies, The American Journal of Political Science, Legislative Studies Quarterly, Election Law Journal, Public Opinion Quarterly, Journal of Elections, Public Opinion and Parties, and PS: Political Science and Politics.

I have published on properties of electoral systems and election administration in Legislative Studies Quarterly, the Australian Journal of Political Science, the British Journal of Political Science, and the Democratic Audit of Australia. I am a Fellow of the Society for Political Methodology and a member of the American Academy of Arts and Sciences.

I am being compensated at a rate of $\$ 250$ per hour.

## 3 Summary

1. Partisan gerrymandering and wasted votes. In two-party, single-member district electoral systems, a partisan gerrymander operates by effectively "wasting" more votes cast for one party than for the other. Wasted votes are votes for a party in excess of what the party needed to win a given district or votes cast for a party in districts that the party doesn't win. Differences
in wasted vote rates between political parties measure the extent of partisan gerrymandering.
2. The efficiency gap $(E G)$ is a relative, wasted vote measure, the ratio of one party's wasted vote rate to the other party's wasted vote rate. EG can be computed directly from a given election's results, without recourse to extensive statistical modeling or assumptions about counter-factual or hypothetical election outcomes, unlike other extant measures of the fairness of an electoral system (e.g., partisan bias).
3. The efficiency gap is an "excess seats" measure, reflecting the nature of a partisan gerrymander. An efficiency gap in favor one party sees it wasting fewer votes than its opponent, thus translating its votes across the jurisdiction into seats more efficiently than its opponent. This results in the party winning more seats than we'd expect given its vote share $(V)$ and if wasted vote rates were the same between the parties. $E G=0$ corresponds to no efficiency gap between the parties, or no partisan difference in wasted vote rates. In this analysis (but without loss of generality) $E G$ is normed such that negative $E G$ values indicate higher wasted vote rates for Democrats relative to Republicans, and $E G>0$ the converse.
4. A districting plan in which $E G$ is consistently observed to be positive is evidence that the plan embodies a pro-Democratic gerrymander; the magnitudes of the $E G$ measures speak to the severity of the gerrymander. Conversely, a districting plan with consistently negative values of the efficiency gap is consistent with the plan embodying a pro-Republican gerrymander.
5. Performance of the efficiency gap in 786 state legislative elections. My analysis of 786 state legislative elections (1972-2014) examines properties of the efficiency gap. $E G$ is estimated with some uncertainty in the presence of uncontested districts (and uncontested districts are quite prevalent in state legislative elections), but this source of uncertainty is small relative to differences in the $E G$ across states and across districting plans.
6. Stability of the efficiency gap. $E G$ is stable in pairs of temporally adjacent elections held under the same districting plan. In 580 pairs of consecutive
$E G$ measures, the probability that each $E G$ measure has the same sign is $74 \%$. In 141 districting plans with three or more elections, $35 \%$ have a better than $95 \%$ probability of $E G$ being negative or positive for the entire duration of the plan; in about half of the districting plans the probability that $E G$ doesn't change sign is above $75 \%$.
7. Recent decades show more pro-Republican gerrymandering, as measured by the efficiency gap. Efficiency gap measures in recent decades show a pronounced shift in a negative direction, indicative of an increased prevalence of districting plans favoring Republicans. Among the 10 most proDemocratic EG measures in my analysis, none were recorded after 2000.
8. The current Wisconsin state legislative districting plan (the "Current Wisconsin Plan"). In Wisconsin in 2012, the average Democratic share of districtlevel, two-party vote $(V)$ is estimated to be $51.4 \%$ ( $\pm 0.6$, the uncertainty stemming from imputations for uncontested seats); recall that Obama won $53.5 \%$ of the two-party presidential vote in Wisconsin in 2012. Yet Democrats won only 39 seats in the 99 seat legislature ( $S=39.4 \%$ ), making Wisconsin one of 7 states in 2012 where we estimate $V>50 \%$ but $S<50 \%$. In Wisconsin in 2014, $V$ is estimated to be $48.0 \%( \pm 0.8)$ and Democrats won 36 of 99 seats ( $S=36.4 \%$ ).
9. Accordingly, Wisconsin's EG measures in 2012 and 2014 are large and negative: -. 13 and -. 10 (to two digits of precision). The 2012 estimate is the largest EG estimate in Wisconsin over the 42 year period spanned by this analysis (1972-2014).
10. Among $79 E G$ measures generated from state legislative elections after the 2010 round of redistricting, Wisconsin's EG scores rank 9th (2012, 95\% CI 4 to 13) and 18th (2014, $95 \%$ CI 14 to 21). Among 786 EG measures in the 1972-2014 analysis, the magnitude of Wisconsin's 2012 EG measure is surpassed by only $27(3.4 \%)$ other cases.
11. Analysis of efficiency gaps measures in the post-1990 era indicates that conditional on the magnitude of the Wisconsin 2012 efficiency gap (the first election under the Current Wisconsin Plan), there is a $100 \%$ probability
that all subsequent elections held under that plan will also have efficiency gaps disadvantageous to Democrats.
12. The Current Wisconsin Plan presents overwhelming evidence of being a proRepublican gerrymander. In the entire set of 786 state legislative elections and their accompanying $E G$ measures, there are no precedents prior to this cycle in which a districting plan generates an initial two-election sequence of $E G$ scores that are each as large as those observed in WI.
13. The Current Wisconsin Plan is generating EG measures that make it extremely likely that it has a systematic, historically large and enduring, proRepublican advantage in the translation of votes into seats in Wisconsin's state legislative elections.
14. An actionable threshold based on the efficiency gap. Historical analysis of the relationship between the first $E G$ measure we observe under a new districting plan and the subsequent $E G$ measures lets us assess the extent to which that first $E G$ estimate is a reliable indicators of a durable and hence systematic feature of the plan. In turn, this let us assess the confidence associated with a range of possible actionable EG thresholds.
15. My analysis suggests that $E G$ greater than .07 in absolute value be used as an actionable threshold. Relatively few plans produce a first election with an $E G$ measure in excess of this threshold, and of those that do, the historical analysis suggests that most go on to produce a sequence of $E G$ estimates indicative of systematic, partisan advantage consistent with the first election $E G$ estimates, At the 0.07 threshold, $95 \%$ of plans would be either (a) undisturbed by the courts, or (b) struck down because we are sufficiently confident that the plan, if left undisturbed, would go on to produce a one-sided sequence of $E G$ estimates, consistent with the plan being a partisan gerrymander. In short, our "confidence level" in the 0.07 threshold is $95 \%$.
16. The Current Wisconsin Plan is generating estimates of the efficiency gap far in excess of this proposed, actionable threshold. In 2012 elections to the Wisconsin state legislature, the efficiency gap is estimated to be -.13 ; in

2014, the efficiency gap is estimated to be -.10 . Both measures are separately well beyond the conservative .07 threshold suggested by the analysis of efficiency gap measures observed from 1972 to the present.

A vivid, graphical summary of my analysis appears in Figure 1, showing the average value of the efficiency gap in 206 districting plans, spanning 41 states and 786 state legislative elections from 1972 to 2014. The Current Wisconsin Plan has been in place for two elections (2012 and 2014), with an average efficiency gap of -.115 . Details on the interpretation and calculation of the efficiency gap come later in my report, but for now note that negative values of the efficiency gap indicate a districting plan favoring Republicans, while positive values indicate a plan favoring Democrats. Note that only four other districting plans have lower average efficiency gap scores than the Current Wisconsin Plan, and these are also from the post-2010 round of redistricting. That is, Wisconsin's current plan is generating the 5th lowest average efficiency gap observed in over 200 other districting plans used in state legislative elections throughout the United States over the last 40 years. The analysis I report here documents why the efficiency gap is a valid and reliable measure of partian gerrymandering and why are confident that the current Wisconsin plan exceeds even a conservative definition of partisan gerrymandering.

## 4 Redistricting plans

A districting plan is an exercise in map drawing, partitioning a jurisdiction into districts, typically required to be contiguous, mutually exclusive and exhaustive regions, and - at least in the contemporary United States — of approximately the same population size. In a single-member, simple plurality (SMSP) electoral system, the highest vote getter in each district is declared the winner of the election. Partisan gerrymandering is the process of drawing districts that favor one party, typically by creating a set of districts that help the party win an excess of seats (districts) relative to its jurisdiction-wide level of support.

What might constitute evidence of partisan gerrymandering? One indication might be a series of elections conducted under the same districting plan in which a party's seat share $(S)$ is unusually large (or small) relative to its vote share $(V)$.


Figure 1: Average efficiency gap score, 206 districting plans, 1972-2014. Plans have been sorted from low average $E G$ scores to high. Horizontal lines cover $95 \%$ confidence intervals. Negative efficiency gap scores are plans that disadvantage Democrats; positive efficiency gap scores favor Democrats. The Current Wisconsin Plan is shown in red. See also Figure 36.

There may be elections where a party wins a majority of seats (and control of the jurisdiction's legislature) despite not winning a majority of votes: $S>.5$ while $V<.5$ and vice-versa. In fact, there are numerous instances of mismatches between the party winning the statewide vote and the party controlling the state legislature in recent decades. I estimate that since 1972 there have been 63 cases of Democrats winning a majority of the vote in state legislative elections, while not winning a majority of the seats, and 23 cases of the reverse phenomenon, where Democrats won a majority of the seats with less than $50 \%$ of the statewide, two-party vote.

Geographic clustering of partisans is typically a prerequisite for partisan gerrymandering. This is nothing other than partisan "packing": a gerrymandered districting plan creates a relatively small number of districts that have unusually large proportions of partisans from party $B$. The geographic concentration of party $B$ partisans might make creating these districts a straightforward task. In other districts in the jurisdiction, party $B$ supporters never (or seldom) constitute a majority (or a plurality), making those districts "safe" for party $A$. This districting plan helps ensure party $A$ wins a majority of seats even though party $B$ has a majority of support across the jurisdiction, or at the very least, the districting plan helps ensures that party $A$ 's seat share exceeds its vote share in any given election.

It is conventional in political science to say that such a plan allows party $A$ to "more efficiently" translate its votes into seats, relative to the way the plan translates party $B$ 's votes into seats. This nomenclature is telling, as we will see when we consider the efficiency gap measure, below.

Assessing the partisan fairness of a districting plan is fundamentally about measuring a party's excess (or deficit) in its seat share relative to its vote share. The efficiency gap is such a summary $m$ easure. To a ssess the properties of the efficiency gap, I first review some core concepts in the analysis of districting plans: vote shares, seat shares, and the relationship between the two quantities in singlemember districts.

### 4.1 Seats-Votes Curves

Electoral systems translate parties' vote shares $(V)$ into seat shares $(S)$. Both $V$ and $S$ are proportions. Plotting the two quantities $V$ and $S$ against one another yields the "seats-votes" curve, a staple in the analysis of electoral systems and districting plans. Two seats-votes curves are shown in Figure 2, one showing a non-linear relationship between seats and votes typical of single-member district systems, ${ }^{1}$ the other showing a linear relationship between seats and votes observed under proportional representation systems.

In pure proportional representation (PR) voting systems, seats-votes curves are 45 degree lines by design, crossing the $(V, S)=(.5, .5)$ point: i.e., under PR, $S=V$ and a party that wins $50 \%$ of the vote will be allocated $50 \%$ of the seats. Absent a deterministic allocation rule like pure PR, seats-votes curves are most usefully thought of in probabilistic terms, due to the fact that there are many possible configurations of district-specific outcomes corresponding to a given jurisdiction-wide $V$, and hence uncertainty - represented by a probability distribution - over possible values of $S$ given $V$.

In single-member, simple plurality (SMSP) systems, we often see non-linear, " S "-shaped seats-votes curves. With an approximately symmetric mix of districts (in terms of partisan leanings), large changes in seat shares ( $S$ ) can result from relatively small changes in votes shares $(V)$ at the middle of the distribution of district types. This presumes a districting plan such that both parties have a small number of "strongholds," with extremely large changes in vote shares needed to threaten these districts, and so the seats-votes curve tends to "flatten out" as jurisdiction-wide vote share ( $V$ ) takes on relatively large or small values. Other shapes are possible too: e.g., bipartisan, incumbent-protection plans generate seats-votes curves that are largely flat for most values of $V$, save for the constraint that the curve run through the points $(V, S)=(0,0)$ and $(1,1)$; i.e., relatively large movements in $V$ generates relatively little change in seats shares.

[^35]

Figure 2: Two Theoretical Seats-Votes Curves

## 5 Partisan bias

Both of the hypothetical seats-votes curves in Figure 2 run through the "50$50 "$ point, where $V=.5$ and $S=.5$. An interesting empirical question is whether actual seats-votes curves run through this point, or more generally, whether the seats-votes curve is symmetric about $V=.5$. Formally, symmetry of the seatsvote curve is the condition that $E(S \mid V)=1-E(S \mid 1-V)$, where $E$ is the expectation operator, averaging over the uncertainty with respect to $S$ given $V$. The vertical offset from the $(.5, .5)$ point for a seats-votes curve is known as partisan bias: the extent to which a party's expected seat share lies above or below $50 \%$, conditional on that party winning $50 \%$ of the jurisdiction-wide vote.

Figure 3 shows three seats-votes curves, with the graph clipped to the region $V \in[.4,6$.$] and S \in[.4, .6]$ so as to emphasize the nature of partisan bias. The blue, positive bias curve "lifts" the seats-votes curve; it crosses $S=.5$ with $V<.5$ and passes through the upper-left quadrant of the graph. That is, with positive bias, a party can win a majority of the seats with less then a majority of the jurisdiction-wide or average vote; equivalently, if the party wins $V=.5$, it can expect to win more than $50 \%$ of the seats. Conversely, with negative bias, the opposite phenomenon occurs: the party can't expect to win a majority of the seats until it wins more than a majority of the jurisdiction-wide or average vote.

### 5.1 Multi-year method

With data from multiple elections under the same district plan, partisan bias can be estimated by fitting a seats-votes curve to the observed seat and vote shares, typically via a simple statistical technique such as linear regression; this approach has a long and distinguished lineage in both political science and statistics (e.g., Edgeworth, 1898; Kendall and Stuart, 1950; Tufte, 1973). Niemi and Fett (1986) referred to this method of estimating the partisan bias of an electoral system as the "multi-year" method, reflecting the fact that the underlying data comes from a sequence of elections.

This approach is of limited utility when assessing a new or proposed districting plan. More generally, it is of no great help to insist that a sequence of elections must be conducted under a redistricting plan before the plan can be properly assessed. Indeed, few plans stay intact long enough to permit reliable analysis in


Figure 3: Theoretical seats-votes curves, with different levels of partisan bias. This graph is "zoomed in" on the region $V \in[.4, .6]$ and $S \in[.4, .6]$; the seatsvotes "curves" are approximately linear in this region.
this way. State-level plans in the United States might generate as many five elections between decennial censuses. Accordingly, many uses of the "multi-year" method pool multiple plans and/or across jurisdictions, so as to estimate average partisan bias. For instance, Niemi and Jackman (1991) estimated average levels of partisan bias in state legislative districting plans, collecting data spanning multiple decades and multiple states, and grouping districting plans by the partisanship of the plan's authors (e.g., plans drawn under Republican control, Democratic control, mixed, or independent).

Assessing the properties of a districting plan after a tiny number of elections - or no elections - requires some assumptions and/or modeling. A single election yields just a single $(V, S)$ data point, through which no unique seats-vote curve can be fitted a nd so partisan bias can't beestimated without further assumptions. Absent any actual elections under the plan, we might examine votes from a previous election, say, with precinct level results re-aggregated to the new districts.

### 5.2 Uniform swing

One approach—dating back to Sir David Butler's (1974) pioneering work on British elections-is the uniform partisan swing approach. Let $\mathbf{v}=\left(v_{1}, \ldots, v_{n}\right)^{\prime}$ be the set of vote shares for party $A$ observed in an election with $n$ districts. Party $A$ wins seat $i$ if $v_{i}>.5$, assuming just two parties (or defining $v$ as the share of two-party vote); i.e., $s_{i}=1$ if $v_{i}>.5$ ) and otherwise $s_{i}=0$. Party $A$ 's seat share is $S={ }_{\underline{n}}^{1} \sum_{i=1}^{n} s_{i} . V$ is the jurisdiction-wide vote share for party $A$, and if each district had the same number of voters $V=\bar{v}=\frac{1}{n} \sum_{i=1}^{n} v_{i}$, the average of the districtlevel $v_{i}$. Districts are never exactly equal sized, in which case we can define $V$ as follows: let $t_{i}$ be the number of voters in district $i$, and $V=\sum_{i=1}^{n} t_{i} v_{i} / \sum_{i=1}^{n} t_{i}$.

The uniform swing approach perturbs the observed district-level results $\mathbf{v}$ by a constant factor $\delta$, corresponding to a hypothetical amount of uniform swing across all districts. For a given $\delta$, let $v_{i}^{*}=v_{i}+\delta$ which in turn generates $V^{*}=V+\delta$ and an implied seat share $S^{*}$. Now let $\delta$ vary over a grid of values ranging from $-V$ to $1-V$; then $V^{*}$ varies from 0 to 1 and a corresponding value of $S^{*}$ can also be computed at every grid point. The resulting set of $\left(V^{*}, S^{*}\right)$ points are then plotted to form a seats-vote curve (actually, a step function). Partisan bias is
simply "read off" this set of results, computed as $S^{*} \mid\left(V^{*}=.5\right)-.5$.
There is an elegant simplicity to this approach, taking an observed set of district-level vote shares $\mathbf{v}$ and shifting them by the constant $\delta$. The observed distribution of district level vote shares observed in a given election is presumed to hold under any election we might observe under the redistricting plan, save for the shift given by the uniform swing term $\delta$.

### 5.3 Critiques of partisan bias

Among political scientists, the uniform swing approach was criticized for its determinism. Swings are never exactly uniform across districts. There are many permutations of observed vote shares that generate a statewide vote share of $50 \%$ other than simply shifting observed district-level results by a constant factor. A less deterministic approach to assessing partisan bias was developed over a series of papers by Gary King and Andrew Gelman in the early 1990s (e.g., Gelman and King, 1990). This approach fits a statistical model to district-level vote shares and, optionally, utilizing available predictors of district-level vote shares - to model the way particular districts might exhibit bigger or smaller swings than a given level of state-wide swing. Perhaps one way to think about the approach is that it is "approximate" uniform swing, with statistical models fit to historical election results to predict and bound variation around a state-wide average swing. The result is a seats-vote curve and an estimate of partisan bias that comes equipped with uncertainty measures, reflecting uncertainty in the way that individual districts might plausibly deviate from the state-wide average swing yet still produce a state-wide average vote of $50 \%$.

The King and Gelman model-based simulation approaches remain the most sophisticated methods of generating seats-votes curves, extrapolating from as little as one election to estimate a seats-votes curve and hence an estimate of partisan bias. Despite the technical sophistication with which we can estimate partisan bias, legal debate has centered on a more fundamental issue, the bypothetical character of partisan bias itself. Recall that partisan bias is defined as "seats in excess of $50 \%$ had the jurisdiction-wide vote split $50-50$." The premise that $V=.5$ is the problem, since this will almost always be a counter-factual or hypothetical scenario. The further $V$ is away from .5 in a given election, the
counter-factual we must contemplate (when assessing the partisan bias of a districting plan) becomes all the more speculative.

In no small measure this is a marketing failure, of sorts. Partisan bias (at least under the uniform swing assumption) is essentially a measure of skew or asymmetry in actual vote shares. Partisan bias garners great rhetorical and normative appeal by directing attention to what happens at $V=.5$; it seems only "fair" that if a party wins $50 \%$ or more of the vote it should expect to win a majority of the districts.

Yet this distracts us from the fact that asymmetry in the distribution of vote shares across districts is the key, operative feature of a districting plan, and the extent to which it advantages one party or the other. Critically, we need not make appeals to counter-factual, hypothetical elections in order to assess this asymmetry.

## 6 The Efficiency Gap

The efficiency gap $(E G)$ is also an asymmetry measure, as we see b elow. But unlike partisan bias, the interpretation of the efficiency gap is not explicitly tied to any counter-factual election outcome. In this way, the efficiency gap provides a way to assess districting plans that is free of the criticisms that have stymied the partisan bias measure.

Stephanopoulos and McGhee (2015) derive the EG measure with the concept of wasted votes. A party only needs $v_{i}=50 \%+1$ of the votes to win district i. Anything more are votes that could have been deployed in other districts. Conversely, votes in districts where the party doesn't win are "wasted," from the perspective of generating seats: any districts with $v_{i}<.5$ generate no seats.

Wasted votes get at the core of what partisan gerrymandering is, and how it operates. A gerrymander against party $A$ creates a relatively small number of districts that "lock up" a lot of its votes ("packing" with $v_{i}>.5$ ) and a larger number of districts that disperse votes through districts won by party $B$ ("cracking" with $\left.v_{i}<.5\right)$. To be sure, both parties are wasting votes. But partisan advantage ensues when one party is wasting fewer votes than the other, or, equivalently, more efficiently translating votes into seats. Note also how the efficiency gap measure is also closely tied to asymmetry in the distribution of $v_{i}$.

Some notation will help make the point more clearly. If $v_{i}>.5$ then party $A$ wins the district and $s_{i}=1$; otherwise $s_{i}=0$. The efficiency gap is defined by McGhee $(2014,68)$ as "relative wasted votes" or

$$
E G=\frac{W_{B}}{n}-\frac{W_{A}}{n}
$$

where

$$
W_{A}=\sum_{i=1}^{n} s_{i}\left(v_{i}-.5\right)+\left(1-s_{i}\right) v_{i}
$$

is the sum of wasted vote proportions for party $A$ and

$$
W_{B}=\sum_{i=1}^{n}\left(1-s_{i}\right)\left(.5-v_{i}\right)+s_{i}\left(1-v_{i}\right)
$$

is the sum of wasted vote proportions for party $B$ and $n$ is the number of districts in the jurisdiction. If $E G>0$ then party $B$ is wasting more votes than $A$, or $A$ is translating votes into seats more efficiently than $B$; if $E G<0$ then the converse, party $A$ is wasting more votes than $B$ and $B$ is translating votes into seats more efficiently than $A$.

### 6.1 The efficiency gap when districts are of equal size

Under the assumption of equally sized districts McGhee $(2014,80)$ re-expresses the efficiency gap as:

$$
\begin{equation*}
E G=S-.5-2(V-.5) \tag{1}
\end{equation*}
$$

recalling that $S=n^{-1} \sum_{i=1}^{n} s_{i}$ is the proportion of seats won by party $A$ and $V=$ $n^{-1} \sum_{i=1}^{n} v_{i}$ is the proportion of votes won by party $A$.

The assumption of equally-sized districts is especially helpful for the analysis reported below, since the calculation of $E G$ in a given election then reduces to using the jurisdiction-level quantities $S$ and $V$ as in equation 1. For the analysis of historical election results reported below, it isn't possible to obtain measures of district populations, meaning that we really have no option other than to rely on the jurisdiction-level quantities $S$ and $V$ when estimating the $E G$.

I operationalize $V$ as the average (over districts) of the Democratic share of the two-party vote, in seats won by either a Democratic or Republican candidate;
this set of seats includes uncontested seats, where I will use imputation procedures to estimate two-party vote share. If districts are of equal size (and ignoring seats won by independents and minor party candidates) then this average over districts will correspond to the Democratic share of the state-wide, two-party vote.

### 6.2 The seats-vote curve when the efficiency gap is zero

This simple expression for the efficiency gap implies that if the efficiency gap is zero, we obtain a particular type of seats-votes curve, shown in Figure 4:

1. the seats-votes curve runs through the 50-50 point. If the jurisdiction wide vote is split 50-50 between party $A$ and party $B$ then with an efficiency gap of zero, $S=.5$.
2. conditional on $V=.5$ (an even split of the vote), the efficiency gap is the same as partisan bias: $V=.5 \Longleftrightarrow E G=S-.5$, the seat share for party $A$ in excess of $50 \%$. That is, the efficiency gap reduces to partisan bias under the counter-factual scenario $V=.5$ that the partisan bias measure requires us to contemplate. On the other hand, the efficiency gap is not premised on that counter-factual holding, or any other counter-factual for that matter; the efficiency gap summarizes the distribution of observed district-level vote shares $v_{i}$.
3. the seats-votes curve is linear through the $50-50$ point with a slope of 2 . That is, with $E G=0, S=2 V-.5$. Or, with a zero efficiency gap, each additional percentage point of vote share for party $A$ generates two additional percentage points of seat share. A zero efficiency gap does not imply proportional representation (a seats-votes that is simply a 45 degree line).
4. a party winning $25 \%$ or less of the jurisdiction-wide vote should win zero seats under a plan with a zero efficiency gap; a party winning $75 \%$ or more of the jurisdiction-wide vote should win all of the seats under a plan with a zero efficiency gap. This is a consequence of the " 2 -to- 1 " seats/vote ratio and the symmetry implied by a zero efficiency gap. A party that wins an extremely low share of the vote $(V<.25)$ can only be winning any seats if it enjoys an efficiency advantage over its opponent.


Figure 4: Theoretical seats-votes curves. The $E G=0$ curve implies that (a) a party winning less than $V=.25$ jurisdiction-wide should not win any seats; (b) symmetrically, a party winning more than $V=.75$ jurisdistion-wide should win all the seats; and (c) the relationship between seat shares $S$ and vote shares $V$ over the interval $V \in[.25, .75]$ is a linear function with slope two (i.e., for every one percentage point gain in vote share, seat share should go up by two percentage points).

Moreover, the efficiency gap is trivial to compute once we have $V$ and $S$ for a given election. We don't need a sequence of elections under a plan in order to compute $E G$, nor do we need to anchor ourselves to a counter-factual scenario such as $V=.5$ as we do when computing partisan bias. For any given observed $V$, the hypothesis of zero efficiency gap tells us what level of $S$ to expect.

### 6.3 The efficiency gap as an excess seats measure

In this sense the efficiency gap can beinterpreted e ven more imply as an "excess seats" measure. Recall that $E G=0 \Longleftrightarrow S=2 V-.5$. In a given election we observe $E G=S-.5-2(V-.5)$. The efficiency gap can be computed by noting how far the observed $S$ lies above or below the orange line in Figure 4.

A positive $E G$ means "excess" seats for party $A$ relative to a zero efficiency gap standard given the observed $V$ in that election; conversely, a negative $E G$ mean a deficit in seats for party $A$ relative to a zero efficiency gap standard given the observed $V$.

## 7 State legislative elections, 1972-2014

We estimate the efficiency gap in state legislative elections over a large set of states and districting plans, covering the period 1972 to 2014 . We begin the analysis in 1972 for two primary reasons: (a) state legislative election returns are harder to acquire prior to the mid-1960s, and not part of the large, canonical data collection we rely on (see below); and (b) districting plans and sequences of elections from 1972 onwards can be reasonably considered to be from the post-malapportionment era.

For each election we recover an estimate of the efficiency gap based on the election results actually observed in that election. To do this, I compute two quantities for each election:

1. $V$, the statewide share of the two-party vote for Democratic candidates, formed by averaging the district-level election results $v_{i}$ (the Democratic share of the two-party vote in district $i$ ) in seats won by major party candidates, including uncontested seats, and
2. $S$, the Democratic share of seats won by major parties.

Recall that these quantities are the inputs required when computing the efficiency gap (equation 1).

The analysis that follows relies on a data set widely used in political science and freely available from the Inter-University Consortium for Political and Social Research (ICPSR study number 34297). The release of the data I utilize covers state legislative election results from 1967 to 2014, updated by Karl Klarner (Indiana State University and Harvard University). I subset the original data set to general election results since 1972 in states whose lower houses are elected via single-member districts, or where single-member districts are the norm. Multimember districts "with positions" are treated as if they are single-member districts.

Figure 5 provides a graphical depiction of the elections that satisfy the selection criteria described above.

- Arizona, Idaho, Louisiana, Maryland, Nebraska, New Hampshire, New Jersey, North Dakota and South Dakota all drop out of the analysis entirely, because of exceedingly high rates of uncontested races, using multi-member districts, non-partisan elections, or the use of a run-off system (Louisiana).
- Alaska, Hawaii, Illinois, Indiana, Kentucky, Maine, Minnesota, Montana, North Carolina, Vermont, Virginia, West Virginia and Wyoming do not supply data over the entire 1972-2014 span; this is sometimes due to earlier elections being subject to exceedingly high rates of uncontestedness, the use of multi-member districts or non-partisan elections.
- Alabama and Mississippi have four-year terms in their lower houses, contributing data at only half the rate of the vast bulk of states with two-year legislative terms.
- Twenty-three states supply data every two years from 1972 to 2014 , including Michigan and Wisconsin.
- Data is more abundant in recent decades. For the period 2000 to 2014, 41 states contribute data to the analysis at two or four year intervals.

In summary, the data available for analysis span 83,269 district-level state legislative contests, from 786 elections across 41 states.


Figure 5: 786 state legislative elections available for analysis, 1972-2014, by state.

### 7.1 Grouping elections into redistricting plans

Districting plans remain in place for sequences of elections. An important component of my analysis involves tracking the efficiency $g$ ap a cross a series of elections held under the same districting plan. A key question is how much variation in the EG do we observe within districting plans, versus variation in the EG between districting plans.

To the extent that the $E G$ is a feature of a districting plan per se, we should observe a small amount of within-plan variation relative to between plan variation. To perform this analysis we must group sequences of elections within states by the districting plan in place at the time.

Stephanopolous and McGhee (2015) provide a unique identifier for the districting plan in place for each state legislative election, for which I adopt here.

Figure 6 displays how the elections available for analysis group by districting plan. Districts are typically redrawn after each decennial census; the first election conducted under new district boundaries is often the " 2 " election (1982, 1992, etc). Occasionally we see just one election under a plan: examples include Alabama 1982, California, Hawaii 1982, Tennessee 1982, Ohio 1992, South Carolina 1992, North Carolina 2002, and South Carolina 2002.

Alaska, Kentucky, Pennsylvania and Texas held just one election under their respective districting plans adopted after the 2010 Census. In each of those states a different plan was in place for 2014 state legislative elections. Alabama's state legislature has a four year term and we observe only the 2014 election under its post-2010 plan. The last election from Mississippi was in 2011 and was held under the plan in place for its 2003 and 2007 elections.

### 7.2 Uncontested races

Uncontested races are common in state legislative elections, and are even the norm in some states. For $38.7 \%$ of the district-level results in this analysis, it isn't possible to directly compute a two-party vote share $\left(v_{i}\right)$, either because the seat was uncontested or not contested by both a Democratic and Republican candidate, or (in a tiny handful of cases) the data are missing.

In some states, for some elections, the proportion of uncontested races is so high that we drop the election from the analysis. As noted earlier, examples


Figure 6: 786 state legislative elections available for analysis, 1972-2014, by state, grouped by districting plan (horizontal line).
include Arkansas elections prior to 1992 and South Carolina in 1972.
Even with these elections dropped from the analysis, the extent of uncontestedness in the remaining set of state legislative election results is too large to be ignored. Of the remaining elections, $31 \%$ have missing two-party results in at least half of the districts.

A graphical summary of the prevalence of uncontested districts appears in Figure 7, showing the percentage of districts without Democratic and Republican vote counts, by election and by state. Uncontested races are the norm in a number of Southern states: e.g., Georgia, South Carolina, Mississippi, Arkansas, Texas, Alabama, Virginia, Kentucky and Tennessee record rates of uncontestedness that seldom, if ever, drop below $50 \%$ for the period covered by this analysis. Wyoming also records a high proportion of districts that do not have Democratic versus Republican contests. States that lean Democratic also have high levels of uncontestedness too: see Rhode Island, Massachusetts, Illinois and, in recent decades, Pennsylvania.

Michigan and Minnesota are among the states with the lowest levels of uncontested districts in their state legislative elections. Over the set of 786 state legislative elections we examine, there are just three instances of elections with Democrats and Republicans running candidates in every district: Michigan supplies two of these cases (2014 and 1996) and Minnesota the other (2008).

## 8 Imputations for Uncontested Races

Stephanopolous and McGhee (2015) note the prevalence of uncontested races and report using a statistical model to impute vote shares to uncontested districts. They write:

We strongly discourage analysts from either dropping uncontested races from the computation or treating them as if they produced unanimous support for a party. The former approach eliminates important information about a plan, while the latter assumes that coerced votes accurately reflect political support.

I concur with this advice, utilizing an imputation strategy for uncontested districts with two distinct statistical models, predicting Democratic, two-party

Percent single-member districts without D and R candidates/vote counts, by state \& election


Figure 7: Percentage of districts missing two-party vote shares, by election, in 786 state legislative elections, 1972-2014. Missing data is almost always due to districts being uncontested by both major parties.
vote share in state legislative districts $\left(v_{i}\right)$.

### 8.1 Imputation model 1: presidential vote shares

The first imputation model relies on presidential election returns reported at the level of state legislative districts. Presidential election returns are excellent predictors of state legislative election outcomes and observed even when state legislative elections are uncontested. I fit a series of linear regressions of $v_{i}$ on the Democratic share of the two-party vote for president in district $i$, as recorded in the most temporally-proximate presidential election for which data is available and for which the current election's districting plan was in place; separate slopes and intercepts are estimated depending on the incumbency status of district $i$ (Democratic, Open/Other, Republican).

The model also embodies the following assumptions in generating imputations for unobserved vote shares in uncontested districts. In districts where a Republican incumbent ran unopposed, we assume that the Democratic share of the two-party vote would have been less than $50 \%$; conversely, where Democratic incumbents ran unopposed, we assume that the Democratic share of the vote would have been greater than $50 \%$.

In most states the analysis predicts 2014 and 2012 state legislative election results $v_{i}$ using 2012 presidential vote shares; 2006, 2008 and $2010 v_{i}$ is regressed on 2008 presidential vote shares, and so on. Some care is needed matching state and presidential election results in states that hold their state legislative elections in odd-numbered years, or where redistricting intervenes. In a small number of cases, presidential election returns are not available, or are recorded with district identifiers that can't be matched in the state legislative elections data. We lack data on presidential election results by state legislative district prior to 2000, so 1992 is the earliest election with which we can match state legislative election results to presidential election results at the district level.

The imputation model generally fits well. Across the 447 elections, the median $r^{2}$ statistic is 0.82 . The cases fitting less well include Vermont in $2012\left(r^{2}=0.29\right)$, with relatively few contested seats and multi-member districts with positions.

We examine the performance of the imputation model in a series of graphs, below, for six sets of elections: Wisconsin in 2012 and 2014, Michigan in 2014


Figure 8: Distribution of $r^{2}$ statistics, regressions of Democratic share of twoparty vote in state legislative election outcomes on Democratic share of the twoparty for president.
(with no uncontested districts), South Carolina in 2012 (with the highest proportion of uncontested seats in the 2012 data), Virginia in 2013 and Wyoming in 2012 (the latter two generating extremely large, negative values of the efficiency gap). Vertical lines indicate $95 \%$ confidence intervals around imputed values for the Democratic share of the two-party vote in state legislative elections (vertical axis). Separate slopes and intercepts are fit for each incumbency type. Note also that the imputed data almost always lie on the regression lines.

Imputations for uncontested districts are accompanied by uncertainty. Although the imputation models generally fit well, like any realistic model they provides less than a perfect fit to the data. Note too that in any given election, there is only a finite amount of data and hence a limit to the precision with which we can make inferences about unobserved vote shares based on the relationship between observed vote shares and presidential vote shares.

Uncertainty in the imputations for $v$ in uncontested districts generates uncertainty in "downstream" quantities of interest such as statewide Democratic vote share $V$ and the efficiency gap measure $E G$. This is key, given the fact that uncontestedness is so pervasive in these data. We want any conclusions about the efficiency gap's properties or inferences about particular levels of the efficiency gap to reflect the uncertainty resulting from imputing vote shares in uncontested districts.


Figure 9: Regression model for imputing unobserved vote shares in 6 selected elections. Vertical lines indicate $95 \%$ confidence intervals around imputed values for the Democratic share of the two-party vote in state legislative elections (vertical axis). Separate slopes and intercepts are fit for each incumbency type. Note also that the imputed data almost always lie on the regression lines.

### 8.2 Imputation model 2

We rely on imputations based on presidential election returns when they are available. But presidential vote isn't always available at the level of state legislative districts (not before 1992, in this analysis). To handle these cases, we rely on a second imputation procedure, one that models sequences of election results observed under a redistricting plan, interpolating unobserved Democratic vote shares given (1) previous and future results for a given district; (2) statewide swing in a given state election; and (3) change in the incumbency status of a given district. This model also embodies the assumption that unobserved vote shares would nonetheless be consistent with what we did observe in a given seat: where a Democrat wins in an uncontested district, any imputation for $v$ in that district must lie above $50 \%$, and where a Republican wins an uncontested district, any imputation for $v$ must lie below $50 \%$.

### 8.3 Combining the two sets of imputations

We now have two sets of imputations for uncontested districts: (1) using presidential vote as a basis for imputation, where available (447 state legislative elections from 1992 to 2014); and (2) the imputation model that relies on the trajectory of district results over the history of a districting plan, including incumbency and estimates of swing, which supplies imputations for uncontested districts in all years.

When there are no uncontested districts, obviously the two imputations must agree, for the trivial reason that are no imputations to perform. As the number of uncontested districts rises, the imputations from the two models have room to diverge. Where the two sets of imputations are available for a given election (elections where presidential vote shares by state legislative districts are available) we generally see a high level of agreement between the two methods.

The two sets of imputations for $V$ correlate at .99 . With only a few exceptions (see Figure 10), the discrepancies are generally small relative to the uncertainty in the imputations themselves. As the proportion of districts with missing data increases, clearly the scope for divergence between the two models increases.

To re-iterate, we prefer the imputations from "Model 1" based on the regressions utilizing presidential vote shares in state legislative districts, and use them
whenever available (i.e., for most states in the analysis, the period 1992-2014). We only rely on "Model 2 " when presidential vote shares are not available. We model the difference between the two sets of imputations, adjusting the "Model 2 " imputations of $V$ to better match what we have obtained from "Model 1 ", had the necessary presidential vote shares by state legislative district been available.


Figure 10: Difference between imputations for $V$ by proportion of uncontested seats. The fitted regression line is constrained to respect the constraint that the imputations must coincide when there are no uncontested seats.

### 8.4 Seat and vote shares in $\mathbf{7 8 6}$ state legislative elections

After imputations for missing data, each election generates a seats-votes $(V, S)$ pair. In Figure 11 we plot all of the $V$ and $S$ combinations over the 786 state elections in the analysis. We also overlay the seats-vote curve corresponding to an efficiency gap of zero. This provides us with a crude, visual sense of how often we see large departures from the zero $E G$ benchmark.

The horizontal lines around each plotted point show the uncertainty associated with each estimate of $V$ (statewide, Democratic, two-party vote share), given the imputations made for uncontested and missing district-level vote shares. Uncontested seats do not generate uncertainty with respect to the party winning the seat, and so the resulting uncertainty is with respect to vote shares, on the horizontal axis in Figure 11.

The efficiency gap in each election is the vertical displacement of each plotted $(V, S)$ point from the orange, zero-efficiency gap line in Figure 1 1. Uncertainty as to the horizontal co-ordinate $V$ (due to imputations for uncontested races) generates uncertainty in determining how far each point lies above or below the orange, zero efficiency gap benchmark.

## 9 The efficiency gap, by state and election

We now turn to the centerpiece of the analysis: assessing variation in the efficiency gap across districting plans.

We have 786 efficiency gap measures in 41 states, spanning 43 election years. These are computed by substituting each state election's estimate of $V$ and the corresponding, observed seat share $S$ into equation 1.

Figure 12 shows the efficiency gap estimates for each state election, grouped by state and ordered by year; vertical lines indicate $95 \%$ credible intervals arising from the fact that the imputation model for uncontested seats induces uncertainty in $V$ and any quantity depending on $V$ such as $E G$ (recall equation 1 ). In many cases the uncertainty in $E G$ stemming from imputation for uncontested seats is small relative to variation in $E G$ both between and within districting plans.

We observe considerable variation in the EG estimates across states and elections. Some highlights:


Figure 11: Democratic seat shares $(S)$ and vote shares $(V)$ in 786 state legislative elections, 1972-2014, in 41 states. Seat shares are defined with respect to single-member districts won by either a Republican or a Democratic candidate, including uncontested districts. Vote shares are defined as the average of districtlevel, Democratic share of the two-party vote, in the same set of districts used in defining seat shares. Horizontal lines indicate $95 \%$ credible intervals with respect to $V$, due to uncertainty arising from imputations for district-level vote shares in uncontested seats. The orange line shows the seats-votes relationship we expect if the efficiency gap were zero. Elections below the orange line have $E G<0$ (Democratic disadvantage); points above the orange line have $E G>0$ (Democratic advantage).

Efficiency gap, by state and year


Figure 12: Efficiency gap estimates in 786 state legislative elections, 1972-2014. Vertical lines cover $95 \%$ credible intervals.

1. estimates of $E G$ range from -0.18 to 0.20 with an average value of -0.005 .
2. The lowest value, -0.18 is from Delaware in 2000. There were 19 uncontested seats in the election to the 41 seat state legislature. Democrats won 15 seats ( $S=15 / 41=36.6 \%$ ). I estimate $V$ to be $52.1 \%$. Via equation 1 , this generates $E G=-0.18$. Considerable uncertainty accompanies this estimate, given the large number of uncontested seats. The $95 \%$ credible interval for $V$ is $\pm 2.03$ percentage points, and the $95 \%$ credible interval for the accompanying $E G$ estimate is $\pm 0.04$.
3. The highest value of $E G$ is 0.20 is from Georgia in 1984. There were 140 uncontested seats in the election to the 180 seat state legislature. Democrats won 154 seats ( $S=154 / 180=85.6 \%$ ). I estimate $V$ to be $57.9 \%$. Again, using equation 1 , this generates $E G=0.2$. Considerable uncertainty also accompanies this estimate, given the large number of uncontested seats. The $95 \%$ credible interval for $V$ is $\pm 1.89$ percentage points, and the $95 \%$ credible interval for the accompanying $E G$ estimate is $\pm 0.04$. Figure 13 contrasts the seats and votes recorded in Georgia against those for the entire data set, putting Georgia's large $E G$ estimates in context.
4. New York has the lowest median $E G$ estimates, ranging from -. 15 (2006) to -. 028 (1984). Statewide $V$ ranges from $53.7 \%$ to $69.2 \%$, but Democrats only win 70 (1972) to 112 (2012) seats in the 150 seat state legislature, so $S$ ranges from .47 to .75 , considerably below that we'd expect to see given the vote shares recorded by Democrats if the efficiency gap were zero. See Figure 15.
5. Arkansas has the highest median $E G$ score by state, .10 ; see Figure 14.
6. Connecticut has the median, within-state median EG score of approximately zero; Figure 16 shows Connecticut's seats and votes have generally stayed close to the $E G=0$ benchmark.
7. Michigan has the third lowest median $E G$ scores by state, surpassed only by New York and Wyoming. Michigan's EG scores range from -. 14 (2012) to .01 (1984). $V$ ranges from $50.3 \%$ to $60.6 \%$, a figure we estimate confidently given low and occasionally even zero levels of uncontested districts
in Michigan state legislative elections. Yet $S$ ranges from $42.7 \%$ (Democrats won 47 out of 110 seats in 2002, 2010 and 2014) to $63.6 \%$ (Democrats won 70 out of 110 seats in 1978). See Figure 17.
8. Wisconsin's EG estimates range from -. 14 (2012) to .02 (1994). Although the $E G$ estimates for WI are not very large relative to other states in other years, Wisconsin has recorded an unbroken run of negative $E G$ estimates from 1998 to 2014 and records two very large estimates of the efficiency gap in elections held under its current plan: -. 13 (2012) and -. 10 (2014). In short, Democrats are underperforming in state legislative elections in Wisconsin, winning fewer seats than a zero efficiency gap benchmark would imply, given, their statewide level of support. See Figure 18.

### 9.1 Are efficiency gap estimates statistically significant?

Recall that $E G<0$ means that Democrats are disadvantaged, with relatively more wasted votes than Republicans; conversely $E G>0$ means that Democrats are the beneficiaries of an efficiency gap, in that Democrats have fewer wasted votes than Republicans. But EG does vary from election to election, even with the same districting plan in place and $E G$ is almost always not measured perfectly, but is estimated with imputations for uncontested seats.

In Figure 19 we plot the imprecision of each efficiency gap estimate (the halfwidth of its $95 \%$ credible interval) against the estimated $E G$ value itself. Points lying inside the cones have $E G$ estimates that are small relative to their credible intervals, such that we would not distinguish them from zero at conventional levels of statistical significance. Not all $E G$ estimates can be distinguished from zero at conventional levels of statistical significance, nor should they. But many estimates of the EG are unambiguously non-zero. Critically, the two most recent Wisconsin EG estimates (-. 13 in 2012, -. 10 in 2014) are clearly non-negative, lying far away from the "cone of ambiguity" shown in Figure 19; the $95 \%$ credible interval for the 2012 estimates runs from -. 146 to -.121 and from -.113 to -.081 for the 2014 estimate.

Democratic seat shares by vote shares, 1972-2014: Georgia in red, 2014 solid point


Figure 13: Georgia, Democratic seat share and average district two-party vote share, 1972-2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts.

Democratic seat shares by vote shares, 1972-2014: Arkansas in red, 2014 solid point


Figure 14: Arkansas, Democratic seat share and average district two-party vote share, 1992-2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts.

Democratic seat shares by vote shares, 1972-2014: New York in red, 2014 solid point


Figure 15: New York, Democratic seat share and average district two-party vote share, 1972-2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts.

Democratic seat shares by vote shares, 1972-2014: Connecticut in red, 2014 solid point


Figure 16: Connecticut, Democratic seat share and average district two-party vote share, 1972-2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts.

Democratic seat shares by vote shares, 1972-2014: Michigan in red, 2014 solid point


Figure 17: Michigan, Democratic seat share and average district two-party vote share, 1972-2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts.

Democratic seat shares by vote shares, 1972-2014: Wisconsin in red, 2014 solid point


Figure 18: Wisconsin, Democratic seat share and average district two-party vote share, 1972-2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts.


Figure 19: Uncertainty in the efficiency gap, against the $E G$ estimate itself. The vertical axis is the half-width of the $95 \%$ credible interval for each $E G$ estimate (plotted against the horizontal axis); points lying inside the cones have $E G$ estimates that are small relative to their credible intervals, such that we would not distinguish them from zero at conventional levels of statistical significance. EG estimates from Wisconsin in 2012 and 2014 are shown as red points in the lower panel. Note the greater prevalence of large, negative and precisely estimated $E G$ measures in recent decades.

### 9.2 Over-time change in the efficiency gap

Are large values of the efficiency gap less likely to be observed in recent decades? This is relevant to any discussion of a standard by which to assess redistricting plans. If recent decades have generally seen smaller values of the efficiency gap relative to past decades, then this might be informative as to how we should assess contemporary districting plans and their corresponding values of the EG.

Figure 20 plots $E G$ estimates over time, overlaying estimates of the smoothed, weighted quantiles ( 25 th, 50 th and 75 th) of the $E G$ measures (the weights capture the uncertainty accompanying each estimate of the $E G$ ). The distribution of $E G$ measures in the 1970s and 1980s appeared to slightly favor Democrats; about two-thirds of all $E G$ measures in this period were positive. The distribution of $E G$ measures trends in a pro-Republican direction through the 1990s, such that by the 2000s, EG measures were more likely to be negative (Republican efficiency advantage over Democrats); see Figure 21.

There is some evidence that the 2010 round of redistricting has generated an increase in the magnitude of the efficiency gap in state legislative elections. For most of the period under study, there seems to be no distinct trend in the magnitudes of the efficiency gap over time; see Figure 22. The median, absolute value of the efficiency gap has stayed around 0.04 over much of the period spanned by this analysis; elections since 2010 are producing higher levels of $E G$ in magnitude.

It is also interesting to note that the estimate of the 75 th percentile of the distribution of $E G$ magnitudes jumps markedly after 2010, suggesting that districting plans enacted after the 2010 census are systematically more gerrymandered than in previous decades. Of the almost 800 EG estimates in the analysis, spanning 42 years of elections, the largest, negative estimates (an efficiency gap disadvantaging Democrats) are more likely to be recorded in the short series of elections after 2010. These include Alabama in 2014 (-.18), Florida in 2012 (-.16), Virginia in 2013 (-.16), North Carolina in 2012 (-.15) and Michigan in 2012 (-.14); these five elections are among the 10 least favorable to Democrats we observe in the entire set of elections. Among the 10 most pro-Democratic EG scores, none were recorded after 2000. The most favorable election to Democrats in terms of $E G$ since 2010 is the 2014 election in Rhode Island ( $E G=.12$ ), which is only the 20th largest (pro-Democratic) $E G$ in the entire analysis.


Figure 20: Efficiency gap estimates, over time. The lines are smoothed estimates of the 25 th, 50 th and 75 th quantiles of the efficiency gap measures, weighted by the precision of each $E G$ measure.


Figure 21: Proportion of efficiency gap measures that are positive, by two year intervals.


Figure 22: Absolute value of efficiency gap measures, over time. The lines are smoothed estimates of the 25 th, 50 th and 75 th quantiles of the absolute value of the efficiency gap measure, weighted by the precision of each $E G$ measure.

### 9.3 Within-plan variation in the efficiency gap

The efficiency gap is measured at each election, with a given districting plan typically generating up to five elections and hence five efficiency gap measures. Efficiency gap measures will change from election to election as the distribution of district-level vote shares varies over elections. Some of this variation is to be expected. Even with the same districting plan in place, districts will display "demographic drift," gradually changing the political complexion of those districts. Incumbents lose, retire or die in office; sometimes incumbents face major opposition, sometimes they don't. Variation in turnout - most prominently, from on-year to off-year - will also cause the distribution of vote shares to vary from election to election, even with the districting plan unchanged. All these electionspecific factors will contribute to e lection-to-election variation in the efficiency gap.

Precisely because we expect a reasonable degree of election-to-election variation in the efficiency gap, we a ssess the magnitude of this "within-plan" variability in the measure. If a plan is a partisan gerrymander - with a systematic advantage for one party over the other - then the "between-plan" variation in $E G$ should be relatively large relative to the "within-plan" variation in $E G$.

About $76 \%$ of the variation in the $E G$ estimates is between-plan variation. The EG measure does vary election-to-election, but there is a moderate to strong "plan-specific" component to variation in the EG s cores. We conclude that the efficiency gap is measuring an enduring feature of a districting plan.

We examine some particular districting plans. The 786 elections in this analysis span 150 districting plans. For plans with more than one election, we compute the standard deviation of the sequence of election-specific $E G$ measures observed under the plan. These standard deviations range from .011 (Kentucky's plan in place for just two elections in 1992 and 1994, or Indiana's plan 1992-2000) to .079 (Delaware's plan between 2002 and 2010).

A bighly variable plan: Deleware 2002-2010. Figure 23 shows the seats, votes and EG estimates produced under the Delaware 2002-2010 plan. This is among the most variable plans we observe with respect to the $E G$ measure. An efficiency gap running against the Democrats for 2002, 2004 and 2006 (the latter election saw Democrats win only 18 seats out of 41 with $54.5 \%$ of the state wide vote) falls to a small gap in $2008(V=0.584, S=25 / 41=.61, E G=-0.058)$ and

Delaware ends the decade with a positive efficiency gap in 2010. The Democratic district-average two-party vote share fell to $V=0.561$ in 2010, but translated into $S=26 / 41=.63, E G=0.012$.

A plan with moderate variability in the EG. The median, within-plan standard deviation of the $E G$ is about .03 . This roughly corresponds to the within-plan standard deviation of the $E G$ observed under the plan in place for five Wisconsin state legislative elections 1992-2000, presented in Figure 24. This was a plan that generated relatively small values of $E G$ that alternated sign over the life of the plan: negative in 1992, positive in 1994 and 1996, and negative in 1998 and 2000.

A low variance plan, Indiana 1992-2000. See Figure 25. The EG measures recorded under this plan are all relatively small and positive, ranging from 0.008 to 0.041 and correspond to an interesting period in Indiana state politics. Democrats won 55 of the 100 seats in the Indiana state house in the 1992 election with what I estimate to be just over $50 \%$ of the district-average vote (29 of 100 seats were uncontested). Democratic vote share fell to about $45 \%$ in the 1994 election ( 38 uncontested seats), and Democrats lost control of the legislature. The 1996 election resulted in a $50-50$ split in the legislature. Democrats won legislative majorities in the 1998 and 2000 elections, while the last election might have been won by Democrats with just less than $50 \%$ of the district-vote; I estimate $V=0.495 \pm .012$ and $E G=0.041$.

Highlighting Delaware plan 4


Figure 23: Seats, votes and the efficiency gap recorded under the Delaware plan, 2002-2010. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts. The inset in the lower right shows the sequence of efficiency gap measures recorded under the plan; vertical lines are $95 \%$ credible intervals.

Highlighting Wisconsin plan 3


Figure 24: Seats, votes and the efficiency gap recorded under the Wisconsin plan, 1992-2000. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts. The inset in the lower right shows the sequence of efficiency gap measures recorded under the plan; vertical lines are $95 \%$ credible intervals.

Highlighting Indiana plan 3


Figure 25: Seats, votes and the efficiency gap recorded under the Indiana plan, 1992-2000. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts. The inset in the lower right shows the sequence of efficiency gap measures recorded under the plan; vertical lines are $95 \%$ credible intervals.

### 9.4 How often does the efficiency gap change sign?

Having observed a particular value of $E G$, how confident are we that:

- the $E G$ measure is distinguishable from zero at conventional levels of statistical significance? That is, how sure are we as to the sign of any particular $E G$ estimate? We addressed this question in section 9.1.
- it will be followed by one or more estimates of $E G$ that are of the same sign?
- over the life of a districting plan, $E G$ remains on one side of zero or the other?

The latter two questions are key. It is especially important that we assess the durability of the sign of the $E G$ measure under a districting plan, if we seek to assert that a districting plan is a partisan gerrymander. We will see that magnitude and durability of the efficiency gap go together: large values of the efficiency gap don't seem to be capricious, but likely to be repeated over the life of a districting plan, consistent with partisan disadvantage being a systematic feature of the plan.

We begin this part of the analysis by considering temporally adjacent pairs of $E G$ estimates. Can we be confident that these have the same sign? In general, yes. Of the full set of 786 elections for which we compute an efficiency gap estimate, 580 are temporally adjacent, within state and districting plan. Figure 26 shows that we usually see efficiency gap measures with the same sign; this probability exceeds $90 \%$ for almost half of the temporally adjacent pairs of efficiency gap measures. Averaged over all pairs, this "same sign" probability is $74 \%$. While the efficiency gap does vary election to election, these fluctuations are not so large that the sign of the efficiency gap is likely to change election to election.

What about over the life of an entire redistricting plan? How likely is it that the efficiency gap retains the same sign over, say, three to five elections in a given state, taking into account election-to-election variation and uncertainty arising from the imputation procedures used for uncontested districts?

We have 141 plans that supply three or more elections with estimate of the efficiency gap. Of these, 17 plans are utterly unambiguous with respect to the sign of the efficiency gap estimates recorded over the life of the plan: for each of these plans we estimate the probability that the $E G$ has the same sign over the life of the plan to be $100 \%$. These plans are listed below in Table 1.

Probabilities that efficiency gap has the same sign as in previous election


Figure 26: Stability in 580 successive pairs of efficiency gap measures

| State | Plan | Start | End | EG avg | EG min | EG max |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Florida | 4 | 2002 | 2010 | -0.112 | -0.136 | -0.084 |
| New York | 4 | 2002 | 2010 | -0.111 | -0.150 | -0.078 |
| Illinois | 3 | 1992 | 2000 | -0.103 | -0.136 | -0.058 |
| Michigan | 4 | 2002 | 2010 | -0.103 | -0.130 | -0.077 |
| New York | 3 | 1992 | 2000 | -0.098 | -0.139 | -0.048 |
| New York | 1 | 1972 | 1980 | -0.097 | -0.108 | -0.079 |
| Missouri | 4 | 2002 | 2010 | -0.091 | -0.142 | -0.061 |
| Ohio | 4 | 2002 | 2010 | -0.090 | -0.143 | -0.049 |
| New York | 2 | 1982 | 1990 | -0.084 | -0.120 | -0.028 |
| Ohio | 3 | 1994 | 2000 | -0.083 | -0.109 | -0.025 |
| Michigan | 3 | 1992 | 2000 | -0.080 | -0.128 | -0.019 |
| Wisconsin | 4 | 2002 | 2010 | -0.076 | -0.118 | -0.039 |
| Colorado | 2 | 1982 | 1990 | -0.075 | -0.117 | -0.055 |
| Colorado | 1 | 1972 | 1980 | -0.041 | -0.067 | -0.018 |
| California | 3 | 1992 | 2000 | -0.041 | -0.057 | -0.018 |
| Pennsylvania | 2 | 1982 | 1990 | -0.033 | -0.056 | -0.020 |
| Florida | 1 | 1972 | 1980 | 0.070 | 0.052 | 0.099 |

Table 1: Plans with no doubt as to the sign of the efficiency gap over the life of the plan (3+ elections).

Interestingly, these plans with an utterly unambiguous history of one-sided EG measures are almost all plans with efficiency gaps that are disadvantagous to Democrats. Michigan's 2002-2010 plan is on this list, as is the plan in place in Wisconsin 2002-2010 (average EG of -.076).

We examine this probability of " $3+$ consecutive $E G$ measures with the same sign" for all of the plans with 3 or more elections in this analysis. $35 \%$ of 141 plans with 3 or more elections have at least a $95 \%$ probability of recording plans with $E G$ measures with the same sign. If we relax this threshold to $75 \%$, then $46 \%$ of plans with 3 or more elections exhibit EG measures with the same sign. Again, there is a reasonable amount of within-plan movement in $E G$, but in a large proportion of plans the efficiency gap appears to be a stable attribute of the plan.

## 10 A threshold for the efficiency gap

We now turn to the question of what might determine a threshold for determining if the EG is a large and enduring characteristic of a plan. We pose the problem as follows:
for a given threshold $E G^{*}>0$, what is the probability that having observed a value of $E G \geq E G^{*}$ we then see $E G<0$ in the remainder of the plan?

To answer this we compute

- if (and optionally, when) a plan has $E G \geq E G^{*}$;
- conditional on seeing $E G \geq E G^{*}$, do we also observe $E G<0$ (a sign flip) in the same districting plan?

For $E G<0$, the computations are reversed: conditional on seeing $E G<E G^{*}$, do we also see $E G>0$ under the same plan?

Figure 27 displays two proportions, plotted against a series of potential thresholds on the horizontal axis. The two plotted proportions are:

- the proportion of plans in which we observe an $E G$ more extreme than the specified threshold $E G^{*}$ (on the horizontal axis);
- among the plans that trip the specified threshold, the proportion in which we see a $E G$ in the same plan with a different sign to $E G^{*}$.

Plans with at least one election with $|E G|>.07$ are reasonably common: over the entire set of plans analyzed here - and again, with the uncertainty in EG estimates taken into account - there is about a $20 \%$ chance that a plan will have at least one election with $|E G|<.07$.

Observing $E G>.07$ is not a particularly informative signal with respect to the other elections in the plan. Conditional on observing an election with $E G>.07$ (an efficency gap favoring Democrats), there is an a $45 \%$ chance that under the same plan we will observe $E G<0$. That is, making an inference about a plan on the basis of one election with $E G>.07$ would be quite risky. Estimates


Figure 27: Proportion of plans that (a) record an efficiency gap measure at least as extreme as the value on the horizontal axis; and (b) conditional on at least one election with $E G$ in excess of this threshold (not necessarily the first election), the proportion of plans where there is another election in the plan with an $E G$ of the opposite sign.
of the "sign flip" rate conditional on a plan generating a relatively large, proDemocratic EG estimates are quite unreliable because there are so few plans generating large, pro-Democratic EG estimates to begin with; note the confidence intervals on the "sign flip" rate get very wide as the data become more scarce on the right hand side of the graph.

This finding is not s ymmetric. The " signal" $E G<-.07$ ( an e fficiency gap disadvantageous to Democrats) is much more informative about other elections in the plan than the opposite signal $E G>.10$ (a pro-Democratic efficiency gap). If any single election in the plan has $E G<-.07$ then the probability that all elections in the plan have $E G<0$ is about .80 . That is, there is a smaller degree of within-plan volatility in plans that disadvantage Democrats. Observing a relatively low value of the $E G$ such as $E G<-.07$ is much more presumptive of a systematic and enduring feature of a redistricting plan than the opposite signal $E G>$.07. Efficiency gap measures that a ppear to indicate a disadvantage for Democrats are thus more reliable signals about the respective districting plan than efficiency gap measures indicating an advantage for Democrats.

We repeat this previous exercise, but restricting attention to more recent elections and plans, with the results displayed in Figure 28. Again we see that plans with pro-Democratic $E G$ measures are quite likely to also generate an election with $E G<0$; and again, note that estimates of the "sign flip" $r$ ate a re quite unreliable because there are so few plans generating large, pro-Democratic EG estimates to begin with.

- Proportion of plans with EG in excess of threshold
- Proportion of plans exceeding threshold that have an EG with opposite sign


Figure 28: Proportion of plans in which (a) the efficiency gap measure is at least as extreme as the value on the horizontal axis; and (b) of these plans with at least one election with $E G$ in excess of this threshold (not necessarily the first election), the proportion of plans in which there is another election in the plan with an EG of the opposite sign. Analysis of state legislative elections in 129 plans, 1991-present.

### 10.1 Conditioning on the first election in a districting plan

We also compute this probability of a sign flip in $E G$ conditional on the magnitude of the $E G$ observed with the first e lection u nder a districting p lan. We perform this analysis twice: (1) for all elections in the data set and (2) for elections held under plans adopted in 1991 or later.

Figures 29 and 30 display the results of these analyses. First, over the full set of data (Figure 29) we observe a roughly symmetric set of EG scores in the first election under a plan. But we s eldom see plans in the 1990s or later that commence with a large, pro-Democratic efficiency gap; the probability of a first election having $E G>.10$ is zero and the probability of a first election having $E G>.05$ (historically, not a large $E G$ ) is only about $11 \%$. Negative efficiency gaps (not favoring Democrats) are much more likely under the first election in the post-1990 plans: almost $40 \%$ of plans open with $E G<-.05$ and about $20 \%$ of plans open with $E G<-10$.

As noted earlier, pro-Democratic efficiency gaps seem much more fleeting than pro-Republican efficiency g aps. C onditional on a p ro-Republican estimate of $E G>0$ in the first election under a plan, the probability of seeing $E G$ change sign over the life of the plan is almost always around $40 \%$ (1972-2014, Figure 29) or 50\% (1991-present, Figure 30).

A very different conclusion holds if the first election observed u nder a plan indicates a sizeable efficiency gap working to disadvantage D emocrats. In fact, the more negative the initial $E G$ observed under a plan, the more confident we can be that we will continue to observe $E G<0$ over the sequence of elections to follow under the plan. Conditional on a first e lection with $E G<-.10$, the probability of all subsequent efficiency gaps being negative is about $85 \%$. Indeed, it is more likely than not that if the first e lection $h$ as $E G<0$ (no matter how small), then so too will all subsequent elections (a $60 \%$ chance of this event).

Note that the Current Wisconsin Plan opens with $E G=-.13$ in the 2012 election. Analysis of efficiency gap measures in the post-1990 era (Figure 30) indicates that conditional on an EG measure of this size and sign, there is a $100 \%$ probability that all subsequent elections held under that plan will also have efficiency $g$ aps d isadvantageous to $D$ emocrats. That is, in the post-1990 e ra, if a plan's first election yields $E G \leq-.13$, we never see a subsequent election under that plan yielding a pro-Democratic efficiency $g$ ap. In short, a signal such as

- Proportion of plans with EG in excess of threshold
- Proportion of plans exceeding threshold that have an EG with opposite sign


Figure 29: Proportion of plans in which the first election (a) has an efficiency gap measure at least as extreme as the value on the horizontal axis; and (b) conditional on the first election having an $E G$ in excess of this threshold, the proportion of those plans in which a subsequent election has an $E G$ of the opposite sign. Analysis of all state legislative elections in all plans with more than one election, 1972-present.

- Proportion of plans with EG in excess of threshold
- Proportion of plans exceeding threshold that have an EG with opposite sign


Figure 30: Proportion of plans in which the first election (a) has an efficiency gap measure at least as extreme as the value on the horizontal axis; (b) conditional on the first election having an $E G$ in excess of this threshold, the proportion of those plans in which a subsequent election has an EG of the opposite sign. Analysis of state legislative elections in 129 plans, 1991-present.
$E G \leq-.13$ is extremely reliable with respect to the districting plan that generated it, at least given the post-1990 record.

### 10.2 Conditioning on the first two elections in a districting plan

The difficulty with conditioning on the first two elections of a districting plan is that the data start to thin out. In the entire data set there simply aren't many districting plans that equal or surpass the two, relatively large values of $E G$ observed in Wisconsin in the first two elections of the current plan. Indeed, the only cases with a similar history of EG measures like Wisconsin's in 2012 and 2014 are contemporaneous cases: Florida, Michigan, and North Carolina in 2012 and 2014.

We relax the threshold of what counts as a similar case to encompass plans whose first two efficiency gap measures are within $75 \%$ of the magnitude of Wisconsin's 2012 and 2014 EG measures; we now pick up 11 roughly comparable cases, 4 of which date from earlier decades. Again, this is testament to how recent decades have seen an increase in the prevalence of larger, negative values of the efficiency gap.

For the four prior cases we plot the sequence of $E G$ estimates in Figure 31. With the exception of the last election in the highly unusual Delaware sequence (among the most volatile observed in the data set; see section 9.3), the other proximate cases all go on to record efficiency gap measures that a re below zero over the balance of the plan. We stress that four cases doesn't provide much basis for comparison, but this only speaks to the fact that the sequence of two large, negative values of the efficiency gap in Wisconsin in 2012 and 2014 are virtually without historical precedent. We have little guidence from the historical record as to what to expect given an opening sequence of $E G$ measures like the ones observed in Wisconsin. But the little evidence we do have suggests that a stream of similarly sized, negative values of the efficiency gap a re quite likely o ver the balance of the districting plan.

### 10.3 An actionable EG threshold?

We now consider a more general question: what is an actionable threshold for the efficiency gap?


Figure 31: Sequence of $E G$ estimates observed over the life of districting plans, for pre-2010 plans with first two EG scores within $75 \%$ of the magnitude of the EG scores observed in Wisconsin in 2012 and 2014.

First, recall that relatively small EG estimates are likely to be swamped by their estimation uncertainty, depending on the proportion of uncontested districts in the given election and the statistical procedures. In every instance though, this is an empirical question; at least in the approach I present here, each $E G$ estimate I generate is accompanied with uncertainty bounds, letting us assess the probability that a given estimate is positive or negative. Figure 19 provides a summary of the relationship between the size of the EG estimate and the "statistical significance" of the estimate (in the sense that the $95 \%$ credible interval for each estimate does not overlap zero).

Second, the distribution of EG statistics in the 1972-2014 period is roughly symmetric around zero. Reference to this empirical distribution might also be helpful in setting actionable thresholds, and answering the question "is the $E G$ measure at issues large relative to those observed in the previous 40 years of state legislative elections?" Double digit EG measures (-. 10 or below; .10 or above) are pushing out into the extremes of the observed distribution of $E G$ estimates: $E G$ estimates of this magnitude are comfortably past the question of "statistical significance." Just $15 \%$ of the 786 E G measures generated in this a nalysis are below -.07 ; fewer than $12 \%$ are greater than .07 .

We do need to be careful when making these kinds of relative assessments about the magnitude of the efficiency $g$ ap. If pro-Republican gerrymandering is widespread, then it will be less unusual to see a large, negative EG estimate, at least contemporaneously; in fact this appears to the case in the post-2010 set of elections, where the longer-term distinctiveness of the Wisconsin numbers is matched and in some cases exceeded by other states also recording unusually large, negative EG estimates (e.g., Florida, Michigan, Virginia and North Carolina). This speaks to the utility of the longer-term, historical analysis in both Stephanopolous and McGhee (2015) and in this report. It it is important to remember that $E G=0$ corresponds to a partisan symmetry in wasted vote rates; we should be wary of arguments that would lead us to tolerate small to moderate levels of the efficiency gap because they appear to be the norm in some period of time, or in some set of jurisdictions.

In any litigation, much will turn on the question of durability in the efficiency gap, and this concern motivates much of the preceeding analysis. We cannot wait until three, four, or more elections have transpired under a plan in order to
assess its properties. Courts will be asked to assess a plan based on only one $E G$ estimate, or two. Analysis of the sort I provide here will be informative in these cases, assessing whether the estimate is so large that the historical record suggests that the first election's $E G$ estimate is a reliable indicator as an enduring feature of the plan, and not an election-specific aberration.

### 10.4 Confidence in a given threshold

Figures 32 and 33 present my estimate of a "confidence rate" associated with a range of possible "actionabale thresholds" for the efficiency $g$ ap. These figures essentially re-package the information shown in Figures 29 and 30. Suppose a court rejects or amends every plan with a first election $E G$ more extreme (further away from zero) than the proposed threshold shown on the horizontal axis of these graphs. A certain number of plans fail to trip this threshold, and so are upheld by the courts if they are challenged. Of those that do trip the threshold and are rejected by a court, what is our confidence that the plan, if left undisturbed, would go on to produce a sequence of $E G$ measures that lie on the same side of zero as the threshold? Combining these two proportions gives us an overall confidence measure a ssociated with a particular threshold.

This analysis points to a benchmark of about -.06 or -.07 as the actionable threshold given a first election with $E G<0$ (Democratic disadvantage) or .08 or .09 when we observe $E G>0$ in the first election under a redistricting plan (Democratic advantage); the asymmetry here reflects the fact that districting plans evincing apparent Democratic advantages are not as durable or as common (in recent decades) as plans presenting evidence of pro-Republican gerrymanders. At these proposed benchmarks the overall confidence rates are estimated to be $95 \%$, with this confidence rate corresponding to a benchmark used widely in statistical decision-making in many fields of science.

Figures 32 and 33 also highlight that $E G<-.07$ or $E G>.07$ would be an extremely conservative threshold. On the pro-Democratic side, $E G>.07$ is a rare event. Districting plans unfavorable to Democrats, with $E G<-.07$ are not unusual; about $10 \%$ of post-1990 plans generate $E G$ measures below -.07; the proportion of these plans that then record a sign flip is only about $10 \%$; see Figure 30. If the presumption was that any plan with a first election showing


Figure 32: Proportion of plans being either (a) undisturbed or (b) if left undisturbed, would continue to produce one-sided partisan advantage (no sign change in subsequent $E G$ measures), as a function of the proposed "first election," efficiency gap threshold (horizontal axis), based on analysis of all multi-election districting plans, 1972-2014. The proportion on the vertical axis is thus interpretable as the "confidence level" associated with intervention at a given first election, $E G$ threshold. Vertical lines indicate $95 \%$ credible intervals.


Figure 33: Proportion of plans being either (a) undisturbed or (b) if left undisturbed, would continue to produce one-sided partisan advantage (no sign change in subsequent $E G$ measures), as a function of the efficiency gap threshold (horizontal axis), based on analysis of post-1990 plans and elections. The proportion on the vertical axis is thus interpretable as the "confidence level" associated with intervention at a given first election, $E G$ threshold. Vertical lines indicate $95 \%$ credible intervals.
$E G<-.07$ would be rejected, then we'd be "wrong" to do so in about $10 \%$ of those cases (in the sense that if left in place, the plan would go on to produce at least one election with $E G>0$ ). The total error rate in this case would be $1 \%$ of all plans. Equivalently, $99 \%$ of all plans would be either left undisturbed or appropriately struck down or amended by a court, given the historical relationship between "first election" $E G$ measures and the sequence of $E G$ measures that follow.

## 11 Conclusion: the Wisconsin plan

Wisconsin has had two elections for its legislature under the plan currently in place, in 2012 and 2014. Both elections were subject to considerable rates of uncontestedness ( 27 of 99 seats in 2012 and 52 of 99 seats in 2014), but these rates are hardly unusual; Wisconsin's rates of uncontested districts in these two elections are low to moderate compared to other states. We use the relationship between state legislative election results and presidential election results in state legislative districts (and incumbency) to impute two-party vote shares in uncontested seats (see section 7.2). With a complete set of vote shares, we then compute average district-level Democratic two-party vote share $(V)$ and note the share of seats (contested and uncontested) won by Democratic candidates $(S)$.

In Wisconsin in 2012, and after imputations for uncontested seats, $V$ is estimated to be $51.4 \%( \pm 0.6)$; recall that Obama won $53.5 \%$ of the two-party presidential vote in Wisconsin in 2012. Yet Democrats won only 39 seats in the 99 seat legislature ( $S=39.4 \%$ ), making Wisconsin one of 7 states in 2012 where we estimate $V>50 \%$ but $S<50 \%$ and where Democrats failed to win a majority of legislative seats despite $V>50$ (the other states are Florida, Iowa, Michigan, North Carolina and Pennsylvania). In 2014, V is estimated to be $48.0 \%( \pm 0.8)$ and Democrats won 36 of 99 seats ( $S=36.4 \%$ ).

This provides the raw ingredients for computing the efficiency gap ( $E G$ ) for these two elections (recalling equation 1). Repeating these calculations across a large set of state elections provides a basis for assessing whether the efficiency gap estimates for Wisconsin in 2012 and 2014 are noteworthy.

Wisconsin's efficiency gap measures in 2012 and 2014 are - .13 and -. 10 (to two digits of precision). These negative estimates indicate the disparity between

Highlighting Wisconsin plan 5


Figure 34: Seats, votes and the efficiency gap recorded under the Wisconsin plan, 2012 and 2014. Orange line shows the seats-votes curve if the efficiency gap were zero; the efficiency gap in any election is the vertical distance from the corresponding data point to the orange line. Gray points indicate elections from other states and elections (1972-2014). Horizontal lines cover a $95 \%$ credible interval for Democratic average district two-party vote share, given imputations in uncontested districts. The inset in the lower right shows the sequence of efficiency gap measures recorded under the plan; vertical lines are $95 \%$ credible intervals.
vote shares and seat shares in these elections, which in turn, is consistent with partisan gerrymandering. The negative EG estimates generated in 2012 and 2014 are unusual relative to Wisconsin's political history (see Figure 35). The 2012 estimate is the largest EG estimate in Wisconsin over the 42 year period spanned by this analysis (1972-2014); the 2014 estimate is the fourth largest (behind 2012, 2006 and 2004, although it is essentially indistinguishable from the 2004 estimate). The jump from the $E G$ values being recorded towards the end of the previous districting plan in Wisconsin (2002-2010) to the 2012 and 2014 values strongly suggests that the districting plan adopted in 2011 is a driver of the change, systematically degrading the efficiency with which Democratic votes translate into Democratic seats in the Wisconsin state legislature.

Wisconsin's 2012 and 2014 EG estimates are also large relative to the EG scores being generated contemporaneously in other state legislative elections. Figure 36 shows $E G$ estimates recorded under plans in place since the post-2010 census round of redistricting; the $E G$ estimates are grouped by state and ordered, with Wisconsin highlighted. We have 78 EG scores from elections held since the last round of redistricting. Among these 79 scores, Wisconsin's EG scores rank eigth (2012, $95 \%$ CI 3 to 12 ) and seventeenth ( $2014,95 \%$ CI 13 to 20).

The historical analysis reported above supports the proposition that Wisconsin's $E G$ scores are likely to endure over the course of the plan. Few states ever record EG scores as large as those observed in Wisconsin; indeed, there is virtually no precedent for the lop-sided, two election sequence of $E G$ scores generated in Wisconsin in 2012 and 2014 in the data I analyze here (1972-2014). The closest historical analogs suggest that a districting plan that generates an opening, two-election sequence of $E G$ scores like those from Wisconsin will continue to do so, generating seat shares for Democrats that are well below those we would expect from a neutral plan.

The Current Wisconsin Plan is generating estimates of the efficiency gap far in excess of the proposed, actionable threshold (see section 10). In 2012 elections to the Wisconsin state legislature, the efficiency gap is e stimated to be-.13; in 2014, the efficiency $g$ ap is e stimated to be - $\mathbf{1 0}$. B oth measures a re separately well beyond the conservative .07 threshold suggested by the analysis of efficiency gap measures observed from 1972 to the present.


Figure 35: History of efficiency gap estimates in Wisconsin, 1972-2014. Vertical lines indicate $95 \%$ credible intervals.


Figure 36: EG estimates in 2012 and 2014, grouped by state and ordered. Horizontal bars indicate $95 \%$ credible intervals.

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# Rebuttal Report 

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## December 21, 2015

## Introduction

In this rebuttal report, I respond to criticisms made by Sean P. Trende and Professor Nicholas Goedert in their respective expert reports. I also conduct new empirical analyses further confirming the validity of the efficiency gap as a measure of partisan gerrymandering and the reasonableness of the proposed 0.07 threshold. More specifically, my principal contributions are the following:

- First, I respond to Goedert's various critiques of the efficiency gap and of the proposed efficiency gap threshold. Among other things, he misunderstands the relevance of efficiency gap data, cherry-picks information from my initial report while ignoring its broader context, and wrongly claims that plaintiffs' test would mandate "hyper-responsiveness" or prevent states from pursuing goals such as competitiveness or proportional representation.
- Second, I calculate several widely accepted prognostic measures-all based on the rates of true positives, false positives, true negatives, and false negatives-with respect to the odds of a district plan's efficiency gap changing signs over the plan's lifetime given a certain efficiency gap value in the plan's first election. Based on these measures, I conclude that the proposed 0.07 threshold is highly conservative. In fact, this threshold sacrifices some accuracy (which would be maximized at a lower threshold) in order to reduce the proportion of false positives.
- Third, I calculate the same prognostic measures with respect to the odds of a district plan's average efficiency gap, over its lifetime, having a different sign than that observed in the first election under a plan, given a certain efficiency gap value in this first election. Under this method, the proposed 0.07 threshold appears even more conservative, driving down the share of false positives to below $5 \%$.
- Fourth, I compare the values of the efficiency gap in the first election under a plan and on average over the plan's lifetime. This relationship is impressively tight $\left(r^{2}=0.73\right)$, indicating that a plan's initial bias is a very good predictor of its overall lifetime bias. For Act 43, this analysis allows us to predict that it will average a pro-Republican efficiency gap of almost $10 \%$ over the 2010 cycle as a whole.
- Fifth, I examine to what extent changes in party control over redistricting are responsible for the pro-Republican trend in the efficiency gap since the 1990s. In the current cycle, about four times more state house plans were designed by Republicans in full control of state government than in the 1990s. Had the distribution of party control over redistricting remained unchanged, essentially all of the pro-Republican movement in the efficiency gap over the last two decades
would not have occurred. It is thus changes in party control, and not changes in the country's political geography, that primarily account for Republicans' growing redistricting advantage over the last generation.
- Sixth, I address recent work by Chen and Rodden (2013), cited by both Trende and Goedert for the proposition that Republicans enjoy a natural geographic advantage over Democrats. Chen and Rodden's simulated maps are not lawful because they ignore the Voting Rights Act and state redistricting criteria; they are based on presidential election results rather than more relevant state legislative election results; they do not constitute a representative sample of the entire plan solution space; and they are contradicted by other recent work (Fryer \& Holden 2011) finding that randomly drawn plans reduce bias and increase electoral responsiveness.
- Lastly, I comment on Trende's analysis of particular state legislative and congressional plans. This analysis is marked by conceptual and methodological errors severe enough to render it useless. For example, Trende ignores two of the three prongs of plaintiffs' proposed test; he calculates congressional efficiency gaps without converting them from percentage points to House seats and for House delegations too small to generate reliable estimates; and he simply substitutes presidential election results for congressional election results whenever the latter are missing due to uncontested races. None of this work meets accepted standards of social science rigor.


## 1 Responses to Goedert's criticisms

In his report, Goedert offers several critiques of the efficiency gap and of the 0.07 threshold I recommended in my initial report, based primarily on the alleged instability of the efficiency gap. None of these critiques have merit. In this section, I respond to Goedert's points relying only on the analysis of my initial report and on the existing literature. My new empirical analyses appear in subsequent sections.

First, Goedert appears to believe that a plan's efficiency gap is only relevant to the extent that it sheds light on the partisan intent (or lack thereof) underlying the plan. He writes that "such intent cannot be inferred" from a large efficiency gap, that "a durable bias . . . is not even a sign of deliberate partisan intent," and that the "efficiency gap [is] a standard to measure partisan intent" (pp. 11, 13, 19). But this is not at all the legal function of the efficiency gap in plaintiffs' proposed test. Rather, partisan intent is its own independent inquiry, and the efficiency gap then comes into play at the second stage of
the test, to determine if a plan's electoral consequences are sufficiently severe that it should be deemed presumptively unconstitutional. To put it simply, the efficiency gap is plaintiffs' measure of partisan effect, not of partisan intent. Goedert's misunderstanding of this basic point infects all of his discussion.

Second, Goedert observes that of all plans, anytime in the decade, with a proDemocratic efficiency gap of greater than 0.07 , a substantial proportion of them switch signs over their lifetimes (p.11). In making this observation, Goedert cherry-picks a single bit of data from my initial report, and an irrelevant piece of data at that. This fact is irrelevant because it applies to plans no matter when their elections were held, while the appropriate universe for plaintiffs, defendants, and courts is limited to the first elections held under plans. It is the first elections that typically will be used in litigation, given Justice Kennedy's admonition in Vieth that plans should not be struck down based on a "hypothetical state of affairs," but rather "if and when the feared inequity arose" (Vieth $v$. Jubelirer (2004), p. 420). And the fact is misleading because it applies only to proDemocratic efficiency gaps above 0.07 , and not to the larger set of pro-Republican efficiency gaps above this threshold.

If we consider only plans that exhibit a pro-Democratic efficiency gap above 0.07 in their first elections, the probability that they will switch signs over their lifetimes drops by about five percentage points (Jackman Report, p. 61). And if we then turn to plans that exhibit a pro-Republican efficiency gap above 0.07 in their first elections-a more sizeable set, for which more accurate estimates are possible-this probability drops all the way to about $15 \%$ (Jackman Report, p. 61). In other words, of plans that open with large pro-Republican efficiency gaps, close to $85 \%$ of them continue to favor Republicans in every election for the remainder of the cycle. This is the most pertinent data point in my report, not the one cherry-picked by Goedert, and it reveals the persistence of many gerrymanders.

Third, Goedert discusses congressional district plans throughout his report, even though this case is exclusively about state legislative redistricting (pp. 7-8, 10, 12, 20). In doing so, he makes some of the same errors as does Trende: namely, not converting the efficiency gap from percentage points to House seats, and improperly handling uncontested races (in his case, by not adjusting for the uncontestedness at all, and simply treating the races as if all of the vote went to one party and none to the other). I discuss these errors in more detail later in this report.

Fourth, Goedert claims that it is "arbitrary" to focus on the first election after redistricting, and that doing so "biases toward a finding of $E G$ durability" by ignoring wave elections (p. 14). As noted above, the first election after redistricting is the critical
one for purposes of litigation, since under Vieth, it is after this election that a lawsuit will typically commence and have to be decided by the courts. Later elections are largely irrelevant for litigation purposes, since it is unreasonable to expect suits to be brought six or eight or even ten years into a cycle. Moreover, my analysis in no way ignored wave elections; to the contrary, I determined the odds that a plan's efficiency gap would switch signs by examining all elections held under the plan, waves and non-waves alike. If anything, the fact that most wave elections over the last forty years have not taken place in the first election after redistricting biases against a finding of durability, since these elections may well cause the efficiency gap to flip signs.

Fifth, Goedert is wrong that an efficiency gap of zero represents "'hyperresponsive' representation" (p.2). In fact, as he has recognized in his own prior work, an efficiency gap of zero corresponds almost exactly to the responsiveness actually displayed by American elections over the course of the twentieth century, under which "a $1 \%$ increase in vote share will produce about a $2 \%$ increase in seat share" (Goedert 2014, p. $3)$. Indeed, this correspondence is one of the efficiency gap's most attractive properties, and it explains why Goedert himself calculated a quantity nearly identical to the efficiency gap in his work (Goedert 2014; Goedert 2015).

And sixth, Goedert is wrong as well that plaintiffs' proposed test might discourage states from pursuing worthwhile goals such as competitiveness or proportional representation (pp. 6-10). If a state's aim in redrawing districts was to make them more competitive or to produce more proportional representation, then the partisan intent required by the first prong of plaintiffs' test would not be present. Even if partisan intent were somehow found, the state would likely be able to show that its plan's large efficiency gap was necessitated by its pursuit of competitiveness or proportional representation. And in any event, competitiveness and proportional representation are extremely rare objectives in American redistricting. Only one state, Arizona, has a competitiveness requirement, and not a single state has a proportional representation criterion. (And needless to say, line-drawers do not tend to seek out either of these goals on their own.)

## 2 Reliability of a district plan's first efficiency gap

Having rebutted Goedert's criticisms using preexisting data, I now provide further analysis of the reliability of the first efficiency gap (EG) observed in the life of a district plan. This played a key role in the determination of the threshold $E G$ value in my initial report. In that report, I focused on the probability of a "sign-flip": that is, given the magnitude of the efficiency gap observed in the first election under a district plan, what
can we infer about the likelihood that all subsequent efficiency gaps observed under that plan will have the same sign as that from the first election.

Under this approach, just one election that produces an efficiency gap with a different sign from the efficiency gap in the first election will generate a "failure," in the sense we would say that the plan has generated an efficiency gap that conflicts with that from the first election. In short, the "constant sign" analysis in my original report considers the most extreme set of efficiency gap estimates produced under a plan and insists that they have the same sign. In this sense, the "constant sign" analysis I performed is a quite stringent and conservative test of what we can or ought to infer from the efficiency gap observed in the first election under the district plan. Another approach would be to inquire as to the average efficiency gap over the life of the district plan. A summary statistic such as the average is-by definition-less sensitive to extreme values. At the same time-and again, by definition-the average measures central tendency or typicality, and is the most widely used summary statistic in existence. I thus consider how well the first $E G$ observed under a district plan predicts the average $E G$ observed over the life of the plan.

But I first provide some additional analysis of the prognostic properties of the first efficiency gap observed under a district plan. In each instance the test is whether the first $E G$ observed under a plan exceeds a given threshold value. The outcome of interest is whether the plan's remaining efficiency gaps have the same sign as the $E G$ from the first election. For purposes of this exercise, plans are classified as "positive" (all $E G$ scores under the plan have the same sign) or "negative" (EG scores differ in sign). With these definitions in place, we can then classify plans according to the accuracy of the prediction implicit in the first $E G$ observed under the plan:

|  | Actual |  |
| :--- | :---: | :---: |
| Test | Positive | Negative |
| Positive | True Positive | False Positive |
| Negative | False Negative | True Negative |

The prognostic measures I rely on are conventional measures of predictive or classification accuracy used throughout the quantitative sciences:

1. sensitivity, or the true positive rate: proportion of positives that test positive, TP/(TP + FN)
2. specificity, or the true negative rate: proportion of negatives that test negative, TN/(TN + FP)
3. balanced accuracy, the average of the sensitivity and the specificity
4. accuracy, the proportion of cases that are true positives or true negatives, (TP + $\mathrm{TN}) /(\mathrm{TP}+\mathrm{FP}+\mathrm{FN}+\mathrm{TN})$.
5. the false positive rate; proportion of negative cases that test positive, 1 minus the specificity or $\mathrm{FP} /(\mathrm{TN}+\mathrm{FP})$.
6. the false discovery rate; proportion of cases testing positive that are actually negative, $\mathrm{FP} /(\mathrm{TP}+\mathrm{FP})$.
7. the false omission rate; proportion of cases that test negative that are actually positive, $\mathrm{FN} /(\mathrm{FN}+\mathrm{TN})$.

Figure 1 shows how these prognostic performance indicators vary as a function of the absolute EG threshold (on the horizontal axis in the figure). That is, as we move to the right in each panel of the graph, the test is becoming increasingly stringent: larger absolute values of the efficiency gap in the first election under a district plan are required to trip the increasingly higher threshold. When the threshold is set to zero, all plans trip the threshold (all first-election EGs are greater than zero in magnitude, by definition) and so all cases test positive; in this case the sensitivity is 1 , while conversely the specificity is 0 and the false positive rate is 1 (all negatives test positive).

The test has better properties as the threshold grows, with the accuracy measures maximized around absolute values of .03 to .04 . Yet accuracy is not all in this context. The rate of false positives is quite high at thresholds where the accuracy is high, as is the false discovery rate. At a threshold of .03 , for example, over half of plans that would go on to exhibit sign flips in their EGs would test positive and be flagged for inspection; of the plans selected for scrutiny, more than a third would turn out to have EG sign flips over the life of the plan. The .07 threshold is thus a conservative standard, the point at which the rate of false positives is becoming reasonably low ( $25 \%$ ), without letting the false omission rate go above $50 \%$.

It is worth noting the weight being put on false discoveries or false alarms versus the weight on false omissions in this context, which in turn reflects the conservatism and caution of the thinking underlying the .07 threshold. We propose accepting twice the rate of false omissions (plans that should have been scrutinized but were not) than the rate of false discoveries (plans that would be flagged for scrutiny given the $E G$ observed in the first election, but would then go on to display sign flips). To reiterate: the proposed standard for judicial scrutiny is cautious and conservative, erring on the side of letting even durably skewed plans stand.


Figure 1: Prognostic performance measures, first efficiency gap under a district plan more extreme than threshold (horizontal axis) as a predictor of whether the subsequent efficiency gaps recorded under the district plan all have the same sign as the first efficiency gap. Vertical lines indicate $95 \%$ confidence intervals. Analysis spans all state legislative elections and district plans as per my initial report, 1972-2014.

Figure 2 repeats this analysis, but only considering the performance of negative values of the first-election efficiency gap threshold, consistent with Republican advantage (and more relevant to the Wisconsin plan at issue). Here the threshold becomes less stringent as we move across the horizontal axis from left to right, from larger negative thresholds to closer to zero at the right hand edge of each panel. With a large negative threshold (left hand edge of each panel), almost all plans test negative and so the sensitivity is close to zero, the specificity is 1 , and the false positive rate is zero. The accuracy measures increase as the threshold becomes less stringent, attaining maxima in the range -.05 to -.02. Again-and consistent with the cautious approach we take-we emphasize that accuracy is not the sole criterion we use to evaluate a decision rule. At low values of the threshold, where accuracy is maximized, the false positive and false discovery rates are relatively high. On the other hand, at the proposed threshold value of .07 , the false positive rate is under $10 \%$ (fewer than $10 \%$ of plans with efficiency gaps changing signs would be scrutinized), and the false omission rate is about $35 \%$ (close to
$35 \%$ of plans would not be flagged despite having EGs of the same sign over their lifetimes). The proposed threshold again errs on the side of restraint, tolerating a higher rate of false omissions than false discoveries.


Figure 2: Prognostic performance measures, first efficiency gap under a district plan more extreme than threshold (horizontal axis) as a predictor of whether the subsequent efficiency gaps recorded under the district plan all have the same sign as the first efficiency gap. Vertical lines indicate $95 \%$ confidence intervals. Analysis examines negative, firstelection threshold values of the efficiency gap, consistent with Republican advantage.

Figure 3 presents the corresponding analysis of positive values of the first-election $E G$ threshold, consistent with Democratic advantage. Here the proposed threshold becomes more stringent as we move to the right of each panel, in the sense that fewer plans trip the threshold. At high values of the threshold (the right hand edge of each panel), no plans trip the threshold and all are classified as "negatives," leading to a specificity of 1 , and false positive and false discovery rates of zero. Once again, accuracy is maximized at a less stringent threshold than the proposed .07 standard, around .03 . The false positive rate is much lower at the proposed threshold of .07 than at the accuracy-maximizing threshold of .03 . Note that the false discovery rates are moderately large but unstable and estimated with considerable imprecision; this is because there are
so few plans exhibiting high (pro-Democratic) levels of $E G$ in their first election. Moreover, of the few plans that do trip a given pro-Democratic threshold in their first election, it is reasonably likely that they will record efficiency gaps that will change sign over the life of the plan; this sign-flip or "false discovery" probability is about $35 \%$ at the proposed threshold of .07 .

Comparing the analyses in Figures 2 and 3, we see an asymmetry in the results. The .07 threshold is more permissive with respect to plans that begin life exhibiting Democratic advantage than it is for plans that initially exhibit Republican advantage. At a +/- .07 threshold, the false discovery rate for plans initially exhibiting Republican advantage is under $10 \%$, but around $35 \%$ for plans initially exhibiting Democratic advantage. As Figure 3 shows, it is difficult to find a threshold for apparently proDemocratic plans that drives the false discovery rate to reliably low levels, if only because the historical record has relatively few instances of these types. We also note that the .07 threshold generates false omission rates of about $30 \%$ for both sets of plans.

Because the preceding discussion is somewhat technical, it is worth restating its principal conclusion: It is that an efficiency gap threshold of 0.07 is quite conservative, in that it sacrifices some accuracy (which would be maximized at a threshold of around 0.03 ) in order to drive down the false positive and false discovery rates. At a threshold of 0.07 , in fact, the false positive and false discovery rates are about half of the false omission rate, indicating that there are about twice as many plans that are not being flagged even though their $E G$ signs would remain one-sided throughout the cycle, than there are plans that are being flagged even though their EG signs would flip. This is further powerful confirmation of the reasonableness of the 0.07 efficiency gap threshold.


Figure 3: Prognostic performance measures, first efficiency gap under a district plan more extreme than threshold (horizontal axis) as a predictor of whether the subsequent efficiency gaps recorded under the district plan all have the same sign as the first efficiency gap. Vertical lines indicate $95 \%$ confidence intervals. Analysis examines positive, firstelection threshold values of the efficiency gap, consistent with Democratic advantage.

## 3 First-election efficiency gap reliability with respect to the plan-average efficiency gap sign

Next we consider a slightly different kind of test; given that the first election under a district plan produces a value of the efficiency gap above or below a given threshold, how likely is it that the average value of the efficiency gap produced over the life of the plan lies on the same side of zero as that of the first election? Recall that the sign of the efficiency gap speaks to the corresponding direction of partisan advantage ( $E G<0$ is consistent with Republican advantage; conversely for $E G>0$ ). We expect that this will be a less strenuous test than asking if any $E G$ has an opposite sign to the first $E G$ observed under a district plan.


Figure 4: Prognostic performance measures, first efficiency gap under a district plan more extreme than threshold (horizontal axis) as a predictor of whether the average efficiency gap recorded under the district plan has the same sign as the first efficiency gap. Vertical lines indicate $95 \%$ confidence intervals. Analysis spans all state legislative elections and district plans as per my initial report, 1972-2014.


Figure 5: Prognostic performance measures, first efficiency gap under a district plan more extreme than threshold (horizontal axis) as a predictor of whether the average efficiency gap recorded under the district plan has the same sign as the first efficiency gap. Vertical lines indicate $95 \%$ confidence intervals. Analysis examines negative, first-election threshold values of the efficiency gap, consistent with Republican advantage.


Figure 6: Prognostic performance measures, first efficiency gap under a district plan more extreme than threshold (horizontal axis) as a predictor of whether the average efficiency gap recorded under the district plan has the same sign as the first efficiency gap. Vertical lines indicate $95 \%$ confidence intervals. Analysis examines positive, first-election threshold values of the efficiency gap, consistent with Democratic advantage.

Figures 4,5 and 6 show the prognostic performance of the first-election $E G$ with respect to the sign of the corresponding plan's average $E G$, looking at the absolute value of the first-election $E G$ (Figure 4), negative first-election efficiency gaps (Figure 5) and positive first-election efficiency gaps (Figure 6). The first thing to observe is the generally superior prognostic performance when it comes to forecasting the sign of the plan-average efficiency gap, relative to the prognostic performance with respect to all of the plan's efficiency gaps having the same sign. As anticipated, the former is better predicted by the plan's first-election efficiency gap than the latter. Second, the accuracy-versus-caution tradeoff noted earlier is also apparent. The proposed threshold of $+/-0.07$ trades away accuracy for very low false positive and false discovery rates, below $5 \%$, at the cost of higher false omission rates, a pattern we observed earlier. Finally, note that at the proposed threshold of $+/-0.07$, almost one-half of all plans with a negative (proRepublican) average EG would not be candidates for scrutiny (right-hand panel of Figure 5); about one-third of plans with a positive (pro-Democratic) average EG also would not trigger the threshold for scrutiny.

## 4 Relationship between the first-election efficiency gap and the plan-average efficiency gap

I next present analysis on a related issue, the relationship between the magnitudes of the first efficiency gap observed under a plan and the average efficiency gap we observe over the life of the plan. Does a larger or smaller first-election efficiency gap portend anything for the average value of the efficiency gap generated over the life of a district plan?

Clearly the first value of the efficiency gap and the plan-average efficiency gap are related; the former contributes to the calculation of the latter, and after the first election under a district plan we observe at most four more elections under the plan (given elections every two years in most states and redistricting once a decade). Accordingly we expect a positive correlation between the two quantities. The interesting empirical question—and one with considerable substantive implications for the issue at hand—is how strong the relationship is between the first-election efficiency gap and the corresponding plan-average efficiency gap. This speaks to the reliability of the firstelection $E G$ measure as a predictor of $E G$ over the life of the plan.

Figure 7 shows the relationship between the first-election $E G$ and the average $E G$ observed over the entire plan. Note that we restrict this analysis to plans with at least three elections, so that the first election does not unduly contribute to the calculation of the average; this restriction has the consequence of omitting elections from the most recent round of redistricting after the 2010 Census, which have contributed at most two elections. The black diagonal line on the graph is a 45-degree line: if the relationship between first-election $E G$ and plan-average $E G$ were perfect, the data would all lie on this line. Instead we see a classic "regression-to-the-mean" pattern, with a positive regression slope of less than one (as indeed we should, given that the first-election $E G$ on the horizontal axis contributes to the average plotted on the vertical axis). But the relationship here is especially strong. The variation in plan-average efficiency gaps explained by this regression is quite large, about $73 \%$; after taking into account the uncertainty in the $E G$ scores (stemming from the imputation procedures used for uncontested districts; see my initial report) a $95 \%$ confidence interval on the variance explained measure ranges from $67 \%$ to $74 \%$ (the uncertainty has the consequence of tending to make the regression fit slightly less well). That is, even given the uncertainty that accompanies $E G$ measures due to uncontestedness, the relationship between firstelection $E G$ and plan-average $E G$ is quite strong.

In particular, at the threshold values of $+/-0.07$ there is very little doubt as to the planaverage value of the efficiency gap. The historical relationship between first-election $E G$ and plan-average EG shown in Figure 7 indicates that a first-election EG of -.07 is typically associated with a plan-average $E G$ of about -0.053 ( $95 \%$ CI -0.111 to 0.004 ); the probability that the resulting, expected plan-average $E G$ is negative is $96.5 \%$. Conditional on a first-election $E G$ of .07 we typically see a plan-average $E G$ of about 0.037 ( $95 \%$ CI -0.021 to 0.093 ); the probability that the resulting, expected plan-average $E G$ is positive is $89.8 \%$. This constitutes additional, powerful evidence that (a) firstelection $E G$ estimates are predictive with respect to the $E G$ estimates that will be observed over the life of the plan; and (b) the threshold values of $+/-0.07$ are conservative, generating high-confidence predictions as to the behavior of the district plan in successive elections.

In the particular case of Wisconsin in 2012-the first election under the plan in question-I estimated the efficiency gap to be -0.133 ( $95 \%$ CI -0.146 to -0.121 ). The analysis of historical data discussed above—and graphed in Figure 7—indicates that the plan-average $E G$ for this plan will be $-0.095(95 \% \text { CI }-0.152 \text { to }-0.032)^{1}$, a quite large value by historical standards, placing the current Wisconsin district plan among the five to ten most disadvantageous district plans for Democrats in the data available for analysis. The probability that the Wisconsin plan—if left undisturbed—will turn out to have a positive, pro-Democratic, average efficiency gap is for all practical purposes zero (less than $0.1 \%$ ).

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Figure 7: Scatterplot of first-election efficiency gap scores (horizontal axis) and planaverage efficiency gap scores (vertical axis). The diagonal black line is a 45 -degree line; the data would lie on this line if first-election efficiency gaps coincided with plan-average efficiency gaps. The solid blue line is a linear regression with slope .64 ( $95 \%$ CI 0.57 to 0.72 ); the shaded region around the blue line is a $95 \%$ confidence interval for the regression line. Vertical and horizontal lines extending from each data point cover $95 \%$ confidence intervals in either direction, summarizing the uncertainty in both first-election $E G$ and plan-average $E G$, stemming from imputations for uncontested districts. Outliers are labeled (state, plan). Analysis restricted to plans with at least three elections (19722010), omitting plans adopted after the 2010 Census. The first-election EG for the current Wisconsin plan is $-0.133(95 \%$ CI -0.146 to -0.121$)$.

## 5 Party control as an explanation for change in the efficiency gap

Both Trende and Goedert point out that, on average, state house plans have exhibited pro-Republican efficiency gaps in recent years (Trende, paragraphs 129-30; Goedert p. 19). They then argue that this pro-Republican mean is attributable to a natural pro-Republican political geography in many states. However, as I found in my initial report, the overall efficiency gap average, over the entire 1972-2014 period, is very close to zero (Jackman Report, p. 35, 45, 57). There is thus no sign of a natural pro-Republican advantage in the dataset as a whole, nor any evidence (despite Trende and Goedert's unsupported assertions to the contrary) that states' political geography is changing in ways that favor Republicans.

In fact, the one historical change that is undeniable is the trend toward unified Republican control over redistricting. As Figure 8 displays, only about $10 \%$ of all state house plans were designed by Republicans in full control of the state government in the 1990s, compared to about $30 \%$ by Democrats in full control and about $60 \%$ by another institution (divided government, a commission, or a court). But in the 2000s, Republicans were fully responsible for slightly more plans than were Democrats (about $20 \%$ versus about $15 \%$ ). And in the 2010 s, the partisan gap jumped again, to about $40 \%$ of plans designed entirely by Republicans, versus less than $20 \%$ designed entirely by Democrats.


Figure 8: Share of all state house plans, by cycle, designed by Democrats in unified control of state government, by Republicans in unified control of state government, or by another institution (divided state government, commission, or court).

To determine the impact of this change in party control on the change in the efficiency gap over the last generation, I carry out three regressions, one for the 1990 redistricting cycle, one for the 2000 cycle, and one for the 2010 cycle. In each case, state house plans' efficiency gaps are the dependent variable, and unified Democratic control over redistricting and unified Republican control over redistricting are the independent variables. (The omitted category is any other institution responsible for redistricting, such as divided government, a court, or a commission.) Figure 9 then displays the actual average efficiency gap for each cycle, as well as the predicted average efficiency gap if the distribution of party control over redistricting had remained unchanged since the 1990s.

As is evident from the chart, state house plans' average efficiency gap in the 2000 cycle would have been substantially less pro-Republican (by about 0.5 percentage points) had Republicans not gained control of more state governments in this cycle relative to the 1990s. And in the current cycle, all of the efficiency gap's movement in a Republican direction would have been erased had the distribution of party control over redistricting not changed since the 1990s. That is, if the same distribution of party control had existed in this cycle as in the 1990s, state house plans' average efficiency gap would have been
very close to zero, not over $3 \%$ in a Republican direction. Accordingly, it is the change in party control that appears to account for essentially all of the pro-Republican trend in the efficiency gap over the past two decades—and not, as claimed by Trende and Goedert, a dramatic alteration of the country's political geography.


Figure 9: Actual and predicted values of state house plans' average efficiency gaps by cycle. Predicted values calculated assuming that the 1990s distribution of party control over redistricting remained constant in subsequent cycles.

## 6 Response to the Chen and Rodden map simulations

Both Trende and Goedert cite a recent article by Chen and Rodden (2013) that purports to find, based on simulations of hypothetical district maps, that random redistricting would benefit Republicans because of their more efficient spatial allocation (Trende, paragraphs 89,126 ; Goedert, pp. 13, 18, 21). While I respect Chen and Rodden's contribution, there are several issues with their work that make it inapplicable here.

First, Chen and Rodden do not even attempt to simulate lawful plans. Rather, they simulate plans "using only the traditional districting criteria of equal apportionment and
geographic contiguity and compactness" (Chen and Rodden, 248). They do not take into account Section 2 of the Voting Rights Act, which often requires majority-minority districts to be constructed. They also do not take into account Section 5 of the VRA, which until 2013 meant that existing majority-minority districts could not be eliminated in certain states. And they do not take into account state-level criteria such as respect for political subdivisions and respect for communities of interest, which are in effect in a majority of states (NCSL 2010, pp. 125-27).

Second, Chen and Rodden only use presidential election results in their analysis, but these outcomes may diverge from state legislative election results due to voter roll-off as well as voter preferences that vary by election level. As Stephanopoulos and McGhee have noted, "If certain voters consistently support Republicans at the presidential level and Democrats at the legislative level, then presidential data may produce more proRepublican estimates than legislative data" (Stephanopoulos \& McGhee, 870). In fact, this is exactly what seems to be occurring; at the congressional level, efficiency gaps are about $6 \%$ more Republican when they are calculating using presidential data than when they are computed on the basis of congressional election results.

Third, Chen and Rodden's simulated maps do not constitute a representative sample of the entire plan solution space. Their simulation algorithm has "no theoretical justification," is "best described as ad-hoc," and is not "designed to yield a representative sample of redistricting plans" (Fifield et al. 2015, pp. 2-3; Altman \& McDonald 2010, p. 108). The explanation for this lack of representativeness is highly technical and involves the details of the particular simulation approach adopted by Chen and Rodden. But its implication is clear: that no conclusions can yet be drawn about the partisan consequences of randomly drawn maps.

Lastly, Chen and Rodden's results are directly contradicted by Fryer and Holden, who also simulated contiguous, compact, and equipopulous districts for multiple states. Unlike Chen and Rodden, Fryer and Holden found that, "[u]nder maximally compact districting, measures of Bias are slightly smaller in all states except [one]" (Fryer \& Holden 2011, p. 514). Fryer and Holden also found that "[i]n terms of responsiveness . . . there are large and statistically significant" increases in all states, sometimes on the order of a fivefold rise (p. 514). Their analysis thus leads to the opposite inference from Chen and Rodden's: that randomly drawn contiguous and compact districts favor neither party and substantially boost electoral responsiveness.

## 7 Trende's analysis of particular plans

Trende devotes a large portion of his report (paragraphs 106-31) to analyzing the efficiency gaps of particular state legislative and congressional plans. He first examines a set of seventeen state legislative plans that had efficiency gaps favoring the same party over their entire lifespans, arguing that not all of these plans were gerrymanders (paragraphs 106-14). He then cites a series of congressional plans, some of which he claims had large efficiency gaps despite not being gerrymanders, and others of which allegedly had small efficiency gaps despite being gerrymanders (paragraphs 115-24). All of this analysis is riddled with conceptual and methodological errors that, in my judgment, renders it unreliable and unhelpful to the court.

Beginning with the set of seventeen state legislative plans that had efficiency gaps of the same sign throughout their lifespans, Trende asserts that they "would be included in the definition of a gerrymander," and are a "list of gerrymandered states" (paragraphs 109-10). But neither plaintiffs nor I argue that these plans should have been held unconstitutional. That is, neither plaintiffs nor I argue that these plans were designed with partisan intent (the first element of plaintiffs' proposed test), that their initial efficiency gaps exceeded a reasonable threshold (the second element), or that their efficiency gaps could have been avoided (the third element). To the contrary, I simply included these plans in my report to illuminate historical cases in which the efficiency gap's direction did not change over the course of a decade. I never stated or implied that these plans should have been deemed unlawful.

However, if we focus on the plans among the seventeen that likely would have failed plaintiffs' proposed test (at least the first two elements), we see that both the test and the efficiency gap perform exceptionally well. Five of the seventeen plans featured unified control by a single party over redistricting (from which, like Goedert (2014) and Goedert (2015), we can infer partisan intent) as well as an initial efficiency gap above $7 \%$ (the threshold I recommended in my initial report): Florida in the 1970s, Florida in the 2000s, Michigan in the 2000s, New York in the 1970 s, and Ohio in the 2000 s. Assuming that these plans' large efficiency gaps were avoidable (a granular inquiry that cannot be carried out here), it would have been quite reasonable for all of these maps to attract heightened judicial scrutiny. In particular:

- Florida's plan in the 1970 s was designed exclusively by Democrats, opened with a $9.9 \%$ pro-Democratic efficiency gap, averaged a $7.0 \%$ pro-Democratic efficiency gap over its lifespan, and never once favored Republicans.
- Florida's plan in the 2000s was designed exclusively by Republicans, opened with a $8.9 \%$ pro-Republican efficiency gap, averaged a $11.2 \%$ pro-Republican efficiency gap over its lifespan, and never once favored Democrats.
- Michigan's plan in the 2000s was designed exclusively by Republicans, opened with a $12.0 \%$ pro-Republican efficiency gap, averaged a $10.3 \%$ pro-Republican efficiency gap over its lifespan, and never once favored Democrats.
- New York's plan in the 1970s was designed exclusively by Republicans, opened with a $10.7 \%$ pro-Republican efficiency gap, averaged a $9.7 \%$ pro-Republican efficiency gap over its lifespan, and never once favored Democrats.
- Ohio's plan in the 2000s was designed exclusively by Republicans, opened with a $8.6 \%$ pro-Republican efficiency gap, averaged a $9.0 \%$ pro-Republican efficiency gap over its lifespan, and never once favored Democrats.

Accordingly, we see that if my report's set of seventeen plans is analyzed properly, the opposite conclusion emerges from the one advocated by Trende. Only a subset of the seventeen plans likely would have failed plaintiffs' proposed test. But every member of this subset turns out to have been an exceptionally severe and durable gerrymander, featuring a very large and consistent efficiency gap over its lifespan. These are precisely the historical cases in which judicial intervention may have been advisable.

After commenting on these seventeen state legislative plans, Trende discusses a series of congressional plans, all from the 2000 and 2010 redistricting cycles. These congressional plans are entirely irrelevant to this case, which deals only with state legislative redistricting. Neither in their complaint nor in their subsequent filings do plaintiffs ever argue that their approach should be applied to congressional plans. And neither Mayer nor I provide any empirical analysis of congressional plans. In my initial report, in particular, I examined state legislative plans from 1972 to the present, but no congressional plans at all.

This state legislative focus has two explanations. First, and more importantly, each congressional delegation is not a legislative chamber in its own right, but rather a portion (often a very small portion) of the U.S. House of Representatives. Methods applicable to entire chambers cannot simply be transferred wholesale to delegations that make up only fractions of Congress. Second, most congressional delegations have many fewer seats than most state houses. The efficiency gap becomes lumpier when there are fewer seats, because each seat accounts for a larger proportion of the seat total, and the efficiency gap thus shifts more as each seat changes hands. This lumpiness is entirely avoided when state legislative plans, which typically have dozens or even hundreds of districts, are at issue.

For these reasons, Stephanopoulos and McGhee make two adjustments when analyzing congressional plans in their work on the efficiency gap. First, they convert the efficiency gap from percentage points to seats by multiplying the raw efficiency gap by each state's number of congressional districts. As they explain their method, "What matters in congressional plans is their impact on the total number of seats held by each party at the national level. Conversely, state houses are self-contained bodies of varying sizes, for which seat shares reveal the scale of parties' advantages and enable temporal and spatial comparability" (Stephanopoulos \& McGhee, 869). Second, they only calculate efficiency gaps for states with at least eight congressional districts. Efficiency gaps are lumpier for states with fewer than eight districts, and additionally, congressional "redistricting in smaller states has only a minor influence on the national balance of power" (Stephanopoulos \& McGhee, 868).

In his report, Trende fails to make either of these necessary adjustments when examining congressional plans. That is, he does not convert the efficiency gap from percentage points to seats, and he calculates the efficiency gap for small congressional delegations with fewer than eight seats. There is no authority in the literature for his methodological choices, and he is unable to cite any. And his flawed methods have serious substantive consequences that render his results entirely untrustworthy.

Take Trende's failure to convert the efficiency gap from percentage points to House seats. He claims that Alabama's congressional plan had an efficiency gap of $12.5 \%$ in 2002, that Arizona's congressional plan had an efficiency gap of $16 \%$ in 2012, that Colorado's congressional plan had an efficiency gap of $-9 \%$ in 2002 and $-10 \%$ in 2012, that Illinois's congressional plan had an efficiency gap of $-9 \%$ in 2002, and that Iowa's congressional plan had an efficiency gap of $-20 \%$ in 2002—all above my suggested $7 \%$ threshold for state legislative plans (paragraphs 115-16, 118-19, 121-22). But when converted to seats, all of these efficiency gaps become quite small, lower in all cases than the two-seat threshold proposed in the literature for congressional plans (Stephanopoulos \& McGhee, 887-88). Specifically, using Trende’s own calculations-which, as I discuss below, are incorrect in any event-Alabama had an efficiency gap of -0.9 seats in 2002, Arizona had an efficiency gap of 1.4 seats in 2012, Colorado had an efficiency gap of -0.6 seats in 2002 and -0.7 seats in 2012, Illinois had an efficiency gap of -1.7 seats in 2002, and Iowa had an efficiency gap of -1.0 seats in 2002 . None of these scores are high enough to rise to presumptive unlawfulness under the literature's suggested two-seat threshold, meaning that we come to exactly the opposite conclusion as Trende after making the necessary adjustment.

Next take Trende's consideration of Alabama's congressional plan in 2002 (which had seven districts), Iowa's congressional plan in 2002 (five districts), and Colorado's congressional plans in 2002 and 2012 (seven districts each) (paragraphs 115-16, 119, 122). All four of these plans have fewer than eight districts, and so, based on the literature, should not be included in any efficiency gap analysis because of the measure's lumpiness when applied to so few seats. Trende nowhere acknowledges this limitation, and indeed appears unaware of its existence.

Moreover, Trende's study of congressional plans is marred by two further flaws, one conceptual and the other methodological. The conceptual defect is that, as in his earlier discussion of state legislative plans, he assumes that a large efficiency gap is all that is necessary to render a plan unconstitutional. He writes that efficiency gaps of $-12.5 \%$, $9 \%,-9 \%,-20 \%$, and $16 \%$ "would invite court scrutiny as a Republican gerrymander" or "would invite court scrutiny as a Democratic gerrymander" (paragraphs 115, 116, 118, $119,121,122)$. But again, this is not plaintiffs' proposed test. A large efficiency gap is only a single prong of the test, and does not result in a verdict of unconstitutionality unless it is paired with a finding of partisan intent and a finding that it could have been avoided. Trende entirely overlooks these other elements.

The methodological defect is that whenever there were uncontested congressional races, Trende simply substituted presidential election results for the missing congressional results. As he put it in his deposition, he "used presidential results" and "imputed those results to the congressional races" whenever the races were uncontested (Trende deposition, p. 83). This is an exceptionally crude method that is guaranteed to produce errors, both because there is voter roll-off from the presidential to the congressional level and because voters may have different presidential and congressional preferences. Of course, presidential results can be used as the inputs to a regression model that predicts the outcomes of uncontested congressional races. Indeed, this is the preferred approach in the literature, and the approach I employed in my initial report. But presidential results cannot simply be plugged in without any adjustment, and no competent social scientist would have done so.

Accordingly, in my judgment, Trende's examination of particular state legislative and congressional plans is unreliable and entitled to no weight by the court. The state legislative analysis ignores the actual elements of plaintiffs' proposed test, and would have led to the opposite conclusion if these elements had been taken into account. Likewise, the congressional analysis ignores the test's prongs, fails to convert the efficiency gap from percentage points to seats, improperly considers states with small House delegations,
improperly substitutes presidential election results whenever congressional results are missing—and deals with federal elections that simply are not part of this case.

Dated December 21, 2015
/s/ Simon Jackman
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Rebuttal Report: Response to Expert Reports of Sean Trende and Nicholas Goedert

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This report presents my responses to the criticisms that Sean Trende and Professor Nicholas Goedert make of my report. ${ }^{1}$

## I. Summary

A. Both Trende and Goedert erroneously argue that Democrats are more geographically concentrated than Republicans in Wisconsin, which creates a natural pro-Republican bias even under a neutrally-drawn district plan. Both arguments are based on unreliable methodologies, flawed measures, and lead to inaccurate conclusions. Trende's methodology for measuring partisan concentration relies on an unorthodox method (the PVI) far more common among political commentators than academics who study spatial patterns of concentration and isolation. Moreover, as he applies it here, Trende relies on fundamentally inaccurate measures of geography that are guaranteed to demonstrate that Democratic wards are closer to one another than are Republican wards.

Goedert's arguments about geographic concentration are analogous to Trende's, and suffer from the same flaws in that they are based on superficial claims that do not rely on actual measures of spatial concentration or isolation. Moreover, Goedert's claims here contradict his own research, in which he finds that even after controlling for urbanization (a proxy for concentration), Republican control of the redistricting process has a large and statistically significant impact on a plan's bias. A model in one of his papers (Goedert 2015) also shows that a court-drawn or bipartisan map in Wisconsin would be expected to produce a pro-Democratic bias. The model generates the same expectation for a court-drawn or bipartisan map in a state that resembles the country as a whole. Accordingly, based on Goedert's own analysis, there is no natural pro-Republican tilt in either Wisconsin or the typical U.S. state.

In contrast to Trende's and Goedert's unorthodox techniques, widely (even universally) accepted measures of spatial distributions, such as Global Moran's I (Cho 2003) and the Isolation Index (Reardon 2004), show that Wisconsin's Republicans and Democrats are equally spatially concentrated and equally spatially isolated from each other, and that in some election years Republicans are more concentrated than Democrats.
B. Trende criticizes my method of estimating the partisanship of uncontested Assembly districts as biased. But his criticism stems from a superficial and erroneous discussion of a single figure in my report (Figure 2), and he erroneously believes that I set the Assembly votes in uncontested districts to the presidential vote in those districts. He does not take notice of the fact that my analysis was based on a comprehensive multiple regression model that controlled for the very factors that he claims create bias, nor that my model produces extraordinarily accurate forecasts of the actual data, using multiple methods.

[^37]C. Trende criticizes my baseline measure of partisanship for not taking into account factors such as incumbency, candidate quality, and spending. This is an inaccurate criticism, because estimating baseline partisanship is designed to control for incumbency, campaign spending, and candidate quality. This is the method preferred in the academic literature on redistricting, which seeks to understand the consequences of hypothetical plans (in which candidate quality, spending, and incumbency are unknown). My approach is identical to the method used by Professor Gaddie, who produced the baseline partisan estimates used by Wisconsin's map drawers in 2011.
D. Goedert challenges my model for estimating baseline partisanship in 2012, contending that I took into account information that the authors of Act 43 did not have (the 2012 election results). However, my baseline estimates of partisanship are nearly identical to those generated by Gaddie in 2011, indicating the same conclusions follow whether 2012 or pre-2012 data are used in the analysis. In addition, pre-2012 election results are highly correlated with 2012 election results, indicating that it would make no difference if I had used earlier election results. Goedert dismisses the convergence between my estimates and Gaddie's estimates as "mostly coincidental," but offers no evidence or data to support his assertion.
E. Geodert also challenges my efficiency gap calculations for ignoring the effects of incumbency, which he asserts that any author of a redistricting plan would incorporate. His criticism fails to acknowledge that controlling for incumbency is the standard methodology for estimating the partisan consequences of a hypothetical district plan. Nevertheless, I recalculated efficiency gap estimates for both Act 43 and my Demonstration Plan, taking incumbency into account. The substantive conclusions are identical: the efficiency gap for my plan increases slightly (but is still well within acceptable limits), as does the efficiency gap for Act 43. The difference between the two plans' efficiency gaps remains enormous.
F. Goedert criticizes my efficiency gap calculations for not including any sensitivity testing to determine whether my results are robust to changes in the statewide electoral environment. I conducted a uniform swing analysis over the range of plausible election results, based on the maximum and minimum statewide Democratic Assembly vote since 1992. This analysis shows that the efficiency gaps of both Act 43 and the Demonstration Plan are robust: Act 43's efficiency gap remains very high across this range, significantly above the plaintiffs' suggested 7\% threshold even in the face of an historic Democratic wave, and the Demonstration Plan's efficiency gap remains very low, and is always well below the threshold. Goedert is simply incorrect in asserting that the plans' respective efficiency gaps are not robust, and, again, offers no data or evidence to support his claim.
G. Throughout their reports, neither Trende nor Goedert has actually done any analysis that identifies problems with my analysis, or that specifically shows where my analysis is incorrect. Trende and Goedert merely offer speculative and unsubstantiated criticism, but never offer any substantive data or evidence that supports their arguments. And, as I will show, when they attempt to analyze Wisconsin's political geography, their conclusions are utterly wrong.

## II. The Claim that Wisconsin's Political Geography Has a Pro Republican Bias

While I will go into more detail on the specific points each report makes, I focus first on a central argument both Trende and Goedert make: that Wisconsin has a natural distribution of Republicans and Democrats that produces an intrinsic pro-Republican bias in a neutrally-drawn redistricting plan. They claim that because Democrats in Wisconsin happen to be (allegedly) naturally concentrated in small pockets of overwhelming Democratic strength, even a neutrallydrawn map would produce a large pro-Republican efficiency gap. As a result, they conclude, it is not possible to consider a large pro-Republican efficiency gap as evidence of gerrymandering.

I begin by noting that both Trende and Goedert ignore the role that political geography already plays in plaintiffs' proposed test. Under the test's first prong, if the state's motive in enacting its plan was simply to follow the contours of the state's geography, then partisan intent would not be present and plaintiffs would proceed no further in their claim. Similarly, under the test's third prong, if the state can show that its plan's large efficiency gap was necessitated by the geographic distribution of the state's voters, then the plan would be upheld. These points mean that geography is already properly incorporated into plaintiffs' proposal.

There are, additionally, two points that fundamentally negate the utility of this line of attack. First, the geographic concentration argument is predicated on the foundational assumption that a neutrally-drawn map would have produced a pro-Republican bias. Even if Trende and Goedert are correct in this assumption (which they are not), they take no position on whether the process in Wisconsin was, in fact, neutral. The record of the federal redistricting trial clearly shows that Act 43 was designed with the predominant purpose of benefiting Republicans and disadvantaging Democrats, and neither Trende nor Goedert contradicts the findings in my report of examples of blatant packing and cracking that are the very DNA of a partisan gerrymander.

And second, even if the state's experts are correct that political geography has produced the pro-Republican bias in Wisconsin's state legislative district plan (which they are not), it is impossible for them to quantify how much of an effect geography has had: is it $5 \%$ ? $10 \%$ ? $90 \%$ ? $100 \%$ ? Neither Trende nor Goedert have actually done any analysis that demonstrates that the alleged concentration of Democrats in Wisconsin will produce a pro-Republican efficiency gap, or any work that quantifies how concentration is related to efficiency gap calculations. They simply assert (incorrectly) that Democrats are more concentrated than Republicans, and therefore that even a neutral map will produce a pro-Republican bias.

But they are also wrong on the facts. Their argument about geographic concentration is based on flawed data and measures, and has no basis in accepted methods of measuring geographic concentration and isolation. Trende, in particular, uses an unorthodox method with no support in the peer-reviewed literature, and one that is guaranteed to produce a biased result that shows Democrats far more concentrated than they actually are. Goedert's argument contradicts his own published work, which shows that partisan control of redistricting generates a substantial bias even after partisan concentration is taken into account. His argument, further, falls victim to the Modified Areal Unit Problem, in that it is based entirely on the analysis of wards, ignoring the fact that wards are aggregated into districts. As I demonstrate, this aggregation process completely changes the applicability of Goedert's conclusions.

When I analyze the geographic distribution of Wisconsin's Democrats and Republicans using widely accepted measures of spatial concentration and isolation (Global Moran's I and the Isolation Index), I find that there is very little evidence of significant disparities in how the parties' voters have been distributed in recent election cycles. Republicans are in fact more concentrated than Democrats when measured by the 2012 Assembly vote.

## A. Trende

Trende spends nearly half of his report (paragraphs 62-105) arguing that Democrats are naturally more concentrated ("clustered") than Republicans in Wisconsin, which creates a natural packing effect. Much of this discussion is entirely irrelevant to Wisconsin (Trende's discussion of patterns in the southern United States, Virginia, and differences between the 1996 and 2008 Democratic coalitions; see paragraphs 62-77). Trende also simply asserts that "there is little doubt that the Democratic vote in Wisconsin is also increasingly concentrated in fewer counties" (paragraph 71). He neither explains the relevance of the county vote to the issue of geographic distribution and legislative redistricting, nor why the county vote pattern in 1988 or 1996 is germane to the environment in 2012.

## 1. The PVI (partisan vote index) is the wrong quantity of interest

As applied to Wisconsin, Trende attempts to demonstrate that over the last 20 years Democrats have become more concentrated. His method relies on a quantity he calls the Partisan Lean Index, which is the party's county or ward vote share minus the party's statewide vote share, and appears to be analogous to the Cook PVI, which is the same quantity calculated using the congressional district vote and the national presidential vote. Trende argues that Democratic wards are closer together than Republican wards, which to him is evidence of geographic clustering that produces a natural pro-Republican redistricting bias.

The PVI (which is how Trende abbreviates the measure) is a quantity that is not commonly used in the academic literature, and when it is, it is used largely as a simple descriptive statistic. What this index does is simply redistribute the ward vote around the statewide average, and thus tells us which areas are more Democratic (or Republican) than the
state as a whole, and which areas are less so. ${ }^{2}$ It tells us little about overall partisan strength, and is useful only in comparing elections at one level (here, counties or wards) to elections at another (the state).

The PVI is used almost exclusively by political commentators to describe congressional districts (the most widely known is the Cook PVI, which compares the average congressional district vote split over two consecutive elections to the average national presidential vote over those same elections). It is used less frequently in academic research, and then largely as a basic descriptive statistic used to classify districts as competitive or not. It is not used in the context of state legislative redistricting (Trende did not cite any studies that support the use of his measure, and could not identify any in his deposition).

Moreover, Trende appears to have made two errors in his calculation of the PVI. ${ }^{3}$ First, while he states that his PVI is based on the top-of-the-ticket race in each year, he uses the gubernatorial elections as his top-of-the-ticket race in 2002, 2010, and 2014, but the U.S. Senate race in 2006, even though there was a gubernatorial race that year. While scholars may differ on whether a gubernatorial or U.S. Senate election is the correct top-ticket race, there is no justification whatsoever for being inconsistent. ${ }^{4}$

Second, in calculating his 2014 PVI, Trende mistakenly subtracted the 2014 statewide percentages from the 2012 ward totals (this is the code he used to generate the PVI for 2014; the error is highlighted, and "map_2012\$r_share" is the ward vote for 2012):
map_2014-readOGR("Wards_Final_Geo_111312_2014_ED.shp", "Wards_Final_Geo_111312_2014_ED")
map_2014=spTransform(map_2014, CRS("+proj=longlat +datum=WGS84"))
map_2014\$r_share=map_2014\$GOVREP14/(map_2014\$GOVREP14 + map_2014\$GOVDEM14)
map_2014\$pvi=map_2012\$r_share - sum(map_2014\$GOVREP14)/(sum(map_2014\$GOVREP14) + sum(map_2014\$GOVDEM14))
map_2014\$pvi[which(is.nan(map_2014\$pvi))]=0
Instead of the PVI, the actual ward level vote (or party vote share) is a much more direct measure of ward partisanship. I used LTSB ward level data from 2002 to 2014 to calculate the average Democratic percentage of the vote in a Democratic ward (all wards that were more than $50 \%$ Democratic in the top-ticket race), and the average Republican vote in wards where Republicans won more than $50 \%$ of the top-ticket vote. A graph of this data shows a very different pattern from what Trende claims (Republicans are in red; Democrats in blue):

[^38]

Here, we see that Democrats and Republicans have moved in almost identical fashion between 2002 and 2014. In 2002, Democrat wards were about $60.8 \%$ Democratic, and Republican wards were about $60.5 \%$ Republican in the top-ticket races. In 2014, similarly, both Democratic and Republican wards became more partisan: Democratic wards were $63.3 \%$ Democratic, and Republican wards 63.6\% Republican.

Trende's claim that Democratic wards have become more Democratic, while Republican wards have not become more Republican (paragraphs 91-95), is simply false.

Trende offers no justification or support for why he is relying on the PVI measure rather than more direct indicators of ward partisanship; he merely asserts that it is a relevant quantity. Given that there are far more widely used and relevant measures of district level partisanship, his reliance on it in this context is unsupportable.

## 2. Trende's "Nearest Neighbor" Method is Inappropriate and Inaccurate

After introducing the PVI, Trende attempts to use it to demonstrate that Democrats have become more closely packed than Republicans (which, he asserts, produces a natural proRepublican gerrymander). Apart from the irrelevance of the PVI, Trende's analysis uses a fundamentally flawed measure that is guaranteed to exaggerate the extent of Democratic concentrations. Instead of his measure, widely used and academically accepted metrics of concentration and isolation show that Democrats and Republicans are both highly segregated, and to about the same extent. Just as there are core areas of high Democratic strength in Milwaukee and Madison, there are similar Republican core areas in the "collar counties" of Waukesha, Ozaukee, and Washington.

The premise of Trende's argument is that pro-Democratic wards are closer to other proDemocratic wards than are pro-Republican wards to other pro-Republican wards. His method, which I infer from his description, is to identify a pro-Democratic or pro-Republican ward of a certain percentage lean, and then to find the distance to the nearest ward with the same partisan lean. He determines the median distance between similar wards, and presents two graphs (about paragraph 98 in his report) showing that the median distance between similar Democratic wards is smaller than for Republican wards, and that as Democratic wards become more Democratic, they become closer to one another.

This is reminiscent of the nearest neighbor method used in the study of populations, but it bears little resemblance to how the concept is actually used in the literature, even in its earliest form (Clark and Evans (1954) used it to study the distribution of plant and animal populations). ${ }^{5}$ His application of this method is highly unorthodox, unsuited to the study of redistricting, and not based on any accepted peer-reviewed academic work (he does not cite a single study in support of his method).

Trende's method is to start with a ward (call it $i$ ), calculate its PVI and assign it to a quantile, and then locate the closest ward that shares this PVI quantile (call it $j$ ). The geographic distance between wards $i$ and $j$ (presumably calculated using the ward centroids, although Trende fails to specify this key detail) is then recorded (paragraph 97). The process is repeated for every ward over every election from 2002 to 2014 , producing for each election a matrix consisting of every ward and the distance to the nearest ward with the same PVI quantile. He then calculates median distances between wards of the same PVI quantiles, which he claims shows that Democratic wards are, and have been continuing to move, closer together than Republican wards.

There are several problems with this approach. First, and most fundamentally, the proximity of similar wards is simply not a measure of geographic concentration or clustering. Trende's method tells us nothing about which wards are actually adjacent to wards of a certain PVI. It only tells us how far these wards tend to be from other wards of the same partisan lean. It is entirely possible for wards of the same partisan makeup to be far apart but still easy to join in the same district (think of a sparsely populated but uniformly partisan area). Likewise, it is entirely possible that wards of the same partisan makeup are close together but quite difficult to combine in the same district (think of a densely populated but politically heterogeneous area). Trende's method cannot distinguish between these scenarios, and as a result it cannot tell us anything about the geographic patterns that actually matter for redistricting.

Second, Trende does not explicitly define in his report what a "similar partisan index" (paragraph 97) means. Clearly, Trende is classifying them in some way, defining "similar" as within some range, as his vague discussion of quantiles indicates (paragraph 98). But without specifying the range, it is impossible to know whether his measure has any meaning. Different

[^39]classification methods -- requiring a match of, say, within 0.1 percentage points, or classifying according to deciles or some other method -- are likely to yield very different results than requiring a match of within 0.5 or 1.0 percentage points or using a larger number of categories. His graphs suggest he is using some type of percentile distribution (the x axis label refers to "(.05\% is the most Democratic [or Republican] Ward)," but he does not explicitly define why he chose this particular scheme or how he calculated the quantiles. On this point alone, his method lacks validity or replicability.

But there are two additional serious - fatal, in fact - flaws in this method. First, in treating the geographic distances between wards as his quantity of interest, Trende does not take into account the fact that wards in Wisconsin are not uniform in area. Ward areas actually vary widely: some are very small, others are moderate in size, and still others are very large (wards are drawn within specified population limits, but their geographic areas are not similarly constrained).

Table A shows the mean and median areas (in square miles) of Wisconsin wards. The average is $8.41 \mathrm{mi}^{2}$, but the range is huge: the smallest ward with a nontrivial population is in the City of Middleton: ward 19, with 690 people in an area of $0.0071 \mathrm{mi}^{2}$. The largest ward in the state is in the Town of Winter: ward 2 (in Sawyer County), with 565 people in an area of 227.7 $\mathrm{mi}^{2}$.

Geographic distances between ward centroids will, obviously, depend on how large the wards are. Although centroid-to-centroid distances will not map perfectly onto area differences (because the distances will vary with the shape and orientation of wards), two large wards - even if they are adjacent - will show up as much farther apart than two smaller wards that might be separated by numerous other wards and municipal boundaries.

The problem is magnified when we observe that ward sizes are correlated with other relevant variables, particularly whether a ward is in a city, and most crucially, whether it is a Democratic or Republican ward:

| Table A 2012 Ward Sizes (square miles) ${ }^{6}$ |  |  |
| :---: | :---: | :---: |
|  | Mean | Median |
| Statewide Average | 8.41 | 1.12 |
| City of Milwaukee | 0.29 | 0.20 |
| Rest of State | 8.83 | 1.27 |
| Democratic Wards | 5.91 | 0.56 |
| Republican Wards | 10.96 | 3.45 |

Wards in the city of Milwaukee have a mean area of only $0.29 \mathrm{mi}^{2}$, which is $3 \%$ of the size of the mean area statewide. Democratic wards (measured by whether the 2012 Democratic presidential vote was above $50 \%$ ) are, on average, only about half the size of Republican wards ( $5.91 \mathrm{mi}^{2}$ vs. $10.96 \mathrm{mi}^{2}$ ).

In relying on the distance between wards, Trende is thus putting his thumb on the scale; all other things equal, this method will always show Democratic wards to be much closer than Republican wards, irrespective of whether this concentration is real or merely an artifact of ward area. To put it most simply, smaller Democratic wards will always appear closer than larger Republican wards.

But a second and equally serious problem lurks. Trende does not use the mean distance between wards as his quantity of interest, but rather the median. He justifies this choice "because outlying wards, such as Menominee County, exert an undue amount of leverage on averages" (paragraph 97).

This is the wrong measure, because the "nearest neighbor" approach is unlikely to pair, say, a ward in Milwaukee with a ward in northwest Wisconsin. Menominee County will not exercise "an undue amount of leverage" because it is an outlying ward. It will exercise an undue amount of leverage because it has a very large area $\left(222.8 \mathrm{mi}^{2}\right)$, which is something Trende should, but does not, correct for.

His use of the median rather than the mean further exaggerates the difference between Republican ward distances and Democratic ward distances. The average Republican ward area is 1.9 times larger than the average Democratic ward area ( 10.96 vs. $5.91 \mathrm{mi}^{2}$ ). But the median Republican ward is 6.2 times larger than the median Democratic ward ( $3.45 \mathrm{mi}^{2}$ vs. $0.56 \mathrm{mi}^{2}$ ).

[^40]Because the disparity is three times larger for the median versus the mean area, Trende is further stacking the deck in favor of his preferred hypothesis.

I was able to replicate Trende's analysis, using LTSB data and the R code he disclosed. When the mean distances between similar wards are included, Figure B is the result for the 2012 Election: ${ }^{7}$

Figure B
Nearest Neighbor Analysis: Mean vs. Median Distance


In this graph, the dotted lines are the median nearest neighbor distances for Democratic (blue) and Republican (red) wards, replicating what Trende did in his median distance graphs around paragraph 98 in his report. Wards become more partisan as we move from right to left.

The mean distances are shown with solid lines. While Republican wards remain farther apart than Democratic wards, the mean distances for both parties are much larger than the median distances. Proportionally, Republican and Democratic wards are much closer together in mean than in median distances (which is what one would expect, given the exaggerated difference between median Democratic and Republican ward sizes). Specifically, the mean distance between Republican wards is only about 70\% larger than the mean distance between Democratic wards, compared to a $180 \%$ difference between the median Republican and Democratic distance.

[^41]More relevant is the shape of the mean distance lines. They show that Republican and Democratic distances move precisely in parallel, and that strongly Democratic wards are significantly farther apart than weaker Democratic wards (as are strongly Republican wards). This is the complete opposite of Trende's claim that stronger Democratic wards are closer together than weaker Democratic wards, and it obliterates the core of Trende's report: the assertion that the pro-Republican bias evident in Act 43 is the natural result of Democrats being more geographically concentrated.

To conclude, Trende's argument about Democratic concentration is based on an irrelevant measure of partisanship (PVI) that is incorrectly calculated, applies a methodology that bears no relationship to any scholarship or actual research on spatial distribution, ignores a key feature of Wisconsin's actual political geography (ward area), relies on an improper distance measure that is enormously biased in favor of his hypothesis, and produces a result that fundamentally misrepresents what the data actually shows. Because of his use of a questionable method and fundamentally flawed measures, Trende's opinions should be regarded as uninformative.

## B. Goedert

Goedert, like Trende, asserts that Wisconsin's natural geography creates an intrinsic proRepublican bias in redistricting (p. 17). He cites his own research that geography produced a pro-Republican bias in the 2012 congressional election (p. 19).

The only analysis Goedert conducts as to Wisconsin is an examination of wards, which he claims shows "the bias inherent in Wisconsin's geography" (p. 21). His analysis is a simple "uniform swing" study of wards in 2012, adjusting the Democratic presidential vote in each ward downward by $3.5 \%$ to determine the overall ward distribution in the event of a tied election (Figure 1, p. 22). He asserts that based on this analysis, "Republicans would win $60.2 \%$ of wards, comprising $54.4 \%$ of the voting population" in a tied election (p.22). This is the extent of his analysis.

This analysis, however, is a non sequitur, because it fails to aggregate wards to the relevant geographic level, which is districts. Goedert's failure to take this into account is an example of the Modified Areal Unit Problem, in which inferences at one level of geography frequently do not hold at other levels of aggregation; see King (1996). In this example, the ward level vote is far less relevant than the district level vote, because it is entirely possible that wards will be aggregated in such a way that the pattern he observes either disappears (or even reverses).

When we examine the distribution of districts, which have a population deviation small enough that we can consider them equal (the deviation under Act 43 is $0.76 \%$ ), we in fact see almost the reverse pattern. The following graph (Figure C) displays Goedert's adjusted ward level presidential vote in a simulated 50-50 election, along with an adjusted baseline forecast for Act 43 districts, using my baseline open seat model, in a simulated tied election. Both wards and districts are weighted based on the number of votes cast in each unit. This allows me to directly compare ward level results to district level results:

Figure C: Distribution of Wards vs. Districts


What this figure demonstrates is that as wards are aggregated into districts, the distribution substantially changes. The red line is a kernel density plot of the ward Democratic vote percentage in a simulated tied election; it is a continuous version of the histogram Goedert presents in his Figure 1. The dotted blue line shows the predicted Democratic vote in Act 43 districts in a simulated tied election - or, what occurs after the wards are aggregated into Assembly districts. The overall shape of the curves, the mode of each distribution, and even the mean vote percentage vary as we aggregate from wards up to districts. Knowing the ward distribution ultimately does not tell us much about what the distribution of districts will look like; the process of aggregation is crucial.

More significantly, the district distribution is much more tilted in a Republican direction than is the ward distribution. The ward distribution is nearly normal in shape, and has a peak very close to $50 \%$ Democratic. In contrast, the district distribution is skewed to the right, and has a much higher peak around $42 \%$ Democratic, meaning that there are many more districts that Republicans win by relatively small margins (indicating that Democrats are cracked), and many more districts where Democrats win by much larger margins (indicating packing). Accordingly, the district distribution does not mirror the underlying distribution of wards. Rather, it reveals that Act 43's designers were able to distort a fairly neutral ward distribution into a far more advantageous district distribution, through gerrymandering.

## 1. Goedert's Published Work Contradicts His Report

Goedert's own prior work indicates that unified party control of state government has an independent and significant effect on the bias of redistricting plans, even after controlling for
population concentration. This work also indicates that if Wisconsin, or a state resembling the country as a whole, had a court-drawn or bipartisan map in 2012, this map would have had a slight pro-Democratic bias. These findings further obliterate the claim that Act 43's extreme partisan tilt resulted from Wisconsin's natural political geography.

In a 2014 article, Goedert analyzes the consequences of different redistricting processes, looking for evidence that partisanship and geography each have an independent effect on the partisan bias of redistricting plans. ${ }^{8}$ Using an unorthodox definition of gerrymandering Goedert defines any redistricting plan created in a state with unified party control of state government as a partisan gerrymander - he finds that in states with more than six congressional districts, both urbanization (a proxy for Democratic concentration) and unified party control have a strong and statistically significant effect on the bias of a district plan $(2014,6)$. Goedert interprets his results as indicating that geography matters, and that higher urban concentration leads to more bias against Democrats $(2014,6)$. But what his results also show is that even after taking urbanization into account, the partisanship of the map drawers introduces a separate and significant bias: Republican-drawn maps are associated with an additional $13.6 \%$ proRepublican bias.

Geodert updated his 2014 article in a more recent manuscript, which incorporated the results of the 2014 midterm elections. Here, he finds that urbanization no longer has a statistically significant effect on the bias of district plans $(2015,6)$. Yet he stills finds evidence that the partisanship of map-drawers has a significant effect on district plans' bias (in 2014, a Republican-drawn plan adds $12.4 \%$ bias, or roughly the same as the $13.6 \%$ estimate for 2012).

So, on the one hand, Goedert's own work comes to different conclusions about the impact of urbanization (or Democratic concentration): sometimes it matters, other times it does not. But his work is consistent about the effect of partisan control: when partisans draw maps, they always do so in ways that dramatically bias plans in their favor. The clear inference is that geography matters much less than partisan control in explaining plans' electoral consequences.

Furthermore, we can use Goedert's regression model to generate a forecast of what would have occurred in 2012 in Wisconsin - as well as in a state resembling the country as a whole under a neutral process (i.e., a court-drawn or bipartisan plan). His regression model includes the following variables $(2015,11)$ :

1. Whether a district plan was drawn by Democrats or Republicans (court-drawn and bipartisan plans are the excluded category)
2. A state's African American population percentage
3. A state's Hispanic population percentage

[^42]4. The percentage of a state that is urbanized (according to the Census)
5. The statewide Democratic vote
6. The number of congressional seats.

With the coefficients of this model, and the appropriate data for Wisconsin (or any other state), we can calculate what the expected bias would be for a plan in $2012 .{ }^{9}$ The dependent variable here is a measure of bias almost identical to the efficiency gap, with positive values indicating a pro-Democratic bias, and negative values a pro-Republican bias. Because this is a linear regression, we can multiply each coefficient by the value of the independent variable, and then sum the results to generate a forecast from any set of data values. In Table B, I set both Democratic and Republic Gerrymanders to 0, simulating a neutrally-drawn plan:

[^43]| Table B <br> Goedert's Regression Model for 2012 <br> Dependent Variable: <br> Pro-Democratic Bias in a District Plan |  |  |  |
| :---: | :---: | :---: | :---: |
| Variable <br> Name | (a) <br> Coefficient Value | (b) <br> Variable value for Wisconsin | Value $\text { (a) } \times(b)$ |
| Democratic Gerrymander | 16.6 | 0 | 0 |
| Republican Gerrymander | -13.6 | 0 | 0 |
| \% Black | -0..29 | 6.6 | -1.914 |
| \% Hispanic | 0.77 | 6.5 | 5.005 |
| \% Urbanized | -0.72 | 70.2 | -50.544 |
| Statewide <br> Democratic Congressional Vote | 0.11 | $\begin{gathered} 50.8 \\ (2012) \end{gathered}$ | 5.588 |
| Number of Seats | -0.16 | 8 | -1.28 |
| Constant | 45.0 | 1 | 45 |
| Total | (sum of | values) | 1.855 |

Goedert's regression model thus predicts that if Wisconsin had a neutrally drawn plan in 2012, the resulting map would have had a pro-Democratic bias of $1.855 \%$. In other words, in the absence of unified Republican control over the redistricting process, Wisconsin's demographic, geographic, and political characteristics would have resulted in a small natural Democratic advantage. And this is no fluke of the state or the election year. We can also use Goedert's model to predict what would happen in a state resembling the United States as a whole (i.e., a state that is $13.2 \%$ black, $17.4 \%$ Hispanic, $80.7 \%$ urbanized, $51 \%$ Democratic, and with
8.7 congressional seats ${ }^{10}$ ). Substituting these values into the regression model shows that in an "average" state, a neutrally-drawn map would have had a pro-Democratic bias of $0.684 \%$ in 2012.

Goedert's 2014 variant of the model $(2015,13)$ further predicts that Wisconsin would have had a pro-Democratic bias of $4.392 \%$ in 2014, and that the average state would have had a pro-Democratic bias of $1.589 \%$. At this point, it is hard to see what is left of the thesis that political geography inherently favors Republicans. If anything, Goedert's own published analysis shows that Wisconsin's political geography slightly favors Democrats.

## C. Accepted Measures of Geographic Concentration and Isolation Show that Democrats and Republicans are Equally Dispersed

In arguing that Republicans in Wisconsin enjoy a natural geographic advantage, both Trende and Geodert use ad hoc, unorthodox measures of concentration that are neither relevant nor accepted by the academic literature. In fact, there exist widely accepted metrics of geographic concentration and dispersion, used by geographers and demographers to study spatial patterns. Two of the most common are Global Moran's I (Anseln 1995; Cho 2003), and the Isolation Index (Glaeser and Vigdor 2012; Reardon 2004). I use these metrics to determine how Democrats and Republicans in Wisconsin are actually distributed.

Moran's I is a measure of spatial autocorrelation, or how values of a variable in space correlate with values in nearby space. It can be calculated for an entire geographic system (Global Moran's I), or for any specific point in space (Local Moran's I). The Isolation Index indicates, for the average member of a group residing in a certain geographic unit (such as a ward), what share of the member's neighbors in the unit belong to the same group (Iceland and Weinberg 2002, 120). It measures how geographically isolated a group is (Reardon 2004, 153), and it can easily be adjusted, by deducting a group's share of the statewide population, to show how much more isolated a group is than we would expect given its statewide size (Glaeser and Vigdor 2012, 2). Both Moran's I and the Isolation Index are widely used in studies of residential segregation and sorting (Chung and Brown 2007; Massey and Denton 1989; Glaeser and Vigdor 2012; Dawkins 2007; Reardon 2004; Iceland and Weinberg 2002), epidemiology (Moore and Carpenter 1999), network effects (Cho 2003), and political geography (Glaeser and Ward 2005). The measures are also used by the U.S. Census Bureau itself (Iceland and Weinberg 2002).

Both Moran's I and the Isolation Index are directly applicable to the issue of measuring the geographic distribution of Democrats and Republicans in Wisconsin. In this context, Global Moran's I tells us how likely Democrats are to live clustered next to other Democrats (and Republicans to Republicans), and the Isolation Index, adjusted as noted above, tells us to what extent the average Democrat (or Republican) lives in a ward that is more heavily Democratic (or Republican) than the state as a whole. I use these indices to directly assess the geographic distribution of Democrats, and, more importantly, to compare it to the geographic distribution of Republicans.

[^44]Global Moran's I is analogous to a correlation coefficient, and ranges from -1 to 1 ; scores close to 1 indicate a very high spatial correlation (i.e., clustering) of Democrats (or Republicans). The Isolation Index ranges from 0 to 1 , and, adjusted as noted above, indicates to what extent the average Democrat or Republican lives in a ward that is more heavily Democratic or Republican than Wisconsin as a whole. In calculating both measures, I use the ward as the basic unit of geography and actual Assembly votes. ${ }^{11}$ Because I only have geodata for the current wards, I only estimate Global Moran's I for 2012 and 2014. For the Isolation Index, I compute scores dating back to 2004. Both Global Moran's I and the Isolation Index are asymmetrical, and so must be calculated separately for Democrats and Republicans.

Table C shows the values of the Isolation Index, adjusted as noted above, for Democrats and Republicans in Wisconsin from 2004 to 2014:

|  | Table C <br> Isolation Index |  |
| :---: | :---: | :---: |
|  | Dem- <br> Rep | Rep- <br> Dem |
| 2014 | 0.23 | 0.20 |
| 2012 | 0.14 | 0.12 |
| 2010 | 0.15 | 0.17 |
| 2008 | 0.15 | 0.14 |
| 2006 | 0.16 | 0.17 |
| 2004 | 0.20 | 0.21 |

As is evident from Table C, Democrats were slightly less isolated than Republicans in 2004, 2006, and 2010, and slightly more so in 2008, 2012, and 2014. In all cases, the differences in isolation were very small, amounting to only one to three percentage points (out of a scale extending from $0 \%$ to $100 \%$ ). In the 2012 election, for instance, the average Democrat lived in a ward whose Democratic vote share was $14 \%$ more Democratic than the state as a whole; analogously, the average Republican lived in a ward whose Republican vote share was $12 \%$ more Republican than the entire state. In the previous election, it was Republican voters who were more isolated than Democratic voters ( $17 \%$ versus $15 \%$ ). This analysis in no way supports the claim that Republicans are more advantageously distributed than Democrats; on the contrary, both parties' supporters are almost identical in their geographic isolation over the last decade, and there is no clear temporal pattern. In some years, Democrats are marginally more isolated than Republicans, and in other years Republicans are marginally more isolated than Democrats.

[^45]The results are very similar with the Global Moran's I, again calculated for Democrats and Republicans in Wisconsin, although only for the two elections (2012 and 2014) for which the geodata is readily available:

|  | Table D <br> Global Moran's I |  |
| :---: | :---: | :---: |
|  | Democrats | Republicans |
| 2014 | 0.75 | 0.68 |
| 2012 | 0.68 | 0.69 |

Here, we see that Democrats were slightly less spatially concentrated than Republicans in 2012, but slightly more spatially concentrated in 2014. The differences in both cases are tiny: 0.01 in 2012 and 0.07 in 2014, on a scale that stretches from -1 to 1 . The message is quite clear: both Democrats and Republicans in Wisconsin tend to live near one another in distinct clusters, but there is no evidence that Democrats are more geographically clustered than Republicans.

Accordingly, two widely used and accepted measures of geographic distribution show no consistent pattern, and no material difference in how Wisconsin's Democrats and Republicans are dispersed spatially. In no sense, therefore, is it an accurate statement that Democrats are much more concentrated than Republicans - the unsubstantiated claim that comprised the core of both Trende's and Geodert's arguments about natural gerrymanders.

## III. Trende's Claim That My Vote Model Is Biased Is Incorrect

Trende claims that there may be "a systematic bias involved in imputing presidential results to state House results" (paragraph 135). As evidence he points to Figures 2 and 3 in my original report, which display the relationship between the ward level presidential vote and the ward level Assembly vote. Trende notes that Figure 2 shows that there is close to a $1: 1$ relationship between Republican presidential and Assembly votes, as the dots on the graph are distributed around the 45-degree line:


However, Trende claims that the relationship is different for Democratic votes (Figure 3 in my original report):


Here, Trende argues, the "dots systematically fall below the line, often creating differences on the order of 10 percent" (paragraph 138). This pattern, he asserts, will "skew the imputation" of votes, resulting in "too many votes [being] imputed in wards reporting a high number of Democratic votes" (paragraph 139).

Trende is completely and unambiguously wrong in this claim, which belies a fundamental lack of understanding of multiple regression and the causes of bias in statistical models. Trende appears to believe that I simply assumed that ward level Democratic Assembly votes are actually equal to ward level Democratic presidential votes, or that in estimating the Assembly vote in uncontested wards I merely used the value of the presidential vote (presumably because that is how he imputes the vote in uncontested districts in his own analysis; deposition page 83).

That is wrong. I displayed this graph merely to show that there is in fact a strong relationship between the two variables. The fact that the Democratic Assembly vote tends to fall below the presidential vote is completely irrelevant to any possible bias. In fact, regression analysis estimates the relationship between the two quantities by identifying the slope of the line that relates them, not how the relationship varies across a 45-degree line.

Below (Figure D) is a graph that plots the data in Figure 3 of my original report along with a fitted line of predicted values from a bivariate regression of the Democratic Assembly vote on the Democratic presidential vote. The red line consists of the predicted values of the Democratic Assembly vote in each ward:


Here, we see that the fitted line runs exactly down the middle of the plotted points. My regression analysis of the Democratic Assembly vote (Table 1 in my original report) shows that the coefficient for the Democratic presidential vote is 0.931 ( $\mathrm{p}<0.0001$ ), which is precisely the pattern than we see in the bivariate relationship above. In a linear model, this coefficient is the
slope of the line that relates the presidential vote to the assembly vote. It is less than 1 (a 45degree line), indicating that the Assembly vote rises more slowly than the presidential vote; i.e., the predicted Assembly vote will lie below the 45-degree line in Figure 2.

And, as is immediately apparent from the actual results of my regression (Figure 4 in my original report, which plots the actual vs. predicted ward level votes), there is no bias in the results. In this graph, the 45 -degree line is where the predicted Assembly vote would fall if it were exactly equal to the actual Assembly vote:

Figure 4: Ward Level Predicted vs. Actual Assembly Vote - 2012


Trende's criticism on this point is utterly misinformed. No one with a solid understanding of quantitative methods or regression analysis would have made it.

## IV. Trende's Claim That My Efficiency Gap Calculations Ignore Incumbency, Candidate Quality, and Campaign Spending

In paragraphs 140-143, Trende criticizes my efficiency gap calculations for failing to take into account factors that can affect election results, such as get-out-the vote drives, candidate quality, recruitment, and campaign spending.

Trende offers no evidence that these factors would actually have a material effect on my estimates if I had more directly taken them into account. And he ignores the fact that any
estimation of the results of a hypothetical district plan utilizes baseline estimates that, in effect, average out the effects of these factors (Gelman and King 1990; 1994). That is to say, my regression model does implicitly incorporate these factors, in its analysis of the relationship between the presidential vote (where none of these variables will affect the vote) and the Assembly vote (where they are all incorporated into the estimates).

Moreover, Trende's criticism overlooks the point that my model is based on precisely the same information that the authors of Act 43 considered in estimating the likely partisan effects of the new districts. In particular, Gaddie's analysis of the partisan effects in the new Act 43 districts was functionally equivalent to mine and based on exactly the same considerations.

Like his complaints about alleged bias in the regression analysis that I discuss above, Trende's criticism is uninformed and betrays a lack of knowledge of how hypothetical district plans are evaluated.

## V. Goedert's Claim That My Efficiency Gap Calculations Incorporate Information Not Available to Act 43's Designers, and Ignore the Effects of Incumbency

Goedert criticizes my analysis for incorporating information that map drawers did not have (2012 election results), and for ignoring information that map drawers would have taken into account (incumbency in particular).

The first criticism is incorrect, as Act 43 's designers in fact had information functionally equivalent to the 2012 election results in their possession, in the form of Gaddie's Act 43 district level estimates. These estimates, like my own, are baseline measures of partisanship, and they correlate almost perfectly with my results ( $\mathrm{r}^{2}=0.96$ ). In his deposition, Gaddie described in detail his method, which like mine assumed that all seats would be contested and that no incumbents would run (Gaddie Deposition, pp. 197, 198, 201, 202, 204):

Let's suppose we have a seat with an incumbent and a seat without an incumbent and
each one has an Assembly election. The party of the incumbent is presumably going to do a little stronger in the district where they have an incumbent than in an open seat. So I can't really take -- Let's suppose I move precincts from the open seat into that incumbent seat. I can't really take those open seat Assembly votes, add them, compare them to the percentage for the incumbent running for the same party, get an accurate estimation of the partisanship and the competitiveness of the district. So we attempt to create a substitute measure. Statewide elections are held in all precincts, they're held in all constituencies, so one thing that we often do is we do what we call reconstituted elections, or proxy elections, where we'll take one election or a composite of elections, like I described previously, and attempt to create some measure of partisan competitiveness, an expected vote or what we call a normal vote, what the vote would usually do without an incumbent in the district." (Gaddie Deposition, pp. 204-5)

To highlight the similarity between Gaddie's pre-2012 estimates and my own estimates using 2012 election results, below is a graph plotting the two sets of data (Figure 7 in my original report, p. 30):


This graph shows that the information the Act 43 authors relied on when drawing their map (the Gaddie estimates) and my estimates, are nearly identical. This is largely because they are both estimates of the same underlying quantity - the baseline partisanship of a hypothetical Assembly district. Goedert dismisses the nearly perfect correlation as "mostly coincidental" (p. 17), but offers no analysis or data to support this conclusion. It is simply an assertion offered without evidence.

And it is an entirely unpersuasive assertion for the additional reason that election results in Wisconsin (and in most states) are extremely highly correlated from one election to the next. For example, Wisconsin's counties remained geographically constant between 2008 and 2012, and Trende supplied information about the presidential vote in each county in each of these years. The 2008 county level presidential vote and the 2012 county level presidential vote are almost perfectly correlated $\left(r^{2}=0.96\right)$, indicating that it would make no difference whether Act 43 was assessed using the former or the latter. ${ }^{12}$ Either way, the same conclusion would follow: that

[^46]the map is an extreme Republican gerrymander, and that the authors of Act 43 had information in their possession that predicted it.

Second, Goedert claims that map drawers do not ignore incumbency when drawing maps. That will generally be true when map drawers are trying to figure out which incumbent should be included in which district. But when it comes to estimating the likely partisanship of the new districts, ignoring incumbency (that is, controlling for it) is precisely what the drawers of Act 43 did, as Gaddie noted in his description of his methods. This approach is sensible since incumbents can be defeated, retire, run for higher office, or switch parties over a plan's decadelong lifespan. A map's authors will typically want to ensure that their projections do not depend on particular incumbents continuing to run in particular districts.

In any event, including incumbency in no way changes my substantive conclusions about Act 43 or the Demonstration Plan. I recalculated the efficiency gap for both maps, using my baseline partisan estimate and then incorporating incumbency into the model. For Act 43, I used the actual incumbents who ran in the plan's districts, with the adjustments noted in my report to account for paired incumbents and those who lost in primaries (p. 18, footnote 14). ${ }^{13}$ For my plan, I geocoded incumbents' home addresses ${ }^{14}$ and then identified which districts had incumbents residing in them using Maptitude for Redistricting. Table E shows the resulting efficiency gap calculations, and compares them to the open seat baseline I generated in my report:

| Table E |  |
| :---: | :---: | :---: |
| Efficiency Gap Calculations |  |
| with Incumbents |  |

The efficiency gap increases marginally for both plans (by $1.69 \%$ for the Demonstration Plan and $2.46 \%$ for Act 43), in large part because there were more Republican (50) than

[^47]Democratic (24) incumbents running in 2012. With twice as many incumbents, Republicans will win more seats than in the open seat baseline even though the Republican vote percentage remains below $50 \%$ in both cases. It is thus apparent that taking incumbency into account has no effect on my conclusion that Act 43 was an egregious partisan gerrymander; the substantive inferences are identical, with or without incumbency. ${ }^{15}$

## VI. Goedert's Claim That I Did Not Perform Sensitivity Testing for Act 43's or the Demonstration Plan's Efficiency Gaps

Goedert criticizes the efficiency gap calculations for both Act 43 and the Demonstration Plan, arguing that I "provide no estimates for the efficiency gap of the demonstration plan under the range of plausible election outcomes facing legislators at the time they were drawing the map" (p. 16), and that I conduct no "sensitivity testing" of my calculations of Act 43's efficiency gap.

I note that Goedert has not provided any actual analysis showing that this sensitivity testing would have materially altered my conclusions, or even any citations showing that such testing is necessary to evaluate the adequacy of my calculations.

Still, it is possible to show that my calculations are robust to significant changes in the electoral environment. Using Jackman's historical estimates of the statewide Assembly vote in Wisconsin, I can determine the plausible variation of the overall vote over the course of a decade. Since 1992, the statewide Democratic percentage of the Assembly vote has ranged from a high of $54.6 \%$ (in 2006) to a low of $46.4 \%$ (in 2010). The Democratic share of the statewide vote in 2012 was $51.2 \%$ in my baseline calculations, which suggests a plausible range of $-5 \%$ to $+3 \%$ in conducting a sensitivity analysis. In effect, this approach asks whether Act 43's and the Demonstration Plan's efficiency gaps would be durable in the face of massive Democratic or Republican waves - an extremely rigorous test that exceeds what is normally found in the literature.

Following Goedert's method of applying a uniform swing (p.21), I can estimate the effects that these swings will have on the efficiency gap, both for Act 43 and for the Demonstration Plan. To maintain consistency and to address his concern that I did not

[^48]incorporate incumbency in my baseline, I estimate the effects while treating as incumbents all of the prevailing candidates in the incumbent baseline (see Efficiency Gap With Incumbency in Table E above). Functionally, this simulates what would happen over the remainder of the decade (2014-2020) if after the 2012 elections Wisconsin experienced a Democratic or Republican wave.

The results are shown in the following two tables, the first for the Demonstration Plan (Table F), and the second for Act 43 (Table G). ${ }^{16}$ For the Demonstration Plan, the efficiency gap remains well below the plaintiffs' suggested 7\% threshold, even when the statewide vote reaches the most extreme values either party has seen over the last three decades. Specifically, the efficiency gap goes to $3.75 \%$ in the event of a Democratic wave akin to that of 2006, and to $0.14 \%$ if a Republican wave like that of 2010 occurs. For Act 43, however, the efficiency gap remains extremely large and above the threshold absent a Republican wave, ranging from $14.88 \%$ in a Democratic wave to $6.09 \%$ in a Republican wave. Moreover, the sensitivity testing shows that even if the Democrats obtained over $54 \%$ of the statewide Assembly vote - equal to their best performance in a generation - they still would not capture a majority of the Assembly, gaining only 45 seats. Act 43 's gerrymandering thus effectively insulates the Republican Assembly majority from all plausible shifts in voter sentiment.

[^49]|  | Table F <br> Efficiency Gap Estimates, Uniform Swing <br> Demonstration Plan |  |  |
| :---: | :---: | :---: | :---: |
|  | D Minus 5 <br> (all incumbents) | My Plan <br> Incumbent <br> Baseline | D Plus 3 <br> (all incumbents) |
| party split (R-D) | $51-48$ | $50-49$ | $43-56$ |
| Rep share of <br> Seats | $52 \%$ | $49 \%$ | $43 \%$ |
| Wasted <br> Republican Votes | 711,621 | 655,733 | 660,706 |
| Wasted <br> Democratic Votes | 707,789 | 766,234 | 767,927 |
| Gap | $(3,833)$ | 110,501 | 107,221 |
| Total Democratic <br> Votes | $1,334,535$ | $1,455,846$ | $1,571,786$ |
| Total Republican <br> Votes | $1,504,285$ | $1,388,087$ | $1,285,480$ |
| Total Votes | $2,838,820$ | $2,843,933$ | $2,857,266$ |
| Efficiency Gap <br> (gap/total votes) | $\mathbf{- 0 . 1 4 \%}$ | $\mathbf{3 . 8 9 \%}$ | $\mathbf{3 . 7 5 \%}$ |


|  | Table G <br> Efficiency Gap Estimates, Uniform <br> Swing <br> Act 43 Districts |  |  |
| :---: | :---: | :---: | :---: |
|  | D Minus 5 <br> (all incumbents) | Act 43 Actual | D Plus 3 <br> (all incumbents) |
| Party Split (R-D) | $60-39$ | $60-39$ | $54-45$ |
| Rep share of <br> Seats | $61 \%$ | $61 \%$ | $55 \%$ |
| Wasted <br> Republican Votes | 622,966 | 509,747 | 500,607 |
| Wasted <br> Democratic Votes | 795,844 | 911,954 | 924,690 |
| Gap | 172,878 | 402,207 | 424,083 |
| Total Democratic <br> Votes | $1,317,061$ | $1,452,132$ | $1,551,205$ |
| Total Republican <br> Votes | $1,520,560$ | $1,391,269$ | $1,299,388$ |
| Total Votes | $2,837,621$ | $2,843,401$ | $2,850,593$ |
| Efficiency Gap <br> (gap/total votes) | $\mathbf{6 . 0 9 \%}$ | $\mathbf{1 4 . 1 5 \%}$ | $\mathbf{1 4 . 8 8 \%}$ |

Figure E below shows these results graphically: the red x's are the efficiency gap estimates for the Demonstration Plan, and the blue diamonds the estimates for Act 43. The dotted line is at plaintiffs' suggested threshold of $7 \%$. The figure clearly demonstrates that even across huge partisan swings, the efficiency gap under Act 43 remains very large, and the efficiency gap for the Demonstration Plan remains very small. In fact, Table G demonstrates the remarkable efficiency of Act 43's gerrymander, in that an additional 5\% of the Republican statewide vote does not add a single seat to the Republican Assembly majority. The important feature here is how well Act 43 protects against a Democratic wave. This is further powerful confirmation of the durability of Act 43's bias - and the durable lack of bias of the Demonstration Plan.

# Figure D: Sensitivity Analysis -5 to +3 Democratic Swing 

15\%-1

* Demonstration Plan
- Act 43 Districts


## VII. Conclusion

In their criticism of my report, both Trende and Goedert offer nothing but supposition, speculation, irrelevant discourse about Wisconsin political history, extraneous discussion of congressional redistricting in other parts of the United States, wildly inapposite and inaccurate conjecture about the geographic concentration of Democrats as a possible source of the proRepublican bias of Act 43, unreliable methodologies, and minor quibbles that have no consequences for my conclusions. Neither Trende nor Goedert has conducted any valid analysis of either Act 43 or the Demonstration Plan - in fact, they make no mention at all of the specifics of the Demonstration plan.

Most significantly, nothing in their reports undercuts my fundamental conclusion that Act 43 constituted an egregious and durable gerrymander, and that it was entirely possible to draw a
neutral map that met or exceeded Act 43 on all legal dimensions. If anything, the sensitivity testing substantially bolsters this conclusion, since it shows that Act 43's large efficiency gap and the Demonstration Plan's small one are durable in the face of enormous changes in Wisconsin's electoral environment.

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Revised: March 31, 2016
/s/ Kenneth R. Mayer
Kenneth R. Mayer, Ph.D.
Department of Political Science
University of Wisconsin-Madison

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## Sensitivity of the Efficiency Gap to Uniform Swing

How sensitive is the efficiency gap to reasonable swings in vote shares? In his report, Goedert asserts that it is extremely sensitive (pp. 11-15), but his claim is based on a small number of examples (pp. 12-13) as well as his own work at the congressional level involving only two elections (Goedert 2015). Sections 1-4 of my rebuttal report show that the first efficiency gap observed under a plan is a reliable indicator of the efficiency gap's magnitude and direction over the remainder of the plan's lifespan. These sections, however, are based on historical efficiency gap data rather than the "sensitivity testing for future results" deemed "crucial" by Goedert (p. 13). Accordingly, we conduct sensitivity testing here of exactly the kind earlier carried out by Stephanopoulos \& McGhee (pp. 889-90, 898-99) and recommended by Goedert. This testing confirms the findings in Sections 1-4 of my rebuttal report, and further corroborates my conclusions therein about the efficiency gap's durability and reliability.

Methodologically, we investigate the behavior of the efficiency gap when we perturb it by mimicking "uniform swing" across a jurisdiction. That is, a given election produces a set of vote shares across districts. A new hypothetical election is considered in which all vote shares move up or down by a predetermined quantity (i.e., the "swing"); since all districts move by the same amount, this technique is known as uniform swing. In real-world elections swings are never precisely uniform, and so this method is widely considered to be a simplification; on the other hand, modeling or predicting swing district by district is quite difficult, especially for state legislative elections where we often lack useful district-level predictors of swing (or, more tellingly, predictors of the way the swing in a given state legislative district might depart from the statewide swing).

We restrict the following exercise to elections since the 2010 round of redistricting. For each election we simulate a series of uniform swings, evenly spaced between $-5 \%$ to $+5 \%$, a quite
large set of swings by the standards of state legislative elections. For instance, swings in Wisconsin state legislative elections from 1972 to 2014 are estimated to range between -7.6 percentage points from 2008 to 2010 (Democratic share of two-party vote, averaged by district) and +5.0 percentage points from 2004 to 2006. Similarly, Stephanopoulos \& McGhee found that a swing of $+/-5.5$ percentage points covered the vast majority of state legislative elections from 1972 to 2012 (p. 874).

At each level of uniform swing, we record the new vote shares and seat shares (some seats change hands if the swing pushes Democratic two-party vote share to the other side of $50 \%$ ) and recompute the efficiency gap. We then examine how much the simulated efficiency gaps-generated under different levels of uniform swing-depart from the efficiency gap observed under the actual election. In particular, if relatively small swings produce large changes in $E G$, we might rightly be concerned about the stability and reliability of the efficiency gap as a characterization of a district plan. Keep in mind that this exercise keeps the district plan as it is and simply shifts vote shares up and down over a range of hypothetical levels of statewide swing, held constant over districts.

Figure 1 shows the relationships between efficiency gaps estimated using actual election results in state legislative elections held since the 2010 round of redistricting, and efficiency gaps estimated using a range of uniform swings. When uniform swing is zero, the simulation exercise leaves the actual election results unperturbed, and we simply recover the original efficiency gap estimates; all the data in the panel labelled "Swing +0.0 " lies on the 45 -degree line. As we increase the magnitude of hypothetical levels of uniform swing, the relationship between the observed efficiency gaps and the simulated efficiency gaps weakens, but only by a moderate amount. Even at high levels of uniform swing (approaching $+/$ - five percentage points), the relationship between observed efficiency gaps and simulated efficiency gaps remains of significant strength; the blue line in each panel of Figure 1 is a regression line and in every case has a large
and unambiguously positive slope, indicating a positive correlation between actual and simulated efficiency gaps.


Figure 1: Actual efficiency gaps from state legislative elections 2012 to 2014 (horizontal axis), and corresponding simulated efficiency gaps generated by varying levels of uniform swing. Vertical lines indicate $95 \%$ confidence intervals. Dark diagonal lines are at forty-five degrees, the fit to the data that would result if actual and simulated efficiency gaps were equal (as is the case when the simulated level of uniform swing is set to zero, as in the middle panel of the second row). The blue line indicates a regression fit. For small to even moderately large values of uniform swing, there is a high degree of correspondence between the actual and simulated efficiency gaps.


Figure 2: Correlation between actual efficiency gaps and simulated efficiency gaps (top row) and proportion of simulated efficiency gaps with same sign as actual efficiency gaps (bottom row), by hypothetical levels of uniform swing (horizontal axis). Vertical lines are $95 \%$ confidence intervals. The three columns correspond to actual efficiency gaps that are low in magnitude (less than .03 in absolute value; left column), medium ( .03 to .07 in absolute value, middle column) and high (above .07 in absolute value, right column). When uniform swing is zero, the simulated efficiency gaps correspond to the actual efficiency gaps, and so the correlation between the two sets of efficiency gaps is exactly 1.0 and $100 \%$ of the simulated efficiency gaps have the same sign as the actual efficiency gaps.

The top row of Figure 2 displays correlations between actual efficiency gaps and simulated efficiency gaps, under different hypothetical levels of uniform swing (horizontal axis), with separate panels for low, medium, and high values of actual efficiency gaps. Note that when uniform swing is zero, the simulated efficiency gaps correspond to the actual efficiency gaps, and so the correlation between the two sets of efficiency gaps is exactly 1.0. As levels of uniform swing increase, the correlation between actual and simulated efficiency gaps diminishes. Small efficiency gaps (less than .03 in absolute value) are less resistant to perturbations from uniform swing; at high levels of uniform swing for small actual efficiency gaps, the correlation between actual efficiency gaps and simulated efficiency gaps approaches zero. However, larger values of the efficiency gap are much more robust to perturbations from uniform swing. In fact, for large actual efficiency gaps (greater than .07 in magnitude), the correlation between actual and simulated efficiency gaps stays impressively large over the entire range of uniform swing levels considered here (top right panel of Figure 2).

The bottom row of Figure 2 displays the proportion of simulated efficiency gaps that have the same sign as actual efficiency gaps, under a range of hypothetical levels of uniform swing (horizontal axis), again with separate panels for low, medium, and high values of actual efficiency gaps. Again we see that small efficiency gaps-less than .03 in magnitude and hence relatively close to zero-are reasonably likely to flip signs under moderate to large values of hypothetical uniform swing: about half of these small efficiency gap estimates flip signs when subjected to reasonably large statewide swings one way or the other. But large efficiency gaps-those greater than .07 in magnitude-show great resistance to flipping signs even in the face of moder- ate or even large hypothetical statewide swings (lower right panel of Figure 2). None of the large efficiency gaps flip signs when swings are below 2.5 percentage points and barely any flip signs even we consider larger statewide swings. Just $11 \%$ of actual efficiency gaps greater than .07 in magnitude flip signs when exposed to a very large, hypothetical statewide swing of minus five percentage points and only $9 \%$ flip signs when we consider a statewide swing of positive five percentage points.

In short, efficiency gap estimates display a high level of resistance to perturbations from even large levels of uniform swing. This further bolsters our confidence that the efficiency gap is measuring a durable property of a district plan. Moreover, the analysis reported here demonstrates that efficiency gaps are especially reliable when they are large, as is the case for the efficiency gaps generated under the Wisconsin plan. The efficiency gap changes if vote totals change, even if the district plan remains constant; this is "hardwired" into the definition and accompanying arithmetic of the efficiency gap. But to reiterate a conclusion from my original report: the amount of election-to-election variation in the efficiency gap is small relative to the var-iation in the efficiency gap across plans.

Exhibit 122 - Average Efficiency Gaps for Wisconsin Plans (1970s-2010s)

| Cycle | Designer | Average Efficiency Gap |
| :---: | :---: | :---: |
| 1970 s | Divided government | $-0.3 \%$ |
| 1980 s | Court | $-1.9 \%$ |
| 1990 s | Court | $-2.4 \%$ |
| 2000 s | Court | $-7.6 \%$ |
| 2010 s | Professor Mayer | $-1.9 \%$ |

The measure of partisanship should exist to establish the change in the partisan balance of the district. We are not in court this time; we do not need to show that we have created a fair, balanced, or even a reactive map. But, we do need to show to lawmakers the political potential of the district.

I have gone through the electoral data for state office and built a partisan score for the assembly districts. It is based on a regression analysis of the Assembly vote from 2006, 2008, and 2010, and it is based on prior election indicators of future election performance.

I am also building a series of visual aides to demonstrate the partisan structure of Wisconsin politics. The graphs will communicate the top-to-bottom party basis of the state politics. It is evident, from the recent Supreme Court race and also the Milwaukee County executive contest, that the partisanship of Wisconsin is invading the ostensibly nonpartisan races on the ballot this year.

EXHIBIT
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| Statewide2_Milwaukee_Gaddie_4_16_11_V1_B |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  | Senate |  |  |  |
| DISTRICT |  |  |  | DISTRICT |  |  |  |
| 2 | 51.25\% | 51.22\% | 0.07\% | 1 | 54.04\% | 53.48\% | -0.56\% |
| 2 | 54.93\% | 53.82\% | 1.11\% |  |  |  |  |
| 3 | 56.10\% | 55.81\% | -0.29\% |  |  |  |  |
| 4 | 53.31\% | 53.76\% | 0.45\% | 2 | 55.44\% | 54.15\% | 1.29\% |
| 5 | 53.74\% | 55.31\% | 1.57\% |  |  |  |  |
| 区 | 59.774 | 53.47\%6 | -5.30\% |  |  |  |  |
| 7 | 48.20\% | 44.42\% | -3.78\% | 3 | 40.52\% | 37.54\% | -2.98\% |
| E | 22.39\% | 21.22\% | -1.17x |  |  |  |  |
| 9 | 36.73\% | 35.67\% | -1.06\% |  |  |  |  |
| 20 | 10.27\% | 16.52\% | 6.25\% | 4 | 17.58\% | 19.41\% | 1.83\% |
| 12 | 11.91\% | 17.63\% | 5.72\% |  |  |  |  |
| 12 | 29.23\% | 24.97\% | 4.31\% |  |  |  |  |
| 13 | 43.67\% | 55.57\% | 11.90\% | 5 | 50.62\% | 54.90\% | 4.28\% |
| 24 | 59.06\% | 54.40\% | -4.66\% |  |  |  |  |
| 15 | 48.21\% | 54.61\% | 6.40\% |  |  |  |  |
| 16 | 14.21\% | 13.02\% | -1.19\% | 5 | 14.12\% | 17.86\% | 3.7446 |
| 27 | 13.21\% | 22.95\% | 9.74 k |  |  |  |  |
| 18 | 15.28\% | 15.86\% | 0.58\% |  |  |  |  |
| 19 | 29.15\% | 26.71\% | -2.44\% | 3 | 41.13\% | 39.65\% | -1.48\% |
| 20 | 43.71\% | 41.73\% | -1.98\% |  |  |  |  |
| 21 | 51.92\% | 52.85\% | 0.93\% |  |  |  |  |
| 22 | 39.05\% | 56.14\% | 17.09\% | 8 | 52.82\% | 62,31\% | 9.49\% |
| 23 | 51.70\% | 61.82\% | 10.12\% |  |  |  |  |
| 24 | 67.29\% | 69.84\% | 2.55\% |  |  |  |  |
| 25 | 52.79\% | 53.33\% | 0.54\% | 9 | 52.96\% | 57.67\% | 4.71\% |
| 20 | 45.42\% | 54.99\% | 9.57\% |  |  |  |  |
| 27 | 59.20\% | 54.23\% | 5.03\% |  |  |  |  |
| 28 | 54.85\% | 54.94\% | 0.09\% | 10 | 53.14\% | 53.30\% | 0.16\% |
| 29 | 51.32\% | 5092\% | -0.40\% |  |  |  |  |
| 30 | 53.29\% | 53.81\% | 0.52\% |  |  |  |  |
| 31 | 67.57\% | 56.05\% | -11.52\% | 11 | 67.64\% | 58.19\% | 9.4556 |
| 32 | 61.06\% | 62.73\% | 1.67\% |  |  |  |  |
| 33 | 72.24\% | 56.31\% | -15.93\% |  |  |  |  |
| $3 \pi$ | 54.51\% | 53.44\% | -1.078 | 12 | 53.37\% | 53.89\% | 0.52\% |
| 35 | 52.30\% | 53.29\% | 0.99\% |  |  |  |  |
| 36 | 53.06\% | 55.07\% | 2.016 |  |  |  |  |
| 37 | 51.33\% | 60.43\% | 9.1056 | 13 | 59.22\% | 61.69\% | 247\% |
| 38 | 65.80\% | 62.52\% | $3.28 \%$ |  |  |  |  |
| 39 | 60.35\% | 62.04\% | 1.69 M |  |  |  |  |
| 30 | 58.50\% | 55.67\% | 2.83\% | 14 | 55.86\% | 55.64\% | -0.220 |
| 4I | 60.60\% | 55.2990 | 5.319 |  |  |  |  |
| 42 | 48.54\% | 55.97\% | 7.43\% |  |  |  |  |
| 43 | 44.14\% | 38.55\% | .5.59\% | 15 | 41.20\% | 38.75\% | -2.45\% |
| 44 | 36.74\% | 37.27\% | 0.53\% |  |  |  |  |
| 45 | 42.39\% | 40.82\% | -1.57\% |  |  |  |  |
| 45 | 42.07\% | 44.57\% | 2.50 m | 16 | 39.06\% | 36.54\% | -252x |
| 47 | 48.69\% | 39.36\% | -9.33\% |  |  |  |  |
| 48. | 28.03\% | 27.248 | -0.7990 |  |  |  |  |
| 43 | 49.68\% | 49.74\% | 0.06 k | 17 | 48.46\% | 49.23\% | 0.77\% |
| 50 | 52.08\% | 51.90\% | -0.18\% |  |  |  |  |
| 31 | $44.01 \%$ | 46.20\% | 2.190 |  |  |  |  |
| 52 | 57.39\% | 57.88\% | 0.49\% | 18 | 54.96\% | 55.05\% | 0.09\% |
| 53 | 62.74\% | 62.788\% | 0.043 |  |  |  |  |
| 54 | 45.08\% | 45.19\% | 0.11\% |  |  |  |  |
| 55. | 49.34\% | 57.94\% | 8.60\% | 19 | 53.32\% | 52.56\% | -0.76\% |
| 56 | 61.05\% | 53.44\% | 7.6189 |  |  |  |  |
| 57 | 47.26\% | 46,45\% | -0.819 |  |  |  |  |
| 58. | 70,90\% | 70.79\% | -0.11\% | 20 | 70.55\% | 68.06M | -2.490 |
| 39 | 72.74\% | 61.52\% | -11.22\% |  |  |  |  |
| 60 | 68.12\% | 72.32\% | 3.2006 |  |  |  |  |
| 5 | 35.98\% | 57.24\% | 21.26\% | 21 | 49.86\% | 57.79\% | 7.93\% |
| 62 | 44,35\% | 59.48\% | 15.13\% |  |  |  |  |
| 63 | 63.09\% | 56.78\% | -6.3140 |  |  |  |  |
| 69 | 35.66\% | 42.16\%6 | $6.50 \%$ | 22 | 47.56\% | 37.34\% | -10.22\% |
| 65 | 45.44\% | 36.0076 | -9.4449 |  |  |  |  |
| 56 | 59.12\% | 33.44\% | -25.68\% |  |  |  |  |
| 67 | 51.72\% | 51.63\% | -0.09\% | 23 | 49.98\% | 51.75\% | 1.77\% |
| 58 | 45.01\% | 50.00\% | $4.99 \%$ |  |  |  |  |
| 3 | 54,06\% | 53.67\% | -0.39\% |  |  |  |  |
| 70 | 49.74\% | 47.54\% | -2.20\% | 24 | 46.72\% | 45.54\% | -0.084 |
| 31 | 41.68\% | 41.01\% | $-0.67 \times$ |  |  |  |  |


| DSTEL |  | $\square$ |  | Disimic |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 12 | 49.03\% | 51.69\% | 2.66\% |  |  |  |  |
| 7 | 39.55\% | 40.05\% | 0.50\% | 25 | 44.88\% | 45.67\% | 0.79\% |
| 74 | 43.78\% | 45.03\% | 1.25\% |  |  |  |  |
| 75 | 51.71\% | 52.31\% | 0.60\% |  |  |  |  |
| 76 | 24.29\% | 20.80\% | -3.499 | 26 | 20.85\% | 20.85\% | 0.00\% |
| 77 | 23.83\% | 24.52\% | 0.64 c |  |  |  |  |
| 25 | 14.09\% | 17,18\% | 3.09\% |  |  |  |  |
| 29 | 37.49\% | 36.70\% | -0.79\% | 27 | 38,38\% | 40.45\% | 2.07\% |
| 80 | 42.15\% | 40.32\% | -1.83\% |  |  |  |  |
| 31 | 36.16\% | 44.54\% | 8.38\% |  |  |  |  |
| 82 | 58.59\% | 55.72\% | -2.87\% | 28 | 64,48\% | 62.49\% | -1.99\% |
| 83 | 69.70\% | 70.15\% | 0.45\% |  |  |  |  |
| 84 | 64.99\% | 61.26\% | 4.73\% |  |  |  |  |
| 55 | 48.91\% | 53.65\% | 4.7466 | 29 | 52.00\% | 54.23\% | 2.23\% |
| 85 | 54.56\% | 55.A78. | 0.91\% |  |  |  |  |
| 87 | 52.16\% | 53,42\% | 1.26\% |  |  |  |  |
| 88 | $44.85 \%$ | 58.65\% | $13.80{ }^{\text {c/ }}$ | 30 | 50.38\% | 52.29\% | 1.918 |
| 89 | 55.76\% | 55.58\% | -0.18\% |  |  |  |  |
| 90 | 49.59\% | - 40, $33 \%$ | -9.46\% |  |  |  |  |
| 91 | 45.87\% | 44,318\% | 1.56\% | 31 | 46.89\% | 44.94\% | 1.95\% |
| 92 | 50.79\% | 39.55\% | -11.24\% |  |  |  |  |
| 93 | 44.73\% | 51.15\% | 6.42\% |  |  |  |  |
| 94 | 51.57\% | 51.93\% | 0.36\% | 32 | 44.43\% | 44.63\% | 0,2086 |
| 95 | 36.02\% | 36.26\% | 0.24\% |  |  |  |  |
| 96 | 45.32\% | 46.4030 | 1.08\% |  |  |  |  |
| 97 | 59.96\% | 62.39\% | 2.43x | 33 | 68.84\% | 67.98\% | -0.8690 |
| 98 | 70.96\% | $67.99 \%$ | -2.97\% |  |  |  |  |
| 99 | 73.35\% | 72.66\% | -0.69x |  |  |  |  |


| Current Map |  |  | Now Map |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Sate GOP ( $55 \%$ \% ) | 27 | 7 | Sate GOP ( $55 \%$ \% 1 | 35 | 10 |
| Lean GOP (52.1-54.9\%): | 13 | 8 | New Lean GOP (52.1-54.99): | 17 | 8 |
| Total GOP Seats (safe - lean): | ${ }^{40}$ | 15 | Total GOP Seats (safe + lean): | 52 | 18 |
| 5wing ( $48.52 \%$ ) | 19 | 5 | New Swing (48-528) | 9 | 2 |
| Lean DEM (45.1-47.9\%): | 7 | 3 | New Lean DEM (45.1-47.994): | 5 | 2 |
| Safe DEM (-45\%): | 33 | 10 | Sate DEM (-45\%) : | 32 | 11 |
| Total DEM Seats (safe + lean) : | 40 | 13 | Total DEM Seats (sate + lean): | 38 | 13 |


| Final Map |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  | Senate |  |  |  |
| DiSTRICT |  |  |  | DIStrict |  |  |  |
| 1 | 51.15\% | 51.22\% | 0.078 | $\pm$ | 54.04\% | 53.73\% | -0.31\% |
| 2 | 54.93\% | 54.84\% | . 0.0996 |  |  |  |  |
| 3 | 56.10\% | 55.58\% | -0.52\% |  |  |  |  |
| 4 | 53.31\% | 53.47x | 0.16 K | 2 | 55.44\% | 55.23\% | 0.212 |
| 5 | 53.74\% | 54.28\% | 0.54\% |  |  |  |  |
| \% | 59.77\% | 58.33\% | 1.444\% |  |  |  |  |
| 7 | 48.20\% | 45.38\% | -2.82\% | 3 | 40.52\% | 38.12\% | -2.40\% |
| 8 | 22.39\% | 30.48\% | $8.09 \%$ |  |  |  |  |
| 9 | 36.73\% | 29.14\% | .7.59\% |  |  |  |  |
| 20 | 10.27\% | 12.5976 | 2.32\% | 4 | 17.58\% | 19.63\% | 2.05\% |
| 11 | 11.918 | 19.58\% | 7.67\% |  |  |  |  |
| 12 | 29.23\% | 27.51\% | -1.72\% |  |  |  |  |
| 12 | 43.67\% | 58.67\% | 15.00\% | 5 | 50.62\% | 57.72\% | 7.10\% |
| 24 | 59.06 N | 58.64\% | -0.42\% |  |  |  |  |
| 15 | 48.21\% | 35.48\% | 7.27\% |  |  |  |  |
| $\frac{15}{15}$ | 14.21\% | 10.54\% | 3.679 | 5 | 14.12\% | 15.55\% | 1.43\% |
| 17 | 13.21\% | 19.84\% | 6.63\% |  |  |  |  |
| 18 | 15.28\% | 14.94\% | -0.3489 |  |  |  |  |
| 19 | 29.15\% | 28.03\% | -1.12\% | 7 | 41.13\% | 40.53\% | -0,600\% |
| 20 | 43.71\% | 43.12\% | -0.59\% |  |  |  |  |
| 21 | 51.92\% | $52.94 \%$ | 1.02\% |  |  |  |  |
| 22 | 39.05\% | 56.82\% | 27.77\% | 8 | 52.82\% | 60.88\% | 8.06\% |
| 23 | 51.70\% | 57.64\% | 5.94\% |  |  |  |  |
| 24 | 67.29\% | 58.49\% | 8.80\% |  |  |  |  |
| 25 | 52.79\% | 53.26\% | 0.47\% | 9 | 52.96\% | 55.19 k | 2.23\% |
| 25 | 45.42\% | 55.97\% | $10.55 \%$ |  |  |  |  |
| 27 | 59.20\%6 | 56.19\% | -3.01\% |  |  |  |  |
| $2 \pi$ | 54.85\% | 55.00\% | 0.15\% | 10 | 53.14\% | 53.32\% | 0.18\% |
| 29 | 51.32\% | 50.97\% | -0.35\% |  |  |  |  |
| 30 | 53.29\% | 53.78\% | 0.49\% |  |  |  |  |
| 38 | 67.57\% | 56.33\% | -11.24\% | 11 | 67.64\% | 60.13\% | -7.51\% |
| 32 | 61.06\% | 62.27\% | 1.219- |  |  |  |  |
| 33 | 72.24\% | 61.81\% | -10.430 |  |  |  |  |
| 34 | 54.51\% | $55.22 \%$ | 0.71\% | 12 | 53.37\% | 54.39\% | $1.02 \%$ |
| 35 | 52.30\% | 52.99\% | 0.69\% |  |  |  |  |
| 36 | 53.06\% | 54.84\% | 1.785 |  |  |  |  |
| 37 | 51.33\% | 58.11\% | $6.78 \%$ | 13 | 59.22x | 60.17\% | 0.95\% |
| 38 | 65.80\% | 60.45\% | -5.35\% |  |  |  |  |
| 39 | 60.35\% | 62.00\% | 1.65\% |  |  |  |  |
| 40 | 58.50\% | 58.07\% | -0.4389 | 14 | 55.86\% | 56.02\% | 0.16\% |
| St | 60.60\% | $55.16 \%$ | -5.4440 |  |  |  |  |
| 42 | 48.54\% | 54.94\% | 5.40\% |  |  |  |  |
| 23 | 44.148\% | 43.06\% | -1.08\% | 15 | 41.20\% | 40.17\% | -2.03\% |
| 4 | 36.74\% | 37.22\% | 0.48\% |  |  |  |  |
| 45 | 42.39\% | - 40.08\% | .2318 |  |  |  |  |
| 46 | 42.076 | - $42.39 \%$ | 0.329 | 16 | 39.06\% | 34.13\% | 4.933 |
| 47 | 48.69\% | - 33.35\% | . $15.34 \% 6$ |  |  |  |  |
| 48 | 28.03\% | - $27.56 \%$ | -0.A7\% |  |  |  |  |
| 69 | 49.63\% | 49.59\% | -0.09\% | 17 | 48.46\% | 49.23\% | 0.778 |
| 50 | 52.08\% | 52.06\% | -0.02\% |  |  |  |  |
| 51 | 44.01\% | 46.23\% | 2.22\% |  |  |  |  |
| 52 | 57.3964 | 59.06\% | $1.67 \times$ | 18 | 54.96\% | 55.01\% | 0.05\% |
| 53 | 52.74\% | 61.85\% | -0.899 |  |  |  |  |
| 54 | 45.08\% | 45.22\% | $0.14 \%$ |  |  |  |  |
| 55 | 49,34\% | 55.06\% | 5.72\% | 19 | 53.32\% | 53.02\% | -0.3080 |
| 56 | 61.05\% | 58.86\% | -2.19\% |  |  |  |  |
| 57 | 47.26\% | 44.50\% | -2.76\% |  |  |  |  |
| 58 | 70.90\% | 70.54\% | -0.36\% | 20 | 70.55\% | 69.46\% | -109\% |
| 59 | 7274\% | 58.31\% | 4.43\% |  |  |  |  |
| 00 | 68.12\% | 69.52\% | 1.40\% |  |  |  |  |
| 61 | 35.98\% | 57.22\% | 21.24\% | 21 | 49.86\% | 57.77\% | 7.91\% |
| 52 | 44.35\% | 56.56\% | 12.21\% |  |  |  |  |
| 63 | 63.09\% | 59.64\% | . $3.45 \%$ |  |  |  |  |
| B4 | 35.66\% | 42.72\% | 7.06\% | 22 | 47.56\% | 35.97\% | -10.59\% |
| 65 | 45.44\% | - 35.92\% | -9.529 |  |  |  |  |
| 66 | 59.12\% | 31.71\% | -27.41\% |  |  |  |  |
| 67 | 51.72\% | 51.67\% | -0.05\% | 23 | 49.98\% | 51.75\% | 1.779 |
| 68. | 45.01\% | 49,38\% | 4.37\% |  |  |  |  |
| 19 | 54.06\% | 54,16\% | 0.10\% |  |  |  |  |
| 70 | 49.74 N | 50.73\% | 0.995 | 24 | 46.72\% | 47.51\% | 0.79\% |
| 72 | 41.68\% | 40.72\% | -0.96\% |  |  |  |  |


| DSTaict |  | - |  | Dstmic, |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 27 | 49.03\% | 51496 | 2.46\% |  |  |  |  |
| 73 | 39.55\% | - $40.16 \%$ | 0.61\% | 25 | 44.88\% | 44.88\% | 0.0096 |
| 74 | 43.78\% | - $42.89 \%$ | -0.89\% |  |  |  |  |
| 75 | 51.71\% | 52.18\% | 0.47\% |  |  |  |  |
| 76 | 24.29\% | 14.4996 | -9,80\% | 26 | 20.85\% | 20.98\% | 0.13\% |
| 77 | 23.88\% | 19.23\% | -4.65\% |  |  |  |  |
| 78 | 14.09\% | 30.84\% | 16.75\% |  |  |  |  |
| 79 | 37.49\% | 41.80\% | 4.31\% | 27 | 3838\% | 41.49\% | 3.118 |
| 50 | 42.15\% | 38.55\% | -3.60\% |  |  |  |  |
| 81 | 36.16\% | 44.56\% | 8.40\% |  |  |  |  |
| 82 | 58.59\% | 57.08\% | -1.51\% | 28 | 64.48\% | 60.93\% | $3.55 \%$ |
| 83 | 69.70\% | 68.31\% | -1.39\% |  |  |  |  |
| 84 | 64.99\% | 57.10\% | -7.89\% |  |  |  |  |
| 55 | 48.91\% | 48.38\% | -0.53\% | 29 | 52.00\% | 52.47\% | 0.47\% |
| 85 | 54.56\% | 55.08\% | 0.52\% |  |  |  |  |
| 97 | 52.16\% | $53.74 \%$ | 1.58\% |  |  |  |  |
| Es | 44.85\% | 53.19\% | 8.346 | 30 | 50,38\% | 50.55\% | 0.1780 |
| 39 | 55.76\% | 55.73\% | -0.03\% |  |  |  |  |
| 90. | 49.59\% | 40.40\% | -9.19x |  |  |  |  |
| 92 | 45.87\% | 39.57\% | -6.30\% | 37 | 46.89\% | 44.94\% | -1.95\% |
| 92 | 50,7996 | 44.30\% | -6.49\% |  |  |  |  |
| 93 | 44,73\% | 51.10\% | 6.37\% |  |  |  |  |
| 94 | 51.57\% | 51.91\% | 0.3496 | 32 | 44.43\% | 44.636 | $0.20 \%$ |
| 95 | 36.02\% | 36.36\% | 0.34\% |  |  |  |  |
| 96 | 45.32\% | 46.40\% | 1.08\% |  |  |  |  |
| 97 | 59.96\% | 62.91\% | 2.958 | 33 | 68.84\% | 68.60\% | . $0.244 \%$ |
| 98 | 70.96\% | 67.02\% | -3.944 |  |  |  |  |
| 99 | 73.35\% | 74.85\% | 1.50\% |  |  |  |  |

SA325

| Current Map |  |  | New Map |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Strong GOP ( $55 \%$ ) | 27 | 7 | Strong GOP (55\%-) | 38 | 12 |
| Lean GOP (52.1-54.999): | 13 | 8 | New Lean GOP (52.1-54.996) | 14 | 5 |
| Total GOP Seats (strong + lean): | 40 | 15 | Total GOP Seats (strong + lean): | 52 | 17 |
| Swing (48-52\%): | 19 | 5 | New Swing (48-52\%) | 10 | 3 |
| Lean DEM (45.1-47.9\%); | 7 | 3 | New Lean DEM (45.1-47.966): | 4 | 1 |
| Strong DEM (-45\%): | 33 | 10 | Strong DEM (-45\%): | 33 | 12 |
| Total DEM Seats (strong + lean): | 40 | 13 | Total DEM Seats (strong + iean); | 37 | 13 |


| Final Map |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  | Senate |  |  |  |
| District |  |  |  | DISTRICT |  |  |  |
| 99 | 7335\% | 74.853 | 1.50\% | 1 | 54.04\% | 53.73\% | -0.31\% |
| 53 | 70.90\% | 70.54K | .0.36\% |  |  |  |  |
| 60 | 68.12\% | 69.52\% | 1.40\% |  |  |  |  |
| 59 | 72.74\% | 68.31\% | 4.43\% | 2 | 55.444 | 55.23\% | -0.218 |
| 83 | 69.70\% | $6831 \%$ | -1.39\% |  |  |  |  |
| 38 | 70.96\% | 67.028 | 4.94\% |  |  |  |  |
| 22 | 39.05\% | 65.82\% | 27.77\% | 3 | 40,52\% | 38.12\% | -2.40\%6 |
| 97 | 59.96\% | 62.919- | 2.95 k |  |  |  |  |
| 32 | 62.06\% | 52.27\% | 1.21\% |  |  |  |  |
| 39 | 60.35\% | ह2.00\% | 1.65\% | 4 | 17.58\% | 19.63\% | 2.05\% |
| 53 | 62.74\% | 61.85\% | -0.899 |  |  |  |  |
| 33 | 72.24\% | 61.81\% | -10.43\% |  |  |  |  |
| 38 | 65.80\% | 60.45\% | -5.35\% 5 | 5 | 50.62\% | 57.72\% | 7.10\% |
| 63 | 63,09\% | 59.64\% | -3.45\% |  |  |  |  |
| 52 | 57,39\% | 59.06\% | 1.67\% |  |  |  |  |
| 56 | 61.05\% | 58.86\% | -2.19\% | 5 | 14.12\% | 15.55\% | 1.439 |
| 13 | 43.67\% | 58.67\% | 15.005 |  |  |  |  |
| 24 | 59.06\% | 58.64\% | -0.42\% |  |  |  |  |
| 24 | 67.29\% | 58.49\% | -8.80\% | 7 | 41.13\% | 40.53\% | -0.60\% |
| 6 | 59.77\% | 58,33\% | - $1.44 \%$ |  |  |  |  |
| 37 | 51.33\% | 58.11\% | 6.78\% |  |  |  |  |
| 40 | 58.50\% | 58.07\% | -0.43\% | 8 | 52.82\% | 60.88\% | 8.06\% |
| 23 | 51.70\% | 57.64\% | 5.94\% |  |  |  |  |
| 61 | 35.98\% | 57.22\% | 21.24\% |  |  |  |  |
| 84 | 54.99\% | 57,10\% | -7.899 | 9 | 52.96\% | 55.19\% | 2.23\% |
| 82 | $58.59 \%$ | 57.08\% | -1.51\% |  |  |  |  |
| O2 | 44.35\% | 56.56\% | 12.21\% |  |  |  |  |
| 31 | 67.57\% | 56.33\% | -11.24\% | 10 | 53.14\% | 53.32\% | $0.18 \%$ |
| 27 | 59.20\% | 56,19\% | -3.01\% |  |  |  |  |
| 25 | 45.42\% | 55.97\% | $10.55 \%$ |  |  |  |  |
| 89 | 55.76\% | 55.73\% | -0.03\% | 11 | 67.64\% | 60.13\% | .751\% |
| 3 | 56.10\% | 55.58\% | -0.52\% |  |  |  |  |
| 15 | 48.21\% | 55.48\% | 7.27x |  |  |  |  |
| 37 | 54.51\% | 55.22\% | 0.71\% | 12 | 53.37\% | 54.39\% | 1.02\% |
| 41 | 60.60\% | 55.16\% | -5.44\% |  |  |  |  |
| 85 | 54.56\% | 55.08\% | 0.52\% |  |  |  |  |
| 55 | 49.34\% | 55.06\% | 5.72\% | 13 | 59.22\% | 60.17\% | 0.9516 |
| 28 | 54.85\% | 55.00\% | 0.15\% |  |  |  |  |
| 42 | 48.54\% | 54.94\% | 6.400 |  |  |  |  |
| 2 | 54.93\% | 54.84\% | -0.098 | 14 | 55.86\% | 56.02\% | 0.16\% |
| 36 | 53.06\% | 54.84\% | 1.78\% |  |  |  |  |
| 5 | 53.74\% | 54.28\% | 0.54\% |  |  |  |  |
| 69 | 54.06\% | 54.16\% | 0,10\% | 15 | 41.20\% | 40.17** | -1.03\% |
| 30 | 53.29\% | 53.78\% | 0.49\% |  |  |  |  |
| 87 | 52,16\% | 53.74\% | 1.58\% |  |  |  |  |
| 5 | 53.31\% | 53.47\% | 0.16\% | 16 | 39.06\% | 34.13\% | -4.93\% |
| 25 | 52.79\% | 53.26\% | 0.47\% |  |  |  |  |
| 88. | 44.85\% | 53.19\% | 8.34\% |  |  |  |  |
| 35 | 52.30\%6 | 52.99\% | 0,69\% | 17 | 48.46\% | 49.73\% | 0.77\% |
| 21 | 51.92\% | 52.94\% | 1.02\% |  |  |  |  |
| 75 | 51.71\% | 52.18\% | 0.A7\% |  |  |  |  |
| 50 | 52.08\% | 52.06\% | -0.029 | 18 | 54.96\% | 55.01\% | 0.05\% |
| 94 | 51.57\% | 51.91\% | 0.3489 |  |  |  |  |
| 57 | 51.72\% | 51.67\% | -0.05\% |  |  |  |  |
| 72 | 49.03\% | 51.A9\% | 2.469 | 19 | 53.32\% | 53.02\% | -0.30\% |
| 1 | 51.15\% | 51.22\% | 0.07\% |  |  |  |  |
| 93 | 44.73\% | 51.10\% | 6.378 |  |  |  |  |
| 29 | 51.32\% | 50.97\% | -0.35\% | 20. | 70.55x | 69.46\% | -1.095 |
| 70 | 49.74\% | 50.73x | 0.9985 |  |  |  |  |
| 49 | 49.68\% | 49.59\% | -0.09\% |  |  |  |  |
| 68 | 45.01\% | 49.38\% | 4.37\% | 21 | 49.86\% | 57.77x | 7.91\% |
| 55 | 48.91\% | 48.38\% | -0.53\% |  |  |  |  |
| 96 | 45.32\% | $46.40 \%$ - | 1.08\% |  |  |  |  |
| 51 | 44.01\% | 46.23\% | 2.22\% | 22 | 47.56\% | 36.97\% | -10.59\% |
| 7 | 48.200 | 45.38\% | -2.82\% |  |  |  |  |
| 54 | 45.08\% | 45.22\% | $0.14 \%$ |  |  |  |  |
| 61 | 36.16\% | 44.56\% | $8.40 \%$ | 23 | 49.98\% | 51.75\% | 1.77\% |
| 57 | 47.26\% | 44,50\% | -2.76\% |  |  |  |  |
| 92 | 50.7986 | 44.30\% | -6.49\% |  |  |  |  |
| 20 | $43.71 \%$ | 43.129\% | -0.5976 | $2 \pi$ | 46.725 | 47.51\% | 0.79\% |
| 43 | 44,14\% | 43.06\% | -1.08\% |  |  |  |  |


| DISTu\|ct | 12M |  |  | Distuiat |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 74 | 43.78\% | 42.89\% | -0.89\% |  |  |  |  |
| 54 | 35.66\% | 42.72\% | 7.06\% | 25 | 44.88\% | 44.88\% | 0.00\% |
| 46 | 42.07\% | 42.39\% | 0.32\% |  |  |  |  |
| 79 | 37.4996 | 41.80\% | 4.31\% |  |  |  |  |
| 21 | 41.68\% | 40.72\% | 0.96\% | 26 | 20.85N | 20.98\% | 0.13\% |
| 90 | 49.59\% | 40.40\% | -9.1999 |  |  |  |  |
| 73 | 39.55\% | 40.16\% | $0.61 \%$ |  |  |  |  |
| 45 | 42.39\% | 40.08\% | -2.31\% | 27 | 38.38\% | 41.498 | 3.11\% |
| 31 | 45.87\% | 39.57\% | -6.30\% |  |  |  |  |
| 00 | 42.15\% | 38.55\% | -3.60\% |  |  |  |  |
| 44 | 36.74\% | 37.22\% | 0.48\% | 28 | 64.48\% | 60.93\% | .3.55\% |
| 95. | 36.02x | 36.36\% | 0.348 |  |  |  |  |
| $5{ }^{515}$ | 45.445\% | 35,92\% | -9.52\% |  |  |  |  |
| 47 | 48.69\% | 33.35\% | -15.34\% | 29 | 52.00\% | 52.47\% | 0.47\% |
| 65. | 59.12\% | 31.71\% | -27.41\% |  |  |  |  |
| 78 | 14.09\% | 30.84\% | 16.75\% |  |  |  |  |
| s | 22.39\% | 30.48\% | 8.09\% | 30 | 50.38\% | 50.55\% | 0.17x |
| 9 | 36.73\% | 29.14\% | 7.59\% |  |  |  |  |
| 19 | 29.15\% | 28.03\% | -1.12\% |  |  |  |  |
| 48 | 28.03\% | 27.56\% | -0.473 | 31 | 46.89\% | 44.94\% | -1.95\% |
| 12 | 29.23\% | 27.51\% | -1.72\% |  |  |  |  |
| 17 | 13.21 K | 19.84\% | 6.63\% |  |  |  |  |
| 11 | 11.915 | 19.58\% | 7.67\% | 32 | 44.43\% | 44.63\% | 0.20\% |
| 77 | 23.88\% | 19.23\% | -4.65\% |  |  |  |  |
| 18 | 1528\% | 14.94\% | -0.34\% |  |  |  |  |
| 76 | 24.2975 | 14.49\% | -9.80\% | 33 | 68.84\% | 68.60\% | -0.24x |
| 10 | 10.276 | 12.59\% | 2.32\% |  |  |  |  |
| 16 | 14.219 6 | 10.54\% | 3.67\% |  |  |  |  |


| Current Map |  |  | New Map |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Strong GOP (55\% + | 27 | 7 | Strong GOP (55\% + | 38 | 12 |
| Lean GOP (52.1-54.9\%): | 13 | 8 | New Lean GOP (52.1.54.996): | 14 | 5 |
| Total GOP Seats (strong + lean): | 40 | 15 | Total GOP Seats (strong + lean): | 52 | 17 |
| Swing (48-523): | 19 | 5 | New Swing (48.52\%) | 10 | 3 |
| Lean DEM (45, 1-47,9\%): | 7 | 3 | New Lean DEM (45.1-47.96): | 4 | 1 |
| Strong DEM (-45\%): | 33 | 10 | Strong DEM (-45\%): | 33 | 12 |
| Totai DEM Seats (strong + iean): | 40 | 13 | Total DEM Seats (strong + lean): | 37 | 13 |


| Kessler Map |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  | Senate |  |  |  |
| DISTRICT |  | ar |  | DISTRICT |  |  |  |
| 1 | 51.15\% | 58.28\% | 7.13\% | 1 | 54.04\% | 55.88\% | 1.84\% |
| 2 | 54.93\% | 48.90\% | -6.03\% |  |  |  |  |
| 3 | 56.10\% | 59.95\% | 3.85\% |  |  |  |  |
| 4 | 53,31\% | 54.91\% | $1.60 \%$ | 2 | 55.44\% | 5784\% | $2.40 \%$ |
| 5 | 53.745 | -58.65\% | 4.919 |  |  |  |  |
| 6 | 59.776 | 60.17\% | 0.405 |  |  |  |  |
| 7 | 48.20\%6 | 48.01\% | -0.19\% | 3 | 40.52\% | 40.00\% | -0.52\% |
| 5 | 22.39\% | - $22.82 \%$ | 0.43x |  |  |  |  |
| 9 | 36.73\% | - $34.52 \%$ | -2.219 |  |  |  |  |
| 10 | 10.27\% | -33.07\% | 22.80\% | 4 | 17.58\% | 31.02\% | 13.44\% |
| 11 | 11.91\% | 30.48\% | 18.57\% |  |  |  |  |
| 12 | 29.23\% | 29.01\% | -0.22\% |  |  |  |  |
| 13 | A3.67\% | 45.28\% | 1.61\% | 5 | 50.62\% | 49.98\% | -0.64\% |
| 14. | 59.06\% | 57 34\% | -1.72\% |  |  |  |  |
| 13 | 48.21* | 47.62\% | -0.59\% |  |  |  |  |
| 16 | 14.21\% | - 14.26\% | 0.05\% | 6 | 14.12\% | 21.34* | 7.228 |
| 17 | 13.21\% | - $24.94 \%$ | 11.73x |  |  |  |  |
| 18 | $15.28 \%$ | 23.19\% | 7.91\% |  |  |  |  |
| 19 | 29.15\% | 31.45\% | 2.30\% | 7 | 41.13\% | 41.45\% | 0.32\% |
| 20 | 43.71\% | 45.14\% | 1.43\% |  |  |  |  |
| 21 | 31.92\% | 49.51\% | -2.41\% |  |  |  |  |
| 22 | 39.05\% | - $25.68 \%$ | -13,37\% | 8 | 52.82\% | 48.86\% | -3.96\% |
| 23 | 51.70\% | 46.50\% | -5.20\% |  |  |  |  |
| 25 | 67.29\% | 71.718 | 4,42\% |  |  |  |  |
| 25 | 52.79\% | 49,48\% | .3.31\% | 9 | 52.96\% | 49.17\% | -3.7996 |
| 26 | 45.42\% | 46.38\% | 0.96\% |  |  |  |  |
| 27 | 59.20\% | 51.22\% | 7.98\% |  |  |  |  |
| 28 | 54.85\% | 55.60\% | 0.75\% | 10 | 53.14\% | 53.19\% | 0.05\% |
| 29 | 51.32\% | 46.68\% | 4.642\% |  |  |  |  |
| 30 | 53.29\% | 57.21\% | 3.92\% |  |  |  |  |
| 31 | 57.57x | 69.18\% | 1.619 | 13 | 67.64\% | 68.08\% | 0.44\% |
| 32 | 61.06\% | 61.62\% | 0.56\% |  |  |  |  |
| 33 | 72.24\% | 71.77\% | 0.47\% |  |  |  |  |
| 34 | 54.51\% | 48.62\% | -5.89\% | 12 | 53.37\% | 51.36\% | -2.01\% |
| 35 | $52.30 \%$ | 50.09\% | -2.21\% |  |  |  |  |
| 36 | 53.06\% | 54.77\% | 1.718 |  |  |  |  |
| 37 | 51.33\% | 49.82\% | -1.51\% | 13 | 59.22\% | 60.12\% | 0.905 |
| 38 | 65.80\% | 67.73\% | 1.93\% |  |  |  |  |
| 39 | 60.35\% | 62.35\% | 2.00\% |  |  |  |  |
| 60 | 58.50\% | 57.79\% | -0.71\% | 14 | 55.86\% | 49.86\% | -6.00\% |
| 41 | 60.60\% | 44.17\% | -16.43\% |  |  |  |  |
| 62 | 48.54\% | 48.23\% | -0.315 |  |  |  |  |
| 43 | 44.14\% | 42.34\% | -1.80x | 15 | 41.20\% | 41.30\% | 0.10\% |
| 44 | 36.74\% | 38.88\% | 2.14\% |  |  |  |  |
| 45 | 42.39\% | 43.02\% | 0.63\% |  |  |  |  |
| 46 | 42.07\% | 42.59\% | 0.528 | 16 | 39.06\% | 38.13* | -0.93x |
| 17 | 48.69\% | 47.09\% | -1.60x |  |  |  |  |
| 48 | 28.03\% | 27,47\% | -0.56\% |  |  |  |  |
| 62 | 49.68\% | 49.88\% | 0.16\% | 27 | 48.46\% | 48.46\% | 0.000 |
| 50 | 52.08\% | 51.88\% | -0.20x |  |  |  |  |
| 51 | 44.018 | 44.09\% | 0.08\% |  |  |  |  |
| 52 | 57.39\% | 57.29\% | -0.10x | 18 | 54.96\% | 54.84\% | -0.12\% |
| 53 | 62.74\% | $62.70 \%$ | -0.04\% |  |  |  |  |
| 54 | 45.08\% | 44.00\% | -1.08\% |  |  |  |  |
| 55 | 49.34\% | 49.95\% | 0.61\% | 13 | 53.32\% | 52.88\% | -0.4466 |
| 56 | 61.05\% | 60.64\% | -0.41\% |  |  |  |  |
| 57 | 47.26\% | 48.31\% | 1.055 |  |  |  |  |
| 59 | 70.90\% | 70.35\% | 0.55\% | 20 | 70.55\% | 69.15\% | -1.4084 |
| 59 | 72.74\% | 69.94\% | -2.80\% |  |  |  |  |
| 60 | 68.12\% | 67,37\% | -0.75\% |  |  |  |  |
| 61 | 35.98\% | $42.56 \%$ | 6.58\% | 21 | 49.86\% | 49.36\% | -0.5006 |
| 62 | 44.35\% | 41.72 N | -2.63\% |  |  |  |  |
| 6 | 63.09\% | 61.66\% | -1.433\% |  |  |  |  |
| 54 | 35.66\% | 36.48\% | 0.82\% | 22 | 47.56\% | 46.30\% | -1.26\% |
| 65 | 45.44\% | 44.02\% | -1.42\% |  |  |  |  |
| 56 | 59.12\% | 58.37\% | -0.755 |  |  |  |  |
| 67 | 51.72\% | 51.10\% | -0.62\% | 23 | 49.98\% | 49.21\% | -0.7780 |
| 5 | 45.01\% | 44.54\% | -0.47\% |  |  |  |  |
| 69 | 54.06\% | 51.90\% | -2.16\% |  |  |  |  |
| 70 | 49.74\% | 49.42\% | 0.32\% | 24 | 46.72\% | 46.56\% | -0.16\% |
| 72 | 41.68\% | 41.4850 | -0.204] |  |  |  |  |


| DSTM\| |  |  |  | DSTRIC | $\square$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 72 | 49.03\% | 48.8776 | -0.16\% |  |  |  |  |
| 73 | 39.55\% | 40.78\% | 1.23\% | 25 | 44.88\% | 4531\% | 0.43\% |
| 74 | 43.78\% | 44.86\% | 1.08\% |  |  |  |  |
| 75 | 51.719 | 50,50\% | -1.21\% |  |  |  |  |
| 76 | 24.29\%6 | 24.20\% | -0.09\% | 26 | 20.85\% | 2136\% | 0.519 |
| $\pi$ | 23.88\% | - 26.21\% | 2.33\% |  |  |  |  |
| 78. | 14.09\% | 13.34\% | -0.75\% |  |  |  |  |
| 73 | 37.49\% | 38.52\% | 1.03\% | 27 | 38.38\% | 38.25\% | -0.13\% |
| 50 | 42.15\% | 41.95\% | -0.20\% |  |  |  |  |
| a1 | $36.16 \%$ | 34.87\% | -1.29\% |  |  |  |  |
| 32 | 58.59\% | $59.54 \%$ | 1.05\% | 28 | 64.48\% | 65.01\% | 0.53* |
| 83 | 69.70\% | 67.79\% | -1.91\% |  |  |  |  |
| EA | 64.99\% | 66.69\% | 1.70\% |  |  |  |  |
| 05 | 48.91\% | 56.47\% | 7.56\% | 29 | 52.00\% | 56.13\% | 4.137010 |
| 55 | 54.56\% | 36.80\% | $2.24 \times$ |  |  |  |  |
| 87 | 52.16\% | 54.92\% | 2.76\% |  |  |  |  |
| 88 | 44.85\% | 45.13\% | 0.28\% | 30 | 50.38\% | 49.62\% | -0,7600 |
| 29 | 55.76\% | 55.33\% | -0.43\% |  |  |  |  |
| 90 | 49.59\% | 47.70\% | -1.89\% |  |  |  |  |
| 21 | 45.87\% | 45.82\% | -0.05\% | 31 | 46.89\% | 46.82\% | 0.07\% |
| 92 | 50.79\% | $49.85 \%$ | -0.944 |  |  |  |  |
| 23 | 44.73\% | 45.409\% | 0.57\% |  |  |  |  |
| 94 | 51.57\% | 47.65\% | -3.92\% | 32 | 44.43\% | 44.43\% | 0.00\% |
| 95 | 36.02\% | 40.4446- | 4.42\% |  |  |  |  |
| 96 | 45.32\% | 45.76\% | 0.44\% |  |  |  |  |
| \% | 59.96\% | 69.88\% | $9.92 \times$ | 33 | 68.84\% | 71.46\% | 262\% |
| 98 | 70.96\% | 72.93\% | 1.97\% |  |  |  |  |
| 99 | 73.35\% | 71.844\% | -1.51\% |  |  |  |  |


| Curtent Map |  |  | New Map |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembily | Senate |  | Assembly | Senate |
| Strong GOP ( $55 \%+$ ) | 27 | 7 | Strong GOP ( $555 \%+1$ | 31 | 8 |
| Lean GOP (52.1-54.9\%) = | 13 | 8 | New Lean GOP ( $52.1-54.96 \%$ ): | 3 | 3 |
| Total GOP Seats (strong + lean): | 40 | 15 | Total GOP Seats (strong + lean): | 34 | 11 |
| 5 wing 148-52\%): | 19 | 5 | New Swing (48.-52\%) | 19 | 9 |
| Lean DEM ( $45.1-47.9 \%$ ) $=$ | 7 | 3 | New Lean DEM ( $45.1-47$ 9\%) | 13 | 4 |
| Strone DEM (-45s) : | 33 | 10 | Strong DEM (-45\%): | 33 | 9 |
| Total DEM Seats (strong e lean): | 40 | 13 | Total DEM Seats (strong + iean): | 46 | 13 |

## Questions and Responses:

Every question can be traced back to the principles that guide redistricting:

1. Equal Population
2. Sensitivity to Minority Concerns
3. Compact and Contiguous districts.

Different choices can be made along the way, but those criteria must be followed. SB 148 meets these criteria.

## Why so many pairings?

Pairings are usually an inevitable consequence of reapportionment and the result of compliance with the principles of equal population, compact and contiguous districts and sensitivity to minority concerns. Legislative districts are reapportioned to be in place for 10 years. Out of 132 legislators, only 35 remain today in the seats they held in 2000. That is about 75 percent turnover for the 10 years that the districts put in place after the last reapportionment were in effect.

Why did you (split, draw, pair) X?
There are a number of ways to reapportion. The reapportionment involves competing principles and choices that have to be made. This legislation represents the choices that have been made that are consistent with the legal standards required.

## Who made the decisions on how these districts were drawn?

We are making that decision right now. Today. The Legislature. Staff developed this bill in consultation with attorneys retained by the Senate and the Assembly to make sure that it conformed with all legal principles. The duty to pass it falls on the Legislature.

## What is the partisan makeup of these districts?

The election data for the last 10 years was made available by the Government Accountability Board to the Legislature. All four caucuses were provided this information along with the hardware and software to use it. Everyone has the ability to draw their own conclusions and interpret how past elections may play out in the new districts. But no one has a crystal ball that will tell you how elections may play out in these districts next year, or 10 years from now when these districts will still be in effect. 10 years ago different experts offered wildly different opinions on how the proposed maps would perform politically.


Who did you talk with about these maps?
Staff consulted with attorneys to make sure that all legal principles were followed in reapportioning the state.

Why are you offering choices on the Hispanic districts, but not on the African American districts?

Given the rapid growth of the Hispanic population during the ten year cycle, which is very different than other minority growth patterns, we simply thought providing a number of alternatives would be appropriate. If there are other alternatives for other minority groups, then those can be proposed and acted on by the Committee and the legislature.
Why were Republican Attorneys hired to draw maps but Democrats were not allowed attorneys to draw maps?

Attorneys did not draw these maps. Staff drew them. Attorneys merely advised on the legal principles that have to be followed. Your staff has had all the same hardware, software and data available to them for over a year. The census data has been available since the end of March. I don't know what your staff has been doing with all that equipment and data. Our staff has been working on this bill.

Why are you not drawing a 50 percent voting age Hispanic seat?
$I$ haven't seen a map that has a Senate seat with a 50 percent voting age Hispanic population. No one has produced one that I'm aware of.

Why are you acting now? Why are you acting before the locals?
Former State Senate leader, Senator Robson, is suing the state in federal court for not acting quickly enough. This is a constitutional duty of the Legislature. There is no reason for us to delay action and let a court do our job for us.

Redistricting is not something that we have discretion on. The Constitution requires the legislature to do this every 10 years. Only when the legislature is unable to agree do the courts step in. When we pass these maps, it will be the first time in nearly 30 years that the Legislature has met its obligation.

Many of you weren't here 10 years ago and most of the Assembly was not here 10 years ago. But because the courts drew a fair map after the last census, we're here today in the majority in both houses. The maps we pass will determine who's here 10 years from now.

Today we're going to walk through the proposed maps and talk about how we got there. We have an opportunity and an obligation to draw these maps that Republicans haven't had in decades.

There are 3 primary principles that go into the drawing of every map.

- Equal population
- Compact and contiguous districts
- Sensitivity to minority concerns.

The process is never pain-free, no matter who draws the maps. There are always tradeoffs as you make decisions when drawing a map. The constitution and the statues lay out the principals we have to follow. What we've come up with is a fair map that meets all these criteria and which we are confident will stand up to any legal scrutiny.

- Three principles
- Disenfranchisement
- Over/Under population map
- Start in Milwaukee
- Hispanic districts

Notes
59 split municipalities
Overall deviation of $0.76(0.37$ to -0.39$)$ in Asm. $0.62(0.27$ to -0.35$)$ in Sen

1. Currently, the urban areas of Racine and Kenosha are paired in 2 senate districts with the more rural parts of each county. This maps pairs the two urban areas in one senate district, and the more rural parts of each county together in another senate district. This results in 2 districts which each share more in common throughout the senate seat.

## PRIVILEGED ATTORNEY-CLEINT COMMUNCATION

## Confidentiality and Nondisclosure Related to Reapportionment

Michael Best \& Friedrich LLP ("MB\&F") is currently engaged to represent the Wisconsin State Senate, by its Majority Leader Scott L. Fitzgerald ("Senate") in connection with matters relating to the reapportionment of the Wisconsin Senate, Assembly and Congressional Districts arising out of the 2010 census (the "Representation"). In connection with the Representation we have instructed certain individuals, working at our direction, to meet with certain members of the Senate for the purpose of discussing matters within the scope of the Representation. Such discussions shall be conducted for the sole purpose of assisting MB\&F in rendering legal advice to the Senate and, therefore, are subject to the attorney-client and attorney work product privileges. Consistent with those privileges, such discussions are and shall remain confidential.

This letter will confirm our understanding that such discussions are and shall remain confidential and that you agree not to disclose the fact and/or contents of such discussions or any draft documents within your possession related to the subject of the Representation with persons outside of the privilege. If you have any questions regarding the foregoing, please feel free to raise those questions with me. Otherwise, in order to confirm the foregoing understanding, please sign on the line indicated below.

Sincerely,
MICHAEL BEST \& FRIEDRICH LLB


Eric M. McLeod

## APPROVED AND AGREED UPON:



## PRIVILEGED ATTORNEY-CLIENT COMMUNICATION

## Confidentiality and Nondisclosure Related to Reapportionment

Michael Best \& Friedrich LLP ("MB\&F") is currently engaged to represent the Wisconsin State Assembly, by its Speaker Jeff Fitzgerald ("Assembly") in connection with matters relating to the reapportionment of the Wisconsin Senate, Assembly and Congressional Districts arising out of the 2010 census (the "Representation"). In connection with the Representation we have instructed certain individuals, working at our direction, to meet with certain members of the Assembly for the purpose of discussing matters within the scope of the Representation. Such discussions shall be conducted for the sole purpose of assisting MB\&F in rendering legal advice to the Assembly and, therefore, are subject to the attorney-client and attorney work product privileges. Consistent with those privileges, such discussions are and shall remain confidential.

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Sincerely,

## MICHAEL BEST \& FRIEDRICH LLP



Eric M. McLeod

## APPROVED AND AGREED UPON:



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| Distric t/COM P | Observed | All_40 | All_41 | All_42 | All_43 | All_44 | All_45 | All_46 | All_47 | All_48 | Compos ite | All_50 | All_51 | All_52 | All_53 | All_54 | All_55 | All_56 | All_57 | All_58 | All_59 | All_60 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 15 | 0.1786 |  |  |  |  |  |  |  |  | ${ }^{0.1676}$ |  |  |  |  |  |  | 0.23 |  |  |  |  | 0.3031 |
| 9 | 0.1941 | 0.1031 | 0.1131 | 0.1231 | 0.1331 | 0.1431 | 0.1531 | 0.1631 | 0.1731 | 0.1831 | ${ }^{0.1941}$ | 0.203 | 0.2131 | 0.2231 | 0.2331 | 0.2431 | 0.2531 | 0.2631 | 0.2731 | 0.2831 | 0.2931 | ${ }^{0.3031}$ |
| 75 | 0.2085 | 0.1175 | 0.1275 | 0.1375 | 0.1475 | 0.1575 | 0.1675 | 0.1775 | 0. 1875 | 0.1975 | 0.2 | 0.2175 | 0.2275 | 0.2375 |  | 0.2575 | 0.2675 | 0.2775 | 0. 2875 | 0.2975 | 0.3075 | 0.3175 |
| 8 | 0.2125 | 0.1215 | 0.1315 | 0.1415 | 0.1515 | ${ }^{0} .1615$ | 0.1715 | 0.1815 | 0.1915 | 0.2015 | 0.2125 | 0.2215 | 0.2315 | 0.2415 | 0.2515 | 0.2615 | 0.2715 | 0.2815 | 0.2915 | 0.3015 | ${ }^{0.3175}$ | 0.3215 |
| 16 | 0.2245 | 0.1335 | -0.1435 | ${ }^{0.1535}$ | 0.1635 0.1683 | 0.1735 | -0.1835 | 0.1935 0.1983 | -0.2035 | 0.2135 0.2183 0.283 | 0.2245 | 0.2335 0.2383 | 0.2435 0.2483 | 0.2535 | 0.2635 0.2683 | 0.2735 0.2783 | -0.2835 | 0.2935 0.2983 | 0.3035 | 0.3135 0.3183 | 0.3235 0.3283 | 0.3335 0.3383 |
| 76 | 0.2293 | 0.1383 0.1612 0.0 | 0.1483 0.1712 | ${ }_{0}^{0.1883}$ | ${ }_{0}^{0.11912}$ | ${ }_{0} .2012$ | 0.2112 | 0.2212 | 0.2312 | ${ }_{0} .2412$ | 0.2522 | 0.2612 | 0.2712 | ${ }_{0} .2812$ | 0.2912 | 0.3012 | 0.3112 | 0.3212 | 0.3312 | 0.3412 | 0.3512 | 0.3383 0.3612 |
| 17 | 0.2911 | 0.2001 | 0.2101 | 0.2201 | 0.2301 | 0.2401 | 0.2501 | 0.2601 | 0.2701 | 0.2801 | 0.2911 | 0.3001 | 0.3101 | 0.3201 | 0.3301 | 0.3401 | 0.3501 | 0.3601 | 0.3701 | 0.3801 | 0.3901 | 0.4001 |
| 77 | 0.304 | 0.213 | 0.223 | 0.233 | 0.243 | 0.253 | 0.263 | 0.273 | 0.283 | 0.293 | 0.304 | 0.313 | 0.323 | 0.333 | 0.343 | 0.353 | 0.363 | 0.373 | 0.383 | 0.393 | 0.403 | 0.413 |
| 74 | 0.3078 | 0.2168 | 0.2268 | 0.2368 | 0.2468 | 0.2568 | 0.2668 | 0.2768 | 0.2868 | 0.2968 | 0.3078 | 0.31 | 0.3268 | 0.3368 | 0.3468 | 0.3568 | 0.3668 | 0.3768 | 0.38 | 0.3968 | 0.40 | 0.4168 |
| 14 | 0.33 | 0.239 | 0.249 | 0.259 | 0.269 | 0.279 | 289 |  | 0.309 | 0.319 |  |  |  | 0.359 |  | 0.3 | 0.3 | 0.3 |  | 0.419 | 0.429 |  |
| 45 | 0.3654 <br> 0.368 | 0.2744 | 0.2844 | 0.2944 | 0.3044 | 0.3144 | 0.3244 | 0.3344 | 0.3444 | 0.3544 | 0.3654 | 0.3744 | 0.3844 | 0.3944 | 0.4044 | 0.4144 | $0.4244$ | 0.4344 | 0.4444 | 0.4544 | 0.4644 | 0.4744 0.477 |
| 73 | 0.3727 | 0.2817 | 0.2917 | 0.3017 | 0.3117 | 0.3217 | 0.3317 | 0.3417 | 0.3517 | 0.3617 | 0.3727 | 0.3817 | 0.3917 | 0.4017 | 0.4117 | 0.4217 | 0.4317 | 0.4417 | 0.4517 | 0.4617 | 0.4717 | 0.4817 |
| 73 63 | 0.3727 0.3734 |  | 0.2924 | 0.3024 | 0.3124 | 0.3224 | 0.3324 | 0.3424 | 0.3524 | 0.3624 | 0.3734 | 0.3824 | 0.3924 | 0.4024 | 0.4124 | 0.4224 | 0.4324 | 0.4424 | 0.4524 | 0.4624 | 0.4724 | 0.4817 0.4824 |
| 63 6 | 0.3734 0.3754 | - 0.28844 | ${ }^{0.29244}$ | - 0.30244 | ${ }_{0}^{0.3124}$ | ${ }_{0}^{0.3224}$ | ${ }_{0}^{0.3324}$ | -0.3424 | -0.3544 | - 0.36244 | (e.3754 | -0.3844 | -0.39244 | - 0.4044 | -0.4124 | 0.42244 | 0.0.4324 | -0.4424 | -0.4524 | 0.4624 | 0.4724 0.4744 | 0.4824 0.4844 |
| 46 | 0.3756 | 0.2846 | 0.2946 | 0.3046 | 0.3146 | 0.3246 | 0.3346 | 0.3446 | 0.3546 | 0.3646 | 0.3756 | 0.3 | 0.3946 | 0.4046 | 0.4146 | 0.4246 | 0.4346 | 0.4446 | 0.4546 | 0.4646 | 0.4746 | 0.4846 |
| 42 | 0.3876 | 0.2966 | 0.3066 | 0.3166 | 0.3266 | 0.3366 | 0.3466 | 0.3566 | 0.366 | 0.3766 | 0.3876 | 0.396 | 0.406 | 0.416 | 0.4 | 0.43 | 0.4466 | 0.4566 | 0.4666 | 0.4766 | 0.4866 |  |
| 18 | 0.3965 | 0.3055 | 0.3155 | 0.3255 | 0.3355 | 0.3455 | 0.3555 | 0.3655 | 0.3755 | 0.3855 | 0.3965 | 0.4055 | 0.4155 | 0.4255 | ${ }^{0.4355}$ | 0.4455 | 0.4555 | 0.4655 | 0.4755 | 0.4855 | 0.4955 | 0.5055 |
| 78 | 0.4045 | 0.3135 | 0.3235 | 0.3335 | 0.3435 | 0.3535 | 0.3635 | 0.3735 | 0.3835 | 0.3935 | 4045 | 0.4135 | 0.4235 | 0.4335 | 0.4435 | 0.4535 | 0.4635 | 0.4735 | 0.4835 | 0.4935 | 0.5035 | 0.5135 |
| 43 | 0.4096 | 0.3186 | 0.3286 | 0.3386 | 0.3486 | 0.3386 | 0.3686 | 0.3786 | 0.3886 | 0.3986 | - 0.4096 | 析 |  |  |  | 0.4586 0.459 | 0.4 | . | 0.48 | 986 | 0.5086 | 0.5186 |
| ${ }_{47}^{64}$ | 0.41 0.4109 | 0.319 0.3199 | 0.329 0.3299 | 0.339 0.3399 | 0.349 0.3499 | 0.359 0.3599 | 0.369 0.3699 | 0.379 0.3799 | 0.389 0.3899 | 0.399 0.3999 | 0.4109 | 0.419 0.4199 | 0.429 0.4299 | 0.439 0.4399 | 0.449 0.4499 | 0.459 0.4599 | 0.469 0.4699 | 0.4799 | 0.489 0.4899 | 0.4999 | 0.5099 | 0.519 0.5199 |
| 89 | 0.413 | 0.322 | 0.332 | 0.342 | 0.352 | 0.362 | 0.372 | 0.382 | 0.392 | 0.402 | 0.413 | 0.422 | 0.432 | 0.442 | 0.452 | 0.462 | 0.472 | 0.482 | 0.492 | 0.502 | 0.512 | 0.522 |
| 44 | 0.4155 | 0.3245 | 0.3345 | 0.3445 | 0.3545 | 0.3645 | 0.3745 | 0.3845 | 0.3945 | 0.4045 | 0.4155 | 0.4245 | 0.4345 | 0.4445 | 0.4545 | 0.4645 | 0.4745 | 0.4845 | 0.4945 | 0.5045 | 0.5145 | 0.5245 |
| 41 | 0.4373 | 0.3463 | 0.3563 | 0.3663 | 0.3763 | 0.3863 | 0.3963 | 0.40 | 0.4163 | 0.4263 | 0.4 | 0.4463 | 0.4563 | 0.4663 | 0.4763 | 0.4863 | 0.4963 | 0.5063 | 0.5163 | 0.5263 | 0.5363 | 0.5463 |
| 70 | 0.4423 | 0.3513 | 0.3613 | 0.3713 | 0.3813 | 0.3913 | 0.4013 | 0.4113 | 0.4213 | 0.4313 | 0.4423 | ${ }^{0.4513}$ | 0.4613 | 0.4713 | 0.4813 | 0.4913 | 0.5013 | 0.5113 | 0.5213 | 0.5313 | ${ }^{0.5413}$ |  |
| 93 | 0.4463 | 0.3553 | 0.3653 | 0.3753 | 0.3853 | 0.3953 | 0.4053 | 0.4153 | 0.4253 | 0.4353 | 0.4463 | 0.4553 | ${ }^{0.4653}$ | 0.4773 | 0.4853 | 0.4953 | 0.5053 | 0.5153 | 0.5253 | 0.5353 | 0.5453 |  |
| 13 | 0.4482 | 0.3572 | 0.3672 | ${ }^{0.3772}$ | 0.3872 | ${ }^{0.3972}$ | 0.4072 | 0.4172 | 0.4272 | 0.4372 | 0.4482 | 0.4572 | 0.4672 | ${ }^{0.4772}$ | 0.4872 | 0.4972 | 0.5572 | ${ }^{0.5172}$ | ${ }^{0.5572}$ | 0.5372 | 0.5472 |  |
| 71 | 0.4488 | 0.3578 | 0.3678 | 0.3778 | 0.3878 | 0.3978 | 0.4078 | ${ }^{0.4178}$ | 0.4278 | 0.4378 | 0.44 | 0.4578 | 0.4678 | 0.4778 | 0.4878 | 0.4978 | 0.5578 | ${ }^{0.5178}$ | ${ }^{0.5278}$ | ${ }^{0.5378}$ | 0.5478 |  |
| 72 90 | 0.4494 0.4494 | 0.3584 <br> 0.3584 <br> 0.354 | 0.3684 0.3684 | 0.3784 0.3784 | 0.3884 0.3884 | 0.3984 0.3984 | 0.4084 0.4084 | 0.4184 0.4184 | 0.4284 <br> 0.4284 | 0.4384 <br> 0.4384 | 0.4494 <br> 0.4494 | 0.4584 | 0.4684 | 0.4784 0.4784 | 0.4884 0.4884 | 0.4984 | 0.5084 | 0.5184 | 0.5284 | 0.5384 0.5384 |  | ${ }_{0}^{0.5}$ |
| 5 | 0.4515 | 0.3605 | 0.3705 | 0.3805 | 0.3905 | 0.4005 | 0.4105 | 0.4205 | 0.4305 | 0.4405 | 0.4515 | 0.4605 | 0.4705 | 0.4805 | 0.4905 | 0.5005 | 0.5105 | 0.5205 | 0.5305 | 0.5405 |  |  |
| 65 | 0.4564 | 0.3654 | 0.3754 | 0.3854 | 0.3954 | 0.4054 | 0.4154 | 0.4254 | 0.4354 | 0.4454 | 0.4564 | 0.4654 | 0.4754 | 0.4854 | 0.4954 | 0.5054 | 0.5154 | 0.5254 | 0.5354 | 0.5454 | 0.5554 |  |
| 92 | 0.4625 | 0.3715 | 0.3815 | 0.3915 | 0.4015 | 0.4115 | 0.4215 | 0.4315 | 0.4415 | 0.4515 | 0.4625 | 0.4715 | 0.4815 | 0.4915 | 0.5015 | 0.5115 | 0.5215 | 0.5315 | 0.5415 |  | 0.5615 | 0.5715 0.5743 |
| 62 | 0.4653 | ${ }^{0.3743}$ | 0.3843 | 0.3943 | 0.4043 | 0.4143 | 0.4243 | 0.4343 | 443 | 0.4543 | 0.4653 | 0.4743 | 0.4843 | 0.4943 | 0.5043 | 0.5143 | 0.5243 | 0.5343 | 0.5443 | 0.5543 | 0.56 |  |
| ${ }_{79}$ | 0.4664 | ${ }^{0.3754}$ | 0.3854 | 0.3954 | 0.4054 | 0.4154 | 0.4254 | 0.4354 | 0.4454 | 0.4554 | 0.4664 | 0.4754 | 0.4854 0.489 | 0.4954 | 0.5054 | 0.5154 | 0.5254 | 0.5354 | - 0.5454 | 0.5554 0.559 0.550 | 0.5654 |  |
| 79 | 0.47 | 0.379 | 0.389 | 0.399 | 0.4093 | 0.4193 | 0.4293 | 0.439 0.4393 | 0.449 0.4493 | 0.459 0.4593 | 0.47 0.4703 | 0.479 0.4793 | 0.489 0.4893 | 0.4999 0.4993 | 0.5093 | 0.519 0.5193 | 0.5293 | 0.5393 | 0.549 0.5493 | 0.559 | 0.569 |  |
| 68 | 0.4703 0.473 | 0.3793 0.382 | 0.3893 0.392 | ${ }^{0.3993}$ | 0.4.412 | ${ }_{0}^{0.422}$ | 0.4293 0.432 | 0.442 | 0.4452 | 0.462 | 0.473 | 0.482 | 0.492 | 0.502 | 0.512 | 0.522 | 0.532 | 0.542 |  | 0.0 .562 | 0.5097 |  |
| ${ }_{91}^{88}$ | 0.4739 | 0.3829 | 0.3929 | 0.4029 | 0.4129 | 0.4229 | 0.4329 | 0.4429 | 0.4529 | 0.4629 | 0.4739 | 0.4829 | 0.4929 | 0.5029 | 0.5129 | 0.5229 | 0.5329 | 0.5429 | 0.5529 | 0.56 | 0.5729 | 0.5829 |
| 11 | 0.4796 | 0.3886 | 0.3986 | 0.4086 | 0.4186 | 0.4286 | 0.4386 | 0.4486 | 0.4586 | 0.4686 | 0.4796 | 0.48 | 0.4986 | 0.5086 | 0.5186 | 0.5286 | 0.5386 | 0.5486 |  |  |  |  |
| 94 | ${ }^{0.4828}$ | ${ }^{0.3918}$ | ${ }^{0.4018}$ | ${ }^{0.4118}$ | 0.4218 | 0.4318 | 0.4418 | ${ }^{0.4518}$ | 0.4418 | ${ }^{0.4718}$ | 0.4828 | -0.4918 | 0.5018 0.5113 | 0.5118 0.5213 | ${ }_{0}^{0.5218}$ | ${ }^{0.5318}$ | 0.5418 0.513 |  | 0.5618 0.5713 | 0.5718 0.5813 | 0.5818 0.5913 | 0.5 |
| 48 | 0.4923 | ${ }^{0.4013}$ | ${ }^{0.4113}$ | ${ }^{0.4213}$ | 0.4313 | 0.4413 | 0.4513 | ${ }^{0.4613}$ | 0.4713 | 0.4813 | 0.4923 | 0.5013 | 0.5113 | 0.5213 | ${ }^{0.5313}$ | 0.5413 |  | 0.56 | 0.5713 | -. 5813 | 0.5913 |  |
| ${ }_{6}^{40}$ | 0.4948 | - | -0.41388 | 0.4238 0.4327 | 0.4338 0.4427 | 0.4438 0.4527 | 0.4538 0.4627 | 0.4638 0.4727 | 0.4738 0.4827 | -0.4838 | - | (e.5127 | 0.5138 0.5227 | 0.5238 0.5327 | -0.5338 | 0.5438 | 0.5538 0.5527 | 0.5638 0.5727 | -0.5827 | ${ }^{0.5838}$ | 0.5938 0.6027 |  |
| 19 | ${ }_{0}^{0.5053}$ | 0.4143 | -0.4243 | 0.4343 | 0.4447 | ${ }_{0}^{0.45543}$ | -0.4643 | 0.47743 0.4 | 0.4843 | -.4943 | 0.5053 | 0.5143 | ${ }_{0.5243}$ | ${ }_{0.5343}^{0.537}$ | 0.5443 | 0.5543 | ${ }_{0.5643}^{0.567}$ | ${ }_{0.5743}^{0.5727}$ | ${ }_{0.5843}^{0.527}$ | 0.5943 | ${ }_{0}^{0.6043}$ | 0.61 |
| 4 | ${ }^{0.5075}$ | 0.4165 | 0.4265 | 0.4365 | 0.4465 | 0.4565 | 0.4665 | 0.4765 | 0.4865 | 0.4965 | 0.5075 | 0.5165 | 0.5265 | 0.5365 | 0.5465 | 0.5565 | 0.5665 | 0.5765 | 0.5865 | 0.5965 | 0.6065 |  |
| 66 | 0.5175 | 0.4265 | 0.4365 | 0.4465 | 0.4565 | 0.4665 | 0.4765 | 0.4865 | 0.4965 | 0.5065 | 0.5175 | 0.5265 | 0.5365 | 0.5465 |  | 0.5665 | 0.5765 | 0.5865 | 0.5965 | 0.6065 | $\frac{0.6165}{0.617}$ |  |
| ${ }^{87}$ | 0.5196 | 0.4276 | ${ }^{0.4386}$ | $\frac{0.447}{0.4486}$ | 0.4586 | 0.4686 | ${ }^{0.4786}$ | 0.4886 | 0.4986 | 0.5086 | 0.5196 | 0.5286 | 0.5386 | 0.5486 |  |  | 0.5 | ${ }^{0.5888}$ |  | 0.607 | $\frac{0.617}{6186}$ |  |
| 53 | 0.5268 | 0.4358 | 0.4458 | 0.4558 | 0.4658 | 0.4758 | 0.4858 | 0.4958 | 0.5058 | ${ }^{0.5158}$ | ${ }^{0.5268}$ | 0.5358 | 0.5458 |  | 0.5658 | 0.5758 | 0.5858 | 0.5958 | 0.6058 | 0.6158 | 0.6258 | 0.63 |
| 54 | 0.527 | 0.436 | 0.446 | 0.456 | 0.466 | 0.476 | 0.486 | 0.496 | 0.506 | 0.516 | ${ }^{0.527}$ | 0.536 | 0.546 |  | 0.56 | 0.576 | 0.586 | 0.595 | 0.606 | 0.616 | ${ }^{0.626}$ |  |
| 27 | 0.533 | 0.442 | 0.452 | 0.462 | 0.472 | 0.482 | 0.492 | 0.502 | 0.512 | 0.522 | 0.533 |  |  |  |  |  |  |  |  |  |  |  |
| 61 | 0.5354 | 0.4444 | 0.4544 | 0.4644 | 0.4744 | 0.4844 | 0.4944 | 0.5044 | 0.5144 | 0.5244 | 0.5354 | 0.5444 | 0.55 | 0.564 | 0.5744 | 0.5844 | 0.5944 | ${ }^{0.6044}$ | 0.6144 | 0.6244 | 0.6344 | 4 |
| 28 | 0.5372 | . 4462 | 0.4562 | 4662 | 0.4762 | 0.4862 | 0.4962 | 0.5062 | 0.5162 | 0.5262 | 0.5372 | ${ }^{0.5462}$ |  |  |  |  |  |  | ${ }^{0.6162}$ | 0.6262 | ${ }^{0.6362}$ |  |
| 3 | 0.5374 | 0.4464 | 0.4564 | 0.4664 | 0.4764 | 0.4864 | 0.4964 | 0.5064 | 0.5164 | 0.5264 | 0.5374 | 0.5464 | 0.5564 | 0.5664 | 0.5764 | 0.5864 | 0.5964 | 0.6064 | 0.6164 | 0.6264 | 0.6364 |  |
| 84 | 0.5401 | 0.4491 | 0.4591 | 0.4691 | 0.4791 | 0.4891 | 0.4991 | 0.5091 | 0.5191 | 0.5291 | 0.5401 | 0.5491 | . 5591 | 691 | 0.5791 | 0.5891 | 0.5991 | 0.6091 | 0.6191 | 0.6291 | 0.6391 |  |
| 1 | 0.5442 | 0.4532 | 0.4632 | 0.4732 | 0.4832 | 0.4932 | 0.5032 | 0.5132 | 0.5232 | 0.5332 | 0.5442 |  | 0.5632 | 0.5732 | 0.5832 | 0.5932 | 0.6032 | 0.6132 | 0.6232 | 0.6332 | 0.6432 |  |
| 52 | 0.5443 | 0.4533 | 0.4633 | 0.4733 | 0.4833 | 0.4933 | 0.5033 | 0.5133 | 0.5233 | 0.5333 | 0.5443 | 0.5533 | 0.5633 | 0.5733 | 0.5833 | ${ }^{0.5933}$ | ${ }^{0.6033}$ | 0.6133 | 0.6233 | 0.6333 | 0.6433 |  |
| 2 | 0.545 | 0.454 | 0.464 | 0.474 | 0.484 | 0.494 | 0.504 | 0.514 | 0.524 | 0.534 |  |  | 0.564 | 0.574 | 0.584 | 0.594 | 0.604 | 0.614 | ${ }^{0.624}$ | 0.634 | 0.644 |  |
| 12 | 0.549 | 0.458 | 0.468 | 0.478 | 0.488 | ${ }^{0.498}$ | ${ }^{0.508}$ | 0.518 | 0.528 | 0.538 | 0.549 |  |  |  |  |  |  |  |  |  |  |  |
| 32 | 0.5501 | 0.4591 | 0.4691 | 0.4791 | 0.4891 | 0.4991 | 0.5091 | 0.5191 | 0.5291 | 0.5391 |  | 0.5 | 0.5691 | 0.5791 | 0.5891 | 0.5991 | 0.6091 | 0.6191 | 0.6291 | 0.6391 | 0.6491 |  |
| 51 <br> 86 | ${ }^{0.5505}$ | 0.4595 | 0.4695 | ${ }^{0.4795}$ | 0.4895 | 0.4995 | ${ }^{0.5095}$ | 0.5195 | 0.5295 | 0.5395 0.5401 | 0.5505 0.5511 | 0.55 | 0.5701 | 0.5795 | 0.5895 | 0.5995 | 0.6095 | 0.6195 | ${ }^{0.6295}$ | 0.6395 | 0.6495 0.6501 |  |
| ${ }_{85}^{86}$ | ${ }_{0}^{0.5513}$ | - 0.4603 | 0.4703 | -0.4803 | ${ }_{0}^{0.49903}$ | ${ }_{0}^{0.5003}$ | ${ }_{0}^{0.5103}$ | ${ }_{0}^{0.5203}$ | ${ }_{0.5303}$ | 0.5403 | ${ }_{0}^{0.55511}$ |  | ${ }^{0.55701}$ | ${ }^{0.5801}$ | ${ }_{0}^{0.5901}$ | ${ }_{0}^{0.60003}$ | ${ }_{0}^{0.610103}$ | ${ }_{0.6203}$ | ${ }_{0}^{0.63303}$ | ${ }_{0}^{0.6403}$ | ${ }^{0.65503}$ |  |
| 33 | 0.5529 | 0.4619 | 0.4719 | 0.4819 | 0.4919 | 0.5019 | 0.5119 | 0.5219 | 0.5319 | 0.5419 | 0.5529 | 0.561 | 0.5719 | 0.5819 | 0.5919 | 0.6019 | 0.6119 | 0.6219 | 0.6319 | 0.6419 | 0.6519 |  |
| 39 | ${ }^{0.5562}$ | ${ }^{0.4652}$ | 0.4752 | ${ }^{0.4852}$ | 0.4952 | 0.5052 | ${ }^{0.5152}$ | 0.5252 | 0.5352 | 0.5452 |  | 0.5 | 0.5752 | 0.5852 | 0.5952 | ${ }^{0.6052}$ | 0.6152 | 0.6252 | 0.6352 | 0.6452 | ${ }^{0.6552}$ |  |
| 50 | 0.5579 | 0.4669 | 0.4769 | 0.4869 | 0.4969 | 0.5069 | 0.5169 | 0.5269 | 0.5369 | 0.5469 |  |  | 0.576 |  |  |  | 0.6169 |  |  |  |  |  |
| 20 | ${ }^{0.5721}$ | 0.4811 | 0.4911 | 0.5011 | 0.5111 | 0.5211 | 0.5311 | 0.5411 |  | 0.5611 <br> 0.515 | 0.5722 | 0.581 | ${ }^{0.5911}$ | ${ }^{0.6011}$ | ${ }^{0.6111}$ | ${ }^{0.6211}$ | 0.66311 | ${ }^{0.6411}$ | ${ }^{0.6511}$ | 0.6611 | 0.6711 |  |
| 83 | 0.5725 | 0.4815 | 0.4915 | 0.5015 | 0.5115 | 0.5215 |  |  | 0.5515 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 34 | 0.5729 | 0.4819 | 0.4919 | 0.5019 | 0.5119 | 0.5219 | 0.5319 | 0.5419 | 0.5519 | 0.5619 |  |  | 0.5919 | 19 | 0.6119 | 0.6219 | 0.6319 | 0.6419 | 0.6519 | 0.6619 | 6719 |  |
| ${ }_{29}^{26}$ | 0.5735 | 0.4825 | 0.4925 | 0.5025 | ${ }_{0}^{0.5125}$ | 0.5225 | ${ }_{0}^{0.5325}$ | 0.5425 | 0.5525 0.5531 | 0.5625 0.5631 0.581 | - |  | 0.5925 | 0.6025 | 0.6125 | ${ }^{0.6225}$ | ${ }^{0.6325}$ | ${ }^{0} 0.6425$ | O.6525 | ${ }_{0}^{0.6625}$ | 0.6725 |  |
| 29 31 | ${ }_{0}^{0.57471}$ | 0.4831 | 0.4931 0.4931 | 0.5031 | ${ }_{0}^{0.5131}$ | 0.5231 0.5231 | 0.5331 0.5311 | 0.5431 0.5431 | ${ }^{0.5531}$ | - 0.5631 | ${ }_{0}^{0.574}$ | [0.5831 | ${ }^{0.5931}$ | ${ }_{0}^{0.6031}$ | ${ }_{0.6131}^{0.6131}$ | ${ }_{0}^{0.6231}$ | ${ }_{0}^{0.63331}$ | ${ }_{0}^{0.6431}$ | ${ }_{0.6531} 0.051$ | ${ }_{0}^{0.6631}$ | ${ }_{0.6731}^{0.671}$ |  |
| 31 38 | 0.5741 0.579 | - | 0.4931 0.4949 | 0.5049 | ${ }_{0}^{0.5141}$ | 0.5231 0.5249 | ${ }_{0.5349}^{0.531}$ | 0.5449 | 0.5549 | ${ }_{0}^{0.5649}$ | 0.575 | 0.584 | 0.5949 | 0.6049 | ${ }_{0.6149}^{0.6131}$ | ${ }_{0}^{0.6249}$ | ${ }_{0.6349}^{0.631}$ | ${ }_{0.6449}$ | ${ }_{0} 0.6$ | ${ }_{0}^{0.6649}$ | 0.6749 |  |
| 24 | 0.5767 | 0.4857 | 0.4957 | 0.5057 | 0.5157 | 0.5257 | 0.5357 | 0.5457 | 0.5557 | 0.5657 | 0.576 | 0.5 | 0.5957 | 0.6057 | 0.6157 | 0.6257 | 0.6357 | 0.6457 | 0.6557 | 0.6657 | 0.6757 |  |
| 80 | ${ }^{0.5769}$ | 0.4859 0.4869 | ${ }^{0.4959}$ | ${ }^{0.5059}$ | 0.5159 0.5159 | 0.5259 0.5269 | ${ }^{0.5359}$ | 0.5459 0.5469 | 0.5559 0.5569 | 0.5659 0.5659 | ${ }_{0}^{0.5769}$ | 0.58 | 0.5959 0.5969 | 0.6059 0.6069 |  | ${ }_{0}^{0.625}$ | 0.6359 0.6369 |  |  | 0.6659 0.6699 | 0.6759 0.6769 |  |
| ${ }_{55}$ | ${ }_{0}^{0.5818}$ | 0.4908 | -0.5008 | ${ }_{0.5108}^{0.5099}$ | ${ }_{0}^{0.5208}$ | ${ }_{0}^{0.5508}$ | 0.5349 0.5408 |  | ${ }_{0}^{0.5669}$ | 0.5699 0.5708 | 0.5819 <br> 0.518 | ${ }^{0.5860}$ | ${ }^{0.56008}$ | ${ }^{0.66108}$ | ${ }_{0.6208}^{0.65}$ | ${ }_{0}^{0.6308}$ | 0.6408 | 0.6508 | 0.6608 | ${ }_{0}^{0.6708}$ | 0.6808 |  |
| 30 | ${ }^{0} .5889$ | ${ }^{0.4909}$ | 0.5009 | 0.5109 | ${ }^{0.5209}$ | 0.5339 | 0.5409 | . 55 | 0.5609 | 0.5709 | 0.58 | 0.5 | 0.6009 | 0.6109 | 0.6209 | 0.6309 | 0.6409 | 0.6509 | ${ }^{0.6609}$ | 0. 6779 | 0.6889 | - $\begin{array}{r}0.690 \\ 0.501\end{array}$ |
| 25 56 | 0.5825 0.5963 | 0.4915 0.5053 | 0.5015 0.5153 | 0.5115 0.5253 | 0.5215 0.5353 | 0.5315 0.5453 | ${ }^{0.5415}$ | 0.5515 0.5653 | 0.5615 0.5753 | 0.5715 0.5853 | ${ }_{0}^{0.5825}$ | ${ }_{0}^{0.591} 0.605$ | ${ }^{0.6015} 0.6153$ | ${ }_{0}^{0.6115}$ | ${ }_{0}^{0.6215}$ | ${ }^{0.6315} 0$ | ${ }_{0}^{0.64553}$ | ${ }^{0.6515}$ | 0.6615 0.6753 | ${ }^{0.67853}$ | ${ }_{0}^{0.6815} 0$ |  |
| 95 | 0.598 | 0.507 | 0.517 | 0.527 | 0.537 | 0.547 | 0.557 | 0.567 | 0.577 | 0.587 | 0.5 |  | 0.617 | 0.627 | 0.637 | 0.647 | 0.657 | 0.667 | 0.677 | 0.687 | 0.697 |  |
| 23 | 0.5985 | 0.5075 | 0.5175 | 0.5275 | 0.5375 | 0.5475 | 0.5575 | 0.5675 | 0.5775 | 0.5875 | 0.5 | 0.60 | 0.6175 | 0.6275 | 0.6375 | 0.6475 | 0.6575 | 0.6675 | 0.6775 | 0.6875 | 0.6975 | 0.70 |
| 35 | ${ }^{0.601}$ | 0.51 | 0.52 | 0.53 | 0.54 |  | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | S | ${ }^{0.62}$ | 0.63 0.63 | 0.64 | 0.65 | 0.66 | 0.67 | 0.68 | 0.69 | -0.7 |  |
| 36 | ${ }_{0}^{0.6169}$ | 0.51 0.5259 | 0.5359 | 0.5459 | 0.54 | 0.5659 | 0.5759 | 0.5859 | 0.5959 | 0.5059 |  |  |  |  |  |  | 0.6759 |  | 0.6959 |  |  |  |
| 22 | 0.621 | 0.53 | 0.54 |  | 0.56 | 0.57 | 0.58 | 0.59 |  | 0.61 | 0.621 | 0.6 | 0.64 | 0.65 | 0.66 | 0.67 | 0.68 | 0.69 | 0.7 | 0.71 | 0.72 | - 0.7 |
| 21 | 0.6231 | 0.5321 | 0.5421 | 0.5521 | 0.5621 | 0.5721 | 0.5821 | 0.5921 | 0.6021 | 0.6121 | 0.623 | 0.632 | 0.6421 | 0.6521 | 0.6621 | 0.6721 | 0.6821 | 0.6921 | 0.7021 | 0.7121 | 0.7221 | 0.732 |
| 82 | ${ }^{0.6243}$ | ${ }^{0.5333}$ | ${ }^{0.5433}$ | 0.5533 | 0.5633 | 0.5733 | 0.5833 | ${ }^{0.5933}$ | 0.6033 | 0.6133 | 0.624 | 0.633 | 0.6433 | ${ }^{0.6533}$ | 0.6633 | 0.6733 | 0.6833 | ${ }^{0.6933}$ | 0.7033 | 0.7133 | 0.7233 | 0.733 |
| 81 59 | 0.6249 0.6345 | 0.5339 0.5435 | 0.5439 0.5535 | 0.5539 | 0.55335 | 0.5739 0.5835 | ${ }^{0.5839}$ | 0.5939 | ${ }^{0.6039} 0$ | ${ }^{0.6139}$ | 0.6249 | ${ }_{0}^{0.633}$ | 0.6439 0.6535 | ${ }^{0.6539}$ | 0.6639 0.6735 | 0.6739 0.6835 | 0.68839 | ${ }^{0.6939} 0$ | 0.7039 0.7135 | 0.0 .71395 | 0.7239 0.7335 | ${ }^{0.7743}$ |
| 5 | 0.6404 | 0.5494 | 0.5594 | 0.5694 | 0.5794 | 0.5894 | 0.5994 | 0.6094 | . 6194 | 0.6294 | 0.6404 | 0.6 | 0.6594 | 0.6694 | 0.6794 | 0.6894 | 0.6994 | 0.7094 | 0.7194 | 0.7294 | 0.7394 | 0.74 |
| 96 | ${ }^{0.6798}$ |  | 0.5988 | 0.6088 | 0.6188 | 0.6288 | 388 396 | 0.6488 | ${ }^{0.6588}$ | . 6688 | 0.6798 |  | 0.6988 | 0.7088 | 0.7188 0.7196 0 | 0.7288 | 0.7388 | 0.7488 0.7496 |  | 0.7688 | 0.7788 0.7796 |  |
| 97 | ${ }^{0} 0.78086$ |  | 0.59226 | ${ }_{0}^{0.6326}$ |  |  | ${ }^{0.6396}$ | ${ }_{0}^{0.64726}$ |  |  |  | ${ }_{0} .8 .7126$ | -0.7226 | 0.7326 | 0.7426 | 0.7526 | 0.7626 | 0.7726 | 0.7826 | 0.7926 | 0.8026 | 0.81 |
| 98 | ${ }_{0}^{0.7266}$ | ${ }_{0}^{0.6356}$ | ${ }_{0}^{0.6456}$ | ${ }_{0}^{0.65556}$ | ${ }_{0}^{0.64656}$ | ${ }_{0.6756}^{0.656}$ | ${ }_{0}^{0.68556}$ | 0.6956 | 0.7056 | 0.7156 | 0.7266 | 0.7356 | 0.7456 | 0.7556 | 0.7656 | 0.7756 | 0.7856 | 0.7956 | 0.8056 | 0.8156 | 0.8256 | 0.83 |
| 99 | 0.7266 | 0.6356 |  |  |  |  |  |  |  |  | 0.726 | 0.73 | 74 | 0.7556 |  | 0.77 | 2.785 | 2,7 |  |  | S2 | 0.8356 |

District/COMP All_40 All_41 All_42 All_43 All_44 All_45 All_46 All_47 All_48 Composite All_50 All_51 All_52 All_53 All_54 All_55 All_56 All_57 All_58 All_59 All_60

|  | II_40 | All_41 | II_42 | II_43 | All_44 | II_45 | All_46 | II_47 | All | posit | All_50 | All_51 | All_52 | All_53 | All_54 | All_5 | All_5 | All_5 | All_5 | All_5 | All_60 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 16 | 0.0502 | 0.0602 | 0.0702 | 0.0802 | 0.0902 | 0.1002 | 0.1102 | 0.1202 | 0.1302 | 0.1412 | 0.1502 | 0.1602 | 0.1702 | 0.1802 | 0.1902 | 0.2002 | 0.2102 | 0.2202 | 0.2302 | 0.2402 | 0.2502 |
| 10 | 0.0848 | 0.0948 | 0.1048 | 0.1148 | 0.1248 | 0.1348 | 0.1448 | 0.1548 | 0.1648 | 0.1 | 0.184 | 0.1948 | 0.2048 | 0.2148 | 0.2248 | 0.2348 | 0.2 | 0.2548 | 0.2648 | 0.2748 |  |
|  | 0.0926 | 0.1026 | 0.1126 | 0.1226 | 0.1326 | 0.142 | 0.1526 | 0.16 | 0.1726 | 0.1 | 0.192 | 0.202 | 0.2 | 0.22 | 0.2 | 0.2 | 0.2 | 0.2626 | 0.2726 | 0.2826 | 0.2926 |
| 17 | 0.115 | 0.125 | 0.135 | 0.145 | 155 | 165 | 175 | 85 | 0.195 | 0.2 | 0.21 | 0.225 | 0.235 | 0.24 | 0.255 | 0.265 | 0.275 | 285 | 0.29 | 305 | 0.315 |
| 76 | 0.117 | 0.12 | 13 | 0.1475 | 0.1575 | 0.1675 | 0.1775 | 0.1 | . 197 | . 20 | 0.21 | . 2 | 0.2375 | 0.2475 | 0.2575 | 0.2675 | 0.2775 | 0.2875 | 0.2975 | 0.3075 | 0.3175 |
|  |  | 0.137 |  |  |  |  |  |  |  |  |  | 0.237 |  | 0.257 | 0.267 |  | 0.287 | 0.29 | 0.30 | 0.317 |  |
| 77 | 0.1685 | 0.1785 | 0.1885 | 0.1985 | 0.2085 | 0.2185 | 0. 2285 | 0.2385 | 0. 2485 | 0.259 | 0.268 | 0.2785 | 0.2885 | 0.2985 | 0.3085 | 0.3185 | 0.3285 | 0.3385 | 0.3485 | 0.3585 | 0.3685 |
| 15 | 793 | 0.1 | 0.1993 | 0.2093 | 219 | 22 | d. 239 | 49 | 0. 2593 | . 270 | 27 | 0.289 | 0.29 | 0.309 | 0.31 | 0.329 | 0.339 | 0.349 | . 35 | 0.36 | 0.3793 |
| 11 | 0.2154 | 0.225 | 0.235 | 0.2 | 0.255 | . 265 | 0.2754 | 0.285 | 0.2954 | 306 | 0.31 | 0.325 | 0.335 | 0.345 | 0.35 | 0.365 | 0.375 | 0.385 | 0.39 | 0.40 | 0.4154 |
| 18 | 0.2205 | 0.2305 | 0.2405 | 0.250 | 0.2605 | 0.2705 | 0.2805 | 0.2905 | 0.3005 | 0.3 | 0.3 | 0.330 | 0.3 | 0. | 0.36 | 0.3 | 0.3 | 0.3 | 0. | 0.4105 | 0.4205 |
| 78 | 0.2211 | 0.2311 | 0.241 | 0.251 | 0.2611 | 0.271 | 0.2811 | 0.2911 | 0.3011 | 0.31 | 0.32 | 0.331 | 0.341 | 0.351 | 0.361 | 0.371 | 0.381 | 0.391 | 0.401 | 0.41 | 0.4211 |
| 75 | 0.2233 | 0.2333 | 0.2433 | 0.2533 | 0.2633 | 0.2733 | 0.2833 | 0.2933 | 0.3033 | 0.314 | 0.323 | 0.333 | 0.343 | 0.353 | 0.363 | 0.373 | 0.383 | 0.393 | 0.403 | 0.413 | 0.4233 |
| 74 | 0.2898 | 0.299 | 0.3098 | 0.319 | 0.3298 | 0.339 | 0.349 | 0.359 | 0. 3698 | 0.38 | 0.38 | 0.399 | 0.409 | 0.419 | 0.429 | 0.439 | 0.449 | 0.4598 | 0.46 | 0.4798 | 0.4898 |
| 79 | 0.2928 | 0.3 |  |  |  |  |  | 0. |  |  | 0.392 | 0.4028 | 0. | 0.422 | 0. 4 |  | 0 | 0.4 | 0.4728 | 0.4828 |  |
| 46 | 0.2996 | 0.3 | 0.3196 | 0.3296 | 0.3396 | 0.349 | 0. | 0.3696 | 0. | 0.390 | 0.399 | 0.4096 | 0.419 | 0.4296 | 0. | 0.449 | 0.459 | 0.4696 | 0.479 | 0.4896 |  |
|  | 0.3142 | 0.3242 | 0.3342 | 0.3442 | 0.3542 | 0.3642 | 0.3742 | 0.3842 | 0.3 | 0.4052 | 0.414 | 42 | 0.4342 | 0.444 | 45 | 0.46 | 47 | 0.4842 | 0.4942 | 0.50 | 0.5142 |
| 44 | 0.3148 | 0.324 | 0.334 | 0.344 | 0.354 | . 364 | 0.374 | 0.384 | 0.3948 | O.405 | 0.414 | 0.42 | 0.43 | 0.44 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.504 | 0.5148 |
| 47 | 0.3157 | 0.325 | 0.335 | 0.345 | 0.355 | . 3657 | 0.375 | 0.38 | 0.39 | 0.4067 | 0.415 | 0. | 0.4 | 0.445 | 0.4 | 0.4657 | 0.4757 | 0.48 | 0.4957 | 0.5057 | 0.5157 |
| 48 | 0.3197 | 0.3297 | 0.33 | 0.34 | 0.3 | 0.3 | 0.3797 | 0.3 | 0.3997 | 0. | 0. | 0. | 0. | 0.4497 | 0.4 | 0.469 | 0. | 0.489 | 0.499 | 0.509 | 0.5 |
| 19 | 0.3203 | 0.3303 | 0.3403 | 0.3503 | 0.3603 | 0.3703 | 0.3803 | 0.3 | 0.4003 | 0.4 | 0.4 | 0.430 | 0.4 | 0.4503 | 0.4603 | 0.4703 | 0.4803 | 0.4903 | 0.500 | 0.5103 | 0.5203 |
| 43 | 0.321 | 0.33 | 0.341 | 0.35 | . 36 | 0.37 | . 381 | 0.39 | 0.401 | . 4 | 0.42 | 0.431 | 0.44 | 0.45 | 0.46 | 0.47 | 0.48 | 0.491 | 0.50 | 11 |  |
| 4 | 0.342 | 0.352 | 0.362 | 0.37 |  | 0.39 | 0.402 | 0.41 | 0.422 |  |  | 0.452 | 0.462 |  |  |  | 0.50 | 0.512 | 0.52 | 0.532 |  |
| 71 | 0.3428 | 0.3528 | 0.3628 | 0.3728 | 0.3828 | 92 | 02 | 0.4128 | 422 | 0.433 | 0.442 | 0.4528 | 0.4628 | 0.4728 | 0.4828 | 0.492 | 50 | 51 | 52 | 0.532 | 0.542 |
| 72 | 0.3501 | 360 | 0.370 | 0.3801 | 0.3901 | 00 | 0.4101 | 0.4201 | 0.4301 | 0.441 | . 45 | 460 | . 47 | . 48 | . 49 | 0.500 | 51 | 0.520 | 0.53 | 0.54 |  |
| 93 | 0.3519 | 0.361 | 0.371 | 0.381 | 0.391 | 0.4019 | 0.411 | . 421 | 0.4319 | 0.442 | . 45 | . 46 | 0.47 | 0.48 | 0.49 | 0.501 | 0.51 | 0.521 | 0.531 | 0.541 |  |
| 94 | 0.3533 | 0.363 | 0.373 | 0.383 | 0.393 | 0.4033 | 0.413 | 0.423 | 0.4333 | 0.444 | 0.453 | 0.4633 | 0. | 0.4833 | 0.4 | 0.503 | 0.51 | 0.523 | 0.53 | 0.5 |  |
| 45 | 0.3541 | 0.3 | 0.3 | 0.3 | 0. | 0.404 | 0.414 | 0. | 0. | 0.445 | 0.454 | 0.464 | 0. | 0.484 | 0. | 0.50 | 0.5 | 0.52 | 0.53 | 0.54 |  |
| 20 | 0.3552 | 0.3 | 0.3 | 0.3 | 0. | 0.4052 | 0. | 0. | 0.4 | 0.446 | 0. | 0.4 | 0.4 | 0. | 0.4 | 0.505 | 0.51 | 0.525 | 0.53 | 0.54 |  |
| 73 | 0. | 0.367 | 0.37 | 0.3 | 0. | 0.4078 | 0. | 0.4 | 0.4378 | 0.448 | 0.4 | 467 | 0.47 | 0.48 | 0.49 | 0.50 | 0.51 | 0.527 | 0.53 | 0.547 |  |
| 14 | 0.3602 | 0.3 | 0.3802 | 0.3902 | 0.4002 | 0.4102 | 0.4202 | 0.430 | 0.4402 | 0.4512 | 0.460 | 0.4702 | 0.480 | 0.490 | 0.500 | 0. | 0.520 | 0.530 | 0.54 | 0.550 |  |
| 12 | 0.3649 | 0.37 |  | 0.3949 |  |  |  |  |  | 0.455 | 0.46 | 0.4749 | 0.484 | 0.494 | 0.50 | 0.5 | 0.52 | 0.53 | 0.54 | 0.55 |  |
| 80 |  | 0.377 |  |  |  |  |  |  |  |  |  | 0.477 | . 48 | 0.497 | 0.50 | 0.51 | 0.527 | 0.537 | 0.5 |  |  |
| 70 | 3762 | 0.386 | . 3962 | 406 | 4162 | . 426 | 0.4362 | 446 | 0.4562 | . 467 | . 476 | 486 | 0.496 | 0.5062 | 0.516 | 0.5262 | 0.536 | 0.5462 | 0.556 | 0.56 | 0.57 |
| 91 | 0.3779 | 0.387 | 0.397 | 407 | 0.417 | . 427 | 0.4379 | 447 | . 4579 | . 468 | . 47 | 48 | 0.497 | 0.507 | 0.517 | 0.5279 | 0.537 | 0.5479 | 0. 55 | 0.56 |  |
| 95 | 0. | 0.388 | 0.3 | 0.408 | 0.418 | . 428 | 0.438 | 0. | 0.4587 | 0.4697 | 0.478 | 0.488 | 0. | 0.5 | 0.5 | 0.52 | 0.53 | 0.5487 | 0.558 | 0.568 |  |
| 21 | 0. | 0.3 | 0.4 | 0.4 | 0.4 | 0.4316 | 0.441 | 0.4 | 0.4616 | 0.4726 | 0. | 0.4916 | 0.5 | 0.51 | 0.52 | 0.53 | 0.54 | 0.5516 | 0.5 | 0.57 |  |
| 64 | 0.38 | 0.3946 | 0.40 | 0.41 | 0.424 | 0.43 | 0.4 | , | 0.4 | 0.4756 | 0.48 | 49 | 0.50 | 0.51 | 0.52 | 0.53 | 0.5446 | 0.5 | 0.5 | 0.5 |  |
| 69 | 0.39 | 0.4 | 0.41 | 0.42 | 0.43 | 0.4 | 0.45 | 0.46 | 0.47 | 0.481 | 0.4 | 0.5 | 0.51 | 0.52 | 0.53 | 0.54 |  | 0.56 | 0.5 | 0.58 |  |
| 49 | 3936 | 03 | 0.4136 | 0.4236 | 0.4336 | 0.4436 | 0.4536 | 0.4636 | 0.4736 | 0.484 | 93 | 036 | 51 | 0.5236 | 53 | 0.543 | 0.55 | 63 | 573 | 0.5836 |  |
| 90 | 0.396 | 0.406 | 0.416 | 0.42 | 0.436 |  | 0.456 | 0.466 | 0.476 | 0.487 | 0.4 | 50 | 0.516 | . 5 | . 53 | 0.54 |  | . 56 | 0.576 | . 58 |  |
| 92 | 0.3976 | 0.407 | 0.4176 | 0.427 | 437 | 0.4476 | 0.4576 | 0.467 | 0.4776 | . 488 | 0.497 | 0.507 | 0.517 | 0.527 | 0.537 | 0.5476 | 0.55 | 0.567 | 0.57 | 0.587 | . 5 |
| 68 | 02 | 0.412 | . 422 | . 432 | . 442 | 0.4521 | 0.462 | . 472 | 0.4821 | , 93 | 0.502 | 0.512 | 0.522 | 0.532 | 0.542 | 521 | 0.56 | 0.572 | 0.58 | 0.59 |  |
|  |  | 0.4 | . | 0.433 |  |  | 0.4636 |  | 0.4836 |  |  | 0.5 | 0.5 | 0.53 | 0.54 |  |  |  |  |  |  |
| 62 | 0.4066 | 0.4 | 0.4 | 0.436 | 0.446 | 0.456 | 0.4666 | 0. | 0.4 | 0.497 |  | 0.5 | 0.5 | 0.53 | 0.5 |  |  | 0.57 |  |  |  |
| 61 | 0.4076 | 0.4 | 0.4 | 0.437 | 0. | 0.4576 | 0.4676 | 0.477 | 0. | 0.498 | 0.5 | 0.517 | 0.52 | 0.53 | 0.54 | 0.55 | 0.56 | 0.57 | 0.58 | 0.5 |  |
| 67 | 0.4088 | 0.4 | 0.4288 | 0.438 | 0.4488 | 0.458 | 0.4688 | 0.478 | 0.4888 | . 49 | 0.508 | 0.5188 | 0 | 0.5 | 0.5 | 0 | 0.56 | 0. | 0.58 | 0.598 |  |
| 63 | 0.4101 | 0.420 | 0.430 |  | 0.4501 | 0.460 | 0.470 | 0.480 | 0.4901 | . 501 | 0.510 | . 52 | 0.530 | 0.5 | 0.55 | - | 0.57 | 0.58 | 0. | 0.6 |  |
| 88 | 0.4128 | 0.422 | 0.4328 | 0.442 | 0.4528 | 0.4628 | 0.4728 | 0.482 | 0.4928 | . 503 | . 51 | . 52 | . 53 | 0.54 | 0.55 | 0.56 | 0.5 | 0.58 | 0.59 | 0.60 |  |
| 89 | 0.4141 | 0.4241 | 0.434 | 0.44 | 0.4541 | 464 | 0.4741 | . 484 | 0.4941 | 0.5051 | 0.514 | 0.5241 | 0.5341 | 0.544 | 0.554 | 0.564 | 0.574 | 0.5841 | 0.594 | 0.604 |  |
| 13 | 0.4152 | 0.4252 | 0.4352 | 0.445 | 0.4552 | 0.4652 | 0.4752 | 0.4852 | 0.4952 | 0.5062 | 0.5152 | 0.5252 | 0.5352 | 0.5452 | 0.5552 | 0.5652 | 0.575 | 0.5852 | 0.595 | 0.6052 |  |
| 86 | 0 | 0.42 | 0.438 |  | 0.458 | 0.468 | 47 | 48 | 498 | . 509 | 0.51 | 0.528 | 0.538 | 0.54 |  |  | 0.5 |  | 0.5 | 0.608 |  |
| 50 | 0.419 | 0. | 0.43 | 0.44 | 0.45 | 0.46 | 0.4 | 0.48 | 0.499 | 0.5 | 0.5 | 0.5 | 0.53 | 0.54 | 0.55 | 0.5 | 0.57 | 0.5 | 0.5 | 0.60 |  |
| 41 | 0.419 | 0.429 | 39 | 0.449 | 0.459 | 0.4696 | 0.479 | . 48 | 0.4996 | . 51 | 0.51 | 0.529 | 0.539 | 0.5496 | 0.55 | . 5 | . 5 |  | . 5 | 0.60 |  |
|  |  | 0.431 |  |  |  | . 47 | . 48 | 0.491 | 0.5014 | . 512 |  | 0.53 | 0.54 | 0.55 | 0.56 | 0.57 |  |  | 0.60 |  |  |
|  | 0.4263 | 0.4363 |  | 0.4563 | , | 0.4763 | 4863 | 0.4963 | 0.5063 | 5173 |  | 0.536 | 0.5463 |  | 0.56 | 0.57 | 0.5 | 0.59 | 0.60 | 0.61 |  |
|  | 0.429 | 0.439 | 0.44 | 0.459 | 0.469 | 0.479 | 0.489 | 0.499 | 509 | 0.52 |  | 0.53 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.60 | 0.6 |  |
|  | . 4296 | . 439 | 0.4 | 0.4596 | 0.4696 | 0.479 | 0.4896 | 0.4996 | . 096 | . 520 | . 52 | . 5396 | 0.549 | 0. 55 | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | 0.61 |  |
| 35 | 0.4309 | 0.440 | 0.45 | 0.460 | 0.470 | 0.480 | 0.4909 | 0.500 | 109 | 0.521 | 0.530 | 0.5409 | 0.550 | . 56 | 0.57 | 0.58 | 0.590 | 0.600 | 0.61 | 0.62 |  |
| 60 | 31 | 0.44 | 0.45 | 0.46 | 0.47 | 0.481 | 0.491 | 0.501 | 0.511 | 0.52 | 0.53 | . 54 | 0.55 | 0.56 | 0.5 | 0.58 | 0.5 | . 60 | 0.6 | 0.62 |  |
| 53 | 0.4335 | 0.443 | 0.4535 | 463 | 0.473 | 483 | 0.493 | 0.503 | 0.5135 | 524 | 0.533 | 0.5435 | 0.55 | 0.56 | 0.57 | 0.583 | 0.59 | 0.60 | 0.61 | 0.62 |  |
|  | 0.4349 | 0.4 |  |  | 0.474 | 0.4849 | 0.4 |  | 0.51 | . 5 | 0.53 | 0.5449 | 0.5 | 0.5 | 0.57 | , | 0. | . | 0.6 | 0.62 |  |
|  |  | 0.447 |  | 0.467 | 崖 | 0.487 | 0.497 | 0.50 | 0.5172 | . 528 | 0.53 | 0.547 | 0.55 | 0.56 | 0.57 | 0.58 | -. 59 | 0.60 | 0.61 | 0.62 |  |
|  |  | 0.448 |  | 0.468 | 0.478 |  |  |  |  |  |  | 0.548 |  |  |  | 0.5086 |  | 0.60 | 0.61 | 0.62 |  |
|  |  | 0.450 | 0.460 | 0.470 | 0.480 | 0.490 | 500 | 0.510 | 0.520 | 0.531 | 54 |  | 0.5 | 0.570 | 0.5 | 0.590 | 0.60 | 0.61 | 0.62 | 0.63 |  |
|  |  | 0.452 |  |  | 0.482 |  | . 502 | 0.512 | . 522 | 533 | 0.542 |  | 0.56 | 0.57 | 58 | 0.592 | , 602 | 0.61 | 0. 62 | 0.63 |  |
| 3 | 0.4427 | 45 | 0.462 | 0.4727 | 0.482 | 0.4927 | 02 | 0.512 | . 522 | . 533 | . 542 | 0.552 | 0.562 | 0.57 | . 5 | 0.592 | 0.60 | 0.61 | . 62 | 63 |  |
| 26 | 0.4452 | 455 | 0.4652 | 0.4752 | 85 | 95 | 0.505 | . 515 | 0.5252 | . 536 | 0.545 | 0. | 0.565 | 0.575 | 0.58 | 0.595 | 0.60 | 0.61 | 0.62 | 0.63 | 0.6452 |
|  | . 446 | 0.456 | 0.46 | 0.476 | 0.486 | 0.49 | . 50 | 51 | 0.526 | . 53 | 0.54 | 0.55 | 0.56 | 0.57 | 0.5 | 0.59 | 0.6 | 0.61 | 0.6 | 0.63 | 0.646 |
|  | 0.4494 | 0.45 | 0.4694 | 0.479 | 0.4 | . 49 | 0.509 | 0.519 | 0.5294 | 0.540 | 0.549 | 0.5 | 0.56 | 0.57 | 0.58 | 0.599 | 0.609 | 0.61 | 0.62 | 0.63 | 0.64 |
|  | 0.4532 | 0.4632 | 0.4732 | . 483 | 4932 | 5032 | 0.5132 | 0.5232 | 0.5332 | 0.5 | 0.5 | 0.5 | 0.573 | 0.58 | 0.59 | 0.603 | 0.61 | 0.62 | 0.63 | 0.64 |  |
| 51 | 0. | 0.466 | , 76 | 0.486 | 496 | . 506 | 0.516 | 0.526 | 0.5364 | 0.5 |  | 0.5 | 0.57 | 0.586 | 0.59 | 0.606 | 0.616 | 0.626 | 0.63 | 0.64 |  |
|  | 0.4571 | 467 | 0.477 | 87 | 497 | 507 | 0.517 | 0.527 | 0.5371 | 5 | 0.55 | 0.567 | 0.577 | 0.58 | 0.59 | 0.607 | 0.61 | 0.627 | 0.63 | 0.64 |  |
|  | 0.4586 | 4686 |  | 88 | 986 | 508 | 0.5186 | 528 | 0.5386 | 0.549 | 0.55 | . 5.568 | 0.578 | 0.588 | 0.598 | 0.608 | 0.618 | 0.628 | 0.63 | 0.64 |  |
| 81 | 0.4589 | 0.4689 | 0.4789 | 889 | 0.4989 | . 508 | 0.5189 | 0.528 | 0.5389 | 0.5499 | 0.55 | 0.568 | 0.578 | 0.588 | 0.598 | 0.608 | 0.618 | 0.628 | 0.638 | 0.648 |  |
| 27 | 0.4618 | 0.4718 | 0.4818 | 4918 | 0.5018 | 0.5118 | 0.5218 | 0.5318 | 0.5418 | 0.55 | 0.561 | 0.5718 | 0.581 | 0.591 | 0.601 | 0.611 | 0.621 | 0.631 | 0.64 | 0.651 | 0.66 |
|  | 0.4634 | 0.47 | 0.4834 | . 49 | 0.5034 | 0.5134 | 0.5234 | 0.5334 | 0.5434 | 0.5 | 0.56 | 0.573 | 0.583 | 0.593 | 0.603 | 0.613 | 0.623 | 0.633 | 0.64 | 0.653 | 0.66 |
| 40 | 0.4676 | 0.4776 | 0.48 | 0.4976 | 0.5076 | 0.5176 | 0.5276 | 0.5376 | 0.5476 | 0.55 | 0.56 | 0.577 | 0.58 | 0.59 | 0.60 | 0.617 | 0.62 | 0.63 | 0.64 | 0.65 | 0.66 |
| 24 | 0.4 | 0.4 | 0.4944 | 0.504 | 0.514 | 0.524 | 0.5344 | 0.5444 | 0.5544 | 0.5 | 0.57 | 0.584 | 0.594 | 0.604 | 0.61 | 0.62 | 0.63 | 0.64 | 0.65 | 0.66 |  |
| 8 | 0. | 0.4 | 0.4988 | 0.508 | 0.518 | 0.528 | 0.5388 | 0.5488 | 0.5588 |  | 0.578 | 0.588 | 0.598 | 0.608 | 0.618 | 0.628 | 0.638 | 0.648 | 0.65 | 0.66 |  |
| 36 | 0.4797 | 0.4897 | 0.4997 | 5097 | 0.5197 | 0.5297 | 0.5397 | 0.5497 | 0.5597 | 0.570 | 0.579 | 0.589 | 0.599 | 0.600 | 0.619 | 0.629 | 0.639 | 0.649 | 0.65 | 0.66 |  |
| 29 | 0.4873 | 4973 | 0.5073 | 5173 | 0.5273 | 0.5373 | 0.5473 |  | 0.5673 | 0.57 | 0.587 | 0.5973 | 0.607 | 0.617 | 0.62 | 0.637 | 0.64 | 0.657 | 0.662 | 0.677 |  |
| 23 | 0.4876 | 0.4976 | 0.5076 | 0.5176 | 0.5276 | 0.5376 | 0.5476 | 0.557 | 0.5676 | . | 0.587 | 0.5976 | 0.607 | 0.617 | 0.62 | 0.637 | 0.64 | 0.657 | 0.66 | 0.67 |  |
| 37 | 0.5012 | 0.5112 | 0.5212 | 0.5312 | 0.5412 | 0.5512 | 0.5612 | 0.5712 | 0.5812 | 0.592 | 0.601 | 0.6112 | 0.621 | 0.6312 | 0.641 | 0.6512 | 0.661 | 0.6712 | 0.681 | 0.691 | 0.70 |
| 39 | 0.5074 | 0.5174 | 0.5274 | 0.5374 | 0.5474 | 0.557 | 0.5674 | 0.5774 | 0.5874 | 0.598 | 0.607 | 0.617 | 0.627 | 0.637 | 0.647 | 0.6574 | 0.667 | 0.677 | 0.687 | 0.697 | 0.70 |
| 56 | 0.5118 | 0.5218 | 0.5318 | 0.5418 | 0.5518 | 0.5618 | 0.5718 | 0.5818 | 0.5918 | 0.602 | 0.611 | 0.6218 | 0.631 | 0.641 | 0.651 | 0.661 | 0.671 | 0.681 | 0.69 | 0.701 | - |
| 96 | 0.5131 | 0.5231 | 0.5331 | 0.5431 | 0.553 | 0.563 | 0.5731 | 0. 583 | 0.5931 | 0.604 | 0.613 | 0.623 | 0.633 | 0.643 | 0.65 | 0.663 | 0.67 | 0.683 | 0.69 | 0.703 | 0.7 |
| 33 | 0.5 | 0.52 | 0.53 | 0.5448 | 0.554 | 0.564 | 0.5748 | 0.584 | 0.5948 | 0.605 | 0.614 | 0.624 | 0.634 | 0.644 | 0.65 | 0.664 | 0.674 | 0.684 | 0.69 | 0.70 |  |
| 30 | 0.5167 | 0.5267 | 0.5367 | 0.5467 | 0.556 | 0.566 | 0.576 | 0.586 | 0.5967 | 0.6 | 0.616 | 0.626 | 0.636 | 0.646 | 0.65 | 0.666 | 0.67 | 0.686 | 0.69 | 0.70 | 0.71 |
| 38 | 0.5273 | 0.5373 | 0.5473 | 0.557 | 0.567 | 0.577 | 0.5873 | 0.597 | 0.6073 | 0.618 | 0.627 | 0.637 | 0.647 | 0.657 | 0.667 | 0.677 | 0.687 | 0.697 | 0.7073 | 0.71 | 0.727 |
| 59 | 0.5317 | 0.5417 | 0.5517 | 0.5617 | 0.5717 | 0.5817 | 0.5917 | 0.6017 | 0.6117 | 0.622 | 0.631 | 0.6417 | 0.651 | 0.661 | 0.671 | 0.681 | 0.691 | 0.701 | 0.711 | 0.721 | 0.73 |
| 8 | 0.5329 | 0.5429 | 0.5529 | 0.5629 | 0.5729 | 0.5829 | 0.5929 | 0.6029 | 0.6129 | 0.623 | 0.632 | 0.6429 | 0.6529 | 0.6620 | 0.672 | 0.6829 | 0.692 | 0.7029 | 0.712 | 0.722 | 0.732 |
| 32 | 0.5401 | 0.5501 | 0.5601 | 0.5701 | 0.5801 | 0.5901 | 0.6001 | 0.6101 | 0.6201 | 0.631 | 0.640 | 0.6501 | 0.6601 | 0.6701 | 0.680 | 0.6901 | 0.700 | 0.7101 | 0.720 | 0.730 | 0. |
| 57 | 0.5532 | 0.5632 | 0.5732 | 0.5832 | 0.5932 | 0.6032 | 0.6132 | 0.6232 | 0.6332 | 0.6442 | 0.653 | 0.6632 | 0.6732 | 0.6832 | 0.6932 | 0.7032 | 0.713 | 0.7232 | 0.733 | 0.7432 | 0.75 |
| 82 | 0.5538 | 0.5638 | 0.5738 | 0.5838 | 0.5938 | 0.6038 | 0.6138 | 0.6238 | 0.6338 | 0.644 | 0.653 | 0.6638 | 0.673 | 0. 683 | 0.693 | 0.703 | 0.713 | 0.723 | 0.733 | 0.748 |  |
|  | 0.5854 | 0.595 | 0.605 | 0.615 | 0.625 | 0.635 | 0.6454 | 0.655 | 0.6654 | 0.676 | 0.685 | 0.695 | 0.705 | 0.715 | 0.725 | 0.735 | 0.745 | 0.755 | 0.765 | 0.775 | 0.7 |
|  | 74 | 0.6074 | 0.617 | 0.627 | 0.637 | 0.647 | 0.657 | 0.667 | 0.6774 | . | 0.697 | 0.707 | 0.71 | 0.72 | 0.73 | 0.7474 | 0.75 | 0.767 | 0.77 | 0.787 | 0. |
|  | 45 | 0.6245 | 0.6345 | 0.6 | 0.6545 | 0.664 | 0.6745 | 0.6845 | 0.6945 | 0.705 | 0.714 | 0.7245 | 0.734 | 0.744 | 0.75 | 0.764 | 0.774 | 0.784 | 0.79 | 0.804 | 0.814 |
| 98 | 304 | 0.6404 | 0.6504 | 0.6604 | 0.6704 | 0.6804 | 0.6904 | 0.7004 | 0.7104 | . 72 | 0.730 | 0.7404 | 0.750 | 0.7604 | 0.71 | 0.7804 | 0.7904 | 0.8004 | 0.810 | 0.820 |  |
|  |  | 0.6525 | 0.6625 | 0.6725 | 0.6825 | 0.6925 | 0.7025 | 0.7125 | 0.7225 | 0.733 | 0.7425 | 0.7525 | 0.7625 | 0.7725 | 0.7825 | 0.7925 | 0.8025 | 0.8125 | 0.822 |  |  |

District/COMFAll_40 All_41 All_42 All_43 All_44 All_45 All_46 All_47 All_48 Composite All_50 All_51 All_52 All_53 All_54 All_55 All_56 All_57 All_58 All_59 All_60


| 40 | All_41 | All_ | All_43 | All_44 | All_45 | All_46 | All_47 | All_48 | Al0410 | All_50 | All_51 | All_52 | All_5 | All_54 | All_55 | All_56 | All_57 | All_ 58 | All_59 | All_60 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.6404 | 0.6504 | 0.6604 | 0.6704 | 0.6804 | 0.6904 | 0.7004 | 0.7104 | 0.7204 | 0.7314 | 0.7404 | 0.7504 | 0.7604 | 0.7704 | 0.7804 | 0.7904 | 0.8004 | 0.8104 | 0.8204 | 0.8304 |  |
| 0.6189 | 0.6289 | 0.6389 | 0.64 | 0.6 | 0.6 |  |  | 0.6989 | 0.7099 | 0.7189 | 0.7 | 0.7 | 0.7 | 0.7 | 0.7689 | 0.7789 | 0.7889 | 0.7989 | 0.8089 |  |
| 0.6166 | 0.626 | 0.6 | 0.646 | 0.65 | 0.6 | 0.6766 | 0.6866 | 0.6966 | 0.70 | 0.7 | 0.7 | 0.7 | 0.7 | 07 | 0.7 | 0.7 | 0.7866 | 0.7966 | 0.8066 |  |
| 0.6075 | 0.61 | 0.62 | 0.6375 | 0.6 | 0.6 | 0.6675 | 0.6775 | 0.6875 |  | 0.707 | 0.717 | 0.7 | 0.7375 | 0.7475 | 0.7575 | 0.7675 | 0.7775 | 0.7875 | 0.7975 |  |
| 0.5914 | 0.6014 | 0.611 | 0.621 | 0.631 | 0.64 | 651 | 0.6614 | 0.67 | 0.6824 | 0.6914 | 0.7014 | 0.71 | 0.72 | 0.7 | 0.741 | . 751 | 761 | . 771 | 0.78 | 0.791 |
| 883 | 0.598 | 0.608 | 0.618 | 0.628 | 0.638 | 0.6483 | 0.658 | 0.6 | 0.679 | 0.688 | 0.698 | 0.70 | 0.718 | 0.728 | 0.738 | 0.748 | 0.758 | 0.768 | 0.778 | 0.788 |
| 553 | 0.5653 | 0.5753 | 0.5853 | 0.5953 | 0.6053 | 0.6153 | 0.6253 | 0.635 | 0.646 | 0.655 | 0.665 | 0.675 | 0.685 | 0.6953 | 0.705 | 0.715 | 0.725 | 0.735 | 0.745 | 0.755 |
| 0.5424 | 0.5524 | 0.5624 | 0.5724 | 0.5824 | 0.5924 | 0.6024 | 0.612 | 0.622 | 0.633 | 0.642 | 0.652 | 0.662 | 0.672 | 0.682 | 0.692 | 0.702 | 0.712 | 0.722 | 0.732 | 0.742 |
| 0.5408 |  | 0.5 | 0.5 | 0.5 | 0.59 | 600 | 0.610 | 0.620 | 0.63 | 0.6 | 0.65 | 0.6 | 0.67 | 0.68 | 0.65 | 0.70 | 0.71 | 0.72 | 0.7308 | 0.7408 |
| 0.5377 | 0.5 |  | 0.567 | 0.577 | 0.5877 | 05977 | 0.6077 |  | 0.628 | 0.63 | 0.64 | 0.65 | 0.66 | 0.67 | 0.68 | 0.69 | 0.70 | 0.71 | 0.7277 | 737 |
| 0.5365 | 0.5465 |  |  | 0.576 |  |  | 0.6065 |  | 0.627 | 0.636 | 0.646 | 0.65 | 0.666 | 0.676 | 0.6865 | 0.6965 | 0.7065 | 0.716 | 0.726 |  |
| 0.5208 | 0.5308 | 0.5408 |  | 0.5608 | 0.5708 |  | 0.5908 | 0.6008 | 0.6118 | 0.620 | 0.630 | 0.64 | 0.65 | 0.66 | 0.6708 | 0.680 | 0.690 | 0.700 | 0.710 |  |
| 0.5157 | 0.5257 | 0.5 | 0.545 | 0.5557 | 0.56 | 0.5757 | 0.5857 | 0.59 | 0.60 | 0.61 | 0.62 | 0.6 | 0.6457 | 0.65 | 0.665 | 0.675 | 0.685 | 0.695 | 0.705 | 0.715 |
| 0.5132 | 0.523 | 533 | 0.5432 | 0.5 | 0.563 | . 732 | 0.5832 | 0.5932 | 0.604 | 0.6132 | 0.6232 | 0.6332 | 0.6432 | 0.6532 | 0.6632 | 0.6732 | 0.683 | 0.693 | 0.7032 | 0.7 |
| 0.51 | 0.5 | 0.53 | 0.54 | 0.55 | 0.56 | 0.57 | 0.58 | 59 | 601 | 0.61 | 0.62 | 0.63 | 0.64 | 0.65 | 0.66 | 0.67 | . 68 | 0.69 | 0.7 | . 71 |
| 0.5006 | 0.5106 | 0.5206 | 0.5306 | 0.5406 | 0.5 | 0.5606 | 0.57 | 0.58 | 0.5 | 0.6 | 0.61 | 0.620 | 0.6306 | 0.6406 | 0.650 | 0.660 | 0.670 | 0.680 | 906 |  |
| 0.4982 | 0.5082 | 0.5182 | 0.5282 | 0.5382 | 0.548 | 0. 5582 | 0.5682 | 0.578 | 0.589 | 0.5982 | 0.608 | 0.6182 | 0.628 | 0.6382 | 0.6482 | 0.6582 | 0.6682 | 0.6782 | 0.6882 | 0.6982 |
| 0.4 | 0.506 | 0.516 | 0.526 |  |  |  |  |  | 0.58 |  | 0.606 | 0.616 | 0.626 | 0.636 | 0.646 | 0.656 | 0.666 | 0.676 | 0.686 |  |
| 0.4946 | 0.5046 | 0.5146 | 0.5246 | 0.5346 |  |  | 0.5646 | 0.5746 |  |  | 0.6046 | 0.6146 | 0.624 | 0.6346 | 0.6446 | 0.6546 | 0.6646 | 0.6746 | 0.6846 | 5946 |
| 0.4932 | 0.5032 |  |  |  |  |  |  |  |  |  |  |  | 0.623 |  | 0.6432 |  | 0.6632 | 0.673 | 0883 |  |
| 0.4887 | 0.4987 | 0.5087 | 0.5187 | 0.5287 | 0.5387 | 0.5487 | 0.5587 | 0.5687 | 0.579 | 0.588 | 0.5 | 0.6 | 0.618 | 0.6287 | 0.6387 | 0.6 | 2.65 | d.66 | 0.678 | 0.688 |
| 0.48 | 0.49 | 0.50 | 0.51 | 0.5 | 0.5 | 54 | 0.5555 | 0.5 | 0.5765 | 0.5 | 0.5 | 0.6 | 0.6 | 0.6 | 0.6 | 0.64 | 0.65 | 0.66 | 0.67 | 0.685 |
| 0.4842 | 0.49 | 0.5 | 0.51 | 0.5 | 0.53 | 0.554 | 0.5 | 0.5 | 0.57 | 0.58 | 0.55 | 0.60 | 0.61 | 0.62 | 0.63 | 0.64 | 0.65 | 0.66 | 0.6 | 0.6842 |
| 0.48 | 0.493 | 0.503 | 0.513 | 0.523 | 0.533 | 0.5 | 0.5 | 0.56 | 0.5 | 0.5 | 0.59 | 0.60 | 0.6 | 0.6 | 0.6 | 0.64 | 0.65 | 0.66 | 0.67 | 0.6831 |
| 0.4817 | 0.491 | 0.501 | 0.511 | 0.521 | 0.5317 | 54 | 0.55 | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | 0.61 | 0.62 | 0.63 | 0.64 | 0.65 | 0.661 | 0.67 | 0.6817 |
| 0.4807 | 0.490 | 0.500 | 0.5 | 0.5 | 0.5307 | 0.5407 | 0.55 | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | 0.6107 | 0.62 | 0.63 | 0.64 | 0.65 | 0.66 | 0.6 |  |
| 0.4759 | 0.4 | 0.4 | 0.5059 | 0 | 0.525 | 0.5359 |  |  |  |  |  |  | 0.605 |  |  | 0.63 | 0.6 | 0.65 | 0.665 |  |
| 0.4732 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | D. 64 |  |  |  |
| 0.4717 | 0. | 0.4917 | 0.5017 | 0.511 | 0.5217 | . 5317 | 0.5417 |  | 0.562 | 0.571 | 0.5 | 0.5 | 0.601 | 0.61 |  | 0.631 | 0.641 | 0.65 | 0.661 | 571 |
| 0.46 | 0.478 | 0.48 | 0.49 | 0.5 | 0.51 | 0.5284 | 0.5384 | 0.5484 | 0.5 | 0.5 | 0.5 | 0.58 | 0.59 | 0.60 | 0.6 | 0.62 | 0.63 | 0.64 | 0.65 | 0.6 |
| 0.46 | 0.47 | 0.4 | 0.4942 | 0.5042 | 0.5 | 0.5242 | 0.5342 | 0.5 | 0.5 | 0.5 | 0.57 | 0.58 | 0.5 | 0.6 | 0.6 | 0.62 | 0.63 | 0.64 | 0.65 | 0.66 |
| 0.46 | 0.473 | 0.483 | 0.493 | 0.503 | 0.513 | 52 | 0.53 | 0.5 | 0.5 | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | 0.6 | 0.6 | 0.6 | 0.64 | 0.6 | 0.663 |
| 0.4598 | 0.4698 | 0.47 | 0.48 | 0.4 | 0.50 | 0.5198 | 0.5298 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.6 | 0. | 0.62 | 0.63 | 0.6 |  |
| 85 | 0.46 | 0.4785 | 0.4885 | 0.4985 | 0.5085 | 0.5185 | 0.5285 | 0.5385 | 0.549 | 0.55 | 0.56 | 0.57 | 0.58 | 0.598 | 0.608 | 0.618 | 0.628 | 0.638 | 0.648 |  |
| 0.4521 | 0.4 | 0.472 | 0.482 | 0.492 | 0.5021 |  |  |  |  |  |  |  |  |  |  |  | 0.62 |  |  |  |
| 0.450 | 0. |  |  |  | 0.50 |  | 0.520 |  |  |  | 0.56 |  |  |  |  |  | 0.62 | 0.63 |  |  |
| 0.45 | 0.46 | 0.4 | 0.480 | 0.490 | 0.5009 | 510 | 0.520 | 0.530 | 0.541 | 0.55 | 0.56 | 0.5 | 0.5 | 0.5 | 0.60 |  | 0.62 | 0.63 | 0.64 | 0.65 |
| 0.450 | 0.460 | 0.470 | 0.4804 | 0.4904 | 0.5004 | 0.5104 | 0.5204 | 0.530 | 0.541 | 0.55 | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | 0.61 | 0.62 | 0.63 | 0.64 | 0.65 |
| 0.449 | 0.45 | 0.46 | 0.4796 | 0.4896 | 0.4996 | 0.5096 | 0.5196 | 0.5296 | 0.5406 | 0.549 | 0.5 | 0.5 | 0.57 | 0.58 | 0.59 | 0.60 | 0.61 | 0.62 | 0.63 | 0.6 |
| 0.44 | 0.45 | 0.468 | 0.47 | 0.48 | 0.49 | 0.5086 | 0.5186 | 0.52 | 0.5396 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.6 | 0.6 | 0.62 | 0.6 | 0.648 |
| 0.447 | 0.457 | 0.46 | 0.47 | 0.48 | 0.49 | 0.5074 | 0.5174 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.5 | 0.6 | 0.61 | 0.62 | 0.6 | 0.647 |
| 0.445 | 0.45 | 0.4 | 0.4758 | 0.485 | 0.4958 | 0.5058 | 0.5158 | 0.525 | 0.536 | 0.54 | 0.55 | - | 0.575 | 0.5 | 0.5 | 0.605 | 0.615 | 0.625 | 0.63 | 0.645 |
| 0.442 | 0.4 | 0.462 | 0.472 | 0.482 | 0.492 | . 502 | 0.512 | 0.522 |  |  |  |  |  |  |  |  | 0.61 |  |  |  |
| 0.4404 | 0.4 | 0.460 | 0.470 | 0.480 | 0.490 | 5004 | 0.510 | . 522 | . 53 |  |  | 0.56 |  |  |  |  | 0.61 | 0.62 | . 6.6 |  |
| 0.43 | 0.44 | 0.45 | 0.46 | 0.4 | 0.48 | 4975 | 0.5075 | 0.517 | 0.528 | 0.5 | 0.54 | 0.5 | 0.5 | 0.5 | 0.5 | 0.59 | 0.60 | 0.61 | 0.62 | 0.637 |
| 0.43 | 0.446 | 0.456 | 0.466 | 0.476 | 0.48 | 0.496 | 0.506 | . 51 | 0.52 | 0.53 | 0.54 | 0.55 | 0.566 | 0.57 | 0.58 | 0.59 | 0.606 | 0.61 | 0.62 | 0.6: |
| 0.43 | 0.446 | 0.456 | 0.466 | 0.476 | 0.48 | 0.496 | 0.50 | 0.51 | 0.52 | 0.53 | 0.54 | 0.55 | 0.56 | 0.57 | 0.58 | 0.59 | 0.60 | 0.61 | 0.62 | 06 |
| 0.4 | 0.44 | 0.454 | 0.46 | 0.47 | 0.48 | 0.494 | 0.5 | 0.51 | 0.52 | 0.53 | 0.5 |  | 0.5 | 0.57 |  |  |  | 0.61 | 0.62 |  |
| 0.4 | 0.44 | 0.454 | 0.4648 | 0.474 | 0.48 | 0.494 | 0.5 | 0.51 | 0.525 | 0.53 | 0.5 | 0.5 | 0.56 | -0.5 | 0.5 | 0.5 | 0.6 | 0.6 | 0.62 |  |
| 0.4338 | 0.4 | 0.4 | 0.4 | 0.4738 | 0.4838 | 0.4938 | 0.5038 | 0.5138 | 0.5248 | 0.5338 | 0.54 |  | 0.56 | 0.57 | 0.58 | 0.59 | 0.603 |  |  |  |
| D.43 | 0.4 | 0.4 | 0.4614 | 0.4714 | 0.4814 | . 4914 | 0.501 | 0.5114 | 0.52 | 0.531 | 0.54 |  | 0.56 |  | 0.58 | 0.59 | 0.601 | 0.611 | 0.621 | . 631 |
| 0.43 | 0.44 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.5 | . 51 | 0.52 | 0.53 | 0.5 |  |  |  |  |  | 0. | 0.6 | 0.6 |  |
| 0.4296 | 0.4396 | 0.4496 | 0.4596 | 0.4696 | 0.47 | 4896 | 0.4996 | 0.5096 | 0.5206 | 0.5 | 0.5 | 0.549 | 0.55 | 0.56 | 0.57 | 0.58 | 0.595 | 0.60 | 0.61 | . 6 |
| 0.4292 | 0.43 | 0.449 | 0.4592 | 0.46 | 0.479 | 0.4892 | 0.4992 | 0.509 | 0.5 | 0.52 | 0.53 | 0.549 | 0.55 | 0.56 | 0.57 | 0.589 | 0.599 | 0.60 | 0.619 | 0.62 |
| 0.4225 | 0.432 | 0.442 | 0.452 | 0.462 | 0.472 | 482 | . 492 | 0.50 | 0.51 | 0.52 | 0.53 | 0.5 | 0.55 | 0.56 | 0.57 | 0.58 | 0.592 | 0.602 | 0.61 | 0.6 |
| 0.4217 | 0.4317 | 0.441 | 0.451 | 0.461 | 0.4717 | 481 | 0.491 | 0.5017 | 0.51 | 0.521 | 0.53 | 0.54 | 0.51 | 0.56 |  | 0.58 | 0.591 | 0.6017 | 0.61 | 0.62 |
| 0.4212 | 0.4312 | 0.441 | 0.4512 | 0.4612 | 0.4712 | 0.481 | 0.4912 | 0.50 | 0.512 | 0.52 | 0.53 | 0.54 | 0.55 | 0.56 | 0.57 | 0.581 | 0.591 | 0.601 | 0.61 | 0.62 |
| 0.405 | 0.41 | 0.4 | 0.435 | 0.445 | 0.455 | 0.465 | 0.475 |  | 0.49 | 0.50 | 0.51 |  | 0.535 |  |  |  | 0.57 |  |  |  |
| 0. 4024 | 0.4 | 0.422 |  |  | 0.452 | 462 | 0.472 |  | 0.493 | 0.50 | 0.51 | 0.52 | 0.53 | 0.542 | 0.55 | 0.562 | D. 5 | 0.582 | . 592 |  |
| 0.399 | 0.409 | 0.419 |  |  |  | . 4596 |  |  |  |  |  |  |  | 0.53 | 0.549 |  |  |  |  |  |
| 0.3876 | 0.3 | 0.4076 | 0.417 | 0.427 | 0.437 | 447 | 0.457 | 0.46 | 0.478 | 0.487 | 0.49 |  | 0.51 | 0.52 | . | 0.54 | 0.55 | 0.5 | 0.57 | 0.58 |
| 0.3871 | 0.397 | 0.407 | 0.417 | 0.42 | 0.43 | 0.4471 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.50 | 0.51 | 0.52 | 0.53 | 0.5471 | 0.557 | 0.56 | 0.57 | 0.58 |
| 0.3784 | 0.388 | 0.398 | 0.408 | 0.418 | 0.428 | 0.43 | 0.4 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.50 | 0.51 | 0.52 | 0.5384 | 0.548 | 0.55 | 0.56 | 0.5 |
| 0.3782 | 0.3882 | 0.3982 | 0.408 | 0.4182 | 0.428 | 0.43 | 0.44 | 0.4582 | 0.469 | 0.47 | 0.48 | 0.49 | 0.5 | 0.51 | 0.52 | 0.53 | 0.548 | 0.5 | 0.56 | 0.57 |
| 0.3113 | 0.3813 | 0.3913 | 0.401 | 0.4113 | 0.421 | 0.4313 | -0.44 | 0.4513 | 0.462 | 0.471 | 0.48 | 0.49 | 0.50 | 0.51 | 0.52 | 0.53 | 0.541 | -.55 | 0.561 | 0.57 |
| 0.3658 | 0.3758 | 0.3858 | 0.3958 | 0.405 | 0.41 | 425 | 43 | 445 | 0.456 | 0.46 | 0.475 | 0.48 | 0.49 | 0.50 | 0.515 | 0.52 | 0.535 | 0.545 | 0.55 | 0.56 |
|  | 0.3 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0. 53 |  |  |  |
| 0.36 | 0.37 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | . 53 | 0.543 |  |  |
| 0.36 | 0.37 | 0.3822 | 0.39 | 0.4 | 0.41 | 0.422 | 0.432 | 0.442 | 0.453 | 0.46 | 0.47 | 0.48 | 0.4 | 0.50 | 0.5 | 0.52 | 0.532 | 0.5422 | 0.552 | 0.56 |
| 0.35 | 0.367 | 0.377 | 0.38 | 0.397 | 0.407 | 0.417 | 0.427 | 0.43 | 0.448 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.507 | 0.517 | 0.527 | 0.5379 | 0.5479 | 0.55 |
| 0.353 | 0.3636 | 0.3736 | 0.383 | 0.3936 | 0.4036 | 0.413 | 0.42 | 0.43 | 0.44 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.50 | 0.51 | 0.52 | 0.53 | 0.5436 | 0.55 |
| 0.34s | 0.359 | 0.369 | 0.379 | 0.3895 | 0.3995 | 0.409 | 0.419 | 0.429 | 0.440 | 0.4 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.50 | 0.51 | 0.52 | 0.5395 | 0.54 |
| 0.3495 | 0.3595 | 0.3695 | 0.3795 | 0.3895 | 0.3995 | 0.4095 | 0.4195 | 0.4295 | 0.440 | 0.4495 | 0.4595 | 0.4695 | 0.4795 | 0.489 | 0.499 | 0.509 | 0.5195 | 0.5295 | 0.5395 | 0.54 |
| 0.3361 | 0.346 | 0.3561 | 0.366 | 0.376 | 0.386 | 0.396 | 0.406 | 0.416 | 0.427 | 0.436 | 0.44 | 0.456 | 0.46 | 0.47 | 0.486 | 0.496 | 0.506 | 0.516 | 0.52 | 0.536 |
| 0.335 | 0.34 | 0.35 | 0.3 | -. | 0.38 | 0.395 | 0.405 | .at | 0.42 | 0.43 | . 4 | 0.455 | 0.46 | 0.4 | 0.485 | . 49 | 0.505 | . 51 | 0.525 |  |
| 0.315 | 0.325 | 0.3 | 0.345 | . 35 | . 3 |  | 0.385 | 0.395 | 0.406 | 0.415 |  |  |  | 0.45 | 0.465 | 0.475 | 0. 48 | 0. 495 | . 50 |  |
| 0.313 | 0.323 | 0.333 | 0.343 | 0.35 | 0.363 | 373 | 0.383 | 0.393 | 0. 40 | 0.41 | 0.42 | 0.43 | 0.44 | 0.45 | 0.46 | 0.473 | 0.4838 | 0.4938 | 0.50 | . 51 |
| 0.3114 | 0.3214 | 0.3314 | 0.341 | 0.3514 | 0.3614 | 0.3714 | 0.3814 | 0.3914 | 0.4024 | 0.411 | 0.421 | 0.431 | 0.441 | 0.4514 | 0.4614 | 0.4714 | 0.4814 | 0.4914 | 0.5014 | 0.51 |
| 0.3043 | 0.3143 | 0.3243 | 0.334 | 0.3443 | 0.3543 | 0.3643 | 0.3743 | 0.3843 | 0.395 | 0.404 | 0.414 | 0.424 | 0.4343 | 0.4443 | 0.454 | 0.4643 | 0.4743 | 0.4843 | 0.4943 | 0.50 |
| 0.2947 | 0.3047 | 0.3147 | 0.324 | 0.3347 | 0.3447 | 0.3547 | 0.3647 | 0.3747 | 0.3857 | 0.3947 | 0.4047 | 0.4147 | 0.4247 | 0.4347 | 0.444 | 0.4547 | 0.4647 | 0.4747 | 0.48 | 0.4 |
| 0.2915 | 0.3015 | 0.3115 | 0.3215 | 0.3315 | 0.3415 | 0.3515 | 0.3615 | 0.3715 | 0.3825 | 0.3915 | 0.4015 | 0.4115 | 0.4215 | 0.4315 | 0.4415 | 0.4515 | 0.4615 | 0.4715 | 0.4815 | 0.4915 |
| 0.2828 | 0.292 | 0.3028 | 0.312 | 0.322 | 0.332 | 0.342 | 0.352 | 0.362 | 0.373 | 0.3828 | 0.392 | 0.402 | 0.4128 | 0.4228 | 0.4328 | 0.4428 | 0.4528 | 0.4628 | 0.4728 | 0.48 |
| 0.2723 | 0.282 | 0.292 | 0.302 | 0.312 | 0.322 | 0.332 | 0.342 | 0.35 | 0.363 | 0.372 | 0.38 | 0.39 | 0.402 | 0.412 | 0.422 | 0.432 | 0.442 | 0.452 | 0.462 | 0.4 |
| 2676 | 0.2776 | 0.2876 | . 2.2976 | 376 | 0.3176 | 0.3276 | 0.3376 | . $34 / 16$ | 0.358 | . 3676 | 0.377 | -0.3876 | -0.3976 | . | . 4176 | . 4276 | . 4376 | 447 | . 457 | 0.467 |
| 0.26 | 0.27 | 0.28 | 0.29 | 0.3 | 0.31 | 0.32 | 0.33 | 0.34 | 0.351 | 0.36 | 0.37 | 0.38 | 0.39 | 0.4 | 0.41 | 0.42 | 0.43 | 0.44 | 0.45 | 0.46 |
| 0.2576 | 0.2676 | 0.2776 | 0.2876 | 0.2976 | 0.3076 | 0.3176 | 0.3276 | 0.3376 | 0.3486 | 0.3576 | 0.3676 | 0.3776 | 0.3876 | 0.3976 | 0.4076 | 0.4176 | 0.4276 | 0.4376 | 0.4476 | 0.4576 |
| 0.2237 | 0.2337 | 0.2437 | 0.2537 | 0.2637 | 0.2737 | 0.2837 | 0.2937 | 0.3037 | 0.314 | 0.323 | 0.333 | 0.343 | 0.3537 | 0.3637 | 0.373 | 0.3837 | 0.393 | 0.403 | 0.4137 | 0.423 |
| 0.2176 | 0.2276 | 0.2376 | 0.2476 | 0.2576 | 0.2676 | 0.2776 | 0.2876 | 0.2976 | 0.308 | 0.317 | 0.327 | 0.3376 | 0.3476 | 0.357 | 0.367 | 0.377 | 0.387 | 0.3976 | 0.4076 | 0.4176 |
| 0.1836 | 0.1936 | 0.2036 | 0.213 | 0.2236 | 0.2336 | 0.2436 | 0.2536 | 0.2636 | 0.2746 | 0.2836 | 0.2936 | 0.3036 | 0.3136 | 0.3236 | 0.3336 | 0.3436 | 0.3536 | 0.3636 | 0.3736 | 0.3836 |
| 0.1727 | 0.182 | 0.192 | 0.202 | 0.212 | 0.222 | 0.232 | 0.242 | 0.252 | 0.263 | 0.272 | 0.2827 | 0.29 | 0.30 | 0.31 | 0.322 | 0.3327 | 0.3427 | 0.3527 | 0.3627 | 0.372 |
| . 164 | 0.17 | 0. 18 | 0.19 | 0.20 | 0.21 | 0.22 | 0.23 | 0.24 | 0.235 | 0.26 | 0.2 | 0.28 | 0.29 | 0.37 | 0.31 | 0.32 | 0.33 | 0.34 | 0.35 | 0.364 |
| 1335 | 0.1435 | 0.1535 | 0.1635 | 0.1735 | 0.1835 | 0.1935 | 0.2035 | 0.2135 | 0.2245 | 0.233 | 0.243 | 0.253 | 0.263 | 0.273 | 0.2835 | 0.2935 | 0.3035 | 0.313 | 0.323 | 0.33 |
| 0.1204 | 0.130 | 0.140 | 0.150 | 0.1604 | 0.1704 | 0.1804 | 0.1904 | 0.2004 | 0.2114 | . 22204 | -0.2304- | 0.2404 | 0.2504 | 0.260 | 0.2704 | 0.2804 | 0.290 | 0.300 | 0.310 | 0.33 |
| 0.1193 | 0.1293 | 0.1393 | 0.1493 | 0.1593 | 0.1693 | 0.1793 | 0.1893 | 0.1993 | 0.2103 | 0.2193 | 0.2293 | 0.2393 | 0.2493 | 0.2593 | 0.2693 | 0.2793 | 0.2893 | 0.2993 | 0.3093 | 0.3193 |
| 0.1013 | 0.1113 | 0.1213 | 0.1313 | 0.1413 | 0.1513 | 0.1613 | 0.1713 | 0.1813 | 0.1923 | 0.2013 | 0.2113 | 0.2213 | 0.2313 | 0.2413 | 0.2513 | 0.2613 | 0.2713 | 0.2813 | 0.2913 | 0.3013 |
| 0.0764 | 0.0864 | 0.0964 | 0.1064 | 0.1164 | 0.1264 | 0.1364 | 0.1464 | 0.1564 | 0.1674 | 0.1764 | 0.1864 | 0.1964 | 0.2064 | 0.2164 | 0.2264 | 0.2364 | 0.2464 | 0.2564 | 0.2664 | 0.2764 |
| 0.0536 | 0.0636 | 0.0736 | 0.0836 | 0.0936 | 0.1036 | 0.1136 | 0.1236 | 0.1336 | 0.1446 | 0.1536 | 0.1636 | 0.1736 | 0.1836 | 0.1936 | 0.2036 | 0.2136 | 0.2236 | 0.2336 | 0.2436 | 0.2536 |
| 0.0324 0.0311 | 0.042 | 0.052 | 0.062 | 0.072 | 0.08 | 0.0 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1924 0.1911 | 0.2 | 0.2124 0.2111 | 0.2224 0.2211 | 0.2324 0.2311 |


| District | Observed | Index_40 | Index_41 | Index_42 | Index_43 | Index_44 | Index_45 | Index_46 | Index_47 | Index_48 | Index_49 | Index_50 | Index_51 | Index_52 | Index_53 | Index_54 | Index_55 | Index_56 | Index_57 | Index_58 | Index_59 | Index_60 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 16 | 0.1555 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 9 | 0.1844 | 0.0934 |  |  | 0.1234 | 0.1334 |  |  |  |  |  |  | 0.20 | 0.2134 | 0.22 | 0.23 | 0.2434 | 0.2534 | 0.26 | 734 |  | 34 |
| 8 | 0.2068 | 1158 | -0.1153 | 0.1253 | -0.1353 | -0.1453 | ${ }_{0}^{0.1658}$ | 0.1758 | 0.1858 | ${ }^{0}$ | 0.195 | 0.2158 | 0.2258 | 0.2358 | 0.245 | 0.25 | 0.2658 | 0.2758 | . 2235 | 0.2853 | -. 0.358 |  |
| 76 | 0.2098 | 0.1188 | 0.1288 | 0.1388 | 0.1488 | 0.1588 | 0.1688 | 0.1788 | 0.1888 | 0.1988 | 0.208 | 0.2188 | 0.2288 | 0.2388 | 0.2488 | ${ }^{0.2588}$ | 0.2688 | 0.2788 | 0.2888 | 0.2988 | 0.3088 | 0.31 |
| 17 | 0.2183 | 0.1273 | 0.1373 | 0.1473 | 0.1573 | 0.1673 | 0.1773 | 0.1873 | 0.1973 | 0.2073 | 0.2173 | 0.2273 | 0.2373 | 0.2473 | 0.2573 | 0.2673 | 0.2773 | 0.2873 | 0.2973 | 0.3073 | 0.3173 | 0.3273 |
| 75 | 0.2 | 0.1662 | 0.1762 | 862 | 0.1962 | 203 | 0.2162 | 0.2262 | 0.2362 | 0.2462 | 0.2562 | 0.2662 | 0.2762 | 0.2862 | 0.2962 | 0.3062 | 0.3162 | 0.3262 | 0.3362 | 0.3462 | 0.3562 | 0.3662 |
| 77 |  | 0 | 0.2174 | 0.2274 | 0.2374 | 2474 | 0.2574 | 0.2674 | 0.2774 | 0.287 | 0.297 | 0.3074 | 0.3174 | 0.3274 | 0.3374 | 0.3474 | 0.3574 | 0.3674 | 0.3774 | 0.3874 | 0.3974 |  |
| 18 | ${ }^{0.2995}$ | ${ }^{0.2085}$ | ${ }^{0.2185}$ | 0.2285 | 0.2385 | 0.2485 | 0.2585 | 0.2685 | 0.2785 | 0.2885 | 0.2985 | 0.3085 | ${ }^{0.3185}$ | ${ }^{0.3285}$ | 0.3385 | 0.3485 | 0.3585 | ${ }^{0.3685}$ | 0.3785 | 0.3885 | -. 3985 |  |
| 15 | ${ }^{0.3141}$ | 0.2231 | ${ }^{0.2331}$ | 0.2431 | ${ }^{0.2531}$ | ${ }^{0.2631}$ | 0.2731 | ${ }^{0.2831}$ | ${ }^{0.2931}$ | 0.3031 | 0.3131 | ${ }^{0.3231}$ | ${ }^{0.3331}$ | ${ }^{0.3431}$ | -.3531 | ${ }^{0.3631}$ | -. 3731 | ${ }^{0.3831}$ | ${ }^{0.3931}$ | 0.4031 | 0.4131 | 0.4231 |
| 74 | 0.3368 | 0.2458 | 0.2558 | 0.2658 | 0.2758 | 0.2858 | 0.2958 | 0.3058 | 0.3158 | 0.3258 | 0.3358 | 0.3458 | 0.3558 | 0.3658 | 0.3758 | 0.3858 | 0.3958 | 0.4058 | 0.4158 | 0.4258 | 0.4358 | 0.4458 |
| 46 | 0.347 | 0.25 | 0.2664 | 0.2764 | ${ }^{0.2864}$ | 0.2964 | 0.3064 | 0.3164 | 0.3264 |  | 0.346 | 0.3564 | 0.3664 | 0.3764 | 0.3864 | 0.3964 | 0.4064 | 0.4464 | 0.4264 | 0.4364 | 0.4464 | 0.4564 |
| 47 | 0.3562 | 0.2652 | 0.2752 | 0.2852 | 0.2952 | 0.3052 | 0.3152 | 0.3252 | 0.3352 | 0.3452 | 0.3552 | 0.3652 | 0.3752 | 0.3852 | 0.3952 | 0.4052 | 0.4152 | 0.4252 | 0.4352 | 0.4452 | 0.4552 | 0.4652 |
| 64 | 0.3697 | 0.2787 | 0.2887 | 0.2987 0.3015 | 0.3087 0.315 | ${ }^{0.3187}$ | 0.3287 0.3315 | 0.3387 0.3415 | 0.3487 0.3515 | 0.3587 0.3615 | 0.3687 0.3715 | 0.3787 0.3815 | 0.3887 0.3915 | 0.3987 0.4015 | 0.4087 0.4115 | 0.4187 0.4215 | 0.4287 0.4315 | 0.4387 0.4415 | 487 | 0.4587 0.4615 | ${ }^{0.4687}$ | 0.4787 0.4815 |
| 78 | 0.3725 | 0.2815 0.2902 | ${ }^{0.2915}$ | ${ }_{0}^{0.3102}$ | ${ }^{0.3202}$ | 0.3302 | 0.3402 | 0.3502 | 0.3602 | 0.3702 | 0.3802 | 0.3902 | 0.4002 | 0.4102 | 0.4202 | 0.4302 | 0.4402 | 0.4502 | 0.4602 | 0.4702 | 0.4802 | 0.4002 |
| 45 | 0.3812 0.3852 | 0.2942 | ${ }^{0.3042}$ | 0.3142 | ${ }_{0}^{0.3242}$ | 0.3342 | 0.3442 | 0.3542 | 0.3642 | 0.3742 | 0.3842 | 0.3942 | 0.4042 | 0.4142 | 0. 02422 | 0.4342 | 0.4442 | 0.4542 | 0.4642 | 0.4742 | 0.4842 | 0.4942 |
| 11 | 0.38 |  |  | 0.3185 | 0.3285 | 0.3385 | 0.3485 | 0.3585 | 0.36 | 0.3 |  | 0.3985 | 0.4085 | 0.4185 | 0.4285 | 0.4385 | 0.4485 | 0.4585 | 0.4685 | 0.4785 | 0.4885 | . 4985 |
| 43 | 0.3937 | 0.3027 | 0.3127 | 0.3227 | 0.3327 | 0.3427 | 0.3527 | 0.3627 | 0.3727 | 0.3827 | ${ }_{0} 0.3927$ | 0.4027 | 0.4127 | 0.4227 | 0.4327 | 0.4427 | 0.4527 | 0.4627 | 0.4727 | 0.4827 | 0.4927 | 0.5027 |
| 44 | 0.4047 | 0.3137 | 0.3237 | 0.3337 | 0.3437 | 0.3537 | 0.3637 | 0.3737 | 0.3837 | 0.3937 | 0.4037 | 0.4137 | 0.4237 | 0.4337 | 0.4437 | 0.4537 | 0.4637 | 0.4737 | 0.4837 | 0.4937 | 0.5037 | 0.5137 |
| 19 | 0.4053 | 0.3143 | 0.3243 | 0.3343 | 0.3443 | 0.3543 | 0.3643 | 0.3743 | 0.3843 | 0.3943 | 0.4043 | 0.4143 | 0.4243 | 0.4343 | 0.4443 | 0.4543 | 0.4643 | 0.4743 | 0.4843 | 0.4943 | 0.5043 | 0.5143 |
| 65 | 0.4063 | 0.3153 | ${ }^{0.3253}$ | 0.3353 | 0.3453 | 0.3553 | 0.3653 | 0.3753 | 0.3853 | 0.3953 | 0.4 | 0.4153 | 0.4253 | 0.4353 | 0.4453 | 0.4553 | 0.4653 | 0.4753 | 0.4853 | 0.4953 | 0.5053 | 0.5153 |
| 90 | 0.4138 | 0.3228 | 0.3328 | 0.3428 | 0.3528 | 0.3628 | 0.3728 | 0.3828 | 0.3928 | 0.4028 | 0.4128 | 0.4228 | 0.4328 | 0.4428 | 0.4528 | 0.4628 | 0.4728 | 0.4828 | 0.4928 | 0.5028 | 0.5128 | 0.5228 |
| 79 | ${ }^{0.4156}$ | ${ }^{0.3246}$ | 0.3346 03348 | ${ }^{0.3446}$ | 0.3546 | 0.3646 <br> 0.3648 | 0.3746 0.3748 | 0.3846 <br> 0.3848 <br>  | 0.3946 0.3948 0 | $\begin{array}{r}0.4046 \\ 0.4048 \\ \hline\end{array}$ | 0.4146 <br> 0.4148 <br> 0 | 0.4246 0.4248 | 0.4346 0.4348 0.4 | 0.4446 0.4448 0 | 0.4546 <br> 0.4548 <br> 0 | 0.4646 0.4648 0.4 | 0.4746 0.4748 | 0.4846 0.4848 | 0.4946 0.4948 | 0.5046 0.5048 0 | 0.5146 0.5148 | 0.5246 .5248 |
| 78 | 0.4158 0.4412 | 0.3248 0.3502 | ${ }_{0}^{0.3602}$ | ${ }_{0}^{0.34702}$ | 0.3802 | 0.3902 | 0.4002 | 0.4102 | 0.4202 | 0.4302 | 0.4402 | 0.4502 | 0.4602 | 0.4702 | 0.4802 | 0.4902 | 0.5002 | 0.5102 | 0.5202 | 0.5302 | 0.5402 |  |
| 94 | 0.4463 | 0.3553 | ${ }_{0}^{0.3653}$ | 0.3753 | 0.3853 | 0.3953 | 0.4053 | 0.4153 | 0.4253 | 0.4353 | 0.4453 | 0.4553 | 0.4653 | 0.4753 | 0.4853 | 0.4953 | 0.5053 | 0.5153 | 0.5253 | 0.5353 | 0.5453 |  |
| 72 | 0.448 | 0.3573 | 0.3673 | 0.3773 | 0.3873 | 0.3973 | 0.4073 | 0.4173 | 0.4273 | 0.437 | 0.447 | 0.4573 | 0.4673 | 0.4773 | 0.4873 | 0.4973 | 0.5073 | 0.5173 | 0.5273 | 0.5373 | 0.5473 |  |
| 73 | 0.4488 | 0.3578 | 0.3678 | 0.3778 | 0.3878 | 0.3978 | 0.4078 | 0.4178 | 0.4278 | 0.4378 | 0.4478 | 0.4578 | 0.4678 | 0.4778 | 0.4878 | 0.4978 | 0.5078 | 0.5178 | 0.5278 | 0.5378 | 0.5478 |  |
| 91 | 0.4494 | 0.3584 | 0.3684 | 0.3784 | 0.3884 | 0.3984 | 0.4084 | 0.4184 | 0.4284 | 0.4384 | 0.4488 | 0.4584 | 0.4684 | 0.4784 | 0.4884 | 0.4984 | 0.5084 | 0.5184 | 0.5284 | 0.5384 | 0.5484 |  |
| 66 | 0.4503 | 0.3593 | 0.3693 | 0.3793 | 0.3893 | 0.3993 | 0.4093 | 0.4193 | 0.4293 | 0.4393 | 0.4493 | 0.4593 | 0.4693 | 0.4793 | 0.4893 | 0.4993 | 0.5093 | 0.5193 | 0.5293 | 0.5393 | 0.5493 |  |
| 42 | 0.4504 | 359 | ${ }^{0.3694}$ | 0.3794 | 0.3894 | 0.3994 | 0.4094 | 0.4194 | 0.4294 | 0.4394 | 0.4 | 0.4594 | 0.4694 | 0.4794 | 0.4894 | 0.4994 | 0.5094 | 0.5194 | 0.5294 | 0.5394 | 0.5494 |  |
| 89 | 0.4543 | 0.3633 | 0.3733 | 0.3833 | 0.3933 | 0.4033 | 0.4133 | 0.4233 | 0.4333 | 0.4433 | 0.4533 | 0.4633 | 0.4733 | 0.4833 | 0.4933 | 0.5033 | 0.5133 | 0.5233 | 0.5333 | 0.5433 |  |  |
| 93 | 0.4624 | 0.3714 | 0.3814 | 0.3914 | 0.4014 | 0.4114 | 0.4214 | 0.4314 | 0.4414 | 0.4514 | 0.4614 | 0.4714 | 0.4814 | 0.4914 | 0.5014 | 0.5114 | 0.5214 | 0.5314 | 0.5414 |  | 0.5614 |  |
| 14 | ${ }^{0.4649}$ | ${ }^{0.3739}$ | 0.3839 | 0.3939 | ${ }^{0.4039}{ }^{-141}$ | 0.4139 | ${ }^{0.4239}$ | 0.4339 | 0.4439 | ${ }^{0.4539}$ | 0.4439 | 0.4739 | 0.4839 | 0.4539 | ${ }^{0.5039}$ | 0.5139 | 0.5239 | ${ }^{0.5339}$ | 0.5439 | 0.5539 | 0.5639 |  |
| 63 | ${ }^{0.4721}$ | 0.3811 | 0.3911 | 0.4011 | 0.4111 | 0.4211 | 0.4311 | 0.4411 | 0.4511 | 0.4611 | 0.4711 | 0.4811 | 0.4911 | 0.5011 | 0.5111 | 0.5211 | 0.5311 | 0.5411 |  | 0.5611 | 0.5711 | n 5 |
| 80 | ${ }^{0.4737}$ | 0.3827 | 0.3927 | 0.4027 | 0.4127 | 0.4227 | 0.4327 | 0.4427 | 0.4527 | 0.4627 | 0.4727 | 0.4827 | 0.4927 | 0.5027 | 0.5127 | 0.5227 | 0.5327 | 0.5427 | 0.5527 | 0.5627 | 0.5727 |  |
| 70 | 0.4751 | 0.3841 | ${ }^{0.3941}$ | 0.4041 | 0.4141 | 0.4241 | 0.4341 | 0.4441 | 0.4541 | 0.4641 | 0.4741 | 0.4841 | 0.4941 | 0.5041 | 0.5141 | 0.5241 | 0.5341 | 0.5441 | 0.5541 | 0. 5641 | 0.5741 |  |
| 6 | ${ }^{0.4788}$ | 0.3878 | 0.3978 | 0.4078 | 0.4178 | 0.4278 | 0.4378 | 0.4478 | 0.4578 | 0.4678 | 0.4778 | 0.4878 | 0.4978 | 0.5078 | ${ }^{0.5178}$ | 0.5278 | 0.5378 | 0.5478 <br> 0509 | 0.5578 | -0.5788 | 0.5778 |  |
| ${ }_{95}^{69}$ | 0.4819 0.4854 | 0.3909 0.3944 | 0.4009 0.4044 | 0.4109 0.4144 | 0.4209 0.4244 | 0.4309 0.4344 | 0.4409 0.4449 | 0.4509 0.4544 | 0.4609 0.4644 | 0.4709 0.4744 | 0.4889 0.4844 | 0.4999 0.4944 | 0.5009 0.5044 | 0.5109 0.5144 | 0.5209 0.5244 | ${ }^{0.5309}$ | 0.5409 0.5444 | 0.5544 | 0.5609 0.5644 | 0.5709 0.5744 | 0.5809 0.5844 |  |
| 49 | 0.4923 | 0.4013 | 0.4113 | 0.4213 | 0.4313 | 0.4413 | 0.4513 | 0.4613 | 0.4713 | 0.4813 | 0.4913 | 0.5013 | 0.5113 | 0.5213 | 0.5313 | 0.541 |  | 0.5613 | 0.5713 | 0.5813 | 0.5913 |  |
| 92 | 0.4926 | 0.4016 | 0.4116 | 0.4216 | 0.4316 | 0.4416 | 0.4516 | 0.4616 | 0.4716 | 0.4816 | 0.4916 | 0.5016 | 0.5116 | 0.5216 | 0.5316 | 0.5416 | 0.5516 | 0.5616 | 0.5716 | ${ }_{0} .5816$ | 0.5916 |  |
| 88 | 0.5055 | 0.4145 | 0.4245 | 0.4345 | 0.4445 | 0.4545 | 0.4645 | 0.4745 | 0.4845 | 0.4945 | 0.5045 | 0.5145 | 0.5245 | 0.5345 | 0.5445 |  | 0.5645 | 0.5745 | 0.5845 | 0.5945 | 0.6045 | 0.6 |
| 41 | 0.5056 | 0.4146 | 0.4246 | 0.4346 | 0.4446 | 0.4546 | 0.4646 | 0.4746 | 0.4846 | 0.4946 | 0.5046 | 0.5146 | 0.5246 | 0.5346 | 0.5446 | 0.55 | 0.5646 | 0.5746 | 0.5846 | 0.5946 | 0.6046 |  |
| 12 | 0.5122 | 0.4212 | 0.4312 | 0.4412 | 0.4512 | 0.4612 | 0.4712 | 0.4812 | 0.4912 | 0.5012 | 0.5112 | 0.5212 | 0.5312 | 0.5412 |  | 0.5612 | 0.5712 | 0.5812 | 0.5912 | 0.6012 | 0.6112 |  |
| 68 | 0.5143 | 0.4233 | 0.4333 | 0.4433 | 0.4533 | 0.4633 | 0.4733 | 0.4833 | 0.4933 | 0.5033 | 0.5133 | 0.5233 | 0.5333 | 0.5433 | 0.5533 | 0.5633 | 0.5733 | 0.5833 | 0.5933 | 0.6033 | 0.6133 | 0.6 |
| 20 | 0.5162 | 0.4252 | 0.4352 | 0.4452 | 0.4552 | 0.4652 | 0.4752 | 0.4852 | 0.4952 | 0.5052 | 0.5152 | 0.5252 | 0.5352 | 0.5452 | 0.5552 | 0.56 | 0.5752 | 0.5852 | 0.5952 | 0.6052 | 0.6152 |  |
| $\stackrel{67}{ }$ | 0.5175 | 0.4265 |  |  | 0.4565 | 0.4665 | ${ }_{0}^{0.4765}$ | 0.4865 | 0.4965 | 0.5065 | 0.5165 | 0.5265 | 0.5365 | 0.5465 0.5537 |  |  | 0.5837 |  |  |  |  |  |
| 85 50 | 0.5247 0.5253 | $\begin{aligned} & 0.4337 \\ & 0.4343 \end{aligned}$ | $\begin{array}{\|l\|l} 0.4437 \\ 0.4443 \end{array}$ | 0.4537 0.4543 | 0.4637 0.4643 | 0.4737 0.4743 | 0.4837 0.4843 | 0.4937 0.4943 | 0.5037 0.5043 | 0.5137 0.5143 | 0.5237 0.5243 | 0.5337 0.5343 | ${ }_{0}^{0.5443}$ | 0.5543 | ${ }_{0}^{0.5637}$ | 0.5737 0.5743 | 0.5837 0.5843 | ${ }_{0}^{0.5937}$ | ${ }^{0.6037}$ | ${ }_{0}^{0.6137}$ | ${ }^{0.6237}$ |  |
| 5 | ${ }_{0.5256}$ | ${ }_{0}^{0.4346}$ | ${ }^{0.44436}$ | ${ }_{0}^{0.4546}$ | - 0.4646 | ${ }_{0.4746}^{0.473}$ | ${ }_{0.4846}$ | ${ }_{0.4946}$ | ${ }_{0}$. | ${ }_{0.5146}$ | - 0.5246 | ${ }_{0.5346}$ | ${ }_{0.5446}$ | ${ }^{0.5543}$ | ${ }_{0}^{0.5646}$ | ${ }^{0.5746}$ | ${ }_{0}$. | . 5.5943 | . 0.6043 | ${ }_{0}^{0.6146}$ | ${ }_{0.06246}^{0.0245}$ |  |
| 55 | 0.5302 | 0.4392 | 0.4492 | 0.4592 | 0.4692 | 0.4792 | 0.4892 | 0.4992 | 0.5092 | 0.5192 | 0.5292 | 0.5392 | 0.5492 | 0.5592 | 0.5692 | 0.5792 | 0.5892 | 0.5992 | 0.6092 | 0.6192 | 0.6292 |  |
| 54 | 0.531 | 0.44 | 0.45 | 0.46 | 0.47 | 0.48 | 0.49 | 0.5 | 0.51 | 0.52 | 0.53 | 0.54 |  | 0.56 | 0.57 | 0.58 | 0.59 | 0.6 | 0.61 | 0.62 | 0.63 |  |
| 28 | 0.5332 | 0.4422 | 0.4522 | 0.4622 | 0.4722 | 0.4822 | 0.4922 | 0.5022 | 0.5122 | 0.5222 | 0.5322 | 0.5422 | 0.5522 | 0.5622 | 0.5722 | 0.5822 | 0.5922 | 0.6022 | 0.6122 |  | 0.6322 |  |
| 62 | 0.5372 | 0.4462 | 0.4562 | 0.4662 | 0.4762 | 0.4862 | 0.4962 | 0.5062 | 0.5162 | 0.5262 | 0.5362 | 0.5462 | 0.5562 | 0.5662 | 0.5762 | 0.5862 | 0.5962 | 0.6062 | 0.6162 | 0.6262 | 0.6362 |  |
| 1 | 0.5373 | 0.4463 | 0.4563 | 0.4663 | 0.4763 | 0.4863 | 0.4963 | 0.5063 | 0.5163 | 0.5263 | 0.5363 | 0.5463 | 0.5563 | 0.5663 | 0.5763 | 0.5863 | 0.5963 | 0.6063 | 0.6163 | 0.6263 | 0.6363 |  |
| 29 | 0.5382 | 0.4472 | 0.4572 | 0.4672 | 0.4772 | 0.4872 | 0.4972 | 0.5072 | 0.5172 | 0.5272 | 0.5372 | 0.5472 | 0.5572 | 0.5672 | 0.5772 | 0.5872 | 0.5972 | 0.6072 | 0.6172 | 0.6272 | 0.6372 |  |
| 84 | 0.5384 | 0.4474 | 0.4574 | 0.4674 | 0.4774 | 0.4874 | 0.4974 | 0.5074 | 0.5174 | 0.5274 | 0.5374 | 0.5474 | 0.5574 | 0.5674 | 0.5774 | 0.5874 | 0.5974 | 0.6074 | 0.6174 | 0.6274 | 0.6374 |  |
| 86 | 0.5402 | 0.4492 | 0.4592 | 0.4692 | 0.4792 | 0.4892 | ${ }^{0.4992}$ | 0.5092 | 0.5192 | 0.5292 | 0.5392 | 0.5492 | 0.5592 | ${ }^{0.5692}$ | 0.5792 | 0.5892 | 0.5992 | 0.6092 | 0.6192 | 0.6292 | ${ }^{0.6392}$ |  |
| 87 53 | 0.5422 0.5424 | $\begin{aligned} & 0.4512 \\ & 0.4514 \end{aligned}$ | 0.4612 0.4614 | 0.4712 0.4714 | 0.4812 0.4814 | 0.4912 0.4914 | 0.5012 0.5014 | 0.5112 0.5114 | 0.5212 0.5214 | 0.5312 0.5314 | 0.5412 0.5414 | 0.5512 0.5514 | ${ }_{0}^{0.5612}$ | ${ }^{0.5712}$ | ${ }^{0.5812}$ | 0.5912 0.5914 | 0.6012 0.6014 | ${ }_{0}^{0.6112}$ | 0.6212 0.6214 | ${ }^{0.6312} 0$ | ${ }_{0}^{0.6412}$ |  |
| 27 | 0.5426 | 0.4516 | 0.4616 | 0.4716 | 0.4816 | 0.4916 | 0.5016 | 0.5116 | 0.5216 | 0.5316 | 0.5416 | ${ }_{0} .5516$ | 0.5616 | 0.5716 | 0.5816 | 0.5916 | 0.6016 | 0.6116 | 0.6216 | 0.6316 | 0.6416 |  |
| 34 | 0.5439 | 0.4529 | 0.4629 | 0.4729 | 0.4829 | 0.4929 | 0.5029 | 0.5129 | 0.5229 | 0.5329 | 0.5429 | 0.5529 | 0.5629 | 0.5729 | 0.5829 | 0.5929 | 0.6029 | 0.6129 | 0.6229 | 0.6329 | 0.6429 | 0.6 |
| 3 | 0.544 | 0.453 | 0.463 | 0.473 | 0.483 | 0.493 | 0.503 | 0.513 | 0.523 | 0.533 | 0.543 |  |  | 0.573 <br> 0.5748 |  | 0.5949 | 0.603 | ${ }^{0.613}$ | 0.623 | 0.633 | 0.643 |  |
| 52 | 0.5458 0.5501 | 0.4548 0.4591 | 0.4648 0.4691 | 0.4748 0.4791 | 0.4848 0.4891 | 0.4948 0.4991 | 0.5048 0.5091 | 0.5148 0.5191 | 0.5248 0.5291 | 0.5348 0.5391 | $\begin{aligned} & 0.5488 \\ & 0.5491 \end{aligned}$ | $\begin{aligned} & 0.5548 \\ & 0.5591 \end{aligned}$ | $\begin{aligned} & 0.5648 \\ & 0.5691 \end{aligned}$ | $\begin{aligned} & 0.5748 \\ & 0.5791 \end{aligned}$ | $0.5848$ | $\begin{aligned} & 0.5948 \\ & 0.5991 \end{aligned}$ | 0.6048 0.6091 | $0.6148$ | 0.6248 0.6291 | 0.6348 0.6391 | $0.6448$ $0.6491$ |  |
| 52 25 | ${ }_{0}^{0.5519}$ | ${ }_{0} .4609$ | ${ }_{0}^{0.4709}$ | 0.4809 | 0.4909 | 0.5009 | 0.5109 | 0.5209 | 0.5309 | 0.5409 |  | 0.5609 | 0.5709 | 0.5809 | 0.5909 | 0.6009 | 0.6109 | 0.6209 | 0.6309 | ${ }_{0.6409}$ | 0.6509 |  |
| 4 | 0.5523 | 0.4613 | 0.4713 | 0.4813 | 0.4913 | 0.5013 | 0.5113 | 0.5213 | 0.5313 | 0.5413 | 0.5513 | 0.5613 | 0.5713 | 0.5813 | 0.5913 | 0.6013 | 0.6113 | 0.6213 | 0.6313 | 0.6413 | 0.6513 |  |
| 35 | ${ }^{0.5528}$ | 0.4618 | 0.4718 | ${ }^{0.4818}$ | 0.4918 | 0.5018 | 0.5118 | 0.5218 | 0.5318 | 0.5418 | 0.551 | 0.5618 0.5567 | 0. 5718 | ${ }^{0.5518}$ | -. 5918 | 0.6018 | 0.6118 | ${ }^{0.6218}$ | 0.6318 | 0.6418 | 0.6518 |  |
| ${ }_{51}^{51}$ | 0.5572 | 0.4662 | 0.4776 | 0.4862 | 0.4962 | 0.5062 | 0.5162 | 0.5262 | 0.5362 | 0.5462 | 0.556 | 0.5662 | ${ }^{0.5762}$ | 0.5862 | 0. 5962 | 0.600 | 0.6162 | 0.6262 | 0. 6362 | 0.6462 | 0.6562 |  |
| 26 40 | 0.5576 0.5602 | 0.4666 0.4692 | 0.4766 0.4792 | 0.4866 0.4892 | 0.4966 0.4992 | 0.5066 0.5092 | 0.5166 0.5192 | 0.5266 0.5292 | 0.5366 0.5392 | 0.5466 0.5492 | ${ }_{0}^{0.556}$ | 0.5666 0.5692 | 0.5766 0.5792 | 0.5866 0.5892 | ${ }^{0.5966}$ | 0.6066 0.6092 | $0.6166^{\prime}$ 0.6192 | ${ }^{0.6266}$ | ${ }_{0}^{0.6366}$ | 0.6466 0.6492 | 0.6566 |  |
| 40 | 0.5602 | 0.4692 | 0.4792 | 0.4892 | 0.4992 | 0.5092 | 0.5192 | 0.5292 | ${ }^{0.5392}$ | 0.5492 | 0.559 | 0.5692 | 0.5792 | 0.5892 | 0.5992 |  | 0.6192 | 0.6292 |  |  |  |  |
| ${ }_{33}^{21}$ | ${ }_{0}^{0.5663}$ | 0.4753 0.4754 | 0.4853 0.4854 | ${ }^{0.4953}$ | 0.5053 0.5054 | ${ }_{0}^{0.5153}$ | ${ }_{0}^{0.5253}$ | 0.5353 0.5354 | ${ }_{0}^{0.5453}$ | 0.55 | 0.565 | 0.5753 0.5754 | . 58854 | ${ }^{0.5953}$ | 0.6053 | ${ }^{0.6153}$ | ${ }_{0}^{0.625}$ | 0.635 | 0.645 | 558 | 665 |  |
| 30 | ${ }_{0}^{0.5731}$ | 0.4821 | 0.4921 | 0.5021 | 0.5121 | 0.5221 | 0.5321 | 0.5421 |  | 0.562 | 0.572 | 0.5821 | ${ }_{0.5921}$ | 0.6021 | ${ }_{0.6121}^{0.654}$ | ${ }_{0}^{0.6621}$ | ${ }_{0}^{0.6321}$ | ${ }_{0}^{0.6421}$ | ${ }^{0.65521}$ |  |  |  |
| 81 | 0.5752 | 0.4842 | 0.4942 | 0.5042 | 0.5142 | 0.5242 | 0.5342 | 0.5442 | 0.5542 | 0.564 | 0.5742 | 0.5842 | 0.5942 | 0.5042 | 0.6142 | 0.6242 | 0.6342 | 0.6442 | 0.6542 | 0.6642 | 0.6742 |  |
| 13 | 0.5772 | 0.4862 | 0.4962 | 0.5062 | 0.5162 | 0.5262 | 0.5362 | 0.5462 | 0.5562 | 0.566: | 0.576 | 0.5862 | 0.5962 | 0.6062 | 0.6162 | 0.6262 | 0.6362 | 0.6462 | 0.6562 | 0.6662 | 0.6762 |  |
| 61 | 0.5777 | 0.4867 | 0.4967 | 0.5067 | 0.5167 | 0.5267 | 0.5367 | 0.5467 | 0.5567 | 0. 566 | 0.576 | 0.5867 | 0.5967 | 0.6067 | 0.6167 | 0.6267 | 0.6367 | 0.6467 | 0.6567 | 0.6667 | 0.6767 |  |
| 36 56 | ${ }^{0.5791}$ | 0.4881 |  | 0.5081 | 0.5181 | ${ }^{0.5281}$ | ${ }^{0.5381}$ | 0.5481 0.5487 | 0.5581 | 0.5683 | 0.5781 | 0.5881 | 0.5981 | 0.6081 | ${ }^{0.6181}$ | ${ }^{0.6281}$ | 0.6381 | ${ }^{0.6481}$ | 0.6581 | ${ }^{0.6681}$ | 0.6781 |  |
| 56 39 | 0.5797 0.5839 | 0.4887 0.4929 | 0.4987 0.5029 | 0.5087 0.5129 | 0.5187 0.5229 | 0.5287 0.5329 | 0.5387 0.5429 | ${ }^{0.5487}{ }^{0.5529}$ | 0.5587 0.5629 | 0.5688 0.5725 | ${ }^{0.5787}$ | 0.5887 0.5929 | 0.5987 0.6029 | ${ }^{0.6087}$ | ${ }_{0}^{0.6187}$ | ${ }^{0.65287}$ | 0.6387 0.6429 | ${ }^{0.6487}$ | ${ }^{0} 0.6587$ | 0.6687 0.6729 | 0.6787 0.6829 |  |
| 39 83 | 0.5839 0.587 | 0.4929 | 0.506 | ${ }_{0}^{0.516}$ | 0.5229 0.526 | 0.5336 | ${ }_{0}^{0.546}$ |  |  |  |  |  | 0.6029 | 1.0127 | -. 0.629 |  | 0.642 |  |  |  | 0.629 |  |
| 83 24 | 0.5924 | 0.5014 | 0.5114 | 0.5214 | 0.526 0.5314 | 0.5414 | 0.546 | 0.556 | 0.5714 | -0.581 | 0.591 | 0.596 | ${ }_{0}^{0.6114}$ | ${ }_{0}^{0.61614}$ | ${ }_{0}^{0.6314}$ | 0.636 0.6414 | 0.6514 | ${ }^{0.65614}$ | 0.666 0.6714 | ${ }_{0}^{0.6814}$ | 0.6914 |  |
| 32 | 0.5934 | 0.5024 | 0.5124 | 0.5224 | 0.5324 | 0.5424 | 5524 | 0.5624 | 0.5724 | - 0.582 | 0.59 | 6024 | 0.6124 | 0.622 | . 324 |  | 0.6524 | 0.6624 | 0.6724 | 0.6824 | 0.6924 |  |
| 23 | 0.5996 | 0.5086 | 0.5186 | 0.5286 | 0.5386 | 0.5486 | 0.5586 | 0.5686 | 0.5786 | 0. 588 | 0.598 | 0.6086 | 0.6186 | 0.6286 | 0.6386 | 0.6486 | 0.6586 | 0.6686 | 0.6786 | 0.6886 | 0.6986 |  |
| 31 | 0.6014 | 0.5104 | 0.5204 | 0.5304 | 0.5404 |  | 0.5604 | 0.5704 | 0.5804 | 0.590 | 0.600 | 0.6104 | 0.6204 | 0.6304 | 0.6404 | 0.6504 | 0.6604 | 0.6704 | 0.6804 | 0.6904 | 0.7004 | 0.71 |
| 37 | 0.6017 | 0.5107 | 0.5207 | 0.5307 | 0.5407 | 0.5507 | 0.5607 | 0.5707 | 0.5807 | 0.590 | 0.600 | 0.6107 | 0.6207 | 0.6307 | 0.6407 | 0.6507 | 0.6607 | 0.6707 | 0.6807 | 0.6907 | 0.7007 |  |
| 38 | 0.6018 | 0.5108 | 0.5208 | 0.5308 | 0.5408 | 0.5508 | 0.5608 | 0.5708 | 0.5808 | h 590 | 0.60 | 0.6108 | 10.6208 | 0.6308 | 0.6408 | 0.6508 | 0.6608 | 0.6708 | 0.6808 | 0.6908 | 0.7008 |  |
| ${ }_{82}^{22}$ | ${ }^{0.6088}$ | ${ }^{0.5178}$ | 0.5278 0.5285 | 0.5378 0.5385 | 0.5478 0.5485 | 0.5578 0.5585 | 0.5678 0.5685 | 0.5778 | 0.5878 | 0.5978 | 0.607 | 0.6178 0.6185 | ${ }^{0.6278}$ | 0.6378 0.6385 | ${ }^{0} 0.6478$ | 0.6578 | 0.6678 | ${ }^{0.6778}$ | ${ }^{0} 0.6878$ | ${ }^{0.6978}$ | 0.7078 | - $\begin{array}{r}0.717 \\ \hline\end{array}$ |
| 82 60 | 0.6095 0.6148 | 0.5185 0.5238 | 0.5285 0.5338 | 0.5385 0.5438 | 0.5485 | 0.5585 0.5638 | 0.5685 0.5738 | 0.5785 0.5838 | ${ }^{0.5885}$ | ${ }^{0.5988}$ | ${ }_{0}^{0.6085}$ | 0.6185 0.6238 | ${ }^{0.6285}$ | 0.6385 0.6438 | 0.6485 0.6538 | ${ }^{0.6585}$ | 0.6685 0.6738 | 0.6785 0.6838 | 0.6885 0.6938 | ${ }^{0.6985}$ | ${ }_{0}^{0.7} \mathbf{0 . 7 1 3 8}$ | ${ }_{0}^{0.72}$ |
| 57 | 0.6163 | 0.5253 | 0.5353 | 0.5453 | 0.5553 | 0.5653 | 0.5753 | 0.5853 | 0.5953 | 0.605 | 0.615 | 0.6253 | 0.6353 | 0.6453 | 0.6553 | 0.6653 | 0.6753 | 0.6853 | 0.6953 | 0.7053 | 0.7153 |  |
| 96 | 0.6275 | 0.5365 | 0.5465 |  | 0.5665 | 0.5765 | 5865 | 0.5965 | 0.6065 | 0.616 | 0.626 | 0.6365 | 0.6465 | 0.6565 | 0.6665 | 0.6765 | 0.6865 | 0.6965 | 0.7065 | 0.7165 | 0.7265 |  |
| 59 | ${ }^{0.6561}$ |  |  | 0.5851 0.5992 | 0.5951 0.6092 | 0.6051 | 0. 0.6151 | 0.6251 | 0. 0.631 | 0.6451 0.6592 | 0.6551 | 0.6651 <br> 0.6792 | 0.6751 0.6892 | ${ }^{0.6851}$ | ${ }^{0.6951}$ | 0.7051 | 0.7151 | -0.7251 | 0.7351 | 0.7451 0.7592 | 0.7551 |  |
| 99 97 | 0.6702 <br> 0.686 <br> 0.6 | 0.5792 0.595 | 0.5892 0.605 | 0.5992 0.615 | ${ }^{0.6092}$ | 0.6192 0.635 | 0.6292 0.645 | 0.6392 0.655 | ${ }_{0}^{0.6492}$ | 0.6592 0.675 | 0.6692 | 0.6792 <br> 0.695 <br> 0.65 | $\begin{gathered} 0.6892 \\ 0.705 \end{gathered}$ | 0.6992 0.715 | ${ }^{0.7092}$ | $\begin{aligned} & 0.7192 \\ & 0.735 \end{aligned}$ | 0.7292 0.745 | $\begin{array}{r} 0.7392 \\ 0.755 \end{array}$ | 0.7492 0.765 | 0.7592 <br> 0.775 | 0.7692 0.785 |  |
| 58 | 0.6946 | 0.6036 | 0.6136 | 0.6236 | 0.6336 | 0.6436 | ${ }_{0.6536}$ | 0.6636 | 0.6736 | 0.683 | 0.693 | 0.7036 | 0.7136 | 0.7236 | 0.7336 | 0.7436 | 0.7536 | 0.7636 | 0.7736 | 0.7836 | 0.7936 |  |
| 98 | 0.7109 | 0.6199 | 0.6299 | 0.6399 | 0.6499 | 0.6599 | 0.6699 | 0.6799 | 0.6899 | 0.6999 | 0.7099 | 0.7199 | 0.7299 | 0.7399 | 0.7499 | 0.7599 | 0.7699 | 0.7799 | 0.7899 | 0.7999 | 0.8099 |  |


|  | M | $\bigcirc$ | の $\mathrm{m}_{\text {O }}^{\text {g }}$ | $\stackrel{\square}{\square}$ |  | $\because ヘ \stackrel{\infty}{\square}$ | $\sim$ | $\sim$ 귝 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\stackrel{\stackrel{\rightharpoonup}{4}}{\stackrel{\rightharpoonup}{4}} \mid$ | － | 7 | $\wedge$ ㅇmN | $\stackrel{\text { ® }}{\square}$ | － | のペ | $\odot$ | $\sim$ \＃$\quad \cdots$ |




Our goal for the total range is $1 \%(-.5$ to +.5$)$
For an Assembly district, that equates to abo
For an Assembly district, that equates to about a range of -287 people to +287 people How many municipalities do we split?
How many municipalities do we split over and above the minimum number necessary?

Same questions asked of municipalities

Split Municipalities:
Split Counties:
Pairings:
Total deviation range:

How does that number compare to census-driven disenfranchisement and previous cycles?
$\tau$
$\tau$
25
6
How many people are moved from an even numbered senate seat to an odd one?

## :ұиәшәs!чวиеגјиәs!の әłеиәs

SA341

$0 \quad 0010$

$\circ 010$
$\circ$
statistical loss = seat that is currently held by GOP that goes to $45 \%$ or below

słu!̣od əłeuop łou op очм \%SS גəло әsоцł = sıouop-uou dO૭
Good outcomes:


## (45\% and over) <br> 훈 $\stackrel{0}{0}$ $\stackrel{0}{0}$ 0 3 <br> DEM incumbent



## 

statistical pickup = seat that is currently held by DEM that goes to $55 \%$ or more (example: if \#13 cullen goes from $44 \%$ to $58 \%$ )

DEM incumbent weakened = positive GOP movement on composite

DEM incumbent strengthened = DEM over $45 \%$ who has negative movement on composite
GOP incumbent weakened $=$ those $55 \%$ and below who have negative movement on composite
non


SA343

| district <br> 13 |  | $\underline{\text { old }}$ |  | new |
| :---: | :---: | :---: | :--- | :--- |
|  | $43.7 \%$ | $58.6 \%$ | (open seat) |  |
| 15 |  | $48.2 \%$ | $55.5 \%$ | (staskunas) |
| 22 |  | $39.1 \%$ | $57.6 \%$ | (ott/pasch) |
| 37 |  | $51.3 \%$ | $58.1 \%$ | (open) |
| 62 |  | $44.4 \%$ | $56.6 \%$ | (mason) |

GOP seats strengthened a lot
Currently held GOP seats that start at $55 \%$ or below that improve by at least $1 \%$

| district | new incumbe | old | new | improvement | Pairings | 11 dems 13 gop |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 21 | (honadel) | 51.9\% | 52.9\% | 1.0\% |  |  |  |
| 23 | (knodl) | 51.7\% | 58.5\% | 6.8\% | 7 | 45.4\% krusick/zepnick dem | dem |
| 26 | (endsley) | 45.5\% | 56.0\% | 10.5\% | 92 | 44.3\% danau/radcliffe dem | dem |
| 36 | (mursau) | 53.1\% | 54.8\% | 1.7\% | 77 | 18.9 hulsey/berceau dem | dem |
| 42 | (ripp) | 48.5\% | 54.9\% | 6.4\% |  |  |  |
| 44 | (knilans) | 36.7\% | 38.1\% | 1.4\% |  |  |  |
| 51 | (marklein) | 44.0\% | 46.2\% | 2.2\% | 22 | 57.6\% ott/pasch gop | dem |
| 55 | (kaufert/litgen | 49.3\% | 56.4\% | 7.1\% | 60 | 66.8\% pridemore/kesslegop | dem |
| 68 | (bernier) | 45.0\% | 49.4\% | 4.4\% | 61 | 57.2\% kerkman/steinbrilgop | dem |
| 72 | (krug) | 49.0\% | 51.5\% | 2.5\% | 14 | 58.8\% kooyenga/cullen gop | dem |
| 87 | (williams) | 52.2\% | 53.7\% | 1.5\% | 33 | 61.2\% nass/jorgenson gop | dem |
| 88 | (klenke/jacqu | 44.9\% | 53.2\% | 8.3\% |  |  |  |
| 93 | (petryk) | 44.7\% | 51.1\% | 6.4\% | 55 | 56.4\% kaufert/litgens gop | gop |
| 96 | (nerison) | 45.3\% | 46.4\% | 1.1\% | 88 | 53.2\% klenke/jacque gop | gop |
|  |  |  |  |  | 89 | 55.7\% van roy/nygren gop | gop |
| Vew 31st:laudenbeck and augu:Amy goes from 42.4\% to 56.4\% but get |  |  |  |  | 31 | 56.4\% august/loudenbe gop | gop |

GOP seats strengthened a little Open Seats
Currently held GOP seats that start at 55\% or below that improve less than 1\%

| district | new incumbe | old | new | improvement | 2 | 54.8\% (jacque) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 4 | (weininger) | 53.3\% | 53.5\% | 0.2\% | 9 | 29.1\% (hispanic) |
| 5 | (steineke) | 53.7\% | 54.3\% | 0.6\% | 12 | 27.5\% (african-american) |
| 28 | (severson) | 54.9\% | 55.0\% | 0.1\% | 13 | 58.7\% |
| 30 | (knudson) | 53.3\% | 53.8\% | 0.5\% | 38 | 58.5\% |
| 34 | (meyer) | 54.5\% | 55.2\% | 0.7\% | 56 | 57.6\% |
| 35 | (tiffany) | 52.3\% | 53.1\% | 0.8\% | 65 | 35.9\% |
| 75 | (rivard) | 51.7\% | 52.1\% | 0.4\% | 32 | 62.1\% (august) |
| 86 | (petrowski) | 54.6\% | 55.1\% | 0.5\% | 80 | 38.9\% |
| 49 | (tranel) | 49.7\% | 49.8\% | 0.1\% | 90 | 40.4\% |
| 69 | (suder) | 54.1\% | 54.2\% | 0.1\% | 91 | 39.5\% |
| 25 | (ziegelbauer) | 52.8\% | 53.3\% | 0.5\% | 48 | 27.6\% |
| (bies, nygren, brooks and larson remain $\checkmark$ |  |  |  |  | 78 | 31.4\% |

GOP seats weakened a little
Currently held GOP seats that start at $55 \%$ or below that decline

| district | new incumbe | old | new | improvement |
| :---: | :---: | :---: | :---: | :---: |
| 2 | (open) | 54.9\% | 54.8\% | -0.1\% |
| 29 | (murtha) | 51.3\% | 51.0\% | -0.3\% |
| 43 | (wynn) | 44.1\% | 41.8\% | -2.3\% |

GOP seats likely lost
Currently held GOP seats that drop below 45\%
district new incumbe old new improvement

| 90 | (open) | $49.6 \%$ | $40.4 \%$ | $-9.2 \%$ |
| :--- | :--- | :--- | :--- | ---: |
| 45 | (ringhand) | $42.4 \%$ | $37.9 \%$ | $-4.5 \%$ |
| 47 | (parisi seat) | $48.7 \%$ | $32.9 \%$ | $-15.8 \%$ |

GOP Donors to the Team
Incumbents with numbers above 55\% that donate to the team

| Tauchen | Spanbauer |
| :---: | :---: |
| Ott | Stroebel |
| Looyenga | Strachota |
| Knodl | LaMaheau |
| Nass | Vos |
| Farrow | Kerkman |
| Kleefisch | Stone |
| Peterson | Kuglitsch |
| Ballweg | Craig |
| Kestell | Litgens |

DEMS weakened
Currently held DEM seats (45\% or better) that become more GOP

| district | new incumbe | old | new | improvement |
| :---: | :---: | :---: | :---: | :---: |
| 46 | (hebl) | 42.4\% | 45.0\% | 2.6\% |
| 54 | (hinz) | 45.1\% | 45.2\% | 0.1\% |
| 70 | (vruwink) | 49.7\% | 50.7\% | 1.0\% |
| 94 | (doyle) | 51.6\% | 51.9\% | 0.3\% |




## Memorandum

To: Representative Garey Bies
CC: Speaker Jeff Fitzgerald; Majority Leader Scott Suder; Rep. Robin Vos
From: Adam Foltz - Assembly Redistricting Coordinator
Date: $\quad 6 / 19 / 2011$
Re: $\quad$ New Map for the 1st District

## District Number \& New District Population

As a result of the redistricting process, your district's number did not change and will remain the 1st Assembly District.

Census results showed your current district being under populated by $-3,255$. The new 1 st District has a population of 57,220 , making it is just -224 people, or $-0.40 \%$, off from the new ideal population of 57,444 .

Comparison of Key Races in Current 1st Assembly District Versus New 1st Assembly District

| Race | Old <br> District \% | New District \% | Change in Percentage | Old <br> District <br> Votes | New District Votes | Change in Votes |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Walker 10 | 53.12\% | 53.38\% | 0.25\% | 13079 | 13795 | 716 |
| JB 10 | 60.58\% | 60.74\% | 0.15\% | 14659 | 15428 | 769 |
| McCain 08 | 42.50\% | 42.59\% | 0.09\% | 13481 | 14240 | 759 |
| JB 06 | 50.71\% | 50.70\% | 0.00\% | 12992 | 13661 | 669 |
| Bush 04 | 52.28\% | 52.87\% | 0.59\% | 16756 | 17678 | 922 |

## Memorandum

| To: | Representative Garey Bies |
| :--- | :--- |
| CC: | Speaker Jeff Fitzgerald; Majority Leader Scott Suder; Rep. Robin Vos |
| From: | Adam Foltz - Assembly Redistricting Coordinator |
| Date: | $6 / 19 / 2011$ |
| Re: | New Map for the 1st District |

## District Number \& New District Population

As a result of the redistricting process, your district's number did not change and will remain the 1st Assembly District.

Census results showed your current district being under populated by $-3,255$. The new 1 st District has a population of 57,220 , making it is just -224 people, or $-0.40 \%$, off from the new ideal population of 57,444.

Comparison of Key Races in Current 1st Assembly District Versus New 1st Assembly District

| Race | Old <br> District \% | New <br> District \% | Change in <br> Percentage | Old <br> District <br> Votes | New <br> District <br> Votes | Change <br> in Votes |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Walker ${ }^{\prime} 10$ | $53.12 \%$ | $53.38 \%$ | $0.25 \%$ | 13079 | 13795 | 716 |
| JB 10 | $60.58 \%$ | $60.74 \%$ | $0.15 \%$ | 14659 | 15428 | 769 |
| MCCain 08 | $42.50 \%$ | $42.59 \%$ | $0.09 \%$ | 13481 | 14240 | 759 |
| JB 06 | $50.71 \%$ | $50.70 \%$ | $0.00 \%$ | 12992 | 13661 | 669 |
| Bush 04 | $52.28 \%$ | $52.87 \%$ | $0.59 \%$ | 16756 | 17678 | 922 |



[^50]

## SA351

McLeod, Eric M (22257)

| From: | Jim Troupis [frtroupis@troupislawoffice.com] |
| :--- | :--- |
| Sent: | Tuesday, June 21, 20116:29 PM |
| To: | McLeod, Eric M (22257) |
| Cc: | AdamFolt¿@gmail.com; tottman@gmail.com; Sarah Troupis |
| Subject: | Experts |

Eric,
Ken Mayer returned from Ireland and is willing to come to work for us. I asked him to reach out to Cannon to also assist as they are co-authors on a number of papers. He is going to do that tonight. Professor Grofman has called back and would like to help, as well. I strongly believe Prof. Grofman is essential to our efforts as he brings to any three judge panel three decades of national and international redistricting work on both sides of the aisle. He has been recognized by courts as perhaps the single most respected political scientist addressing matters of redistricting. There is no doubt we will end up in Court of whatever is passed, and so having a stable of powerful experts is essential. Without Grofman in 2001 we would not have succeeded in getting the map we did get as Easterbrook followed his direction in drawing the map. We will need to put everyone under confidentiality and retention agreements which will require retainers. Let's discuss this tomorrow and get folks under contract before the map becomes public. They will want to review it ASAP.
I hope things went well today at the caucus. I am sorry I was here in Washington. I did meet with the General Counsel at the RNC and reported on this and other issues. Unfortunately I missed Dr. Hofeller.
Jim

James R. Troupis
Troupis Law Office LLC
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| Tad MayQandD |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  | Senate |  |  |  |
| DISTAICT | Tr | Asser | 0 | OISTRICT | a |  |  |
| 1 | -51.15\% | [ $51.22 \%$ | $0.07 \times 1$ | 1 | 54.04\% | 53,35\% | -0.69\% |
| 2 | 54.93\% | = $54.14 \%$ | 0.79\% |  |  |  |  |
| 3 | 56.10\% | [ $55.08 \%$ | -1.02x |  |  |  |  |
| 4 | 53.31\% | = 53.144 | -0.17x | 2 | 55.44 K | 51.91\% | -1.53\% |
| 5 | 53.74\% | = 5717\% | 3.43* |  |  |  |  |
| 5 | 59.77\% | = $51.27 \%$ | 8.50\% |  |  |  |  |
| 7 | - $48.20 \%$ | = $46.23 \%$ | -1.97x | 3 | 40.52\% | 38.89\% | -1.63\% |
| 8 | - $22.39 \%$ | e $22.45 \%$ | 0.06\% |  |  |  |  |
| 9 | 36.73\% | - 35.10\% | -1.63\% |  |  |  |  |
| 10 | 10.27\% | 2. $21.14 \%$ | 10.87\% | 4 | 17.58\% | 2239\% | 4.81\% |
| 11 | 11918 | - $21.03 \%$ | 9.12\% |  |  |  |  |
| 12 | - 29.23\% | E $2550 \%$ | .3.73\% |  |  |  |  |
| 23 | 43.67\% | $=60.10 \%$ | 15.43\% | 5 | 50.62\% | 57.88\% | 7.26\% |
| 24 | 59.06\% | - $13552 \%$ | .3.54\% |  |  |  |  |
| 15 | 48.21\% | 5797\% | 9.76\% |  |  |  |  |
| 16 | - 24.215 | 12.21\% | -2.00\% 6 | 6 | 14.12\% | 13.84\% | -0.28\% |
| 17 | - 13.21\% | 1234\% | -0.87\% |  |  |  |  |
| 18 | - $1528 \%$ | \% 16.74\% | 1.46\% |  |  |  |  |
| 19 | 29.15x | - 26.37\% | -2.78\% | 7 | 41.13\% | 39.25\% | .1.88\% |
| 20 | 43.718 | $=40.48 \%$ | -3.23\% |  |  |  |  |
| 21 | 51.92\% | - $5285 \%$ | 0.936 |  |  |  |  |
| 22 | 39.05\% | - $57.27 \%$ | 18.22\% | 8 | 52.82\% | 61.73\% | 8.91\% |
| 23 | 51.70\% | - $60.42 \%$ | 8.72\% |  |  |  |  |
| 24 | 67.29\% | 67.93\% | $0.64 \times$ |  |  |  |  |
| 25 | 52.79\% | = $53.35 \%$ | 0.56\% | 9 | 52.96\% | 56.39\% | 3.43* |
| 26 | 45.42\% | = $52.24 \%$ | 6.82\% |  |  |  |  |
| 27 | 59.20\% | - $62.87 \%$ | 3.67\% |  |  |  |  |
| 28 | 54.85\% | - $54.31 \times$ | .054\% | 10 | 53.14\% | 53.70\% | 0.56\% |
| 29 | 51.32\% | - $54.19 \%$ | 2.87\% |  |  |  |  |
| 30 | - $53.29 \%$ | - $52.59 \%$ | 0.7060 |  |  |  |  |
| 11 | 67.57\% | - $61.18 \%$ | .639\% | 11 | 67.64\% | 63.05\% | -4.59\% |
| 32 | 61.06\% | = $1.68 .24 \%$ | 718\% |  |  |  |  |
| 33 | 72.245 | - $59.16 \%$ | -13.08\% |  |  |  |  |
| 34 | $54.51 \times$ | - $52.73 \%$ | 1.78\% | 12 | 53378 | 54.01\% | 0.64\% |
| 35 | 5230\% | - 53.96\% | $1.66 \%$ |  |  |  |  |
| 36 | 53.06x | = $51.49 \%$ | 2.43\% |  |  |  |  |
| 37 | 5133\% | $=56.27 \%$ | 4.948 | 13 | 59.228 | 57.17\% | 2.205\% |
| 38 | 6588\% | = $56.42 \%$ | 9389 |  |  |  |  |
| 39 | 60.35\% | - $58.92 \%$ | -1.43\% |  |  |  |  |
| 6 | $58.50 \%$ | - $58.42 \%$ | -0.08\% | 14 | 55.86\% | 56.03\% | $0.17 \times$ |
| 41 | 60.60\% | - $57.65 \%$ | -2.95\% |  |  |  |  |
| 42 | 48.54\% | $=52.10 \%$ | 356\% |  |  |  |  |
| 43 | 44.148 | - $4271 \%$ | 1.43\% | 15 | 41.20\% | ${ }^{41} 60 \%$ | $0.40 \%$ |
| 44 | 36.74\% | - $38.57 \%$ | 1.83\% |  |  |  |  |
| 45 | 42.39\% | - $44.05 \%$ | 1.66\% |  |  |  |  |
| 45 | 42.07\% | = $45.32 \%$ | 3.25\% | 16 | 39.06\% | 36.39\% | -2.67x |
| 47 | 48.69\% | $=3738 \%$ | -11.31\% |  |  |  |  |
| 48 | 28.03x | - 27,46\% | -0.57\% |  |  |  |  |
| 69 | 49.68\% | - 47.81\% | -1.87\% | 17 | 48.46\% | 48.93\% | $0.47 \times$ |
| 50 | 52.08x | - $52.02 \%$ | -0.06\% |  |  |  |  |
| 51 | 44.018 | - $46.94 \%$ | 293\% |  |  |  |  |
| 52 | -5739\% | = $57.41 \%$ | 0.028 | 18 | 54.96\% | 55,30\% | $0.34 \times$ |
| 53 | 62.74\% | - $63.34 \%$ | $0.60 \%$ |  |  |  |  |
| 54 | 45.08\% | - $44.89 \%$ | -0.19x |  |  |  |  |
| 55 | 49.34\% | - $58.56 \%$ | 9.22\% | 19 | 53.32\% | 52.76\% | .0.56\% |
| 56 | 61.05k | - $52.77 \%$ | 8.28\% |  |  |  |  |
| 57 | 47.26x | = $46.92 \%$ | -0.34\% |  |  |  |  |
| 58 | 70.90\% | -64.63\% | -627x | 20 | 70.55\% | 68.815 | -1.74\% |
| 59 | - 72748 | - $70.76 \%$ | 1.988 |  |  |  |  |
| 50 | 68.12\% | - $30.99 \%$ | 2.878 |  |  |  |  |
| 61 | 35.98\% | - $\quad 157.52 \%$ | 21.54\% | 21 | 49.86\% | 57.45\% | 7.599 |
| 62 | - $44.35 \%$ | - $54.06 \%$ | 9.71\% |  |  |  |  |
| 3 | 63.09\% | -60.67\% | 2.420 |  |  |  |  |
| 54 | 35.66\% | - $35.86 \%$ | 0.20\% | 22 | 47.56\% | 37.61\% | .995\% |
| 65 | 45.44x | $=44.46 \%$ | -0.98\% |  |  |  |  |
| 66 | 59.128 | $=31.47 \%$ | -27.65\% |  |  |  |  |
| 67 | 51.72\% | = $51.35 \%$ | .0.37\% | 23 | 49.98\% | 50.93\% | 0.95\% |
| 58 | 45.018 | - $47.86 \%$ | 2.855 |  |  |  |  |
| 69 | 54.06\% | = 53.688 | -0.38\% |  |  |  |  |
| 70 | 49.74\% | $=40.24 \%$ | -950\% | 24 | 46.72\% | 47.05\% | 0.33x |
| 71 | . $41.68 \%$ | = $52.58 \%$ | 10.90\% |  |  |  |  |
| 72 | 49.03\% | - $49.05 \%$ | 0.03\% |  |  |  |  |
| 73 | 39,55\% | = $40.61 \%$ | 1.06\% | 25 | 44.88\% | 45.048 | $0.16 \%$ |
| 74 | 43.78\% | = $42.60 \%$ | -1.18\% |  |  |  |  |
| 75 | -51.71\% | - $52.48 \%$ | 0.77x |  |  |  |  |
| 76 | 24.29\% | = $14.46 \%$ | 9.83\% | 25 | 20.85\% | 20.98\% | $0.13 \%$ |
| 年78 | - 23.888 | - $19.23 \%$ | 4.65\% |  |  |  |  |
| 78 | - 14.09\% | = 30.35\% | 16.77\% |  |  |  |  |
| 79 | 37.49\% | $=38.25 \%$ | 0.76\% | 27 | 3838\% | 39.01\% | 0.63\% |
| 30 | 42.15\% | = 34.86\% | -7.29\% |  |  |  |  |
| 81 | 36.16\% | - $44.05 \%$ | 789\% |  |  |  |  |
| 32 | 5859\% | - $54.19 \%$ | $4.40 \%$ | 28 | 64.48\% | 58.50\% | 5.98\% |
| 83 | 699.70\% | - $62.75 \%$ | 6.95\% |  |  |  |  |
| 84 | 64.99\% | - $58.70 \%$ | 6.29\% |  |  |  |  |
| 85 | 4891\% | = 49,60\% | 0.69\% | 29 | 52.006 | 52.79\% | 0.79\% |
| $\frac{86}{87}$ | 54.56x | - $54.95 \%$ | 0.39\% |  |  |  |  |
| 87 | 52.16x | - $53.84 \times$ | 1.68\% |  |  |  |  |
| 88 | 44.85\% | - $56.69 \%$ | 11.84\% | 30 | 50.38\% | 53,21x | 2.83\% |
| $\frac{39}{} 90$ | 55.76\% | - $55.94 \%$ | 0.18\% |  |  |  |  |
| 90 | -49.5974 | - 45.68 x | -3.91\% |  |  |  |  |
| 92 | 45.879 | = $39.53 \%$ | -6.34\% | 31 | 46.899 | 44.85\% | -205\% |
| 92 | 50.79\% | = $45.44 \%$ | 535\% |  |  |  |  |
| 93 | 44.73\% | - $49.34 \%$ | $4.61 \times$ |  |  |  |  |
| 94 | - ${ }^{51.57 x}$ | $=52.06 \%$ | 0.49\% | 32 | 44.43\% | 44.15 | .0.02\% |
| 95 | - $36.02 \%$ | $=36.33 \%$ | 0.318 |  |  |  |  |
| 97 | - $45.372 \%$ | = $45.53 \%$ | $0.21 \times$ |  |  |  |  |
| 98 | - $70.96 \%$ | - 63.88 x | 7.788\% | 35 | 68.84x | 6887\% | 0.03 x |
| 99 | 73.35\% | -73.14\% | 0.21\% |  |  |  |  |


| Current Map |  |  | NewMap |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Sate GOP (55\% ${ }^{\text {a }}$ ) | 27 | 7 | Safe GOP ( $55 \%+$ ) | 33 | 11 |
| Lean GOP (52.1-54.9\%): | 13 | 8 | New Lean GOP (52.1-54.9\%): | 21 | 7 |
| Total GOP Seats (safe + lean): | 40 | 15 | Total GOP Seats (safe + lean): | 54 | 18 |
| Swing (48-52\%): | 19 | 5 | New Swing (48.52\%) | 6 | 2 |
| (ean DEM ( $45.1-479 \%$ ) | 7 | 3 | New Lean DEM (45.1-47.99): | 9 | 2 |
| Safe DEM (-45\%) | 33 | 10 | Sofe DEM (-45\%): | 30 | 11 |
| Total DEM Seats (safe + lean): | 40 | 13 | Total DEM Seats (safe + lean): | 39 | 13 |


| Joe Assertive |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assembly |  |  |  | Senate |  |  |  |
| Oistict | Tura | A |  | DISTRICT |  |  |  |
| 1 | 5115\%\% | 51.43* | 0.28\% | 1 | 54.04\%/ | 53.95\% | -0.09\% |
| 2 | 54.93\% | 55.01\% | 0.08\% |  |  |  |  |
| 3 | 56.10\% | 55.82\% | -0.28\% |  |  |  |  |
| 4 | 5331\% | -52.98\% | 0.33\% | 2 | 55.44\% | 54.51\% | -0.936 |
| 5 | 53.74\% | 53.07\% | -0.67\% |  |  |  |  |
| 6 | 59.77\% | 57.76\% | -2.01\% |  |  |  |  |
| 7 | 48.20\% | 45.41\% | -279\% | 3 | 40.52\% | 38.26\% | -2.26\% |
| 8 | 22.39\% | - 22.306 | -0.098 |  |  |  |  |
| 9 | 36.73\% | - 35.13\% | -1.60\% |  |  |  |  |
| 10 | 1027\% | - 12.82\% | 2.55\% | 4 | 1758\% | 19.36\% | 1.78\% |
| 11 | 1191\% | - 19.63\% | 7.728 |  |  |  |  |
| 12 | - 29.23\% | 2656\% | -2.67\% |  |  |  |  |
| 13 | 43.67\% | -59,22\% | 15.55\% | 5 | 50.62\% | 57.58\% | 6.96\% |
| 12 | 59.06\% | 57,74* | -132\% |  |  |  |  |
| 15 | 48.21\% | 55.34\% | 7.13\% |  |  |  |  |
| 15 | 1 14.21\% | - 11.67\% | 2.54\% | 6 | 14.12\% | 16.03\% | 1.91\% |
| 17 | 13218 | 19.87\% | $5.66 \%$ |  |  |  |  |
| 18 | - 15.28\% | - 15.35\% | 0.07\% |  |  |  |  |
| 19 | 29.15\% | 28.31\% | -0.84\% | 7 | 41.13\% | 40.81\% | -0.3260 |
| 20 | 43.718 | 43.69\% | -0.02\% |  |  |  |  |
| 23 | - $51.92 \%$ | 5 $5286 \%$ | 094\% |  |  |  |  |
| 22 | 39.05\% | 55.96\% | 16.919 | 3 | 52.82\% | 60.68\% | 788\% |
| 23 | 51.70\% | 59.30\% | 7.60\% |  |  |  |  |
| 29 | 67.29\% | . $6737 \%$ | $0.08 \%$ |  |  |  |  |
| 25 | 52.79\% | 53.05\% | 0.26\% | 9 | 52.96\% | 54.74\% | 1.78\% |
| 26 | 45.42\% | -54.67\% | 9.25\% |  |  |  |  |
| 27 | 59.20\% | 56.34\% | 286\% |  |  |  |  |
| 28 | 54.85\% | 56.43\% | 1.58\% | 10 | 53.14\% | 53.46\% | 032\% |
| 29 | 51.32x | 50.64\% | .0.68\% |  |  |  |  |
| 30 | 53.29\% | 53.16\% | .0.13\% |  |  |  |  |
| 31 | 67.57\% | -61.04\% | .6.53m | 11 | 67.64\% | 59.65\% | -7.99\% |
| 32 | 61.06\% | - $58.28 \%$ | -2.78\% |  |  |  |  |
| 33 | - 2.25 * | 59.90\% | -12.34\% |  |  |  |  |
| 34 | 54.51\% | 54.59\% | 0.08\% | 12 | 53.37\% | 54.76\% | 1.39\% |
| 35 | 52.30\% | 53.05\% | 0.75\% |  |  |  |  |
| 36 | 53.06\% | - 56.57 x | 3.51\% |  |  |  |  |
| 37 | 51.33\% | - 53.20 K | 1.87\% | 13 | 59.22\% | 5890\% | -0.32\% |
| 38 | 65.80\% | -63.61\% | -2.19\% |  |  |  |  |
| 39. | 60.35\% | - $59.17 \%$ | .118k |  |  |  |  |
| 40 | 58.50\%. | 58.26\% | -0.245 | 14 | 55.86\% | 55.30\% | 0.56x |
| 41 | 60.60\% | 59944\% | 0.66\% |  |  |  |  |
| 42 | 48.54\% | 47.51\% | -1.03k |  |  |  |  |
| 43 | - $44.14 \%$ | 38.57\% | 5.578 | 15 | 41.20\% | 44.25\% | 3.05\% |
| 44 | 36.74\% | 33.69\% | 195\% |  |  |  |  |
| 45 | 42.39\% | . $54.68 \%$ | 12.299 |  |  |  |  |
| 46 | 42.07\% | 44.35\% | 2.28\% | 16 | 39.06\% | 34.658 | 4.41\% |
| 47 | 48.69\% | 27.46\% | $21.23 \%$ |  |  |  |  |
| 48 | - 28.03\% | 33.77\% | 5.74x |  |  |  |  |
| 49 | 49.68\% | 49991\% | 0.23\% | 17 | 48.46\% | 49.07\% | 0.618 |
| 50 | 5208\% | - 52.11\% | 0.03\% |  |  |  |  |
| 51 | 44.018 | 45.59\% | $1.58 \%$ |  |  |  |  |
| 52 | 5739\% | 58.85\% | 146\% | 18 | 54.96\% | 55.51\% | 0.55\% |
| 53 | . $62.74 \%$ | . $61.50 \%$ | 1.24\% |  |  |  |  |
| 54 | 45.08\%. | 45.52\% | 0.445 |  |  |  |  |
| 55 | 49.34\% | 56.48\% | 7.14\% | 19 | 5332\% | 53.02\% | -0.30\% |
| 56 | 61.05\% | 57.21\% | 3.8450 |  |  |  |  |
| 51 | 47.26\% | . $44.70 \%$ | 2.56\% |  |  |  |  |
| 58 | 70.90\% | -7092\% | 0.026 | 20 | 70.55\% | 69.708 | 0.85\% |
| 59 | 7274\% | -70.63\% | 2.11\% |  |  |  |  |
| 60 | 68812\% | -677776 | .035k |  |  |  |  |
| 61 | 35.98\% | 33.17\% | -281\% | 21 | 49.86\% | $50.10 \%$ | 0.24\% |
| ${ }^{2}$ | 44.35\% | 53.15\% | 8.808 |  |  |  |  |
| 33 | 63.09\% | - 59.05\% | 4.04\% |  |  |  |  |
| 64 | 35.66\% | 34.82\% | .0.84k | 22 | 4756\% | 46.778 | -0.79\% |
| 65 | 45.44\% | 46.67\% | 1.23\% |  |  |  |  |
| 65 | 59.12x | -57.61\% | 1.51\% |  |  |  |  |
| 67 | 51.72\% | - 52.04\% | 0.32\% | 23 | 49.98\% | 51.78\% | 1.80\% |
| 68 | 45.01\% | 50.12\% | 5.11\% |  |  |  |  |
| 9 | 54.06\% | 53.44\% | -0.62\% |  |  |  |  |
| 70 | 49.74\% | 51.03\% | 1.29\% | 24 | 46.72\% | 46.69\% | -0.03\% |
| 71 | 41.68\% | 39.96\% | -1.72\% |  |  |  |  |
| 72 | 49.03\% | . $49.40 \%$ | 0.37\% |  |  |  |  |
| 83 | 39.55\% | 40.17\% | 0.62\% | 25 | 44.88\% | 44.44\% | -0.44\% |
| 74 | 43.78\% | 41.67\% | -2.11\% |  |  |  |  |
| 75 | 51.71\% | 51.95\% | 0.25\% |  |  |  |  |
| 76 | 24.29x | . $23.03 \%$ | 1.26\% | 26 | 20.85\% | 2098\% | 0.13\% |
| 77 | 23.88\% | - 26.37\% | 2.49x |  |  |  |  |
| 78 | - 14.096 | - 14.70\% | 0.61¢ |  |  |  |  |
| 79 | 37.49\% | - $35.63 \%$ | -1.86\% | 27 | 3838\% | 39.87\% | 1.495 |
| 30 | 42.15\% | 41.59\% | -0.56\% |  |  |  |  |
| 81 | 36.16\% | 42.57\% | 6.415 |  |  |  |  |
| 32 | 58.59\% | -57.22\% | -1.37\% | 28 | 64.48\% | 62.10\% | -338\% |
| 33 | 69.70\% | 6837\% | -1.33\% |  |  |  |  |
| 35 | 64.99\% | 57.52\% | -7.47\% |  |  |  |  |
| 85 | 48.91\% | 48.61\% | -0.30\% | 29 | $52.00 \%$ | 52.67\% | $0.67 \%$ |
| 86 | 54.56\% | 55.56\% | 1.006 |  |  |  |  |
| 87 | 52.16\% | 53.65\% | 1.49\% |  |  |  |  |
| 38 | 44.85\% | 51.35\% | 6.50\% | 30 | 50.38\% | 50.56\% | 0.18\% |
| 89 | 55.76\%. | - $56.06 \%$ | 0.30\% |  |  |  |  |
| 90 | 49.59\% | 42.92\% | 6.67\% |  |  |  |  |
| 31 | 45.87\% | 48.89\% | 3.02x | 31 | 4689\% | 44.89\% | 2.006 |
| 92 | 50.79\% | 46.85\% | 3.3946 |  |  |  |  |
| 33 | 44.73\% | - 39.70\% | 5.03x |  |  |  |  |
| 94 | 51.57\% | - $51.59 \%$ | 0.028 | 32 | 44.43\% | 44.58\% | $0.15 \%$ |
| 95 | 36,02\% | 36.53\% | 0.518 |  |  |  |  |
| 96 | 45.32\% | 46.29\% | 0.97\% |  |  |  |  |
| 97 | 59.96\% | . $61.54 \%$ | 1.58\% | 33 | 68.84\% | 68.82\% | -0.028 |
| 98 | 70.96\% | 74.18\% | 322\% |  |  |  |  |
| 9 | 7335\% | 699.58\% | 3.77x |  |  |  |  |


| Curment Map |  |  | NewMas |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Safe GOP ( $55 \%$ +) | 27 | 7 | Sate GOP (55\%\%) | 36 | 9 |
| Lean GOP ( $52.1-54.9 \%$ ): | 13 | 8 | New Lean GOP (52.1-54.98) = | 15 | 7 |
| Total GOP Seats (safe + lean): | 40 | 15 | Totat GOP Seats (sale + lean): | 51 | 15 |
| Swing (48-52\%): | 19. | 5 | New Swing (48-52\%) | 11 | 4 |
| Lean DEM ( $45.1-479 \%$ ): | 7 | 3 | New Lean DEM (45.1-47.9\%): | 7 | 2 |
| Safe DEM (-45\%): | 33 | 10 | Sate DEM ( $45 \%$ ): | 30 | 11 |
| Total DEM Seats (safe + iean): | 40 | 13 | Total DEM Seats (safe + lean): | 37 | 13 |

## DRAFT

## Proposed Map Room Access Policy

## All access pass:

Adam Foltz, Tad Ottman, Speaker Fitzgerald, Majority Leader Fitzgerald, Eric McLeod, Jim Troupis, any legal staff determined by Eric or Jim.

Full map access (Ability to look at complete maps. Access to office when Tad, Adam, or a member of the legal team is present):

Senator Rich Zipperer, Representative?
John Hogan, Andrew Gustafson (once non-disclosure statements are signed)
A limited number of other staff (2 to 4) may be requested to assist in drawing at a later date. No staff allowed in the office without signing a non-disclosure statement.

Joe Handrick
Other consultants designated by the legal team.

## Limited access

Legislators will be allowed into the office for the sole purpose of looking at and discussing their district. They are only to be present when an All Access member is present. No statewide or regional printouts will be on display while they are present (with the exception of existing districts). They will be asked at each visit to sign an agreement that the meeting they are attending is confidential and they are not to discuss it (not a formal non-disclosure statement).

Question: Should legislators who have access to the full map be asked to sign a non-disclosure statement? Is there any advantage or disadvantage in their doing so?

| Team Map |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly |  |  | Senate |  |  |  |
| DISTRICT | Current | New | Deta | DIITRICT | Current | New | Delta |
| 1 | 51.15\% | 51.22\% | 0.07\% | 1 | 54.04\% | 53.73\% | -0.31\% |
| 2 | 54.93\% | 54.84\% | -0.09\% |  |  |  |  |
| 3 | 56.10\% | 55.58\% | -0.52\% |  |  |  |  |
| 4 | 53.31\% | 53.47\% | 0.16\% | 2 | 55.44\% | 55.23\% | 0.21\% |
| 5 | 53.74\% | 54.28\% | 0.54\% |  |  |  |  |
| 6 | 59.77\% | 58.33\% | -1.44\% |  |  |  |  |
| 7 | 48.20\% | 45.38\% | -2.82\% | 3 | 40.52\% | 38.12\% | -2.40\% |
| 8 | 22.39\% | 30.48\% | 8.09\% |  |  |  |  |
| 9 | 36.73\% | 29.14\% | -7.59\% |  |  |  |  |
| 10 | 10.27\% | - 12.59\% | - $2.32 \%$ | 4 | 17.58\% | 19.63\% | 2.05\% |
| 11 | 11.91\% | -19.58\% | 7.67\% |  |  |  |  |
| 12 | 29.23\% | 27.51\% | -1.72\% |  |  |  |  |
| 13 | 43.67\% | 58.67\% | 15.00\% | 5 | 50.62\% | 57.72\% | 7.10\% |
| 14 | 59.06\% | 58.64\% | -0.42\% |  |  |  |  |
| 15 | 48.21\% | 55.48\% | -7.27\% |  |  |  |  |
| 16 | 14.21\% | - 10.54\% | -3.67\% | 6 | 14.12\% | 15.55\% | 1.43\% |
| 17 | 13.21\% | 19.84\% | 6.63\% |  |  |  |  |
| 18 | 15.28\% | - 14.94\% | -0.34\% |  |  |  |  |
| 19 | 29.15\% | 28.03\% | -1.12\% | 7 | 41.13\% | 40.53\% | -0.60\% |
| 20 | 43.71\% | 43.12\% | -0.59\% |  |  |  |  |
| 21 | 51.92\% | 52.94\% | 1.02\% |  |  |  |  |
| 22 | 39.05\% | 57.64\% | 18.59\% | 8 | 52.82\% | 60.88\% | 8.06\% |
| 23 | 51.70\% | 58.49\% | 6.79\% |  |  |  |  |
| 24 | 67.29\% | 66.82\% | -0.47\% |  |  |  |  |
| 25 | 52.79\% | 53.26\% | 0.47\% | 9 | 52.96\% | 55.19\% | 2.23\% |
| 26 | 45.42\% | 55.97\% | 10.55\% |  |  |  |  |
| 27 | 59.20\% | 56.19\% | - $-3.01 \%$ |  |  |  |  |
| 28 | 54.85\% | 55.00\% | 0.15\% | 10 | 53.14\% | 53.32\% | 0.18\% |
| 29 | 51.32\% | 50.97\% | -0.35\% |  |  |  |  |
| 30 | 53.29\% | 53.78\% | 0.49\% |  |  |  |  |
| 31 | 67.57\% | 56.41\% | -11.16\% | 11 | 67.64\% | 60.14\% | -7.50\% |
| 32 | 61.06\% | 62.07\% | 1.01\% |  |  |  |  |
| 33 | 72.24\% | 61.92\% | -10.32\% |  |  |  |  |
| 34 | 54.51\% | 55.22\% | 0.71\% | 12 | 53.37\% | 54.39\% | 1.02\% |
| 35 | 52.30\% | 52.99\% | 0.69\% |  |  |  |  |
| 36 | 53.06\% | 54.84\% | 1.78\% |  |  |  |  |
| 37 | 51.33\% | 58.11\% | 6.78\% | 13 | 59.22\% | 60.17\% | 0.95\% |
| 38 | 65.80\% | 60.45\% | -5.35\% |  |  |  |  |
| 39 | 60.35\% | 62.00\% | 1.65\% |  |  |  |  |
| 40 | 58.50\% | 58.07\% | - $-0.43 \%$ | 14 | 55.86\% | 56.02\% | 0.16\% |
| 41 | 60.60\% | 55.16\% | -5.44\% |  |  |  |  |
| 42 | 48.54\% | 54.94\% | 6.40\% |  |  |  |  |
| 43 | 44.14\% | 41.82\% | -2.32\% | 15 | 41.20\% | 39.37\% | $-1.83 \%$ |
| 44 | 36.74\% | 38.06\% | 1.32\% |  |  |  |  |
| 45 | 42.39\% | 37.89\% | - $4.50 \%$ |  |  |  |  |
| 46 | 42.07\% | 44.95\% | 2.88\% | 16 | 39.06\% | \% 34.74\% | -4.32\% |
| 47 | 48.69\% | 32.92\% | - $15.77 \%$ |  |  |  |  |
| 48 | 28.03\% | 27.56\% | -0.47\% |  |  |  |  |
| 49 | 49.68\% | 49.59\% | -0.09\% | 17 | 48.46\% | 49.23\% | 0.77\% |
| 50 | 52.08\% | 52.06\% | -0.02\% |  |  |  |  |
| 51 | 44.01\% | 46.23\% | 2.22\% |  |  |  |  |
| 52 | 57.39\% | 59.06\% | 1.67\% | 18 | 54.96\% | 55.01\% | 0.05\% |
| 53 | 62.74\% | 61.85\% | -0.89\% |  |  |  |  |
| 54 | 45.08\% | 45.22\% | 0.14\% |  |  |  |  |
| 55 | 49.34\% | 56.43\% | 7.09\% | 19 | 53.32\% | 53.02\% | -0.30\% |
| 56 | 61.05\% | 57.59\% | - $-3.46 \%$ |  |  |  |  |
| 57 | 47.26\% | 44.50\% | -2.76\% |  |  |  |  |
| 58 | 70.90\% | 70.54\% | -0.36\% | 20 | 70.55\% | 69.46\% | -1.09\% |
| 59 | 72.74\% | 68.31\% | -4.33\% |  |  |  |  |
| 60 | 68.12\% | 69.52\% | 1.40\% |  |  |  |  |
| 61 | 35.98\% | 57.22\% | 21.24\% | 21 | 49.86\% | 57.77\% | 7.91\% |
| 62 | 44.35\% | 56.56\% | -12.21\% |  |  |  |  |
| 63 | 63.09\% | 59.64\% | - $3.45 \%$ |  |  |  |  |
| 64 | 35.66\% | 42.72\% | 7.06\% | 22 | 47.56\% | 36.97\% | -10.59\% |
| 65 | 45.44\% | 35.92\% | -9.52\% |  |  |  |  |
| 66 | 59.12\% | 31.71\% | -27.41\% |  |  |  |  |
| 67 | 51.72\% | 51.67\% | - $-0.05 \%$ | 23 | 49.98\% | 51.75\% | 1.77\% |
| 68 | 45.01\% | 49.38\% | 4.37\% |  |  |  |  |
| 69 | 54.06\% | 54.16\% | 0.10\% |  |  |  |  |
| 70 | 49.74\% | 50.73\% | 0.99\% | 24 | 46.72\% | 47.51\% | 0.79\% |
| 71 | 41.68\% | 40.72\% | -0.96\% |  |  |  |  |
| 72 | 49.03\% | 51.49\% | 2.46\% |  |  |  |  |
| 73 | 39.55\% | 40.16\% | 0.61\% | 25 | 44.88\% | 44.88\% | 0.00\% |
| 74 | 43.78\% | 42.89\% | -0.89\% |  |  |  |  |
| 75 | 51.71\% | 52.18\% | 0.47\% |  |  |  |  |
| 76 | 24.29\% | 14.48\% | -9.81\% | 26 | 20.85\% | 20.98\% | 0.13\% |
| 77 | 23.88\% | - $18.90 \%$ | -4.98\% |  |  |  |  |
| 78 | 14.09\% | 31.38\% | 17.29\% |  |  |  |  |
| 79 | 37.49\% | 40.62\% | 3.13\% | 27 | 38.38\% | 41.56\% | 3.18\% |
| 80 | 42.15\% | 39.90\% | - $-2.25 \%$ |  |  |  |  |
| 81 | 36.16\% | 44.56\% | - 8.40\% |  |  |  |  |
| 82 | 58.59\% | 57.08\% | -1.51\% | 28 | 64.48\% | 60.95\% | -3.53\% |
| 83 | 69.70\% | 68.32\% | - $1.38 \%$ |  |  |  |  |
| 84 | 64.99\% | 57.10\% | -7.89\% |  |  |  |  |
| 85 | 48.91\% | 48.38\% | - -0.53\% | 29 | 52.00\% | 52.47\% | 0.47\% |
| 86 | 54.56\% | 55.08\% | 0.52\% |  |  |  |  |
| 87 | 52.16\% | 53.74\% | 1.58\% |  |  |  |  |
| 88 | 44.85\% | 53.19\% | 8.34\% | 30 | 50.38\% | 50.55\% | 0.17\% |
| 89 | 55.76\% | 55.73\% | -0.03\% |  |  |  |  |
| 90 | 49.59\% | 40.40\% | -9.19\% |  |  |  |  |
| 91 | 45.87\% | 39.57\% | -6.30\% | 31 | 46.89\% | - 44.94\% | -1.95\% |
| 92 | 50.79\% | 44.30\% | -6.49\% |  |  |  |  |
| 93 | 44.73\% | 51.10\% | 6.37\% |  |  |  |  |
| 94 | 51.57\% | 51.91\% | 0.34\% | 32 | 44.43\% | 44.63\% | 0.20\% |
| 95 | 36.02\% | 36.36\% | 0.34\% |  |  |  |  |
| 96 | 45.32\% | 46.40\% | 1.08\% |  |  |  |  |
| 97 | 59.96\% | 62.92\% | 2.96\% | 33 | 68.84\% | 68.60\% | -0.24\% |
| 98 | 70.96\% | 74.85\% | 3.89\% |  |  |  |  |
| 99 | 73.35\% | 67.02\% | -6.33\% |  |  |  |  |


| Current Map |  |  | New Map |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Strong GOP (55\%+) | 27 | 7 | Strong GOP (55\%+) | 38 | 12 |
| Lean GOP (52-54.9\%): | 13 | 8 | New Lean GOP (52-54.9\%): | 14 | 5 |
| Total GOP Seats (safe + lean): | 40 | 15 | Total GOP Seats (safe + lean): | 52 | 17 |
| Swing (48-52\%): | 19 | 5 | New Swing (48-52\%) | 10 | 3 |
| Lean DEM (45.1-47.9\%): | 7 | 3 | New Lean DEM (45.1-47.9\%): | 4 | 1 |
| Safe DEM (-45\%): | 33 | 10 | Safe DEM (-45\%): | 33 | 12 |
| Total DEM Seats (safe + lean): | 40 | 13 | Total DEM Seats (safe + lean): | 37 | 13 |


|  | Assembly |  |  |  | Senate |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | DIITRIICT | Current | New | Delta | DISTRICT | Current | New | Delta |
|  | 98 | 70.96\% | 74.83\% | 3.87\% | 33 | 68.84\% | 68.56\% | -0.28\% |
|  | 58 | 70.90\% | 70.55\% | -0.35\% | 20 | 70.55\% | 68.46\% | -2.09\% |
|  | 24 | 67.29\% | 69.52\% | 2.23\% | 8 | 52.82\% | 61.64\% | 8.82\% |
|  | 83 | 69.70\% | 68.35\% | -1.35\% | 28 | 64.48\% | 60.90\% | -3.58\% |
|  | 59 | 72.74\% | 68.26\% | -4.48\% | 13 | 59.22\% | 60.34\% | 1.12\% |
|  | 99 | 73.35\% | 66.88\% | -6.47\% | 11 | 67.64\% | 60.14\% | -7.50\% |
|  | 60 | 68.12\% | 66.82\% | -1.30\% | 5 | 50.62\% | 57.79\% | 7.17\% |
|  | 97 | 59.96\% | 63.05\% | 3.09\% | 21 | 49.86\% | 57.77\% | 7.91\% |
|  | 32 | 61.06\% | 62.06\% | 1.00\% | 14 | 55.86\% | 55.96\% | 0.10\% |
| 10 | 39 | 60.35\% | 62.00\% | 1.65\% | 2 | 55.44\% | 55.23\% | -0.21\% |
| 11 | 33 | 72.24\% | 61.92\% | -10.32\% | 9 | 52.96\% | 55.19\% | 2.23\% |
| 12 | 53 | 62.74\% | 61.81\% | -0.93\% | 18 | 54.96\% | 55.01\% | 0.05\% |
| 13 | 37 | 51.33\% | 60.58\% | 9.25\% | 12 | 53.37\% | 54.43\% | 1.06\% |
| 14 | 63 | 63.09\% | 59.64\% | -3.45\% | 1 | 54.04\% | 53.73\% | -0.31\% |
| 15 | 52 | 57.39\% | 59.06\% | 1.67\% | 10 | 53.14\% | 53.31\% | 0.17\% |
| 16 | 14 | 59.06\% | 58.76\% | -0.30\% | 19 | 53.32\% | 53.02\% | -0.30\% |
| 17 | 13 | 43.67\% | 58.68\% | 15.01\% | 29 | 52.00\% | 52.52\% | 0.52\% |
| 18 | 23 | 51.70\% | 58.51\% | 6.81\% | 23 | 49.98\% | 51.69\% | 1.71\% |
| 19 | 38 | 65.80\% | 58.46\% | -7.34\% | 30 | 50.38\% | 50.55\% | 0.17\% |
| 20 | 6 | 59.77\% | 58.33\% | -1.44\% | 17 | 48.46\% | 49.23\% | 0.77\% |
| 21 | 22 | 39.05\% | 57.63\% | 18.58\% | 24 | 46.72\% | 47.39\% | 0.67\% |
| 22 | 56 | 61.05\% | 57.55\% | -3.50\% | 31 | 46.89\% | 44.93\% | -1.96\% |
| 23 | 40 | 58.50\% | 57.51\% | -0.99\% | 25 | 44.88\% | 44.84\% | -0.04\% |
| 24 | 61 | 35.98\% | 57.23\% | 21.25\% | 32 | 44.43\% | 44.63\% | 0.20\% |
| 25 | 82 | 58.59\% | 57.13\% | -1.46\% | 27 | 38.38\% | 40.98\% | 2.60\% |
| 26 | 84 | 64.99\% | 56.94\% | -8.05\% | 7 | 41.13\% | 40.53\% | -0.60\% |
| 27 | 62 | 44.35\% | 56.56\% | 12.21\% | 15 | 41.20\% | 39.51\% | -1.69\% |
| 28 | 55 | 49.34\% | 56.43\% | 7.09\% | 3 | 40.52\% | 38.12\% | -2.40\% |
| 29 | 31 | 67.57\% | 56.41\% | -11.16\% | 22 | 47.56\% | 36.97\% | -10.59\% |
| 30 | 27 | 59.20\% | 56.13\% | -3.07\% | 16 | 39.06\% | 35.24\% | -3.82\% |
| 31 | 26 | 45.42\% | 56.03\% | 10.61\% | 26 | 20.85\% | 20.98\% | 0.13\% |
| 32 | 41 | 60.60\% | 55.79\% | -4.81\% | 4 | 17.58\% | 19.63\% | 2.05\% |
| 33 | 89 | 55.76\% | 55.73\% | -0.03\% | 6 | 14.12\% | 15.57\% | 1.45\% |
| 34 | 3 | 56.10\% | 55.58\% | -0.52\% |  |  |  |  |
| 35 | 15 | 48.21\% | 55.52\% | 7.31\% |  |  |  |  |
| 36 | 34 | 54.51\% | 55.22\% | 0.71\% |  |  |  |  |
| 37 | 86 | 54.56\% | 55.11\% | 0.55\% |  |  |  |  |
| 38 | 28 | 54.85\% | 54.99\% | 0.14\% |  |  |  |  |
| 39 | 2 | 54.93\% | 54.84\% | -0.09\% |  |  |  |  |
| 40 | 36 | 53.06\% | 54.84\% | 1.78\% |  |  |  |  |
| 41 | 42 | 48.54\% | 54.63\% | 6.09\% |  |  |  |  |
| 42 | 5 | 53.74\% | 54.28\% | 0.54\% |  |  |  |  |
| 43 | 87 | 52.16\% | 53.92\% | 1.76\% |  |  |  |  |
| 硡 | 69 | 54.06\% | 53.85\% | -0.21\% |  |  |  |  |
| 45 | 30 | 53.29\% | 53.80\% | 0.51\% |  |  |  |  |
| 46 | 4 | 53.31\% | 53.47\% | 0.16\% |  |  |  |  |
| 47 | 25 | 52.79\% | 53.27\% | 0.48\% |  |  |  |  |
| 48 | 88 | 44.85\% | 53.19\% | 8.34\% |  |  |  |  |
| 咗 | 35 | 52.30\% | 53.08\% | 0.78\% |  |  |  |  |
| 50 | 21 | 51.92\% | 52.94\% | 1.02\% |  |  |  |  |
| 51 | 75 | 51.71\% | 52.14\% | 0.43\% |  |  |  |  |
| 52 | 94 | 51.57\% | 51.91\% | 0.34\% |  |  |  |  |
| 53 | 50 | 52.08\% | 51.87\% | -0.21\% |  |  |  |  |
| 54 | 67 | 51.72\% | 51.67\% | -0.05\% |  |  |  |  |
| 55 | 1 | 51.15\% | 51.22\% | 0.07\% |  |  |  |  |
| 56 | 93 | 44.73\% | 51.15\% | 6.42\% |  |  |  |  |
| 57 | 29 | 51.32\% | 50.94\% | -0.38\% |  |  |  |  |
| 58 | 72 | 49.03\% | 50.42\% | 1.39\% |  |  |  |  |
| 59 | 70 | 49.74\% | 50.17\% | 0.43\% |  |  |  |  |
| 60 | 49 | 49.68\% | 49.75\% | 0.07\% |  |  |  |  |
| 61 | 68 | 45.01\% | 49.66\% | 4.65\% |  |  |  |  |
| 62 | 85 | 48.91\% | 48.31\% | -0.60\% |  |  |  |  |
| 63 | 96 | 45.32\% | 46.40\% | 1.08\% |  |  |  |  |
|  | 51 | 44.01\% | 46.28\% | 2.27\% |  |  |  |  |
| 65 | 7 | 48.20\% | 45.35\% | -2.85\% |  |  |  |  |
| 66 | 54 | 45.08\% | 45.26\% | 0.18\% |  |  |  |  |
| 67 | 46 | 42.07\% | 45.19\% | 3.12\% |  |  |  |  |
| 68 | 81 | 36.16\% | 44.56\% | 8.40\% |  |  |  |  |
| 69 | 57 | 47.26\% | 44.49\% | -2.77\% |  |  |  |  |
| 70 | 92 | 50.79\% | 44.30\% | -6.49\% |  |  |  |  |
| 71 | 20 | 43.71\% | 43.06\% | -0.65\% |  |  |  |  |
| 72 | 64 | 35.66\% | 42.72\% | 7.06\% |  |  |  |  |
| 73 | 74 | 43.78\% | 42.43\% | -1.35\% |  |  |  |  |
| 74 | 43 | 44.14\% | 41.96\% | -2.18\% |  |  |  |  |
| 75 | 71 | 41.68\% | 41.92\% | 0.24\% |  |  |  |  |
|  | 73 | 39.55\% | 40.52\% | 0.97\% |  |  |  |  |
|  | 90 | 49.59\% | 40.40\% | -9.19\% |  |  |  |  |
| 78 | 79 | 37.49\% | 40.09\% | 2.60\% |  |  |  |  |
| 79 | 91 | 45.87\% | 39.53\% | -6.34\% |  |  |  |  |
| 80 | 80 | 42.15\% | 38.65\% | -3.50\% |  |  |  |  |
| 81 | 45 | 42.39\% | 38.18\% | -4.21\% |  |  |  |  |
| 82 | 44 | 36.74\% | 38.07\% | 1.33\% |  |  |  |  |
| 83 | 95 | 36.02\% | 36.36\% | 0.34\% |  |  |  |  |
| 84 | 65 | 45.44\% | 35.92\% | -9.52\% |  |  |  |  |
| 85 | 47 | 48.69\% | 34.47\% | -14.22\% |  |  |  |  |
| 86 | 66 | 59.12\% | 31.71\% | -27.41\% |  |  |  |  |
| 87 | 78 | 14.09\% | 30.86\% | 16.77\% |  |  |  |  |
| 88 | 8 | 22.39\% | 30.48\% | 8.09\% |  |  |  |  |
| 89 | 9 | 36.73\% | 29.14\% | -7.59\% |  |  |  |  |
| 90 | 19 | 29.15\% | 28.03\% | -1.12\% |  |  |  |  |
| 91 | 12 | 29.23\% | 27.51\% | -1.72\% |  |  |  |  |
| 92 | 48 | 28.03\% | 27.45\% | -0.58\% |  |  |  |  |
| 93 | 17 | 13.21\% | 19.88\% | 6.67\% |  |  |  |  |
| 94 | 11 | 11.91\% | 19.58\% | 7.67\% |  |  |  |  |
| 95 | 77 | 23.88\% | - 19.23\% | -4.65\% |  |  |  |  |
| 96 | 18 | 15.28\% | 14.94\% | -0.34\% |  |  |  |  |
| 97 | 76 | 24.29\% | 14.46\% | -9.83\% |  |  |  |  |
| 98 | 10 | 10.27\% | 12.59\% | 2.32\% |  |  |  |  |
| 99 | 16 | 14.21\% | 10.54\% | -3.67\% |  |  |  |  |


|  | Current Map |  |  |  |
| :--- | ---: | ---: | :--- | :--- | ---: | ---: | ---: |



| Current Map |  |  | New Map |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Assembly | Senate |  | Assembly | Senate |
| Strong GOP (55\%+) | 27 | 7 | Strong GOP (55\%+) | 38 | 12 |
| Lean GOP (52-54.9\%): | 13 | 8 | New Lean GOP (52-54.9\%): | 14 | 5 |
| Total GOP Seats (safe + lean): | 40 | 15 | Total GOP Seats (safe + lean): | 52 | 17 |
| Swing (48-52\%): | 19 | 5 | New Swing (48-52\%) | 10 | 3 |
| Lean DEM (45.1-47.9\%): | 7 | 3 | New Lean DEM (45.1-47.9\%): | 4 | 1 |
| Safe DEM (-45\%): | 33 | 10 | Safe DEM (-45\%): | 33 | 12 |
| Total DEM Seats (safe + lean): | 40 | 13 | Total DEM Seats (safe + lean): | 37 | 13 |

## SA358



Wisconsin Sensitivity Testing



## 2002 Assembly Districts - Over/Under Population





[^0]:    ${ }^{1}$ In making this analysis, the Mayer Report used 2012 election results and further assumed that all districts had been contested and no incumbents had run. These are both standard assumptions made by political scientists to determine a plan's underlying partisanship.

[^1]:    Calumet, Fond du Lac, Manitowoc and Sheboygan Counties:

[^2]:    ${ }^{1}$ It would seem that the number of wasted votes for the winner should be one vote less than what plaintiffs' calculation suggests for each district. In this example, the party would need 51 votes to win, so Party A would have nine rather than ten

[^3]:    wasted votes for each district it won (60-51=9) and Party B would have 29 rather than 30 wasted votes for the districts it won (80$51=29$ ). Regardless, the parties do not discuss this potential discrepancy, so we need not consider it.

[^4]:    ${ }_{2}$ They have not conceded, however, that the plaintiffs could meet a more demanding showing of partisan intent. Trans., dkt. \#89, at 88.

[^5]:    Plaintiffs' proposed element regarding intent requires them to prove that defendants disadvantaged plaintiffs intentionally on the basis of political affiliation. Plts.' Br., dkt. \#68, at 58. In their opening brief, defendants limited their discussion of this element to an argument that a requirement to prove intent did not help to overcome the alleged problems with the efficiency gap as a measure of discriminatory effect. They did not challenge the validity of the element itself. However, in their reply brief, defendants argued for the first time that "plaintiffs' intent element is inconsistent with Supreme Court precedent" because it is not sufficiently demanding. Dfts.' Reply Br., dkt. \#73, at 3. In particular, defendants rely on the plurality opinion in Vieth, 541 U.S. at 286, for the view that "partisan districting is a lawful and common practice," so that any successful partisan gerrymandering claim must show an "excess" of a partisan motive.

[^6]:    ${ }^{1}$ The federal requirements are equal population, compliance with Section 2 of the Voting Rights Act, and the ban on racially gerrymandered districts. The state requirements are contiguity, compactness, and respect for political subdivisions (counties, towns, cities, and villages).

[^7]:    ${ }^{2}$ Winners are determined by which candidate receives $>50 \%$ of the vote in a two party race. Seats votes curves depend on the number of seats a party wins in an election (determined by the number of races in which that party received $>50 \%$ of the vote) and the statewide vote totals in legislative races or some other set of statewide races

[^8]:    ${ }^{3}$ Legislative Technology Services Bureau data show 6,592 wards in Wisconsin, of which 66 are unpopulated and another 50 have fewer than 10 people. The average populated ward contains 869 people. Wisconsin statutes 5.15 (2)(b) specifies a permissible population range for wards of 300-4,000, depending on a municipality's size, with exceptions allowed in certain circumstances (for example, when single blocks exceed a permitted ward size, or when a municipality is divided into multiple counties or school districts, contains islands, or has wards that must be altered to match district boundaries).

[^9]:    ${ }^{4}$ The larger n also means that OLS is an accurate method of estimating the underlying relationships, whereas more complicated techniques may be required with smaller sample sizes (Afshartous and de Leeuw 2005).
    ${ }^{5}$ The Census Bureau uses the term "Voting Tabulation District" (VTD). Most states call VTDs precincts. In Wisconsin these units are called "wards."
    ${ }^{6}$ These are known as FIPS (Federal Information Processing Standard) codes. http://www.census.gov/geo/reference/ansi.html.

[^10]:    ${ }^{7}$ The files are available at http://legis.wisconsin.gov/gis/data. The 2012 election results are in the file Wards_111312_ED_110612.xlsx.
    ${ }^{8}$ As I note in the Annex, I was not able to allocate $0.21 \%$ of the vote in 2012 because of inconsistencies between electoral data reported by the GAB and the geographic redistricting data reported by the LTSB. This small number of votes will not change any of my analysis or conclusions, and such errors are inevitable when working with large data sets.

[^11]:    ${ }^{9}$ When using dummy variables (which take binary values of either 0 or 1 ) to measure effects in units or conditions across the full population, one unit must be excluded, as otherwise perfect collinearity prevents estimation (Greene 1990, 240-241).

[^12]:    ${ }^{10}$ The reliance on actual numbers of voters eliminates the Modified Areal Unit Problem, which results when group statistics such as vote percentages or demographic fractions are aggregated into different geographic units levels. All of my variables and measures are scale invariant (see King 1996).
    ${ }^{11}$ The number of votes cast in Assembly races varies considerably even in in contested races. In 2012, the number of major party votes cast in the highest turnout Assembly election in the $23^{\text {rd }}$ Assembly district, 36,205 , was almost twice the number cast in the $90^{\text {th }}$ Assembly district, 18,735 , and almost 5 times the number cast in the uncontested $8^{\text {th }}$ district, 7,869 (numbers taken from GAB figures).
    ${ }^{12}$ The voting eligible population (VEP) adjusts the voting age population by removing adults who are not eligible to vote. In Wisconsin, the two largest categories of ineligible adults that can be identified geographically are noncitizens and adults in prison for felonies. Noncitizens were removed using the 2008-2012 5 year American Community Survey county level noncitizen estimates (available at http://www.census.gov/acs/www/data_documentation/2012_release/. Institutionalized prison populations were identified using Census Bureau "Advanced Group Quarters" files for Wisconsin, available at .http://www2.census.gov/census_2010/02Advance Group Quarters/, and described in http://www.census.gov/newsroom/releases/archives/2010_census/cb11-tps13.html. There are individuals on probation or extended supervision who are also ineligible to vote. I was not able to systematically identify their locations, but they are dispersed enough that they will not have a material effect on my resulting estimates or conclusions. All regression results and district estimates are materially unchanged when the unadjusted data are used.

[^13]:    ${ }^{13}$ In the political science literature, the incumbency advantage has been attributed to the political skills and campaign experience of officeholders, higher name recognition, fundraising advantages, constituency service, redistricting, and the ability to scare off quality challengers. ${ }^{14}$ Incumbents were identified using 2012 election data in the 2013 Wisconsin Blue Book. In the $43^{\text {rd }}$ and $61^{\text {st }}$ Assembly districts two incumbents were paired against each other; these districts were coded as having no incumbent, since the advantage cancels. In the $7^{\text {th }}$ Assembly district, the Democratic incumbent lost in the primary election and ran a write in campaign in the general election. Because the incumbent was not on the ballot, this district is also coded as having no incumbent.

[^14]:    ${ }^{15}$ This major-party contested definition is standard. It counts as uncontested four districts where one major party candidate was not on the ballot but received votes as a write in (districts 7, 17, 48 , and 57), and one district (district 95) where one major party candidate was on the ballot but did not campaign and received only 50 votes (or $0.24 \%$ ). This is consistent with methods used in the literature, which often uses a $95 \%$ threshold for the winning candidate as a standard (Gelman and King 1990, 274).
    ${ }^{16}$ Standard errors were adjusted to reflect the aggregation (or clustering) of wards into districts. The full set of variables is included in an appendix to this report.

[^15]:    ${ }^{17}$ All of the votes for a losing candidate are defined as wasted, whereas only those votes in excess of the number required to win are wasted for the winner.
    ${ }^{18}$ The vote percentages were calculated using the actual and predicted vote totals.

[^16]:    ${ }^{19}$ Uncontested districts were not included in the analysis for reasons specified in section $B(1)(f)$ above.
    ${ }^{20}$ In twelve districts (districts $1,6,34,35,36,49,68,74,75,93,94$ and 96 ), at least one county was entirely contained in a single district, making it impossible to estimate the fixed effect coefficient value for that county. Consequently, when the out-of-sample predictions were calculated, a variable was missing. An accurate test involves districts for which it was possible to estimate the full model.

[^17]:    ${ }^{21}$ Deposition, January 20, 2012, p. 196.
    ${ }^{22}$ I generated this data by calculating predicted values for my model in Act 43 districts, setting all incumbency variables to zero.

[^18]:    ${ }^{23}$ The identifier is a combination of state, county, Census tract, and block FIPS codes.

[^19]:    ${ }^{24}$ Act 43 figures are taken from the Joint Final Pretrial Report filed in Baldus et al. vs Brennan et al.11-CV-562, filed February 24, 2012.

[^20]:    ${ }^{1}$ Wisconsin Statutes $5.15(6)(b)$ allows municipalities with a population under 35,000 to combine wards for purposes of using a common polling place, and allows for the tabulation and reporting of combined ward vote totals.
    ${ }^{2}$ In 2012 the reporting units were Wards 1-2; 5-6; 7-8; 9-10; 11-12; 13-14; 15-16; 3, 4, and 22; and 17-18, 21, and 23-25.

[^21]:    ${ }^{3}$ The voting eligible population (VEP) adjusts the voting age population by removing adults who are not eligible to vote. In Wisconsin, the two largest categories of ineligible adults are noncitizens and adults in prison for felonies. Noncitizens were removed using the 2008-2012 5 year American Community Survey county level noncitizen estimates (available at http://www.census.gov/acs/www/data_documentation/2012_release/. Institutionalized prison populations were identified and removed using Census Bureau "Advanced Group Quarters" files for Wisconsin, available at http://www2.census.gov/census_2010/02-Advance_Group_Quarters/, and described in http://www.census.gov/newsroom/releases/archives/2010_census/cb11tps13.html.
    ${ }^{4}$ Table: Vote for President and Vice President by Ward, November 6, 2012 General Election, 938.

[^22]:    ${ }^{1}$ "Analysis of the Efficiency Gaps of Wisconsin's Current Legislative District Plan and Plaintiff's Demonstration Plan," July 3, 2015.

[^23]:    ${ }^{2}$ The Cook Political Report notes that it "introduced the Partisan Vote Index (PVI) as a means of providing a more accurate picture of the competitiveness of each of the 435 congressional districts." http://cookpolitical.com/story/5604

[^24]:    ${ }^{3}$ These occurred in the R file "Wisconsin_clustering_computation.R" that Trende disclosed.
    ${ }^{4}$ This inconsistency could well affect Trende's results, as the vote percentages were vastly different in the two races in Wisconsin. Democrats garnered $53.8 \%$ of the two-party vote in the gubernatorial election, but $60.5 \%$ in the Senate race (GAB data).

[^25]:    ${ }^{5}$ Byers and Raferty (1998) use a near neighbor method to estimate the statistical relationship between points in space and how they differ from random distributions, or "clutter," in the context of distinguishing landmines from other objects during aerial reconnaissance. Neither their work nor Clark and Evans (1954) supports Trende's use of the method.

[^26]:    ${ }^{6}$ Calculated directly from the LTSB shape files of 2012 wards, obtained from http://legis.wisconsin.gov/gis/data.

[^27]:    ${ }^{7}$ The pattern Trende identifies is largely constant across all elections; adding the additional cycles will not change the results.

[^28]:    ${ }^{8}$ Goedert's definition of bias is essentially identical to the efficiency gap. He "compare[s] the mean vote share with the expected seat share under a 'fair' map with zero bias and a historically average seats-votes curve" $(2014,3)$. In the "historically average seats-votes curve," "a $1 \%$ increase in vote share will produce about a $2 \%$ increase in seat share," which is the same seat-vote relationship implied by a zero efficiency gap $(2014,3)$. Goedert's bias estimates are thus largely indistinguishable from the efficiency gap calculations of Stephanopoulos and McGhee (2015).

[^29]:    ${ }^{9}$ Goedert generated two models, one for states with fewer than 6 congressional districts, and another for states with more than six. As Wisconsin has 8 districts, I use the latter.

[^30]:    ${ }^{10}$ Calculated as 435/50.

[^31]:    ${ }^{11}$ I calculated Global Moran's I using the method in Bivand and Piras (2015) and the R module spdep available at https://cran.r-project.org/web/packages/spdep/index.html. I calculated the isolation index using a Stata module (seg), available at http://econpapers.repec.org/software/bocbocode/s375001.htm.

[^32]:    ${ }^{12}$ Ward level 2008 and 2012 results cannot easily be compared because ward boundaries were redrawn after the 2010 Census.

[^33]:    ${ }^{13}$ I recalculated vote estimates using predicted values of Democratic and Republican Assembly votes when one of the parties had an incumbent running.
    ${ }^{14}$ This information was provided to me by counsel.

[^34]:    ${ }^{15}$ We can use these calculations to determine how many more Democratic legislators would have been elected in 2012 if either the Demonstration Plan, or a plan with an efficiency gap of exactly zero, had been in place. Under the open-seat baseline, $9.49 \%$ more Democrats would have been elected under the Demonstration Plan ( $11.69 \%-2.20 \%$ ), and $11.69 \%$ more under a plan with an efficiency gap of exactly zero. Similarly, under the incumbent baseline, $9.33 \%$ more Democrats would have been elected under the Demonstration Plan ( $13.04 \%-3.71 \%$ ), and $13.04 \%$ more under a plan with an efficiency gap of exactly zero. In all cases, these are very large differences, amounting to anywhere from nine to thirteen Assembly seats.

[^35]:    ${ }^{1}$ The curve labeled "Cube Law" in Figure 2 is generated assuming that $S /(1-S)=[V /(1-V)]^{3}$, an approximation for the lack of proportionality we observe in single-member district systems, though hardly a "law."

[^36]:    1 It is also worth stressing that the confidence interval is computed so as to take into account uncertainty from all known sources: in the underlying efficiency gap scores themselves, the fact that the 2012 EG scores for Wisconsin are large by historical standards, and in the regression relationship between first-election $E G$ and plan-average $E G$.

[^37]:    ${ }^{1}$ "Analysis of the Efficiency Gaps of Wisconsin's Current Legislative District Plan and Plaintiff's Demonstration Plan," July 3, 2015.

[^38]:    ${ }^{2}$ The Cook Political Report notes that it "introduced the Partisan Vote Index (PVI) as a means of providing a more accurate picture of the competitiveness of each of the 435 congressional districts." http://cookpolitical.com/story/5604
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    ${ }^{4}$ This inconsistency could well affect Trende's results, as the vote percentages were vastly different in the two races in Wisconsin. Democrats garnered $53.8 \%$ of the two-party vote in the gubernatorial election, but $60.5 \%$ in the Senate race (GAB data).

[^39]:    ${ }^{5}$ Byers and Raferty (1998) use a near neighbor method to estimate the statistical relationship between points in space and how they differ from random distributions, or "clutter," in the context of distinguishing landmines from other objects during aerial reconnaissance. Neither their work nor Clark and Evans (1954) supports Trende's use of the method.

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[^48]:    ${ }^{15}$ We can use these calculations to determine how many more Democratic legislators would have been elected in 2012 if either the Demonstration Plan, or a plan with an efficiency gap of exactly zero, had been in place. Under the open-seat baseline, $9.49 \%$ more Democrats would have been elected under the Demonstration Plan ( $11.69 \%-2.20 \%$ ), and $11.69 \%$ more under a plan with an efficiency gap of exactly zero. Similarly, under the incumbent baseline, $10.26 \%$ more Democrats would have been elected under the Demonstration Plan ( $14.15 \%-3.89 \%$ ), and $14.15 \%$ more under a plan with an efficiency gap of exactly zero. In all cases, these are very large differences, amounting to anywhere from nine to thirteen Assembly seats.

[^49]:    ${ }^{16}$ There were some minor discrepancies in the underlying data used in my earlier report. The updates are reflected in the March 31, 2016 revision. The discrepancies caused no material difference in the results.

[^50]:    SA350

